

Innlevering 3

Oppgave 1:

$$\lambda_A = \frac{3}{73} = 0.0411$$

$$\lambda_B = \frac{1}{73} = 0.0137$$

$X_A \rightarrow$ antall feil av typen A i 365 dager

$X_B \rightarrow$ antall feil av typen B i 365 dager

a) $E[X_A] = 365 \cdot 0.0411 \approx 15$

$$P(X_A \leq 15) = 0.5681 \quad (\text{Fra tabell})$$

$$P(X_A \leq 15, X_B \leq 10) = P(X_A \leq 15) \cdot P(X_B \leq 10) = 0.5681 \cdot 0.9863 = 0.5603$$

$\lambda_B \cdot 365 = 5$
 $X = 10$
 $\mu = 5$

A og B er uavhengige

$$P(X_A \leq 15 | X_B \leq 10) = P(X_A \leq 15) = 0.5681$$

b) $Z = X_A + X_B$

$$M_A(t) = e^{\lambda_A(e^t - 1)}$$

$$M_B(t) = e^{\lambda_B(e^t - 1)}$$

Momentgenererende funksjoner for
summen av uavhengige stokastiske
variabler er et produkt av deres
momentgenererende funksjoner

$$\begin{aligned} M_Z(t) &= M_A(t) \cdot M_B(t) \\ &= e^{\lambda_A(e^t - 1)} \cdot e^{\lambda_B(e^t - 1)} \\ &= e^{\lambda_A(e^t - 1) + \lambda_B(e^t - 1)} \\ &= e^{(\lambda_A + \lambda_B)(e^t - 1)} \\ &= e^{\lambda_Z(e^t - 1)} \end{aligned}$$

$$\text{da } \lambda_Z = \lambda_A + \lambda_B //$$

c) $f_{X_A|Z}(x_A|z) = P(X_A = x_A | Z = x)$

Siden vi ønsker et gitt antall feil av typen A
gitt en mengde feil z følger denne
en betinget binomisk fordeling.

Oppgave 2:

X er fødselsvekt [g] i Norge

$$\mu = E[X] = 3500$$

$$\sigma = SD[X] = 570$$

$$\begin{aligned} a) \quad P(X \geq 3000) &= 1 - P(X \leq 3000) \\ &= 1 - P\left(\frac{X - \mu}{\sigma} \leq \frac{3000 - 3500}{570}\right) \\ &= 1 - P\left(\frac{X - \mu}{\sigma} \leq -0.88\right) \\ &= 1 - 0.1894 = 0.8106 \end{aligned}$$

$$\begin{aligned} P(3200 \leq X < 4000) &= P(X < 4000) - P(X \leq 3200) \\ &= P\left(\frac{X - \mu}{\sigma} < 0.88\right) - P\left(\frac{X - \mu}{\sigma} \leq -0.53\right) \\ &= 0.8106 - 0.2981 = 0.5125 \end{aligned}$$

$$\begin{aligned} P(X \geq 3500 | X \geq 3000) &= \frac{P(X \geq 3500 \cap X \geq 3000)}{P(X \geq 3000)} = \frac{1 - P(X \leq 3500)}{P(X \geq 3000)} \\ &= \frac{0.5}{0.8106} = 0.6168 \end{aligned}$$

$$\begin{aligned} P(X \leq 4000 | X \geq 3200) &= \frac{P(X \leq 4000 \cap X \geq 3200)}{P(X \geq 3200)} \\ &= \frac{P(X \leq 4000) - P(X \leq 3200)}{1 - P(X \leq 3200)} \\ &= \frac{P\left(\frac{X - \mu}{\sigma} \leq 0.88\right) - P\left(\frac{X - \mu}{\sigma} \leq -0.53\right)}{1 - P\left(\frac{X - \mu}{\sigma} \leq -0.53\right)} \\ &= \frac{0.8106 - 0.2981}{1 - 0.2981} = 0.7302 \end{aligned}$$

6)

$$P(X < c) = 0.01$$

$$n = 100$$

Y er antall undersøkte ved fødsel

$$P(Z > z_\alpha) = \alpha$$

$$P(Z > 2.326) = 0.01$$

↓

$$2.326 \cdot 570 = 1325.82$$

$$C = E[X] - 1325.8 = 3500 - 1325.8 = 2174.2 //$$

For at Y skal være binomial fordelt

må barnes vekter være uavhengige av hverandre.

