

$$X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2) \text{ iid.}$$

unknown

Goal: Test with level α
 $H_0: \mu \geq 0$ $H_1: \mu < 0$

① Estimate μ, σ^2

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i, \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2$$

② Test statistic / pivot

$$\psi = \mathbb{1}\{T_n > s\}$$

③ Adjust s

Non-asymptotic hypothesis test for mean of Gaussians (t-test)

① MLE. Likelihood: $f(\mu, \sigma^2, X_1, \dots, X_n) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{1}{2\sigma^2} (X_i - \mu)^2}$

$$\begin{aligned} \ell(\mu, \sigma^2) &= \log f = \sum_{i=1}^n \left[-\frac{1}{2} \log 2\pi - \frac{1}{2} \log \sigma^2 - \frac{1}{2\sigma^2} (X_i - \mu)^2 \right] \\ &= -\frac{n}{2} \log 2\pi - \frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2 \end{aligned}$$

$$\begin{aligned} \partial_{\mu} \ell(\mu, \sigma^2) &= -\frac{1}{2\sigma^2} \sum_{i=1}^n 2 \cdot (X_i - \mu) \cdot (-1) \stackrel{!}{=} 0 \\ \Rightarrow \hat{\mu} &= \frac{1}{n} \sum_{i=1}^n X_i \end{aligned}$$

$$\begin{aligned} \partial_{\sigma^2} \ell(\mu, \sigma^2) &= -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (X_i - \mu)^2 \stackrel{!}{=} 0 \\ \Rightarrow \hat{\sigma}^2 &= \frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2 \end{aligned}$$

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③ Adjust s

② $\tilde{T}_n^{(1)} = \frac{\sqrt{n}(\mu - \hat{\mu})}{\sqrt{\hat{\sigma}^2}} \sim \mathcal{N}(0, 1)$. Problem: σ^2 unknown, replace by $\hat{\sigma}^2$!

What is distribution of $\hat{\sigma}^2$?

Review: Orthogonal matrices: $[v_1 | \dots | v_n] = V \in \mathbb{R}^{n \times n}$

$$v_i^T v_j = 0, \quad i \neq j, \quad v_i^T v_i = \|v_i\|_2^2 = 1 \Rightarrow \{v_i\}_{i=1, \dots, n} \text{ orthonormal}$$

$$V^T V = \begin{bmatrix} v_1^T \\ \vdots \\ v_n^T \end{bmatrix} \begin{bmatrix} v_1 & \dots & v_n \end{bmatrix} = \begin{bmatrix} 1 & & 0 \\ & \ddots & \\ 0 & & 1 \end{bmatrix} = I_n = V V^T$$

V orthogonal

$$x \in \mathbb{R}^n, \quad \|Vx\|_2^2 = (Vx)^T (Vx) = x^T \underbrace{V^T V}_{=I_n} x = x^T x = \|x\|_2^2$$

$$W = \begin{bmatrix} w_1 & \dots & w_k \end{bmatrix} \in \mathbb{R}^{n \times k}, \quad \{w_i\}_{i=1, \dots, k} \text{ orthonormal}$$

$$\|W W^T x\|_2^2 = (W W^T x)^T (W W^T x) = x^T \underbrace{W W^T W W^T}_{=I_k} x = \|W^T x\|_2^2$$

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③ Adjust s

$$\textcircled{2} \sigma^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$

$$n\hat{\sigma}^2 = \left\| \begin{pmatrix} X_1 \\ \vdots \\ X_n \end{pmatrix} - \begin{pmatrix} \hat{\mu} \\ \vdots \\ \hat{\mu} \end{pmatrix} \right\|_2^2$$

$$= \|X - v_1 v_1^T X\|_2^2$$

$$= \|X - \mu \cdot \mathbb{1} - (v_1 v_1^T X - \mu \mathbb{1})\|_2^2$$

$$= v_1 v_1^T (X - \mu \mathbb{1})$$

$$\Rightarrow \frac{n\hat{\sigma}^2}{\sigma^2} = \|Y - v_1 v_1^T Y\|_2^2$$

$$= \left\| \sum_{i=1}^n v_i v_i^T Y - v_1 v_1^T Y \right\|_2^2$$

$$= \left\| \sum_{i=2}^n v_i v_i^T Y \right\|_2^2$$

$$\hat{\mu} = \frac{1}{n} \frac{\mathbb{1}^T X}{\mathbb{1}^T \mathbb{1}}, \begin{pmatrix} \hat{\mu} \\ \vdots \\ \hat{\mu} \end{pmatrix} = \frac{1}{n} \frac{\mathbb{1} \mathbb{1}^T X}{\mathbb{1}^T \mathbb{1}}$$

$$= \frac{1}{n} \mathbb{1} \left(\frac{\mathbb{1}^T X}{\mathbb{1}^T \mathbb{1}} \right) = v_1 v_1^T X$$

$$Y = X - \mu \mathbb{1} = \begin{pmatrix} X_1 - \mu \\ \vdots \\ X_n - \mu \end{pmatrix} \sim \mathcal{N}(0, I_n)$$

$$V = [v_1 | \dots | v_n] \in \mathbb{R}^{n \times n} \text{ orthogonal}$$

$$I_n = V V^T = [v_1 | \dots | v_n] \begin{bmatrix} v_1^T \\ \vdots \\ v_n^T \end{bmatrix}$$

$$= \left(\sum_{i=1}^n v_i e_i^T \right) \left(\sum_{j=1}^n v_j e_j^T \right)^T = \sum_{i,j=1}^n v_i e_i^T e_j v_j^T$$

$$i \mapsto \begin{pmatrix} 0 \\ \vdots \\ 1 \\ \vdots \\ 0 \end{pmatrix} = \sum_{i=1}^n v_i v_i^T = \sum_{i,j=1}^n \delta_{ij} v_i v_j^T$$

