Exploratory Data Analysis with R

Visualizing Experimental Designs

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Outline

- Design of Experiments
- Visualizing Completely Randomized Designs
- Visualizing Randomized Complete Block Designss
- Visualizing Factorial Designs
- Visualizing Variance Components
- Visualizing Multiple Comparisons

Design of Experiments

- Two ways of data collection
 - observational study
 - experimental design
- Experiment: people apply some treatment and then observe its effects on the subjects (subjects in an experiments generally are called experimental units)
 - ▶ An experiment requires random assignment of subjects to treatments.
 - If done correctly, experiments provide most compelling evidence that a treatment causes an observed outcome
 - ▶ For example, in a randomized clinical study patients in the experimental groups receive the drug while patients in the control groups receive a placebo or sugar pill. The patients do not know if they are receiving the experimental treatment or placebo.

Design of Experiments

- Some terminologies:
 - An experimental unit is the object on which a measurement (or measurements) is taken.
 - ▶ The response is the variable being measured by the experimenter.
 - A factor is an independent variable whose values are controlled and varied by the experimenter.
 - A level is the intensity setting of a factor.
 - ► A treatment is a specific combination of factor levels.
- We need several statistical (probabilistic) models in this section.

- Completely randomized design:
 - one factor only; one-way classification.
 - ▶ The treatments are assigned completely randomly to the experiment units.
- The data will be like

Trt_1	Trt ₂		Trt_k
<i>y</i> ₁₁	<i>y</i> ₂₁		y _{k1}
<i>y</i> ₁₂	<i>y</i> ₂₂		y_{k2}
:	:	:	:
y_{1,n_1}	y_{2,n_2}	• • •	y_{k,n_k}

• Statistical model: By regarding each n_i observations as a random sample from an infinite population, the probability model is:

$$Y_{ij} = \mu_i + \varepsilon_{ij}, j = 1, 2, \dots, n_i, i = 1, 2, \dots, k,$$

where $\varepsilon_{ij} \sim N(0, \sigma^2)$ are independent normal random errors with common variance σ^2 .

- Are the k population means the same, or is at least one mean different from the others?
- We are testing $H_0: \mu_1 = \mu_2 = \cdots = \mu_k$ versus $H_a:$ At least two means are different from each other or "Not all means are equal". If H_0 is rejected, which treatments are better.

Example: Is the attention span of children affected by whether or not they had
a good breakfast? Twelve children were randomly divided into three groups
and assigned to a different meal plan. The response was attention span in
minutes during the morning reading time.

No Breakfast	Light Breakfast	Full Breakfast
8	14	10
7	16	12
9	12	16
13	17	15

- The response variable is attention span.
- The experimenter chooses 3 levels of a single factor breakfast
- Each level of the factor is a treatment
- The experiment is replicated 4 times

Data

```
resp=c(8,7,9,13,14,16,12,17,10,12,16,15);
Trt=c(rep(1,4),rep(2,4),rep(3,4));
Breakfast=as.data.frame(cbind(resp, Trt));
Breakfast$Trt =factor(Breakfast$Trt); #make "Trt" a factor
str(Breakfast)

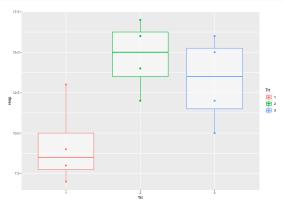
## 'data.frame': 12 obs. of 2 variables:
## $ resp: num 8 7 9 13 14 16 12 17 10 12 ...
## $ Trt : Factor w/ 3 levels "1","2","3": 1 1 1 1 2 2 2 2 3 3 ...
```

• Data (or import the data)

breakfast\$treatment=as.factor(breakfast\$treatment)

Box plot and dot plot

```
library(ggplot2)
ggplot(data=Breakfast, aes(x=Trt, y=resp,color=Trt))+
  geom_boxplot(alpha = 0.5)+ #set transparency
  geom_point();
```



- The Blocking Principle
 - Blocking is a technique for dealing with nuisance factors
 - A nuisance factor is a factor that probably has some effect on the response, but it's of no interest to the experimenter... however, the variability it transmits to the response needs to be minimized
- Typical nuisance factors include batches of raw material, operators, pieces of test equipment, time (shifts, days, etc.), different experimental units

