

Optimization

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Gradient Descent

- In general used to find a minimum of any function
- In ML mostly used to find a minimum of the error function (typically MSE) dependent on model parameters

General formulation : $\vec{x} = (x_1, \dots, x_n)$ $f(x_1, \dots, x_n)$

ML formulation (special case of the above) :

$$\min_{\vec{\theta} = (\theta_1, \dots, \theta_n)} \sum_{(\vec{x}, y) \in D} (y - f(\vec{x} | \vec{\theta}))^2$$

↑
training data

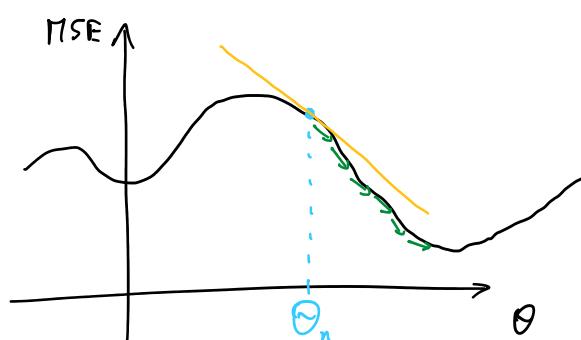
Example

$$\min_{(\theta_0, \theta_1)} \sum_{(x, y) \in D} (y - (\theta_1 x + \theta_0))^2$$

$$f(x | \theta) = \theta_1 x + \theta_0$$

Idea of GD

1. Start with any $\vec{\theta}$
2. Iteratively move $\vec{\theta}$ in the direction opposite to the derivative



The slope of the tangent line is equal to the derivative of the function with respect to θ

$$\frac{\partial \text{MSE}}{\partial \theta}$$

Derivative calculation for MSE

For simplicity we assume a linear model

$$f(x | \theta_0, \theta_1) = f(x, \theta_0, \theta_1) = \theta_1 x + \theta_0$$

$$MSE = \frac{1}{|D|} \sum_{(x,y) \in D} (y - (\theta_1 x + \theta_0))^2$$

Since $(f + g)'(x) = f'(x) + g'(x)$ and $(\alpha f(x))' = \alpha f'(x)$ we can make the derivative calculations for a single element in the sum (single datapoint) and then average.

$$\begin{aligned} \frac{\partial}{\partial \theta_0} (y - (\theta_1 x + \theta_0))^2 &= 2(y - (\theta_1 x + \theta_0)) \cdot \frac{\partial}{\partial \theta_0} (y - (\theta_1 x + \theta_0)) \\ &= 2(y - (\theta_1 x + \theta_0))(-1) \\ &= -2(y - (\theta_1 x + \theta_0)) \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial \theta_1} (y - (\theta_1 x + \theta_0))^2 &= 2(y - (\theta_1 x + \theta_0)) \cdot \frac{\partial}{\partial \theta_1} (y - (\theta_1 x + \theta_0)) \\ &= 2(y - (\theta_1 x + \theta_0))(-x) \\ &= -2x(y - (\theta_1 x + \theta_0)) \end{aligned}$$

Updating rule

$$\boxed{\vec{\theta}^n = \vec{\theta}^{n-1} - \alpha \frac{\partial}{\partial \theta} MSE(\vec{\theta})}$$

$$\begin{aligned} \theta_0^n &= \theta_0^{n-1} - \alpha \frac{\partial}{\partial \theta_0} MSE(\theta_0^{n-1}, \theta_1^{n-1}) \\ &= \theta_0^{n-1} - \alpha \frac{1}{|D|} \sum_{(x,y) \in D} (-2)(y - (\theta_1 x + \theta_0)) \end{aligned}$$

$$\begin{aligned}\theta_1^{(n)} &= \theta_1^{(n-1)} - \alpha \frac{\partial}{\partial \theta_1} \text{MSE}(\theta_0^{(n-1)}, \theta_1^{(n-1)}) \\ &= \theta_1^{(n-1)} - \alpha \frac{1}{|D|} \sum_{(x,y) \in D} (-2x)(y - (\theta_1 x + \theta_0))\end{aligned}$$

