$$M_0: \underline{\theta} = \{\alpha, \sigma^2\} \implies H_0: \beta = 0$$

$$M_1: \underline{\theta} = \{\alpha, \beta, \sigma^2\} \implies H_1: \beta \neq 0$$

$$B = \pi(M_0 \mid x) P(M_1)$$

$$B_{0} = \frac{\pi(M_0 \mid x)}{\pi(M_4 \mid x)} \frac{P(M_4)}{P(M_0)}$$

$$\frac{1}{\Pi(M_1 \mid x)} \frac{1}{P(M_0)}$$

THEORETICAL REMARKS

P(M<sub>3</sub>), P(
$$\Phi$$
| M<sub>3</sub>)

T( $\Phi$ , M<sub>3</sub>|x)  $\propto f(x|\Phi$ , M<sub>3</sub>) P( $\Phi$ | M<sub>3</sub>) P(M<sub>3</sub>)

NOTE P(M) IS A JOINT

PRISE P(M, 8) WHICH

CAN & FACTIRISED IN

$$T(M_3/x)$$
 is "sinfly" THE MARGINAL POSTERIOR DISTRIBUTION

 $T(M_3/x) \propto \int f(x|\theta, M_3) P(\theta|M_3) P(M_3) d\theta$ 

$$= P(M_3) \int f(x|\theta, M_3) P(\theta|M_3) d\theta = P(M_3) f(x|M_3)$$

Plus Lens

OFTEN ANALYTICALLY INTERCTABLE

BF, AND POSTERING MODELS
PROBABILITIES ARE
COMPUTATIONALLY CHALLENGING

SIMPLE MC AMPROACH

TARGET  $\hat{\pi}(H_1(x)) \propto \hat{f}(x|H_1) P(H_1)$ 

(x1m3) = E0 [f(x10, M3)]

=> Î(XIM) BY DRAWING K SAMPLES FROM THE PRIOR OF THE
L'ARAMETERS IN MODEL J AND THEN USE THEM IN THE MC
ESTIMATE

f(RIM) = 1 E f(x 1 0 M)

ONCE WE HAVE  $\hat{\pi}(\eta_j|x)$ , then  $\hat{\pi}(\eta_j|x)$  -0  $\pi(\eta_j|x)$  AS K-0  $\infty$ 

· REPEAT THIS PLACESS FOR ALL 2 MODELS AND RENDAMANGE THE ESTIMATES

$$\frac{1}{\pi}(M_3) \times = \frac{\hat{f}(x|M_3) p(M_3)}{\hat{f}(x|M_3) p(M_3)}$$

THE ESTIMATE I(XIM) ARE GENERALLY VERY UNSTABLE

- · IMPORTANCE SAMPLING (NEXT WEEK)
- · REVERSIBLE JUMP MCMC
- · BAYESIAN VARNABLE SELECTION

AS A RESULT OF THE PARAMETERS BEING DRAWN FROM THE PRIOR!

· MODEL COMPARISON CRITERIA (U.S. DIC & WAIC) - NEXT LECTURE!