Randomized Complete Block Design

- There are two factors: factor of interest and nuisance factor.
- A block is a specific level of the nuisance factor
- The design uses blocks such that units in each block are relatively similar or homogeneous, with one unit within each block randomly assigned to each treatment (all runs within a block are randomized).
- If the design involves a treatments within each of b blocks, then the total number of observations is ab.
- Variability between blocks can be large, variability within a block should be relatively small
- The purpose of blocking is to isolate the block-to-block variability that might hide/affect the effect of the treatments.
- The design must be balanced: the number of units in each block must be equal to the number of treatments.

 Example: We want to investigate the effect of 3 methods of soil preparation on the growth of seedlings. Each method is applied to seedlings growing at each of 4 locations and the average first year growth is recorded.

Table 1: Seedling Problem

	Location						
Soil Prep	1	2	3	4			
А	11	13	16	10			
В	15	17	20	12			
C	10	15	13	10			

- Treatment = soil preparation (a = 3)
- Block = location (b = 4)
- Is the average growth different for the 3 soil preps?

- Suppose that there are a treatments (factor levels) and b blocks
- Let y_{ij} be the response for the i-th treatment applied to the j-th block. $i=1,2,\ldots,a, j=1,2,\ldots,b$
- The data will be look like this

Table 2: Data from Block Design

	Treatments						
Blocks	1	2		а			
1	<i>y</i> ₁₁	<i>y</i> ₂₁		y _{a1}			
2	y_{11}	<i>y</i> ₂₂	• • •	y_{a2}			
:	÷	:	:	÷			
Ь	<i>y</i> _{1<i>b</i>}	У2ь	13	Уab			

Statistical model:

$$Y_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}, j = 1, 2, \dots, b, i = 1, 2, \dots, a,$$

where $\varepsilon_{ij} \sim N(0, \sigma^2)$ are independent normal random errors with common variance σ^2 .

• The relevant null hypothesis for comparing the a treatment means

$$H_0: \mu_1 = \mu_2 = \ldots = \mu_a$$

Data

```
growth=c(11,13,16,10,15,17,20,12,10,15,13,10);
location=rep(c(1,2,3,4),3);
soil=c(rep("A",4),rep("B",4),rep("C",4));
seedlings=cbind(growth,location,soil);
seedlings=as.data.frame(seedlings)
seedlings$growth=as.numeric(seedlings$growth)
seedlings$location=as.factor(seedlings$location)
seedlings$soil=as.factor(seedlings$soil)
str(seedlings)
```

```
## 'data.frame': 12 obs. of 3 variables:
## $ growth : num 11 13 16 10 15 17 20 12 10 15 ...
## $ location: Factor w/ 4 levels "1","2","3","4": 1 2 3 4 1 2 3 4
## $ soil : Factor w/ 3 levels "A","B","C": 1 1 1 1 2 2 2 2 3 3
```

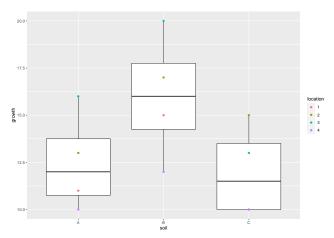
• Data (or import the data)

```
library(readr)
seedlings = read_csv("../data/Soil.csv")
## Rows: 12 Columns: 3
## -- Column specification -----
## Delimiter: ","
## chr (1): soil
## dbl (2): growth, location
##
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this
seedlings$location=as.factor(seedlings$location)
seedlings$soil=as.factor(seedlings$soil)
str(seedlings)
## spec_tbl_df [12 x 3] (S3: spec_tbl_df/tbl_df/tbl/data.frame)
##
   $ growth : num [1:12] 11 15 10 13 17 15 16 20 13 10 ...
##
   $ soil : Factor w/ 3 levels "A", "B", "C": 1 2 3 1 2 3 1 2 3 1 ...
```

##

\$ location: Factor w/ 4 levels "1","2","3","4": 1 1 1 2 2 2 3 3 3 4 ...

```
ggplot(data=seedlings, aes(x=soil, y=growth))+
geom_boxplot()+
geom_point(aes(color=location));
```



Two-factor Factorial Design

- A two-way classification which involves two factors, both of which are of interest to the experimenter.
- There are a levels of factor A and b levels of factor B the experiment is replicated n times at each factor-level combination.
- The effect of a factor is defined to be the change in mean response produced by a change in the level of the factor (generally when the factor is changed from low to high)
- The replications in the design allow the experimenter to investigate the interaction between factors A and B.