Stochastic Gradient Descent (SGD)

1. Start with any $\vec{\theta}$
 2. Iteratively:
 - a) take a datapoint $(\vec{x}, y) \in D$
 - b) calculate the derivative of MSE on this single datapoint with respect to $\vec{\theta}$
 - c) shift $\vec{\theta}$ in the direction opposite to the derivative

Problem : SGD can be unstable and diverge

Mini-batch Gradient Descent

1. Start with any $\vec{\theta}$
 2. Iteratively:
 - a) take a mini-batch $\{(x_i, y_i)\}_{i=1}^m$
 - b) calculate the derivative of MSE on this mini batch with respect to $\vec{\theta}$
 - c) shift $\vec{\theta}$ in the direction opposite to the derivative

Remark: most of the time
when people say SGD
they mean Mini-batch GD

There are many other variants of SGD used in practice:

- SGD with momentum
 - RMSProp

- SGD
- RMSprop
- NAG
- Adam (the most popular)
- AdaGrad
- AdaDelta

Stochastic Gradient Descent for matrix factorization

$$\min_{\{\rho_u, q_i\}} \underbrace{\sum_{(u,i) \in K} (r_{ui} - q_i^T \rho_u)^2 + \lambda (\|q_i\|^2 + \|\rho_u\|^2)}_{\text{error}}$$

For simplicity consider embedding dim = 2

$$\rho_u = (\rho_{u1}, \rho_{u2})$$

$$q_i = (q_{i1}, q_{i2})$$

$$\rho_{u1}^{(n)} = \rho_{u1}^{(n-1)} - \alpha \frac{\partial}{\partial \rho_{u1}} \text{error}$$

$$\begin{aligned} \frac{\partial}{\partial \rho_{u1}} \text{error} &= \frac{\partial}{\partial \rho_{u1}} \left(\sum_{(u,i) \in K} (r_{ui} - (q_{i1} \rho_{u1} + q_{i2} \rho_{u2}))^2 + \lambda (q_{i1}^2 + q_{i2}^2 + \rho_{u1}^2 + \rho_{u2}^2) \right) \\ &= \sum_{(u,i) \in K} \left[\frac{\partial}{\partial \rho_{u1}} (r_{ui} - (q_{i1} \rho_{u1} + q_{i2} \rho_{u2}))^2 + \frac{\partial}{\partial \rho_{u1}} \lambda (q_{i1}^2 + q_{i2}^2 + \rho_{u1}^2 + \rho_{u2}^2) \right] \\ &= \sum_{(u,i) \in K} \left[2(r_{ui} - (q_{i1} \rho_{u1} + q_{i2} \rho_{u2}))(-q_{i1}) + 2\lambda \rho_{u1} \right] \end{aligned}$$

Denote

$$e_{ui} = r_{ui} - (q_{i1} \rho_{u1} + q_{i2} \rho_{u2})$$

Then

$$\frac{\partial}{\partial \rho_{u1}} \text{error} = \sum_{(u,i) \in K} (-2e_{ui} q_{i1} + 2\lambda \rho_{u1})$$

Analogously for $\rho_{u2}, q_{i1}, q_{i2}$

$$\frac{\partial}{\partial \rho_{u2}} \text{error} = \sum_{(u_i:i) \in U} (-2e_{ui} q_{i2} + 2\lambda \rho_{u2})$$

$$\frac{\partial}{\partial q_{u1}} \text{error} = \sum_{(u_i:i) \in U} (-2e_{ui} \rho_{i1} + 2\lambda q_{u1})$$

$$\frac{\partial}{\partial q_{u2}} \text{error} = \sum_{(u_i:i) \in U} (-2e_{ui} \rho_{i2} + 2\lambda q_{u2})$$

