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Question 1

(a)

We should pay attention in angrist90 dataset that which factors also influenced, we may consider adding something related to both mother and father's own sides: father's and mother's race; father's and mother's education; mother's age

(b)

```
. reg p kid3 if yearschm!=. //model1
```

Source	SS	df	MS	Number of obs	=	614,389
Model	1831.06681	1	1831.06681	F(1, 614387)	=	8755.88
Residual	128483.283	614,387	.209124351	Prob > F	=	0.0000
				R-squared	=	0.0141
				Adj R-squared	=	0.0140
Total	130314.349	614,388	.212104321	Root MSE	=	.4573

p	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
kid3	-.115594	.0012353	-93.57	0.000	-.1180152	-.1131728
_cons	.7334874	.0007159	1024.60	0.000	.7320843	.7348905

```
est store model1
```

```
reg p kid3 yearschm //model2
```

Source	SS	df	MS	Number of obs	=	614,389
Model	5230.56842	2	2615.28421	F(2, 614386)	=	12845.74
Residual	125083.781	614,386	.203591522	Prob > F	=	0.0000
				R-squared	=	0.0401
				Adj R-squared	=	0.0401
Total	130314.349	614,388	.212104321	Root MSE	=	.45121

p	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
kid3	-.0965608	.0012278	-78.65	0.000	-.0989671	-.0941544
yearschm	.0303201	.0002346	129.22	0.000	.0298602	.03078
_cons	.3984259	.0026874	148.25	0.000	.3931586	.4036932

```
est store model2
```

	(1)	(2)
	model1	model2
VARIABLES	p	p
kid3	-0.116*** (0.00124)	-0.0966*** (0.00123)
yearschrn		0.0303*** (0.000235)
Constant	0.733*** (0.000716)	0.398*** (0.00269)
Observations	614,389	614,389
R-squared	0.014	0.040

Standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

We can see that the effect of kid3 is decreasing(down) when add/control the mother's education. Maybe it is mainly because there is a positive relationship ($\gamma_1 > 0$) between mother labor supply and mother's education.

Due to $\beta_1 = \beta_2 + \gamma_1 \frac{cov(kid3, edum)}{var(kid3)} < 0$ and since β_2 also < 0 ; $\gamma_1 > 0$ and **absolute**

magnitude $\beta_1 > \beta_2$, we can predict that **$cov(kid3, edum) < 0$** . So, **the sign of the correlation between mother's education and the kid3 is negative.**

(c)

. reg yearschrn kid3

Source	SS	df	MS	Number of obs	=	614,389
Model	54000.4999	1	54000.4999	F(1, 614387)	=	8971.93
Residual	3697887.94	614,387	6.01882517	Prob > F	=	0.0000
Total	3751888.44	614,388	6.10670853	R-squared	=	0.0144
				Adj R-squared	=	0.0144
				Root MSE	=	2.4533

yearschrn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
kid3	-.6277433	.0066273	-94.72	0.000	-.6407327 -.614754
_cons	11.0508	.0038405	2877.42	0.000	11.04327 11.05833

It shows that our prediction from (b) is correct.

(d)

```
. reg kid3 samesex, robust //see if it's OK to proceed as IV
```

Linear regression	Number of obs	=	614,389
	F(1, 614387)	=	2509.06
	Prob > F	=	0.0000
	R-squared	=	0.0041
	Root MSE	=	.47131

kid3	Robust		t	P> t	[95% Conf. Interval]	
	Coef.	Std. Err.				
samesex	.0602308	.0012024	50.09	0.000	.0578741	.0625876
_cons	.3056107	.0008324	367.14	0.000	.3039792	.3072422

```
. ivregress 2sls p (kid3=samesex), robust //IV method
```

We can see that t=50, so it's okay to let samesex be IV

```
. ivregress 2sls p (kid3=samesex), robust //IV method
```

