# Metodi Numerici dell'Informatica

Singular Value Decomposition and Principal Component Analysis

Emanuele Rodolà rodola@di.uniroma1.it



# Motivation

#### Isometries

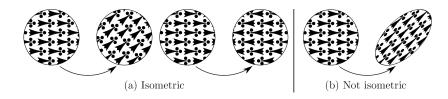
We have seen that orthogonal matrices preserve lengths:

$$\|\mathbf{Q}\mathbf{x}\|_2^2 = \mathbf{x}^{\top}\mathbf{Q}^{\top}\mathbf{Q}\mathbf{x} = \mathbf{x}^{\top}\mathbf{I}\mathbf{x} = \mathbf{x}^{\top}\mathbf{x} = \|\mathbf{x}\|_2^2$$

...and they also preserve angles (i.e. inner products):

$$\langle \mathbf{Q}\mathbf{x}, \mathbf{Q}\mathbf{y} \rangle = \mathbf{x}^{\top}\mathbf{Q}^{\top}\mathbf{Q}\mathbf{y} = \mathbf{x}^{\top}\mathbf{I}\mathbf{y} = \mathbf{x}^{\top}\mathbf{y}$$

By these properties, the map  $\mathbf{x}\mapsto\mathbf{Q}\mathbf{x}$  is an isometry of  $\mathbb{R}^n$ .



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Do we have a similar interpretation for arbitrary matrices  $\mathbf{A} \in \mathbb{R}^{m \times n}$ ?

It turns out that any matrix can be factorized as

$$\mathbf{A} = \mathbf{U} \sum_{m \times n} \mathbf{V}^{\top}$$

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- U and V are orthogonal matrices
- $\Sigma$  is a rectangular diagonal matrix, e.g.  $\left(egin{array}{ccc} \sigma_1 & 0 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & 0 & 0 \\ 0 & 0 & \sigma_3 & 0 & 0 \end{array}\right)$

Do we have a similar interpretation for arbitrary matrices  $\mathbf{A} \in \mathbb{R}^{m \times n}$ ?

$$\mathbf{A}_{m imes n} = \mathbf{U}_{m imes m} \mathbf{\Sigma}_{m imes n} \mathbf{V}^{ op}$$

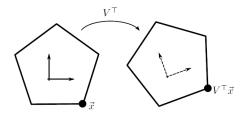
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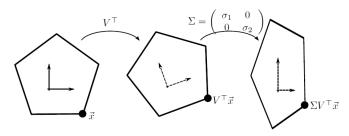
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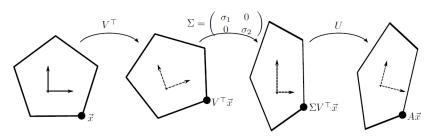
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# Singular Value Decomposition (SVD)

The factorization

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is called the singular value decomposition of matrix  ${\bf A}.$ 

#### SVD

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$$\mathbf{A} = \underbrace{\mathbf{U}}_{\substack{\text{left} \\ \text{singular} \\ \text{vectors}}} \underbrace{\mathbf{\Sigma}}_{\substack{\text{values} \\ \text{values} \\ \text{singular} \\ \text{vectors}}} \mathbf{V}^{\top}$$

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This can also be written as:

$$AV=\mathbf{U}\boldsymbol{\Sigma}$$

which looks quite similar to the eigenvalue equation.



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This can also be written as:

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Algorithms to compute the SVD are variants of those used for eigenvalues. We will not study them in detail.

We can equivalently rewrite the decomposition  $\mathbf{A} = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^{\top}$  as

$$\mathbf{A} = \sum_{i=1}^{\ell} \sigma_i \mathbf{u}_i \mathbf{v}_i^{\top}$$

where  $\mathbf{u}_i$  and  $\mathbf{v}_i$  are the *i*-th columns of  $\mathbf{U}$  and  $\mathbf{V}$ .

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If we round small values of  $\sigma_i$  to zero, we are approximating  ${\bf A}$  with fewer terms:

$$\mathbf{A} \approx \mathbf{U} \mathbf{\tilde{\Sigma}} \mathbf{V}^{\top}$$

where  $\tilde{\Sigma}$  has the small  $\sigma_i$  truncated to zero.