For SGD (error on a single datapoint):

$$(\rho_{u1}^{(n)}, \rho_{u2}^{(n)}) = (\rho_{u1}^{(n-1)}, \rho_{u2}^{(n-1)}) - \lambda \left(\frac{\partial}{\partial \rho_{u1}} \text{error}, \frac{\partial}{\partial \rho_{u2}} \text{error} \right)$$

$$= (\rho_{u1}^{(n-1)}, \rho_{u2}^{(n-1)}) - \lambda (-2e_{ui} q_{i1} + 2\lambda \rho_{u1}, -2e_{ui} q_{i2} + 2\lambda \rho_{u2})$$

$$= (\rho_{u1}^{(n-1)}, \rho_{u2}^{(n-1)}) + 2\lambda (e_{ui} q_{i1} - \lambda \rho_{u1}, e_{ui} q_{i2} - \lambda \rho_{u2})$$

Analogously for q_{i1}, q_{i2}

Rewriting the above equations in vector form gives the following formulation (renaming λ to be $\lambda\mathbf{I}$)

$$\rho_u^{(n)} = \rho_u^{(n-1)} + \lambda (e_{ui} q_i - \lambda \rho_u)$$

$$q_i^{(n)} = q_i^{(n-1)} + \lambda (e_{ui} \rho_u - \lambda q_i)$$

Least Squares for a linear model

$$y = \alpha_1 x_1 + \dots + \alpha_k x_k + \varepsilon \quad \text{where } \varepsilon \sim N(0, \sigma^2)$$

in other words

$$\hat{y} = \alpha_1 x_1 + \dots + \alpha_k x_k$$

$$y = \alpha_1 x_1 + \dots + \alpha_k x_k$$

Problem: given a dataset $D = \{(\vec{x}_n, y_n)\}$
 where $\vec{x}_n = (x_{n1}, \dots, x_{nk}) \in \mathbb{R}^k$
 find $\alpha_1, \dots, \alpha_k \in \mathbb{R}$ which
 minimize the squared error

$$SE = \sum_n (y_n - (\alpha_1 x_{n1} + \dots + \alpha_k x_{nk}))^2$$

SE is a quadratic function of $\alpha_1, \dots, \alpha_k$
 and the coefficient in front of the quadratic term
 is positive hence the minimum is in the point
 $(\alpha_1, \dots, \alpha_k)$ where

$$\nabla \frac{\partial}{\partial \alpha_i} SE(\alpha_1, \dots, \alpha_k) = 0$$

For simplicity of notation consider the 2D case
 where we need to find (α_1, α_2)

We need to find α_1, α_2 where

$$\begin{cases} \frac{\partial}{\partial \alpha_1} \sum_n (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))^2 = 0 \\ \frac{\partial}{\partial \alpha_2} \sum_n (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))^2 = 0 \end{cases}$$

we have

$$\begin{aligned} \frac{\partial}{\partial \alpha_1} \sum_n (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))^2 &= \sum_n \frac{\partial}{\partial \alpha_1} (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))^2 \\ &= \sum_n 2(y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))(-x_{n1}) \end{aligned}$$

$$= -2 \sum_n x_{n1} (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))$$

$$\frac{\partial}{\partial \alpha_2} \sum_n (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))^2 = -2 \sum_n x_{n2} (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2}))$$



$$\left\{ \begin{array}{l} \sum_n x_{n1} (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2})) = 0 \\ \sum_n x_{n2} (y_n - (\alpha_1 x_{n1} + \alpha_2 x_{n2})) = 0 \end{array} \right.$$



$$\left\{ \begin{array}{l} \sum_n x_{n1} y_n = \sum_n x_{n1}^2 \alpha_1 + \sum_n x_{n1} x_{n2} \alpha_2 \\ \sum_n x_{n2} y_n = \sum_n x_{n1} x_{n2} \alpha_1 + \sum_n x_{n2}^2 \alpha_2 \end{array} \right.$$