$$\begin{aligned} P(Y \geq 1) &= 1 - P(Y < 1) \\ &= 1 - (1 - P(X < c))^N \\ &= 1 - (1 - 0.01)^{100} = 1 - 0.99^{100} \\ &= 1 - 0.366 \\ &= 0.634 // \end{aligned}$$

$$P(Y \geq 2 | Y \geq 1) = \frac{P(Y \geq 2 | Y \geq 1)}{P(Y \geq 1)} = \frac{P(Y \geq 2)}{P(Y \geq 1)} = \frac{1 - P(Y=0) - P(Y=1)}{P(Y \geq 1)}$$

$$= \frac{1 - 0.344 - 0.366}{0.634} = 0.264 //$$

$$P(Y=0) = (1-0.01)^{100} = 0.344$$

$$P(Y=1) = 100 \cdot 0.01 \cdot (1-0.01)^{99} = 0.99^{99} = 0.366 //$$

Oppgave 3

$$X_k = Z_k \quad \text{og for } k=2,3,\dots \quad X_k = Z_k - Z_{k-1}$$

$$a) \quad P(X_1 \geq 2) = 1 - F(2) = 1 - (1 - e^{-1.2}) = e^{-2.4} = e^{-2 \cdot 0.5} = e^{-1} = 0.3679 //$$

$$P(X_1 + X_2 \geq 4)$$

X_1 : tid til 1. hendelse

X_2 : tid mellom X_1 og X_2

$X_1 + X_2$: tid til 2 hendelser

$N(t)$: kille antall hendelser i $[0, t]$

$$N(t) \sim \text{Poi}(\lambda t)$$

$$= \frac{(\lambda t)^n}{n!} e^{-\lambda t}$$

$$P(X_1 + X_2 \geq 4)$$

$$= P(N(4) = 0 \cup N(4) = 1)$$

$$= P(N(4) = 0) + P(N(4) = 1)$$

$$= \frac{(4 \cdot 1)^0}{0!} e^{-4 \cdot 1} + \frac{(4 \cdot 1)^1}{1!} e^{-4 \cdot 1} = e^{-4} (1 + 4 \cdot 1) = e^{-4 \cdot 0.5} (1 + 4 \cdot 0.5)$$

$$= e^{-2} (1 + 2) = 0.4060 //$$

$$P(X_1 + X_2 \geq 4 | X_1 \geq 2)$$

$$= 1 - P(X_1 + X_2 \leq 4 | X_1 \geq 2)$$

b) $\alpha = 1$
 $\beta = \frac{1}{\lambda}$ vi setter dette inn i sannsynlighetstetthetsfunksjonen (pdf) gamma fordelingen

$$f_Y(y) = \frac{1}{\Gamma(1)(\frac{1}{\lambda})^1} y^{1-1} e^{-(\frac{1}{\lambda}y)}$$

$$\Gamma(1) = 1! = 1 \quad \text{og} \quad y^{1-1} = 1$$

$$f_Y(y) = \frac{1}{\frac{1}{\lambda}} e^{-\frac{1}{\lambda}y} = \lambda e^{-\lambda y}$$

Dette er identisk med eksponentialfunksjonen med parameter λ //

c) Vi antar Z_{n-1} er gammafordelt med parameter $\alpha = n-1$ og $\beta = \frac{1}{\lambda}$

$Z_n = Z_{n-1} + X_n$, Z_{n-1} og X_n er uavhengige

pdf for Z_n kan uttrykkes slik

$$f_{Z_n}(z) = \int_0^z f_{Z_{n-1}}(y) f_{X_n}(z-y) dy$$

Vi setter inn for $f_{Z_{n-1}}(y)$ som er gammafordelt med $\alpha = n-1$ og $\beta = \frac{1}{\lambda}$
 og for $f_{X_n}(z-y)$ som er eksponentialfordelt

$$f_{Z_n}(z) = \int_0^z \underbrace{\frac{1}{\Gamma(n-1)(\frac{1}{\lambda})^{n-1}} y^{n-2} e^{-\lambda y}}_{f_{Z_{n-1}}(y)} \underbrace{\lambda e^{-\lambda(z-y)}}_{f_{X_n}(z-y)} dy$$

