

Advanced Probabilistic Machine Learning SSY316

Basics of Probability Theory

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Notation

- Vectors: Bold lowercase, e.g., \mathbf{x}
- Matrices: Bold uppercase, e.g., \mathbf{X}
- Random variables, vectors, and matrices: Sansserif font, e.g. x , \mathbf{x} , and \mathbf{X}
- Sets: Caligraphic letters, e.g., \mathcal{X}

Discrete random variables

- Probability mass function: $p_x(x) = \Pr(x = x) = p(x)$, with

$$0 \leq p(x) \leq 1 \quad \text{and} \quad \sum_{x \in \mathcal{X}} p(x) = 1$$

- Joint distribution:

$$p_{x,y}(x, y) = p(x, y)$$

- Conditional distribution

$$p_{x|y}(x|y) = p(x|y)$$

Discrete random variables

- Marginal probability:

$$p(x) = \sum_y p(x, y)$$

In general:

$$p(x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n) = \sum_{x_i} p(x_1, \dots, x_n)$$

- Bayes' theorem:

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)}$$

Continuous random variables

Probability density function: Describes the probability of the value of a continuous random variable x falling within a given interval.

The probability that x falls in an interval $[a, b]$ is

$$p(a \leq x \leq b) = \int_a^b p(x)dx$$

We have

$$p(x) \geq 0 \quad \text{and} \quad \int_{-\infty}^{\infty} p(x)dx = 1$$

- **Marginalization** of $p(x, y)$ with respect to y :

$$p(x) = \int_y p(x, y)dy$$

Expectation, variance, and covariance

- **Expectation** (average value of $f(x)$ under probability distribution $p(x)$):

$$\mathbb{E}_x[f(x)] = \mathbb{E}[f(x)] = \sum_x p(x)f(x) \quad (\text{discrete})$$

$$\mathbb{E}[f(x)] = \int p(x)f(x)dx \quad (\text{continuous})$$

- **Sample mean**: Given N points drawn from $p(x)$,

$$\mathbb{E}[f(x)] \simeq \frac{1}{N} \sum_{i=1}^N f(x_i)$$

with

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N f(x_i) = \mathbb{E}[f(x)]$$

Expectation, variance, and covariance

- For expectations of functions of several variables, we will keep the subscript to indicate the variable averaged over, e.g.,

$$\mathbb{E}_x[f(x, y)] \quad \text{or} \quad \mathbb{E}_{x \sim p(x)}[f(x, y)]$$

- Conditional expectation:

$$\mathbb{E}_{x \sim p(x|y)}[f(x)|y] = \mathbb{E}_{x|y}[f(x)|y] = \sum_x p(x|y)f(x)$$

- Variance (how much variability there is in $f(x)$ around its expected value):

$$\text{Var}[f(x)] = \mathbb{E} \left[(f(x) - \mathbb{E}[f(x)])^2 \right]$$

- Variance of x :

$$\text{Var}[x] = \mathbb{E} [x^2] - \mathbb{E}[x]^2$$

Expectation, variance, and covariance

- **Covariance** of x and y (the extent to which x and y vary together):

$$\begin{aligned}\text{Cov}[x, y] &= \mathbb{E}_{x,y} [(x - \mathbb{E}[x])(y - \mathbb{E}[y])] \\ &= \mathbb{E}_{x,y}[xy] - \mathbb{E}[x]\mathbb{E}[y]\end{aligned}$$

- **Covariance** of **two random vectors**:

$$\text{Cov}[\mathbf{x}, \mathbf{y}] = \mathbb{E}_{\mathbf{x}, \mathbf{y}} [(\mathbf{x} - \mathbb{E}[\mathbf{x}]) (\mathbf{y}^T - \mathbb{E}[\mathbf{y}^T])]$$

Probabilities: Frequentist vs Bayesian view

- **Frequentist interpretation:** relative frequency of occurrence of an outcome after repeating an experiment a large number of times

$$p = \lim_{n \rightarrow \infty} \frac{k}{n}$$

- **Bayesian interpretation:** quantifies the uncertainty of events happening

Probabilistic reasoning: Example

- 90% of people with Kreuzfeld-Jacob (KJ) disease ate hamburgers
 - The probability of an individual to have KJ is one in 100000
1. Assuming half of the population eat hamburgers, what is the probability that a hamburger eater will have KJ disease?