$$\left\{ \begin{array}{l} \sum_n x_{n1} y_n = \alpha_1 \sum_n x_{n1}^2 + \alpha_2 \sum_n x_{n1} x_{n2} \\ \sum_n x_{n2} y_n = \alpha_1 \sum_n x_{n1} x_{n2} + \alpha_2 \sum_n x_{n2}^2 \end{array} \right. (*)$$



$$\left\{ \begin{array}{l} \sum_n x_{n1} y_n \sum_n x_{n2}^2 = \alpha_1 \sum_n x_{n1}^2 \sum_n x_{n2}^2 + \alpha_2 \sum_n x_{n1} x_{n2} \sum_n x_{n2}^2 \\ \sum_n x_{n2} y_n \sum_n x_{n1} x_{n2} = \alpha_1 \sum_n x_{n1} x_{n2} \sum_n x_{n1} x_{n2} + \alpha_2 \sum_n x_{n2}^2 \sum_n x_{n1} x_{n2} \end{array} \right.$$



$$\alpha_1 = \frac{\sum_n x_{n2} y_n \sum_n x_{n1} x_{n2} - \sum_n x_{n1} y_n \sum_n x_{n2}^2}{(\sum_n x_{n1} x_{n2})^2 - \sum_n x_{n1}^2 \sum_n x_{n2}^2}$$

Denoting $x_1 = (x_{11}, x_{21}, \dots, x_{n1})$ — the first coordinate of every datapoint
 $x_2 = (x_{12}, x_{22}, \dots, x_{n2})$ — the second coordinate of every datapoint

Noting $x_1 = (x_{11}, x_{21}, \dots, x_{n1})$ — the first coordinate of every datapoint
 $x_2 = (x_{12}, x_{22}, \dots, x_{n2})$ — the second coordinate of every datapoint

we can write the formula for α_1 in a concise way

$$\alpha_1 = \frac{(x_2 \cdot y)(x_1 \cdot x_2) - (x_1 \cdot y) \|x_2\|^2}{x_1 \cdot x_2 - \|x_1\|^2 \|x_2\|^2}$$

and for α_2 :

$$\alpha_2 = \frac{(x_1 \cdot y)(x_1 \cdot x_2) - (x_2 \cdot y) \|x_1\|^2}{x_1 \cdot x_2 - \|x_1\|^2 \|x_2\|^2}$$

The system of equations (*) can be written in a more generic way

$$\begin{cases} \sum_n x_{n1} y_n = \alpha_1 \sum_n x_{n1}^2 + \alpha_2 \sum_n x_{n1} x_{n2} \\ \sum_n x_{n2} y_n = \alpha_1 \sum_n x_{n1} x_{n2} + \alpha_2 \sum_n x_{n2}^2 \end{cases}$$

$\uparrow \downarrow$
 $X^T y = X^T X A$

where

$$X = \begin{bmatrix} x_{11} & x_{12} \\ x_{21} & x_{22} \\ \vdots & \vdots \\ x_{n1} & x_{n2} \end{bmatrix} \quad y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} \quad A = \begin{bmatrix} \alpha_1 \\ \alpha_2 \end{bmatrix}$$

Then

$$X^T y = X^T X A \mid \cdot (X^T X)^{-1}$$

\uparrow
 $A = (X^T X)^{-1} X^T y$

Thus form holds also for k -dimensional x vectors

$X^T X$ is a $k \times k$ dimensional matrix
 \Rightarrow if k is small, inverting $X^T X$ is not costly

ALS - Alternating Least Squares

Recall that the matrix factorization problem is given by

$$\min_{\beta_u, q_i \in \mathbb{R}^d} \sum_{(u, i) \in K} (r_{ui} - q_i^T \beta_u)^2$$

If we fix the item representations q_i , then the problem becomes

$$\min_{\beta_u \in \mathbb{R}^d} \sum_i (r_{ui} - (q_{i1} \beta_{u1} + \dots + q_{id} \beta_{ud}))^2$$

where this expression has to be minimized for every user over all possible values of $\beta_u = (\beta_{u1}, \beta_{u2}, \dots, \beta_{ud})$

This is the Linear Least Squares problem!