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② Test statistic / pivot
 $\tilde{T}_n^{(1)} = \frac{(\mu - \hat{\mu}) \sqrt{n}}{\sqrt{\sigma^2}} \sim \mathcal{N}(0, 1)$
 $T_n = \frac{-(n-1) \cdot \hat{\mu}}{\sqrt{\hat{\sigma}^2}}$

$$\psi = \mathbb{1}_{\{T_n > s\}}$$

③ Adjust s

② $\frac{n\hat{\sigma}^2}{\sigma^2} = \left\| \sum_{i=2}^n v_i v_i^T Y \right\|_2^2$
 $= \|W W^T Y\|_2^2 = \left\| \underbrace{W^T Y}_{\in \mathbb{R}^{n-1}} \right\|_2^2$

$$\begin{aligned} \text{Cov}(W^T Y) &= W^T \cdot \text{Cov}(Y) \cdot (W^T)^T \\ &= W^T \cdot I_n \cdot W = I_{n-1} \\ \Rightarrow \frac{n\hat{\sigma}^2}{\sigma^2} &\sim \chi_{n-1}^2 \end{aligned}$$

$$\begin{aligned} \text{Cov}\left(\underbrace{W^T Y}_{\in \mathbb{R}^{n-1}}, \underbrace{v_1^T Y}_{\in \mathbb{R}}\right) &= E[W^T Y Y^T v_1] \\ &= W^T \underbrace{E[Y Y^T]}_{I_n} v_1 = W^T v_1 = \begin{pmatrix} 0 \\ \vdots \\ 0 \end{pmatrix} \in \mathbb{R}^{n-1} \\ \Rightarrow W^T Y &\perp v_1^T Y \Rightarrow \hat{\sigma}^2 \perp \hat{\mu} \end{aligned}$$

$$Y = \begin{pmatrix} \frac{x_1 - \mu}{\sigma} \\ \vdots \\ \frac{x_n - \mu}{\sigma} \end{pmatrix} \sim \mathcal{N}(0, I_n)$$

$$V = [v_1 | \dots | v_{n-1}] \text{ orthogonal}$$

$= W \in \mathbb{R}^{n \times (n-1)}$

$$\sum_{i=2}^n v_i v_i^T = W W^T$$

$$\hat{\mu} = \frac{1}{n} v_1^T \cdot X = \mu + \sigma \cdot \frac{1}{n} v_1^T \cdot Y$$

$$\begin{aligned} \tilde{T}_n^{(2)} &= \frac{(\mu - \hat{\mu}) \sqrt{n}}{\sqrt{\sigma^2}} \\ &= \frac{\sqrt{n} \cdot \sqrt{\frac{\partial^2}{\partial \mu^2}}}{\sqrt{\hat{\sigma}^2}} = \frac{(\mu - \hat{\mu}) \sqrt{n-1}}{\sqrt{\hat{\sigma}^2}} \sim t_{n-1} \end{aligned}$$

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② Test statistic/pivot

$$\frac{\hat{\mu} - \mu}{\sqrt{\hat{\sigma}^2/n}} \sim N(0, 1)$$

$$T_n = \frac{-\sqrt{n-1} \cdot \hat{\mu}}{\sqrt{\hat{\sigma}^2}}$$

$$\psi = \mathbb{1}\{T_n > s\}$$

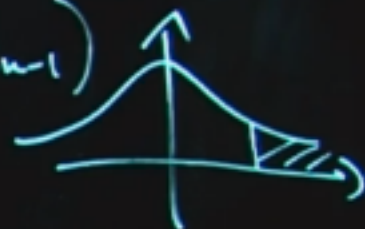
③ Adjust s

$$s = q_{1-\alpha}(t_{n-1})$$

③ $T_n^{(2)} = \frac{(-\hat{\mu} + \mu) \sqrt{n-1}}{\sqrt{\hat{\sigma}^2}} \sim t_{n-1} \quad \Bigg| \quad T_n = \frac{-\hat{\mu} \sqrt{n-1}}{\sqrt{\hat{\sigma}^2}}$

For $\mu = 0$: $\mathbb{P}_{(0, \sigma^2)}(\psi = 1) = \mathbb{P}_{(0, \sigma^2)}(T_n > s)$

$$= \mathbb{P}_{t_{n-1}}(Z > s) \stackrel{!}{=} \alpha \Rightarrow s = q_{1-\alpha}(t_{n-1})$$



For $\mu > 0$: $\mathbb{P}_{(\mu, \sigma^2)}(\psi = 1) = \mathbb{P}_{(\mu, \sigma^2)}(T_n > s)$

$$= \mathbb{P}_{(\mu, \sigma^2)}\left(\frac{-\hat{\mu} \sqrt{n-1}}{\sqrt{\hat{\sigma}^2}} > s\right) = \mathbb{P}_{(\mu, \sigma^2)}\left(\underbrace{\frac{(\mu - \hat{\mu}) \sqrt{n-1}}{\sqrt{\hat{\sigma}^2}}}_{= Z \sim t_{n-1}} > s + \underbrace{\frac{\mu \sqrt{n-1}}{\sqrt{\hat{\sigma}^2}}}_{> 0}\right)$$

$$\textcircled{\leq} \mathbb{P}_{(\mu, \sigma^2)}(Z > s) = \alpha$$

