### Statistics for the Sciences

**Multiple Comparisons** 

Xuemao Zhang East Stroudsburg University

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## **Outline**

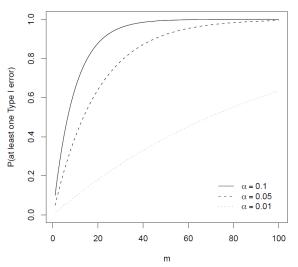
- Multiple Comparisons
  - Bonferroni test
  - Tukey test
  - Contrasts
  - Scheffé Comparison
  - ▶ Dunnett test: Multiple comparisons with a control
- Kruskal-Wallis Test
- Lab

## **Multiple Comparisons**

After conducting one-way ANOVA, if  $H_0$  is rejected, there are several informal methods for determining which means are different:

- Construct boxplots of the different samples to see if one or more of them is very different from the others.
- Construct confidence interval estimates of the means for the different samples, then compare those confidence intervals to see if one or more of them does not overlap with the others (pairwise comparison) or conduct pairwise hypotheses. The method is called LSD (Least Significant Difference).
  - LSD problem: In hypotheses test problems involving a single null hypothesis  $H_0$  the statistical tests are often chosen to control the Type I error rate of incorrectly rejecting  $H_0$  at a pre-specified significance level  $\alpha$ . In general, when testing m null hypotheses using independent test statistics, the probability of committing at least one Type I error is  $1 (1 \alpha)^m$ .

# **Multiple Comparisons**



**Figure 1:** Probability of committing at least one Type I error for different significance levels  $\alpha$  and number of hypotheses m.

- Many multiple comparison methods use the idea of a contrast.
- For example, testing

$$H_0: \mu_3 = \mu_4 \text{ vs. } H_1: \mu_3 \neq \mu_4$$

is equivalent to testing

$$H_0: \mu_3 - \mu_4 = 0$$
 vs.  $H_1: \mu_3 - \mu_4 \neq 0$ 

testing

$$H_0: \mu_1 + \mu_2 = \mu_3 + \mu_4$$
 vs.  $H_1: \mu_1 + \mu_2 \neq \mu_3 + \mu_4$ 

is equivalent to testing

$$H_0: \mu_1 + \mu_2 - \mu_3 - \mu_4 = 0$$
 vs.  $H_1: \mu_1 + \mu_2 - \mu_3 - \mu_4 \neq 0$ 

• In general, a contrast is a linear combination of the parameters of the form

$$\Gamma = \sum_{i=1}^k c_i \mu_i,$$

where the contrast constants  $c_1, \ldots, c_k$  sum to zero,  $\sum_{i=1}^k c_i = 0$ .

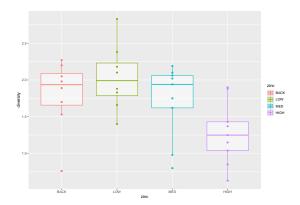
Both the above hypotheses can be expressed in terms of contrasts:

$$H_0: \sum_{i=1}^k c_i \mu_i = 0 \text{ vs. } H_1: \sum_{i=1}^k c_i \mu_i \neq 0$$

 Testing hypotheses involving contrasts can be done using the general t-test since we are testing a linear combination of independent population means.

• For the Medley example, we have

```
## zinc mean
## 1 BACK 1.797500
## 2 LOW 2.032500
## 3 MED 1.717778
## 4 HIGH 1.277778
```



- Suppose we test  $H_0: \mu_1 + \mu_3 = \mu_2 + \mu_4$  or  $H_0: \mu_1 \mu_2 + \mu_3 \mu_4 = 0$ 
  - ightharpoonup We check the p-value of the test to decide if  $H_0$  can be rejected

```
##
     Simultaneous Tests for General Linear Hypotheses
##
##
  Multiple Comparisons of Means: User-defined Contrasts
##
##
## Fit: aov(formula = Model1)
##
## Linear Hypotheses:
          Estimate Std. Error t value Pr(>|t|)
##
## 1 == 0 0.2050
                       0.3203
                                 0.64
                                         0.527
## (Adjusted p values reported -- single-step method)
```

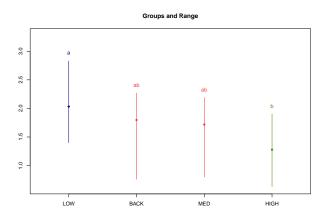
# Scheffé Comparison

- Scheffé test was devised in order to be able to test all the possible contrasts a posteriori
  while maintaining the overall Type I error for the family at a reasonable level, as well as
  trying to have a relatively powerful test.
- Scheffé Comparison is based on F tests

```
##
    Posthoc multiple comparisons of means: Scheffe Test
##
##
      95% family-wise confidence level
##
## $zinc
##
                   diff
                            lwr.ci
                                       upr.ci pval
## LOW-BACK 0.23500000 -0.4549775 0.92497753 0.7973
## MED-BACK -0.07972222 -0.7502599 0.59081541 0.9886
## HIGH-BACK -0.51972222 -1.1902599 0.15081541 0.1769
## MED-I.OW -0.31472222 -0.9852599 0.35581541 0.5929
## HTGH-LOW -0.75472222 -1.4252599 -0.08418459 0.0223 *
## HIGH-MED -0.44000000 -1.0905171
                                   0.21051705 0.2809
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ', 1
```