$$= \frac{1}{\Gamma(n-1)(\frac{1}{\lambda})^{n-1}} \lambda e^{-\lambda z} \int_0^z y^{n-2} dy$$

$$\int_0^z y^{n-2} dy = \frac{1}{n-1} y^{n-1} \Big|_0^z = \frac{z^{n-1}}{n-1} - \frac{0^{n-1}}{n-1} = \frac{z^{n-1}}{n-1}$$

$$= \frac{\lambda^n}{\Gamma(n-1)} e^{-\lambda z} \frac{z^{n-1}}{(n-1)} \quad (n-1) \Gamma(n-1) = \Gamma(n)$$

$$f_{Z_n}(z) = \frac{\lambda^n}{\Gamma(n)} e^{-\lambda z} z^{n-1}$$

gammafordelingen med $\alpha = n$ og $\beta = \frac{1}{\lambda}$ //

Oppgave 4

$$X \sim N(\mu, \sigma^2)$$

$$Y = aX + b$$

a) $E[Y] = aE[X] + b = a\mu + b$

$$\text{Var}[Y] = a^2 \text{Var}[X] = a^2 \sigma^2$$

```
import numpy as np
import matplotlib.pyplot as plt
from scipy.stats import norm

def sim_Y(n, mu, sigma, a, b):
    X = np.random.normal(mu, sigma, n)
    Y = a * X + b
    return Y

N = int(1e5)
mu = 1
sigma = 2
a = 2
b = 0.5

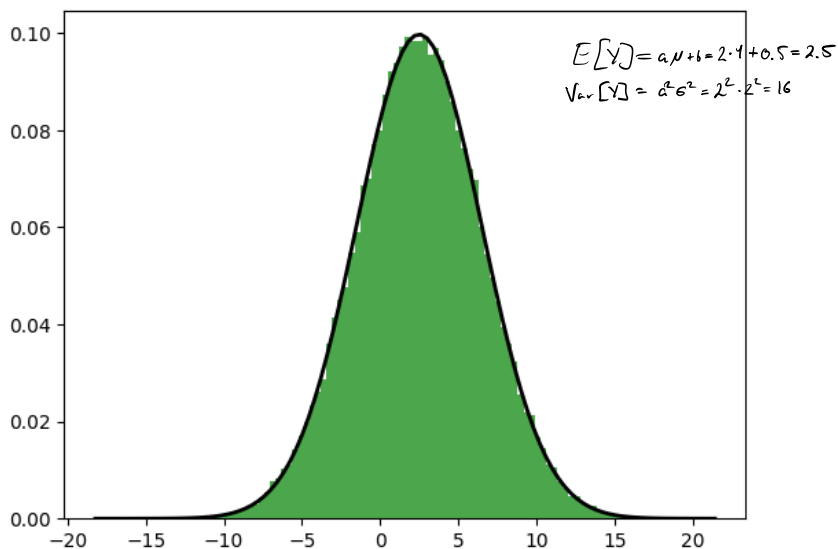
# Simulerer Y
Y = sim_Y(N, mu, sigma, a, b)

# Beregn forventningsverdi og varians for Y
E_Y = a * mu + b
Var_Y = (a * sigma)**2

# Histogram
plt.hist(Y, bins=100, density=True, alpha=0.7, color='g')

# Analytisk fordeling
xmin, xmax = plt.xlim()
x = np.linspace(xmin, xmax, 100)
p = norm.pdf(x, E_Y, np.sqrt(Var_Y))
plt.plot(x, p, 'k', linewidth=2)
```

✓ 2.3s



6)

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left\{-\frac{(x-\mu)^2}{2\cdot\sigma^2}\right\}$$

$$\begin{aligned} f_Y(y) &= \frac{1}{\sqrt{2\pi}\sigma_Y} \exp\left\{-\frac{(y-E[Y])^2}{2\cdot\sigma_Y^2}\right\} \\ &= \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left\{-\frac{(y-(\mu+\sigma))^2}{2\cdot\sigma^2}\right\} \end{aligned}$$