KJ \equiv Having Kreuzfeld-Jacob disease

H \equiv Eating Hamburger

$$p(\text{KJ} = \text{yes} | \text{H} = \text{yes}) ?$$

Probabilistic reasoning

Interpretation (frequentist approach)

Consider a population of 1M people:

- $p(KJ = \text{yes}) \Rightarrow 1000000 \cdot (1/10000) = 10$
- $p(H = \text{yes} | KJ = \text{yes}) \Rightarrow 10 \cdot 0.9 = 9$
- $p(H = \text{yes}) = 0.5 \Rightarrow 500000$

	H = yes	H = no
KJ = yes	9	1
KJ = no	499991	499 999

Now:

- $p(KJ = \text{yes} | H = \text{yes}) \equiv$ proportion of hamburger eaters having KJ:

$$p(KJ = \text{yes} | H = \text{yes}) = \frac{9}{9 + 499991} = 1.8 \cdot 10^{-5}$$

But this can be written as

$$\begin{aligned} p(KJ = \text{yes} | H = \text{yes}) &= \frac{p(KJ = \text{yes}, H = \text{yes})}{p(KJ = \text{yes}, H = \text{yes}) + p(KJ = \text{no}, H = \text{yes})} \\ &= \frac{p(H = \text{yes} | KJ = \text{yes})p(KJ = \text{yes})}{P(H = \text{yes})} \end{aligned}$$

Probabilistic reasoning

Interpretation (Bayesian approach)

x : our hypothesis (e.g. patient has a disease or not)

y : data (e.g., test results or patient symptoms)

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)}$$

- $p(x)$: prior belief in the hypothesis before looking at any data
- $p(y|x)$: likelihood of the data if the hypothesis were true
- $p(y)$: marginal likelihood (commonness of the data)
- $p(x|y)$: posterior belief on a hypothesis given the data

Bayesian (probabilistic) modeling

Two types of variables:

- $\mathcal{D} = \{x_1, \dots, x_N\}$: Observed variables (the data)
- $\theta = \{\theta_1, \dots, \theta_M\}$: Latent variables (we want to learn)

Probabilistic modeling: Treat both observed and latent variables as random variables.

Can model relationship between \mathcal{D} and θ via $p(\mathcal{D}, \theta)$

Usually we will be interested in $p(\theta|\mathcal{D})$.

Bayesian (probabilistic) inference

Many inference problems are of the form:

$$p(\theta|\mathcal{D}) = \frac{p(\mathcal{D}|\theta)p(\theta)}{p(\mathcal{D})}$$

- \mathcal{D} : Observed data
- θ : parameters of some model explaining the data
- $p(\theta)$: **prior** belief of the parameters before collecting any data
- $p(\mathcal{D}|\theta)$: **likelihood** of the data in view of the parameters
- $p(\mathcal{D})$: **marginal likelihood**
- $p(\theta|\mathcal{D})$: **posterior** belief of the parameters after observing the data

Bayesian (probabilistic) inference

$$p(\theta|\mathcal{D}) = \frac{p(\mathcal{D}|\theta)p(\theta)}{p(\mathcal{D})} = \frac{p(\mathcal{D}|\theta)p(\theta)}{\int_{\theta} p(\mathcal{D}|\theta)p(\theta)d\theta}$$

- Seeing quantities as functions of θ , $p(\mathcal{D})$ can be viewed as a **normalization constant** and we can write

$$p(\theta|\mathcal{D}) \propto p(\mathcal{D}|\theta)p(\theta)$$

- Most probable a posteriori (maximum a posteriori (**MAP**)) setting:

$$\theta_{\text{MAP}}^* = \arg \max_{\theta} p(\theta|\mathcal{D})$$

- If $p(\theta)$ is **constant**, **MAP** is **equivalent** to **maximum likelihood**,

$$\theta_{\text{ML}}^* = \arg \max_{\theta} p(\mathcal{D}|\theta)$$

Example: Tossing a biased coin

- $x \in \{0, 1\}$: Outcome of a coin flip ($0 \equiv \text{tail}$, $1 \equiv \text{head}$)

$$p(x = 1) = \mu, \quad p(x = 0) = 1 - \mu$$

Goal: Given a data set $\mathcal{D} = \{x_1, \dots, x_N\}$, estimate μ , i.e., the probability that a toss coin will be a head, $p(\mu|\mathcal{D})$.