Instrumental variables (2SLS) regression	Number of obs	=	614,389
	Wald chi2(1)	=	16.67
	Prob > chi2	=	0.0000
	R-squared	=	0.0127
	Root MSE	=	.45762

p	Robust		z	P> z	[95% Conf. Interval]	
	Coef.	Std. Err.				
kid3	-.0791607	.0193863	-4.08	0.000	-.1171571	-.0411643
_cons	.7212525	.0065336	110.39	0.000	.7084468	.7340582

Instrumented: kid3
Instruments: samesex

VARIABLES	(1)	(2)
	modell	modell_2
	p	p
kid3	-0.116*** (0.00124)	-0.0792*** (0.0194)
Constant	0.733*** (0.000716)	0.721*** (0.00653)

Observations	614,389	614,389
R-squared	0.014	0.013

Standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

With IV equation model 1_2 , it is meaning that mother have children that first two are same gender can reduce 0.0792 working hours.

(e)

$$plim(\hat{\beta}_{1,OLS}) = \beta_1 + \frac{\sigma_u}{\sigma_x} \cdot Corr(x, u)$$

$$plim(\hat{\beta}_{1,IV}) = \beta_1 + \frac{\sigma_u}{\sigma_x} \cdot \frac{Corr(z, u)}{Corr(z, x)}$$

```
. corr kid3 yearschm
(obs=614,389)
```

	kid3 yearschm	
kid3	1.0000	
yearschm	-0.1200	1.0000

```
. corr samesex kid3
(obs=614,389)
```

	samesex	kid3
samesex	1.0000	
kid3	0.0638	1.0000

```
. corr samesex yearschm
(obs=614,389)
```

	samesex yearschm	
samesex	1.0000	
yearschm	-0.0005	1.0000

```
. di -0.0005/0.0638
-.00783699
```

So, the bias from IV (-0.078) is smaller than the bias from OLS(-0.12). But we also should pay attention corr(z,x), the bias of IV could be blown up by it.

(f)

We use Hausman method to test whether is endogenous. We test model 1 and model 2.

```
. hausman IV modell, constant sigmamore
```

Note: the rank of the differenced variance matrix (1) does not equal the number of coefficients being tested (2); be sure this is what you expect, or there may be problems computing the test. Examine the output of your estimators for anything unexpected and possibly consider scaling your variables so that the coefficients are on a similar scale.

	Coefficients		(b-B) Difference	sqrt(diag(V_b-V_B)) S.E.
	(b) IV	(B) modell		
kid3	-.0791607	-.115594	.0364333	.0193334
_cons	.7212525	.7334874	-.0122349	.0064925

b = consistent under Ho and Ha; obtained from ivregress
B = inconsistent under Ha, efficient under Ho; obtained from regress

Test: Ho: difference in coefficients not systematic

chi2(1) = (b-B)'[(V_b-V_B)^(-1)](b-B)
= 3.55
Prob>chi2 = 0.0595

Model 1 $p=0.0595>0.05$, cannot reject H_0 at 5%, so seems like no endogenous problem.

But it is on the edge of level. And we know we need at least one outside exogenous variable,so it may be not trusted. **We need to use more precise model, model2 to test.**

```
. hausman IV2 model2, constant sigmamore //hausman 检验是否内生性，原假设无内生  
> , 拒绝则有内生
```

Note: the rank of the differenced variance matrix (1) does not equal the number of coefficients being tested (3); be sure this is what you expect, or there may be problems computing the test. Examine the output of your estimators for anything unexpected and possibly consider scaling your variables so that the coefficients are on a similar scale.