Interaction

- There is an interaction between two factors if the effect of one of the factors changes for different levels of the other factor.
- Interaction describes the effect of one factor on the behavior of the other. If there is no interaction, the two factors behave independently.
- Example: The data in the table are categorized with two factors:
 - Sex: Male or Female
 - Blood Lead Level: Low, Medium, or High
- The subcategories are called cells, and the response variable is IQ score. The data are presented on the next slide:

Measures of Performance IQ

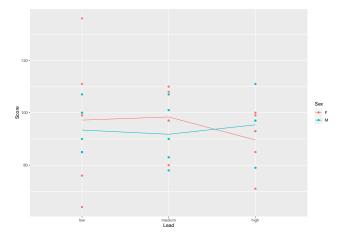
	Blood Lead Level						
	Low	Medium	High				
	85	78	93				
Male	90	107	97				
	107	90	79				
	85	83	97				
	100	101	111				
	64	97	100				
	111	80	71				
Female	76	108	99				
	136	110	85				
	99	97	93				

• Let's explore the IQ data in the table by calculating the mean for each cell and constructing an interaction graph.

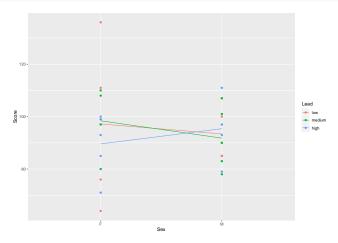
Or import the data

```
library(readr)
IQ = read_csv("../data/IQPerformance.csv")
IQ$Lead=as.factor(IQ$Lead)
IQ$Sex=as.factor(IQ$Sex)
str(IQ)
```

```
IQ$Lead=factor(IQ$Lead, levels=c("low", "medium", "high"))
library(ggplot2);
ggplot(IQ, aes(x=Lead, y=Score, group=Sex, col=Sex)) +
  geom_point() + stat_summary(fun= mean, geom = "line");
```



```
ggplot(IQ, aes(x=Sex, y=Score, group=Lead, col=Lead)) +
  geom_point() + stat_summary(fun= mean, geom = "line");
```



Numerical summary

```
library(dplyr);
IQ%>%group_by(Lead, Sex)%>%summarise(mvalue=mean(Score))->IQ_mean;
IQ_mean
```

```
# A tibble: 6 x 3
  # Groups: Lead [3]
##
##
   Lead
         Sex
              mvalue
## <fct> <fct> <dbl>
## 1 low F
           97.2
  2 low M
           93.4
  3 medium F
              98.4
           91.8
## 4 medium M
  5 high
           89.6
## 6 high M
                95.4
```

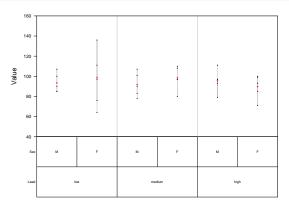
2-Way ANOVA for Factorial Design

Interaction

- An interaction effect is suggested if the line segments are far from being parallel.
- No interaction effect is suggested if the line segments are approximately parallel.
- For the IQ scores, it appears there is an interaction effect:
 - ► Females with high lead exposure appear to have lower IQ scores, while males with high lead exposure appear to have high IQ scores.

- When there are two or more factors, we may want to check the spread of data due to the factors - Variance Components
- The R-package VCA is used to perform variance component analysis.