- **Solution:** Apply Bayes',

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})} \propto p(\mathcal{D}|\mu)p(\mu)$$

Example: Tossing a biased coin

- **Observation:** A single coin toss corresponds to a **Bernoulli RV**,

$$p(x|\mu) = \text{Bern}(x; \mu) = \mu^x (1 - \mu)^{1-x}$$

with

$$\mathbb{E}[x] = \mu, \quad \text{Var}[x] = \mu(1 - \mu)$$

- For N coin tosses,

$$\begin{aligned} p(\mathcal{D}|\mu) &= \prod_{i=1}^N p(x_i|\mu) = \prod_{i=1}^N \mu^{x_i} (1 - \mu)^{1-x_i} \\ &= \mu^{\sum_i x_i} (1 - \mu)^{N - \sum_i x_i} = \mu^h (1 - \mu)^{N-h} \end{aligned}$$

where $h = \sum_{i=1}^N x_i$ is the number of heads

Example: Tossing a biased coin

Frequentist approach:

Can estimate μ by maximizing $p(\mathcal{D}|\mu)$ or, equivalently $\ln p(\mathcal{D}|\mu)$,

$$\begin{aligned}\ln p(\mathcal{D}|\mu) &= \sum_{i=1}^N \ln p(x_i|\mu) = \sum_{i=1}^N \ln (\mu^{x_i} (1 - \mu)^{1-x_i}) \\ &= \sum_{i=1}^N x_i \ln \mu + (1 - x_i) \ln(1 - \mu)\end{aligned}$$

Differentiating and equating to zero we obtain the ML estimator

$$\mu_{\text{ML}} = \frac{1}{N} \sum_{i=1}^N x_i = \frac{h}{N}$$

h : number of heads within data set

Example: Tossing a biased coin

Bayesian approach:

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})} \propto p(\mathcal{D}|\mu)p(\mu)$$

with

$$p(\mathcal{D}|\mu) = \mu^h (1 - \mu)^{N-h}$$

- Specify a prior for $p(\mu)$!
- Assume $\mu \in \{0.1, 0.5, 0.8\}$ with

$$p(\mu = 0.1) = 0.15, \quad p(\mu = 0.5) = 0.8, \quad p(\mu = 0.8) = 0.05$$

$N = 10$ with 2 heads and 8 tails

$$p(\mu = 0.1|\mathcal{D}) = 0.4525 \quad p(\mu = 0.5|\mathcal{D}) = 0.5475 \quad p(\mu = 0.8|\mathcal{D}) = 0.00001$$

$N = 100$ with 20 heads and 80 tails

$$\begin{aligned} p(\mu = 0.1|\mathcal{D}) &= 0.99999807 & p(\mu = 0.5|\mathcal{D}) &= 1.93 \cdot 10^{-6} \\ p(\mu = 0.8|\mathcal{D}) &= 2.13 \cdot 10^{-35} \end{aligned}$$

Example: Tossing a biased coin

And if we consider a continuum of parameters?

A flat (uniform) prior $p(\mu) = k$:

- For continuous variables, we require

$$\int p(\mu) d\mu = 1 \implies \int_0^1 p(\mu) d\mu = k = 1$$

- Now:

$$p(\mu|\mathcal{D}) \propto p(\mathcal{D}|\mu)p(\mu) = \mu^h(1-\mu)^{N-h}$$

We want $p(\mu|\mathcal{D})$ to be a **distribution**,

$$p(\mu|\mathcal{D}) = \frac{1}{c} p(\mathcal{D}|\mu)p(\mu) = \frac{1}{c} \mu^h(1-\mu)^{N-h}$$

where constant c is obtained as

$$c = \int_0^1 \mu^h(1-\mu)^{N-h} d\mu \equiv \text{B}(h+1, N-h+1)$$

The Beta distribution

- Beta function:

$$B(a, b) = \int_0^1 \mu^{a-1} (1 - \mu)^{b-1} d\mu$$

- Beta distribution:

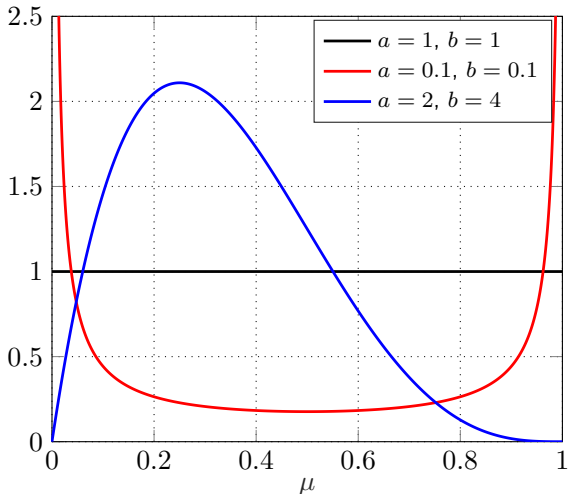
$$\begin{aligned} \text{Beta}(\mu; a, b) &= \frac{1}{B(a, b)} \mu^{a-1} (1 - \mu)^{b-1} \\ &= \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1 - \mu)^{b-1} \end{aligned}$$