Oppgave 5:

$$M_X(t) = E[e^{tx}]$$

$$a) \quad M_{X_i}(t) = P(X_i=1)e^t + P(X_i=-1)e^{-t} = \frac{1}{2}e^t + \frac{1}{2}e^{-t}$$

$$M'_{X_i}(t) = \frac{1}{2}e^t - \frac{1}{2}e^{-t}$$

$$M'_{X_i}(0) = \frac{1}{2}e^0 - \frac{1}{2}e^0 = \frac{1}{2} - \frac{1}{2} = 0 \Rightarrow E[X_i] = 0$$

$$M''_{X_i}(t) = \frac{1}{2}e^t + \frac{1}{2}e^{-t}$$

$$M''_{X_i}(0) = \frac{1}{2} + \frac{1}{2} = 1 \Rightarrow E[X_i^2] = 1$$

$$Var[X] = E[X^2] - E[X]^2 = 1$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \Rightarrow E[\bar{X}] = E\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n} \sum_{i=1}^n E[X_i] = \frac{1}{n} \cdot n \cdot 0 = 0$$

$$Var[\bar{X}] = Var\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n^2} \sum_{i=1}^n Var[X_i] = \frac{1}{n^2} \cdot n \cdot 1 = \frac{1}{n}$$

$$b) \quad M_{\bar{X}}(t) = E[e^{t\bar{X}}] = E\left[e^{t\left(\frac{1}{n} \sum_{i=1}^n X_i\right)}\right] = E\left[\prod_{i=1}^n e^{t \cdot \frac{1}{n} X_i}\right] = \prod_{i=1}^n E[e^{t \cdot \frac{1}{n} X_i}]$$

$$V: \text{vel at } E[e^{t \cdot \frac{1}{n} X_i}] = \frac{1}{2}e^{t \cdot \frac{1}{n}} + \frac{1}{2}e^{-t \cdot \frac{1}{n}}$$

$$M_{\bar{X}}(t) = \left(M_{X_i}\left(\frac{t}{n}\right)\right)^n = \left(\frac{1}{2}e^{\frac{t}{n}} + \frac{1}{2}e^{-\frac{t}{n}}\right)^n$$

$$U = \frac{\bar{X} - E[\bar{X}]}{\sqrt{Var[\bar{X}]}} \xrightarrow[n \rightarrow \infty]{\text{med kl}} U = \frac{\bar{X} - 0}{\sqrt{\frac{1}{n}}} = \sqrt{n} \bar{X}$$

$$M_U(t) = E[e^{tU}] = E[e^{t\sqrt{n}\bar{X}}] = M_{\bar{X}}(t\sqrt{n})$$

$$M_U(t) = \left(\frac{1}{2}e^{\frac{t\sqrt{n}}{n}} + \frac{1}{2}e^{-\frac{t\sqrt{n}}{n}}\right)^n = \left(\frac{1}{2}e^{\frac{t}{\sqrt{n}}} + \frac{1}{2}e^{-\frac{t}{\sqrt{n}}}\right)^n$$

$$\ln M_U(t) = \ln\left(\frac{1}{2}e^{\frac{t}{\sqrt{n}}} + \frac{1}{2}e^{-\frac{t}{\sqrt{n}}}\right)^n = n \ln\left(\frac{1}{2}e^{\frac{t}{\sqrt{n}}} + \frac{1}{2}e^{-\frac{t}{\sqrt{n}}}\right)$$

c)

```

# Definerer funksjonen for M_U(t) basert på den utledede formelen
def M_U(t, n):
    return np.exp(n * np.log((np.exp(t/np.sqrt(n)) + np.exp(-t/np.sqrt(n))) / 2))

# Definerer funksjonen for M_Z(t) basert på den momentgenererende funksjonen for en standard normalfordeling
def M_Z(t):
    return np.exp(t**2 / 2)

# Definerer t-verdier i intervallet [0, 1]
t_values = np.linspace(0, 1, 100)

# Lager plott for ulike n-verdier
n_values = [1, 2, 5]
plt.figure(figsize=(12, 8))
for n in n_values:
    plt.plot(t_values, M_U(t_values, n), label=f'M_U(t) for n={n}')

# Plotter M_Z(t)
plt.plot(t_values, M_Z(t_values), label='M_Z(t)', linestyle='--')
plt.title('Sammenligning av M_U(t) for ulike n og M_Z(t)')
plt.xlabel('t')
plt.ylabel('Momentgenererende funksjon')

```

✓ 0.1s

