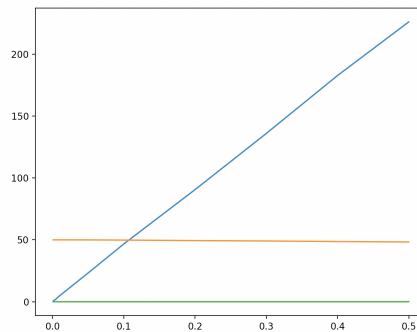
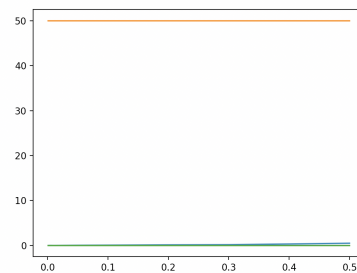


Problem 1

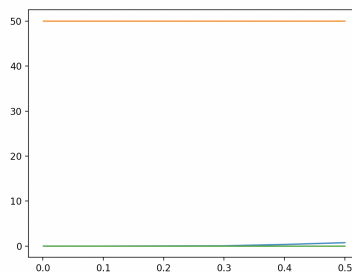
- Given $\frac{1}{N} \sum_{i=1}^N w_i = 1$ then $\sum_{i=1}^N w_i = N$ and therefore $w_i < N$.
Assuming $\alpha \in (0, 1)$, the quantity $\alpha \frac{w_i}{N}$ is less than 1.
Using fact that if P_i is null then, $P(P_i \leq u) = u$ for $u \in [0, 1]$.
So, $u = \frac{\alpha w_i}{N}$ and $P(P_i \leq u) = \frac{\alpha w_i}{N}$.
Since $\frac{w_i}{N} < 1$ then $\alpha \frac{w_i}{N} < \alpha$ and the procedure controls the probability of at least one false discovery under level α .
- $w_i = 1$ implies Bonferroni



- (a) Uncorrelated Testing

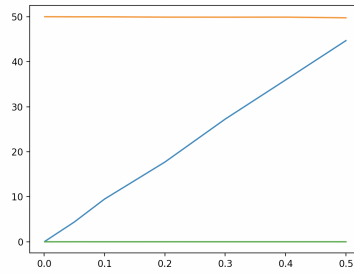


- (b) Bonferroni

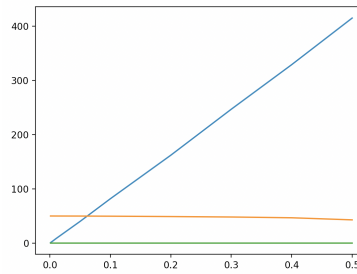


- (c) Benjamini-Hochberg

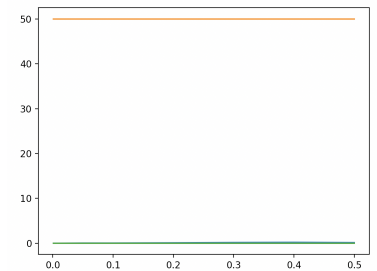
(d) Weighted with $w_i = \frac{2i}{501}$



(e) Weighted with $w_i = \frac{2(501-i)}{501}$



(f) Weighted with $w_i = 0.5$ for $1 \leq i \leq 450$ and $w_i = 5.5$ for $i > 450$

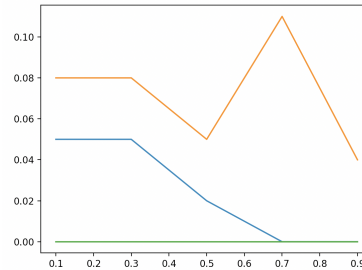
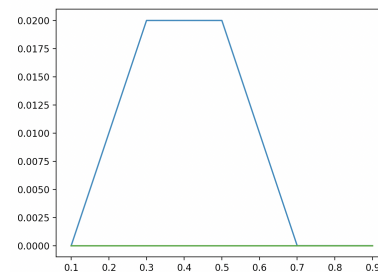


Problem 2

1.