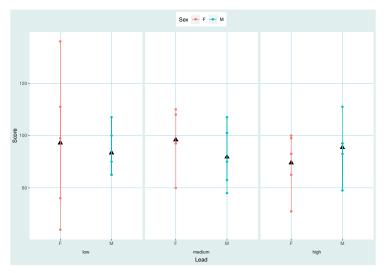
```
library(VCA);
IQ=as.data.frame(IQ) #the data must be a dataframe
varPlot(form=Score~Lead+Sex, Data=IQ)
```



Let's try using ggplot

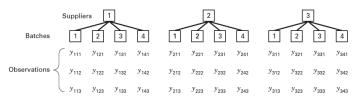
```
IQ%>%group_by(Lead,Sex) %>% mutate(y_mean=mean(Score)) %>%
ggplot(aes(x=Sex,y=Score,color=Sex))+
 geom_point()+
 geom_point(aes(x=Sex,y=y_mean), pch=17,color="black",size=3)+
  geom_line()+
 facet_grid(~Lead,scales="free_x",space="free_x",switch="x")+
 labs(x="Lead")+
  theme(legend.position = "top",
   plot.background = element_rect(fill="azure2"),
   panel.background = element_rect(fill = "white"),
   panel.grid.major = element line(size = 0.2,
              linetype = 'solid',colour = "lightblue"))+
 theme(strip.placement = "outside",
        # Place facet labels outside x axis labels
        strip.background = element_rect(fill = "azure2"))
```

• Let's try using ggplot



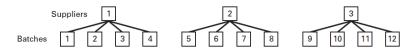
Nested Designs

- In certain multifactor experiments, the levels of one factor (e.g., factor B) are similar but not identical for different levels of another factor (e.g., A). Such an arrangement is called a **nested**, **or hierarchical**, design, with the levels of factor B nested under the levels of factor A.
- For example, consider a company that purchases its raw material from three different suppliers. The company wishes to determine whether the purity of the raw material is the same from each supplier.
 - There are four batches of raw material available from each supplier, and three determinations of purity are to be taken from each batch.



Nested Designs

- This is a two-stage nested design, with batches nested under suppliers.
 - ► The number of stages can be more than two. See VCAdata1 in the VCA package of a three-stage nested design
- Batch 1 from supplier 1 has no connection with batch 1 from any other supplier, batch 2 from supplier 1 has no connection with batch 2 from any other supplier, and so forth.
- To emphasize the fact that the batches from each supplier are different batches, we may renumber the batches as 1, 2, 3, and 4 from supplier 1; 5, 6, 7, and 8 from supplier 2; and 9, 10, 11, and 12 from supplier 3,



supplier 1 supplier 2 supplier 3

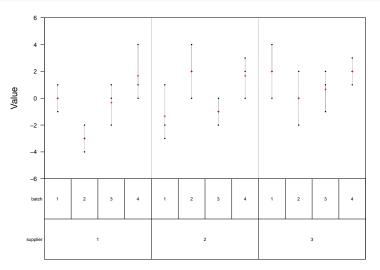
Nested Designs

			Supplier 1				Supplier 2			Supplier 3			
	Batches	1	2	3	4	1	2	3	4	1	2	3	4
		1	-2	-2	1	1	0	-1	0	2	-2	1	3
		-1	-3	0	4	-2	4	0	3	4	0	-1	2
		0	-4	1	0	-3	2	-2	2	0	2	2	1
Batch totals	$y_{ij.}$	0	-9	-1	5	-4	6	-3	5	6	0	2	6
Supplier totals	у _і		-5					4				14	

```
y=c(1, -1, 0, 1, -2, -3, 2, 4, 0, ## batch 1
   -2, -3, -4, 0, 4, 2, -2, 0, 2, ## batch 2
   -2, 0, 1, -1, 0, -2, 1, -1, 2, ## batch 3
   1, 4, 0, 0, 3, 2, 3, 2, 1); ## batch 4
purity=data.frame(y=y, batch=factor(rep(1:4, each=9)),
               supplier=factor(rep(rep(1:3, each=3), 4)));
str(purity);
## 'data.frame': 36 obs. of 3 variables:
##
   $ v : num 1 -1 0 1 -2 -3 2 4 0 -2 ...
   $ batch : Factor w/ 4 levels "1","2","3","4": 1 1 1 1 1 1 1 1 1 2 ...
##
   $ supplier: Factor w/ 3 levels "1","2","3": 1 1 1 2 2 2 3 3 3 1 ...
##
```

Nested Designs

varPlot(form=y~supplier/batch, Data=purity);#batch is nested in supplier



There are several informal methods for determining which means are different:

- Construct boxplots of the different samples to see if one or more of them is very different from the others.
- Construct confidence interval estimates of the means for the different samples, then compare those confidence intervals to see if one or more of them does not overlap with the others (pairwise comparison) or conduct pairwise hypotheses.
 The method is called LSD (Least Significant Difference).
 - LSD problem: In hypotheses test problems involving a single null hypothesis H_0 the statistical tests are often chosen to control the Type I error rate of incorrectly rejecting H_0 at a pre-specified significance level α . In general, when testing m null hypotheses using independent test statistics, the probability of committing at least one Type I error is $1 (1 \alpha)^m$.

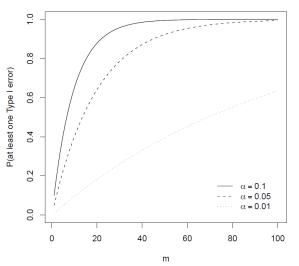


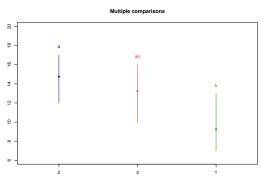
Figure 1: Probability of committing at least one Type I error for different significance levels α and number of hypotheses m.

Several Methods for multiple comparisons:

- Bonferroni test: run t-test on all pairs of hypothesis tests, calculate the p-values and apply a p-value correction for multiple testing problems.
 - Holm procedure: a step-down procedure, which basically consists of repeatedly applying Bonferroni's method while testing the hypotheses in a data-dependent order.
- Tukey's test: also referred to as the Tukey HSD (Honest Significance Difference) test, controls for the Type I error rate across multiple comparisons.
 - ▶ Tukey's test is based on the studentized range. In essence, the Tukey test takes the maximum over the absolute values of all pairwise test statistics.
- Dunnett test(Many-to-one): the standard method for the classical many-to-one problem of comparing several groups with a common control group.

- ullet To visualizing multiple comparisons at significance level lpha, we must first fit a statistical linear model.
- We use package agricolae
 - For the breakfast problem

```
library(agricolae);
Model1=lm(resp ~ Trt, data= Breakfast);
compmeans=HSD.test(aov(Model1),"Trt",alpha=0.05,group=TRUE);
plot(compmeans, main="Multiple comparisons"); box();
```

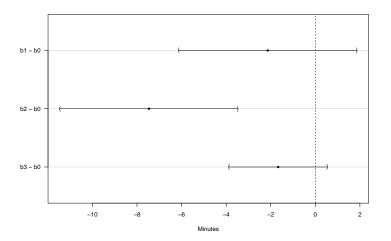


Dunnett test

- Example: A company developed specialized heating blankets designed to help the body heat following a surgical procedure. Four types of blankets b_0 , b_1 , b_2 and b_3 were tested on surgical patients to assess recovery times. The blanket b_0 was a standard blanket already in use at various hospitals. The primary outcome of interest was recovery time in minutes of patients allocated randomly to one of the four treatments. Lower recovery times would indicate a better treatment effect.
- The glht function from multcomp takes the fitted response model to perform the multiple comparisons
- We can use the confint method associated with the glht function to visualize the confidence intervals.

```
library(multcomp)
data(recovery)
siglevel= 0.05
recovery$blanket <- relevel(recovery$blanket, ref = "b0");
# b0 is set as reference level
recovery.aov <- aov(minutes ~ blanket, data = recovery)
recovery.mc <- glht(recovery.aov, linfct = mcp(blanket="Dunnett"),
                    alternative = "two.sided")
#the mcp function for the linfct argument is used
#to specify the comparisons type
summary(recovery.mc);
##
##
     Simultaneous Tests for General Linear Hypotheses
##
## Multiple Comparisons of Means: Dunnett Contrasts
##
##
## Fit: aov(formula = minutes ~ blanket, data = recovery)
##
## Linear Hypotheses:
##
                Estimate Std. Error t value Pr(>|t|)
```

```
recovery.ci <- confint(recovery.mc, level = 1-siglevel)
plot(recovery.ci, main = "", ylim = c(0.5, 3.5), xlab = "Minutes")</pre>
```



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