Therefore

$$\beta_u = (X^T X)^{-1} X^T y$$

where

$$X = \begin{bmatrix} q_{11} & q_{12} & \dots & q_{1d} \\ q_{21} & q_{22} & \dots & q_{2d} \\ \vdots & & & \vdots \\ q_{m1} & q_{m2} & \dots & q_{md} \end{bmatrix} \quad y = \begin{bmatrix} r_{u1} \\ r_{u2} \\ \vdots \\ r_{um} \end{bmatrix}$$

ALS

1. Initialize all user and item representation vectors p_u and q_i with random values
2. Iterate until convergence
(i.e. changing of representations less than ϵ)
 - 2.a. Set all item representations q_i and solve the Linear Least Squares problem for user representations p_u
 - 2.b. Set all user representations p_u and solve the Linear Least Squares problem for item representations q_i

Maximum Likelihood Estimation (MLE)

Consider again the following linear model

$$y = \alpha_1 x_1 + \dots + \alpha_k x_k + \varepsilon \quad \text{where } \varepsilon \sim N(0, \sigma^2)$$

Assuming this model is true the likelihood of observing a datapoint $(x_1, x_2, \dots, x_k, y)$ in the data is equal to

$$L(\varepsilon) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{\varepsilon}{\sigma} \right)^2} = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{y - (\alpha_1 x_1 + \alpha_2 x_2 + \dots + \alpha_k x_k)}{\sigma} \right)^2}$$

The idea behind MLE is that for a given set of observed datapoints $\{(\vec{x}_n, y_n)\} = \{(x_{n1}, x_{n2}, \dots, x_{nk}, y_n)\}$ we want to find such model parameters $\alpha_1, \alpha_2, \dots, \alpha_k$ that the likelihood of observing such dataset is maximal, i.e. we want to solve

$$\max \prod \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{y - (\alpha_1 x_1 + \alpha_2 x_2 + \dots + \alpha_k x_k)}{\sigma} \right)^2}$$

$$\max_{\alpha_1, \dots, \alpha_n} \prod_n \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(y - (\alpha_1 x_1 + \dots + \alpha_n x_n))^2}{\sigma^2}}$$

This expression can be further simplified since:

$$\begin{aligned}
 & \arg \max_{\alpha_1, \dots, \alpha_n} \prod_n \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(y - (\alpha_1 x_1 + \dots + \alpha_n x_n))^2}{\sigma^2}} \\
 = & \arg \max_{\alpha_1, \dots, \alpha_n} \prod_n e^{-\frac{1}{2} \frac{(y - (\alpha_1 x_1 + \dots + \alpha_n x_n))^2}{\sigma^2}} \\
 = & \arg \max_{\alpha_1, \dots, \alpha_n} e^{-\frac{1}{2} \sum_n \frac{(y - (\alpha_1 x_1 + \dots + \alpha_n x_n))^2}{\sigma^2}} \quad \left(\text{since } e^a e^b = e^{a+b} \right) \\
 = & \arg \min_{\alpha_1, \dots, \alpha_n} \sum_n \frac{(y - (\alpha_1 x_1 + \dots + \alpha_n x_n))^2}{\sigma^2} \quad \left(\text{since } e^{-x} \text{ is decreasing} \right) \\
 = & \arg \min_{\alpha_1, \dots, \alpha_n} \sum_n (y - (\alpha_1 x_1 + \dots + \alpha_n x_n))^2
 \end{aligned}$$

But this is exactly Least Squares!

MLE and Least Squares are equivalent
if the noise in the data is normal (Gaussian)

Note

MLE is a powerful and general method which can be used with any probability distribution, for instance Bernoulli, Binomial, Poisson, Exponential, Gamma, Beta