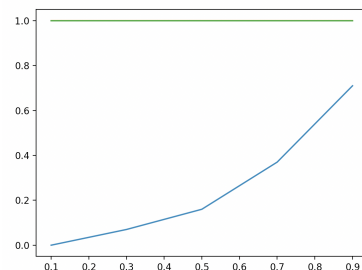
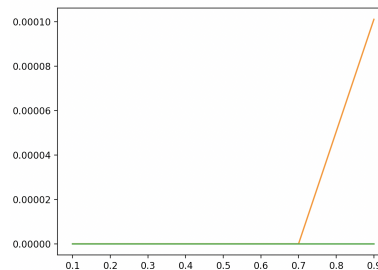
Problem 3

1. Using LORD procedure

(a) Average FDP over 100 trials at each π (b) Average Sensitivity over 100 trials at each π

(c) Highest average sensitivity came from scenario (i).

2. Using Benjamini-Hochberg:

(a) Average FDP over 100 trials at each π (b) Average Sensitivity over 100 trials at each π

(c) Sensitivity in BH for scenarios (ii) and (iii) are similar to the LORD procedure in that they are very close to 0.

Problem 4

1. Given that all P-values are independent, the case of removing a random P_i from the set to obtain $P^{(-i)}$, results in the events:

$\{P_i \leq \frac{\alpha}{N}r, R = r\}$ and $\{P_i \leq \frac{\alpha}{N}r, R^{(-i)} = r - 1\}$ being equal.

The case breaks down into two parts. Given these two cases, the argument should hold.

- (a) If the randomly removed P-value, P_i , is greater than or less than the maximum sorted P-value such that BH holds, then the set of rejected hypotheses will not change between P and $P^{(-i)}$.

- (b) If the randomly removed P-value, P_i , is equal to the would be maximum, R . In this case the set of rejected hypotheses also remains the same because from the definition of $R^{(-i)}$, the condition that

$$P_{(j)}^{(-i)} \leq \frac{\alpha}{N}(j+1)$$

accounts for the size difference of 1 (single missing P-value) by requiring the right side of the comparison to be multiplied by the index + 1 - i.e. $(j+1)$.

$$\begin{aligned} 2. \quad & \frac{1}{R} \sum_{i \in H_0} 1\{P_i \leq \frac{\alpha}{N}R, R > 0\} = \frac{1}{R} \sum_{i \in H_0} 1\{P_i \leq \frac{\alpha}{N}R\} 1\{R > 0\} \\ & = \frac{1}{R} \sum_{i \in H_0} 1\{P_i \leq \frac{\alpha}{N}R\} \sum_{r=1}^N 1\{R = r\} = \frac{1}{R} \sum_{i \in H_0} \sum_{r=1}^N 1\{P_i \leq \frac{\alpha}{N}R\} 1\{R = r\} \\ & = \sum_{i \in H_0} \sum_{r=1}^N \frac{1}{r} 1\{P_i \leq \frac{\alpha}{N}r, R = r\} \text{ and finally from part (a)} \\ & = \sum_{i \in H_0} \sum_{r=1}^N \frac{1}{r} 1\{P_i \leq \frac{\alpha}{N}r, R^{(-i)} = r - 1\} \end{aligned}$$

3. Using the fact that the expectation of an indicator is just its probability:

$$\text{FDR} = \mathbb{E}(\text{FDP})$$

$$\mathbb{E}(\text{FDP}) = \mathbb{E}\left(\sum_{i \in H_0} \sum_{r=1}^N \frac{1}{r} 1\{P_i \leq \frac{\alpha}{N}r, R^{(-i)} = r - 1\}\right)$$

$$= \sum_{i \in H_0} \sum_{r=1}^N \frac{1}{r} P(P_i \leq \frac{\alpha}{N}r, R^{(-i)} = r - 1) = \text{FDR}$$

4.

5.