Construct the matrix:

$$\mathbf{\tilde{A}} \equiv \mathbf{U}\mathbf{\tilde{\Sigma}}\mathbf{V}^{\top}$$

by truncating all but the first  $\boldsymbol{k}$  largest singular values to zero.

Construct the matrix:

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by truncating all but the first k largest singular values to zero.

**Theorem (Eckart-Young)** The matrix above minimizes the error  $\|\mathbf{A} - \tilde{\mathbf{A}}\|_F$  subject to the constraint that the column space of  $\tilde{\mathbf{A}}$  has at most dimension k.

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The rank of a matrix is the dimension of its column space.

Then, truncating the singular values gives a low-rank approximation (i.e. rank at most k) of the initial matrix  $\mathbf{A}$ .

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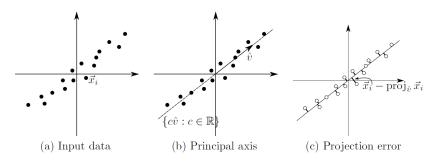
The rank of a matrix is the dimension of its column space.

Then, truncating the singular values gives a low-rank approximation (i.e. rank at most k) of the initial matrix A.

Low-rank approximations have numerous applications!

# Principal component

Consider the two-dimensional data in this plot:



 ${f Q}$ : Find the vector  ${f v}$  such that each data point  ${f x}_i$  can be written as

$$\mathbf{x}_i = c_i \mathbf{v}$$

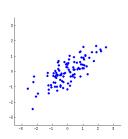
where each  $\mathbf{x}_i$  has its own  $c_i$ 

### Another perspective

Let us be given n data points stored in matrix  $\mathbf{X} \in \mathbb{R}^{d \times n}$ :

$$\mathbf{X}^{\top} = \begin{pmatrix} \mathbf{x}_1^{\top} & \mathbf{\dots} \\ & \vdots \\ \mathbf{x}_n^{\top} & \mathbf{\dots} \end{pmatrix}$$

.

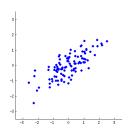


### Another perspective

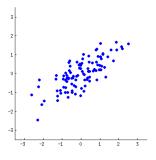
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$$\mathbf{X}^{\top} = \begin{pmatrix} \mathbf{x}_1^{\top} & \mathbf{\dots} \\ \vdots & \vdots \\ \mathbf{x}_n^{\top} & \mathbf{\dots} \end{pmatrix} \approx \begin{pmatrix} - & \tilde{\mathbf{x}}_1^{\top} & - \\ \vdots & \vdots \\ - & \tilde{\mathbf{x}}_n^{\top} & - \end{pmatrix} = \tilde{\mathbf{X}}^{\top}$$

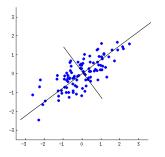
We want to replace them with a lower-dimensional approximation  $\tilde{\mathbf{X}} \in \mathbb{R}^{k \times n}$ , with  $k \ll d$ .



Regard our data as n points in  $\mathbb{R}^d$ :



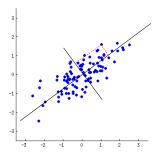
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#### Overall idea:

• Find  $k \le d$  orthogonal directions with the most variance. These span a k-dimensional subspace of the data.

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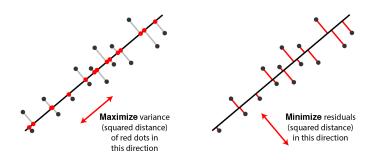
- ullet Find  $k \leq d$  orthogonal directions with the most variance. These span a k-dimensional subspace of the data.
- Project all the data points onto these directions.
   This is lossy, but can be done with the smallest possible error.

We seek the direction w that:

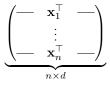
- Minimizes the projection/reconstruction error.
- Maximizes the variance of the projected data.

