

Relevant statistics for the method of Least Squares

We assume that the observed values y_i can be written on the following form

$$y_i = \hat{y}_i + \epsilon_i$$

where \hat{y}_i is the prediction and ϵ_i is the error. We assume that the ϵ_i 's are random variables that are independently and identically distributed from a normal distribution $N(\mu = 0, \sigma^2)$. This makes the observed values y_i independent random variables as well. The mean of y_i is

$$\langle y_i \rangle = \hat{y}_i$$

and the variance of y_i is

$$\text{Var}(y_i) = \sigma^2$$

See Exercise 3.pdf. Since the parameters $\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ depend on \mathbf{y} whose components are random variables, the components of $\hat{\beta}$ are random variables also. We denote the expectation value of $\hat{\beta}$ as β . The covariance matrix of the components of $\hat{\beta}$ is

$$\begin{aligned} \text{Var}(\hat{\beta}) &= \mathbb{E}[(\hat{\beta} - \mathbb{E}[\hat{\beta}])(\hat{\beta} - \mathbb{E}[\hat{\beta}])^T] \\ &= \mathbb{E}[(\hat{\beta} - \beta)(\hat{\beta} - \beta)^T] = \mathbb{E}[(\hat{\beta} - \beta)(\hat{\beta}^T - \beta^T)] \\ &= \mathbb{E}[\hat{\beta}\hat{\beta}^T - \hat{\beta}\beta^T - \beta\hat{\beta}^T + \beta\beta^T] = \mathbb{E}[\hat{\beta}\hat{\beta}^T] - \mathbb{E}[\hat{\beta}]\beta^T - \beta\mathbb{E}[\hat{\beta}^T] + \beta\beta^T \\ &= \mathbb{E}[\hat{\beta}\hat{\beta}^T] - \beta\beta^T - \beta\beta^T + \beta\beta^T = \mathbb{E}[\hat{\beta}\hat{\beta}^T] - \beta\beta^T \\ &= \mathbb{E}\left[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \left\{ (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \right\}^T\right] - \beta\beta^T \\ &= \mathbb{E}\left[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \mathbf{y}^T \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1}\right] - \beta\beta^T \end{aligned}$$

In the last line it has been used that $(\mathbf{ABC})^T = \mathbf{C}^T \mathbf{B}^T \mathbf{A}^T$ and that

$\left\{ (\mathbf{X}^T \mathbf{X})^{-1} \right\}^T = \left\{ (\mathbf{X}^T \mathbf{X})^T \right\}^{-1} = (\mathbf{X}^T \mathbf{X})^{-1}$ (the transpose of the inverse of a matrix is equal to the inverse of the transpose of a matrix). Now we have

$$\text{Var}(\hat{\boldsymbol{\beta}}) = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbb{E}[\mathbf{y} \mathbf{y}^T] \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} - \boldsymbol{\beta} \boldsymbol{\beta}^T$$

Where the expectation value $\mathbb{E}[\mathbf{y} \mathbf{y}^T]$ is

$$\begin{aligned} \mathbb{E}[\mathbf{y} \mathbf{y}^T] &= \mathbb{E}[(\mathbf{X} \boldsymbol{\beta} + \boldsymbol{\epsilon})(\mathbf{X} \boldsymbol{\beta} + \boldsymbol{\epsilon})^T] \\ &= \mathbb{E}[(\mathbf{X} \boldsymbol{\beta} + \boldsymbol{\epsilon})(\boldsymbol{\beta}^T \mathbf{X}^T + \boldsymbol{\epsilon}^T)] = \mathbb{E}[\mathbf{X} \boldsymbol{\beta} \boldsymbol{\beta}^T \mathbf{X}^T + \mathbf{X} \boldsymbol{\beta} \boldsymbol{\epsilon}^T + \boldsymbol{\epsilon} \boldsymbol{\beta}^T \mathbf{X}^T + \boldsymbol{\epsilon} \boldsymbol{\epsilon}^T] \\ &= \mathbf{X} \boldsymbol{\beta} \boldsymbol{\beta}^T \mathbf{X}^T + \mathbf{X} \boldsymbol{\beta} \mathbb{E}[\boldsymbol{\epsilon}^T] + \mathbb{E}[\boldsymbol{\epsilon}] \boldsymbol{\beta}^T \mathbf{X}^T + \mathbb{E}[\boldsymbol{\epsilon} \boldsymbol{\epsilon}^T] = \mathbf{X} \boldsymbol{\beta} \boldsymbol{\beta}^T \mathbf{X}^T + \mathbb{E}[\boldsymbol{\epsilon} \boldsymbol{\epsilon}^T] \end{aligned}$$

The last term is the covariance matrix of $\boldsymbol{\epsilon}$ since

$$\mathbb{E}[\boldsymbol{\epsilon} \boldsymbol{\epsilon}^T] = \mathbb{E}[(\boldsymbol{\epsilon} - \mathbf{0})(\boldsymbol{\epsilon} - \mathbf{0})^T] = \mathbb{E}[(\boldsymbol{\epsilon} - \mathbb{E}[\boldsymbol{\epsilon}])(\boldsymbol{\epsilon} - \mathbb{E}[\boldsymbol{\epsilon}])^T] = \text{Var}(\boldsymbol{\epsilon})$$

The diagonal elements of $\text{Var}(\boldsymbol{\epsilon})$ are

$$\text{Var}(\boldsymbol{\epsilon})_{ii} = \mathbb{E}[(\epsilon_i - \mathbb{E}[\epsilon_i])^2] = \text{Var}(\epsilon_i) = \sigma^2$$

The non-diagonal elements of $\text{Var}(\boldsymbol{\epsilon})$ are

$$\begin{aligned} \text{Var}(\boldsymbol{\epsilon})_{ij} &= \mathbb{E}[(\epsilon_i - \mathbb{E}[\epsilon_i])(\epsilon_j - \mathbb{E}[\epsilon_j])] \\ &= \mathbb{E}[(\epsilon_i - 0)(\epsilon_j - 0)] = \mathbb{E}[\epsilon_i \epsilon_j] \\ &= \mathbb{E}[\epsilon_i] \mathbb{E}[\epsilon_j] = 0 \end{aligned}$$

where the last line follows from the assumption that each ϵ_i are independent (but identically distributed) variables. Therefore $\text{Var}(\boldsymbol{\epsilon}) = \sigma^2 \mathbf{I}$ and

$$\mathbb{E}[\mathbf{y} \mathbf{y}^T] = \mathbf{X} \boldsymbol{\beta} \boldsymbol{\beta}^T \mathbf{X}^T + \sigma^2 \mathbf{I}$$

$$\text{Var}(\hat{\boldsymbol{\beta}}) = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{X} \boldsymbol{\beta} \boldsymbol{\beta}^T \mathbf{X}^T + \sigma^2 \mathbf{I}) \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} - \boldsymbol{\beta} \boldsymbol{\beta}^T$$

$$\begin{aligned}
&= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X} \beta \beta^T \mathbf{X}^T \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} + \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} - \beta \beta^T \\
&= \beta \beta^T + \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1} - \beta \beta^T = \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}
\end{aligned}$$

In particular the variance of the components of $\hat{\boldsymbol{\beta}}$ are the diagonal elements of the covariance matrix

$$\text{Var}(\hat{\beta}_i) = \sigma^2 \left[(\mathbf{X}^T \mathbf{X})^{-1} \right]_{ii}$$