	Coefficients		(b-B) Difference	sqrt(diag(V_b-V_B)) S.E.
	(b) IV2	(B) model2		
kid3	-.0778864	-.0965608	.0186744	.0190936
yearschr	.0307483	.0303201	.0004282	.0004378
_cons	.3875135	.3984259	-.0109125	.0111575

b = consistent under Ho and Ha; obtained from ivregress
B = inconsistent under Ha, efficient under Ho; obtained from regress

Test: Ho: difference in coefficients not systematic

chi2(1) = (b-B)'[(V_b-V_B)^(-1)](b-B)
= 0.96
Prob>chi2 = 0.3281

But model2 with added edum variable, hausman test $p > 0.05$, that prove that we cannot reject kid3 is exogenous(无法拒绝外生原假设). We use model 2 and believe model2 better because this model is more fitted with higher R-squared. [Kid3 is exogenous]

(1)	
Hausman_model2	
VARIABLES	p
kid3	-0.0779*** (0.0191)
yearschr	0.0307*** (0.000497)
Constant	0.388*** (0.0115)
Observations	614,389
R-squared	0.040

Standard errors in parentheses

*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$

VARIABLES	(1)	(2)
	model1	model2
	p	p
kid3	-0.116*** (0.00124)	-0.0966*** (0.00123)
yearschr		0.0303*** (0.000235)
Constant	0.733*** (0.000716)	0.398*** (0.00269)

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```
reg kid3 yearschm sameex
predict v, resid
reg p kid3 yearschm v, r
outreg2 using "/Users/eric/Desktop/学业/1研一上/5103/Assignments/3/Yang Huan A0224968N
Assignment3/Output/edogenoustest_model2.doc" , word excel replace
//using method learnt from class to test model2, the result is same
```

Linear regression	Number of obs	=	614,389
	F(3, 614385)	=	8445.31
	Prob > F	=	0.0000
	R-squared	=	0.0401
	Root MSE	=	.45121

p	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
kid3	-.0778864	.0191328	-4.07	0.000	-.1153861	-.0403867
yearschm	.0307483	.0004982	61.72	0.000	.0297718	.0317248
v	-.0187515	.019173	-0.98	0.328	-.05633	.0188269
_cons	.3875135	.0114994	33.70	0.000	.3649751	.4100519

using method learnt from class to test model2, the result is same, i.e. cannot reject kid3 is exogenous. [kid3 is exogenous]

Q2

(a)

$$\begin{aligned} & \Pr(\text{accurate} = 1 \mid \text{high_temp} = 1, \text{duration} = 180.9) \\ &= \frac{e^{-0.396 - 0.386 - 0.001 \times 180.9}}{1 + e^{-0.396 - 0.386 - 0.001 \times 180.9}} = 0.2762979424 \approx 0.277 \\ & \quad \text{(excel)} \end{aligned}$$

$$\begin{aligned} & \Pr(\text{accurate} = 1 \mid \text{high_temp} = 0, \text{duration} = 180.9) \\ &= \frac{e^{-0.396 - 0.001 \times 180.9}}{1 + e^{-0.396 - 0.001 \times 180.9}} = 0.359646 \approx 0.360 \\ & \therefore 0.277 - 0.360 = -0.083 \end{aligned}$$

Hence, we can see when temperature is higher, probability of the accuracy of umpire reduces by 0.083

(b)

from part (a) the Partial Effect is -0.083 ,
in this linear model, the effect of high temp is -0.0843 ,
 \therefore the difference is very small, not surprisingly.

(c)

$$\begin{aligned} & \Pr(\text{accurate}=1 \mid \text{high-temp}=1, \text{duration}=180.9, \text{year}=2017) \\ &= \frac{e^{-1.003 - 0.425 - 0.003 \times 180.9 + 1.181}}{1 + e^{-1.003 - 0.425 - 0.003 \times 180.9 + 1.181}} = 0.312233 \approx 0.312 \end{aligned}$$

$$\Pr(\text{accurate}=1 \mid \text{high-temp}=1, \text{duration}=180.9, \text{year}=2008)$$

$$= \frac{e^{-1.003 - 0.425 - 0.003 \times 180.9}}{1 + e^{-1.003 - 0.425 - 0.003 \times 180.9}} = 0.122313 \approx 0.122$$

$$\text{Difference} \Rightarrow \therefore 0.312 - 0.122 = 0.190$$

So the probability of the accuracy for the umpire.

is 0.190 higher in 2017 than in 2008