$\Gamma(\cdot)$: Gamma function

$B(a, b)$ A normalization constant to ensure

$$\int_0^1 \text{Beta}(\mu; a, b) d\mu = 1$$

The Beta distribution



- a and b control the distribution of μ (hyperparameters)

Example: Tossing a biased coin

Observation: If prior **proportional** to **powers** of μ and $1 - \mu$, then posterior distribution will have the **same functional** form as the prior.

A **conjugate** prior (posterior will be of same functional form as prior):

- Choose **Beta distribution** for the prior:

$$p(\mu) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1}$$

Then

$$\begin{aligned} p(\mu|\mathcal{D}) &\propto p(\mathcal{D}|\mu)p(\mu) \\ &\propto \mu^h (1-\mu)^{N-h} \mu^{a-1} (1-\mu)^{b-1} \\ &= \mu^{h+a-1} (1-\mu)^{N-h+b-1} \end{aligned}$$

The posterior is also a **Beta distribution**!

$$p(\mu|\mathcal{D}) = \text{Beta}(\mu; a', b')$$

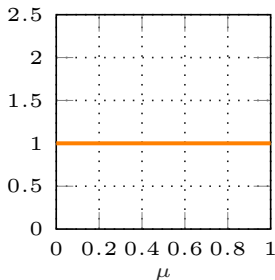
with $a' = a + h$ and $b' = b + N - h$

Example: Tossing a biased coin

prior

$$p(\mu) = \text{Beta}(\mu; a, b)$$

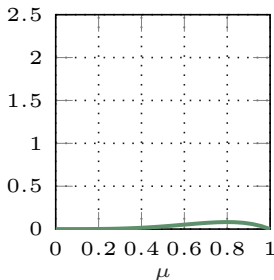
$$a = 1, b = 1$$



likelihood

$$p(\mathcal{D}|\mu) = \mu^h(1-\mu)^{N-h}$$

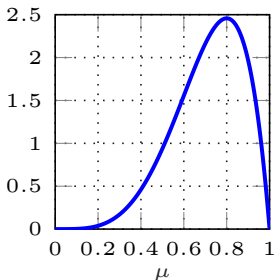
$$N = 5, h = 4$$



posterior

$$p(\mu|\mathcal{D}) = \text{Beta}(\mu; a', b')$$

$$a' = 5, b' = 2$$



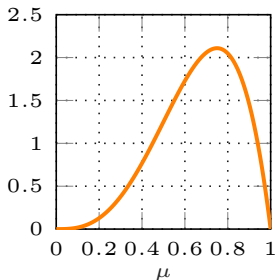
- If we do not know anything about the coin: **uniform prior**
- $N = 1, \mathcal{D} = (1)$ $N = 5, \mathcal{D} = (1, 1, 1, 0, 1)$

Example: Tossing a biased coin

prior

$$p(\mu) = \text{Beta}(\mu; a, b)$$

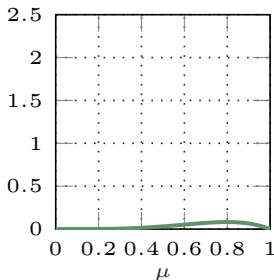
$$a = 4, b = 2$$



likelihood

$$p(\mathcal{D}|\mu) = \mu^h (1 - \mu)^{N-h}$$

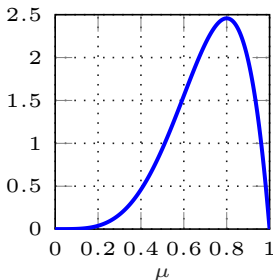
$$N = 5, h = 4$$



posterior

$$p(\mu|\mathcal{D}) = \text{Beta}(\mu; a', b')$$

$$a' = 5, b' = 2$$



- $N = 5, \mathcal{D} = (1, 1, 1, 0, 1)$
- **Sequential inference:** Posterior can act as prior if we subsequently observe additional data!
- $\mathcal{D} = (1, 1, 1, 0, 1)$

Bayesian inference and machine learning

Probabilistic (Bayesian) machine learning:

- Treats model and its parameters as random variables
- Learning does not provide a single model, but a distribution of likely models
- Can incorporate prior knowledge on model and parameters