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Assuming  $\mathbf{W}^{\top}\mathbf{W} = \mathbf{I}$ , for k = d we get:

$$\mathbf{X}^{\top}\mathbf{W} = \mathbf{Z}^{\top}$$
$$\mathbf{X} = \mathbf{W}\mathbf{Z}$$

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Assuming  $\mathbf{W}^{\top}\mathbf{W} = \mathbf{I}$ , for k < d we get:

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We call the columns of W principal components.

They are unknown and must be computed.

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For a given  $\mathbf{w}$ , the projection of all n points onto  $\mathbf{w}$  is  $\mathbf{X}^{\top}\mathbf{w}$ .

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The variance to maximize is  $\|\mathbf{X}^{\top}\mathbf{w}\|_{2}^{2}$ :

$$(\mathbf{X}^{\top}\mathbf{w})^{\top}(\mathbf{X}^{\top}\mathbf{w})$$

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For a given w, the projection of all n points onto w is  $X^Tw$ .

The variance to maximize is  $\|\mathbf{X}^{\top}\mathbf{w}\|_{2}^{2}$ :

$$(\mathbf{X}^{\top}\mathbf{w})^{\top}(\mathbf{X}^{\top}\mathbf{w}) = \mathbf{w}^{\top}(\mathbf{X}\mathbf{X}^{\top})\mathbf{w}$$

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$$(\mathbf{X}^{\top}\mathbf{w})^{\top}(\mathbf{X}^{\top}\mathbf{w}) = \mathbf{w}^{\top}\underbrace{(\mathbf{X}\mathbf{X}^{\top})}_{\mathbf{C}}\mathbf{w}$$

where  $\mathbf{C} \in \mathbb{R}^{d \times d}$  is the symmetric covariance matrix.

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We want to solve the problem:

$$\max_{\mathbf{w}} \mathbf{w}^{\top} \mathbf{C} \mathbf{w}$$
 s.t.  $\|\mathbf{w}\|_2 = 1$ 

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The solution is  $\mathbf{w} = \text{principal eigenvector of } \mathbf{C}$  (Courant minmax principle), and the value  $\mathbf{w}^{\top}\mathbf{C}\mathbf{w}$  is the corresponding eigenvalue.

After solving the problem:

$$\mathbf{w}_1 = \underset{\mathbf{w}}{\operatorname{arg}} \underset{\mathbf{w}}{\operatorname{max}} \ \mathbf{w}^{\top} \mathbf{C} \mathbf{w}$$
  
s.t.  $\|\mathbf{w}\|_2 = 1$ 

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The successive orthogonal direction can be found by solving:

$$\mathbf{w}_2 = \arg\max_{\mathbf{w}} \ \mathbf{w}^{\top} \mathbf{C} \mathbf{w}$$
  
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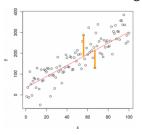
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which is the second eigenvector of C, and so on for all  $\mathbf{w}_{i=2...k}$ .

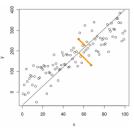
The principal components are thus the first  $k \ll d$  eigenvectors of  ${\bf C}$ , sorted by decreasing eigenvalue.

#### PCA is not linear regression

With linear regression we measure the error along the y coordinate:



With PCA we measure the error orthogonal to the principal direction:



#### PCA as a generative model

Given the  ${\bf W}$  satisfying, for the observations  ${\bf X}$ :

$$\mathbf{X}^{\top}\mathbf{W} = \mathbf{Z}^{\top}$$
 projection  $\mathbf{X} \approx \mathbf{W}\mathbf{Z}$  reconstruction

We can generate new data just by sampling  $\mathbf{z}_{\mathrm{new}} \in \mathbb{R}^k$  and computing:

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#### Example:



Data point  $\mathbf{x}_1$ 



 $\begin{aligned} & \textbf{Generated} \\ & \mathbf{x}_{new} = \frac{1}{2} \mathbf{W} (\mathbf{z}_1 + \mathbf{z}_2) \end{aligned}$ 



Data point  $\mathbf{x}_2$ 

#### Suggested reading

Read Sections 7.1, 7.2.2, 7.2.5 of the book:

J. Solomon, "Numerical Algorithms"