# **Scheffé Comparison**

• Visualization of the grouping



# Many-to-one comparisons: Dunnett test

- Dunnett test is the standard method for the the classical many-to-one problem of comparing several groups with a common control group. Suppose there are m+1 treatment and let  $\mu_0$  be the mean of the control group. Then we are testing  $H_{0i}: \mu_0 = \mu_i, i = 1, \ldots, m$  against one of the three alternatives  $(>, <, \neq)$ .
  - ▶ Rejecting any of the null hypotheses thus ensures that at least one of the alternative is supported at a given confidence level  $1-\alpha$ , if suitable multiple comparison procedures are employed.
- The one-sided Dunnett test takes the minimum (or the maximum, depending on the sideness of the test problem) of the m, say, pairwise t tests

$$t_i = rac{ar{y}_i - ar{y}_0}{s\sqrt{rac{1}{n_i} + rac{1}{n_0}}}, i = 1, \dots, m.$$

# Many-to-one comparisons: Dunnett test

- Each test statistic  $t_i$  is univariate t distributed. The vector of test statistics  $t = (t_1, \ldots, t_m)$  follows an m-variate t distribution with degrees of freedom  $\sum_{i=0}^m n_i (m+1)$  (and a correlation matrix).
- In the Medley example, suppose we compare diversity from all other zinc concentrations to that from the BACK level at significance level 0.05.

```
##
    Simultaneous Tests for General Linear Hypotheses
##
  Multiple Comparisons of Means: Dunnett Contrasts
##
##
## Fit: aov(formula = diversity ~ zinc, data = medley)
##
## Linear Hypotheses:
##
                  Estimate Std. Error t value Pr(>|t|)
## LOW - BACK == 0 0.23500 0.23303 1.008 0.6195
## MED - BACK == 0 -0.07972 0.22647 -0.352 0.9701
## HIGH - BACK == 0 -0.51972  0.22647 -2.295  0.0728 .
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
## (Adjusted p values reported -- single-step method)
```

##

# Many-to-one comparisons: Dunnett test

95% confidence intervals

```
## Estimate lwr upr
## LOW - BACK 0.23500000 -0.3400227 0.81002269
## MED - BACK -0.07972222 -0.6385438 0.47909939
## HIGH - BACK -0.51972222 -1.0785438 0.03909939
## attr(,"conf.level")
## [1] 0.95
## attr(,"calpha")
## [1] 2.467581
```

## **Nonparametric Statistics**

- Parametric tests have requirements about the nature or shape of the populations involved.
- Nonparametric tests do not require that samples come from populations with specified distribution assumptions, for example normal distributions or any other particular distributions, and homogeneous variances. Consequently, nonparametric tests are called distribution-free tests.
- Nonparametric methods tend to waste information because exact numerical data are often reduced to a qualitative form.
  - Ranks of data are often used.
- Nonparametric tests are not as efficient (powerful) as parametric tests (if
  assumptions for parametric tests are satisfied), so with a nonparametric test
  we generally need stronger evidence (such as a larger sample or greater
  differences) in order to reject a null hypothesis.

### Rank

- Data are sorted when they are arranged according to some criterion, such as smallest to the largest or best to worst.
- A rank is a number assigned to an individual sample item according to its order in the sorted list. The first item is assigned a rank of 1, the second is assigned a rank of 2, and so on.
- Nonparametric methods use the rank information in a data set.

#### Rank

 Handling Ties in Ranks: Find the mean of the ranks involved and assign this mean rank to each of the tied items.

**Example** The numbers 4, 5, 5, 5, 10, 11, 12, and 12 are given ranks of 1, 3, 3, 3, 5, 6, 7.5, and 7.5, respectively. The table below illustrates the procedure for handling ties.

Sorted Data	Preliminary Ranking	Rank
4	1	1
<sup>5</sup> )	2)	3
5 }	3 Mean is 3.	3
5	<sub>4</sub> J	3
10	5	5
11	6	6
12 }	7 Mean is 7.5.	7.5
12 }	8 Mean is 7.5.	7.5

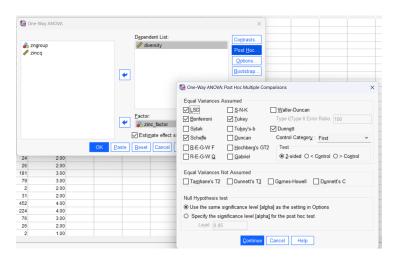
#### Kruskal-Wallis Test

- The Kruskal-Wallis H Test is a nonparametric procedure that can be used to compare more than two populations in a completely randomized design.
- All  $n = n_1 + n_2 + \ldots + n_k$  measurements are jointly ranked.
- We use the sums of the ranks of the k samples (k > 2) to compare the distributions.
- For the Medley example

```
## BACK LOW MED HIGH
## 1.935 1.990 1.940 1.250

##
## Kruskal-Wallis rank sum test
##
## data: diversity by zinc
## Kruskal-Wallis chi-squared = 8.7367, df = 3, p-value = 0.033
```

- Consider the data medley.csv with response diversity and fixed-effect factor zinc.
- ullet Analyze o Compare Means and Proportions o One-Way ANOVA ....



- We read the results of various Post Hoc Tests
- For example, the results of Tukey tests

Multiple Comparisons

Dependent Variable: diversity

			Mean Difference			95% Confid	ence Interval
	(I) zinc_factor	(J) zinc_factor	(I-J)	Std. Error	Sig.	Lower Bound	Upper Bound
Tukey HSD	1.00	2.00	23500	.23303	.746	8686	.3986
		3.00	.07972	.22647	.985	5361	.6955
		4.00	.51972	.22647	.122	0961	1.1355
	2.00	1.00	.23500	.23303	.746	3986	.8686
		3.00	.31472	.22647	.515	3011	.9305
		4.00	.75472*	.22647	.012	.1389	1.3705
	3.00	1.00	07972	.22647	.985	6955	.5361
		2.00	31472	.22647	.515	9305	.3011
		4.00	.44000	.21970	.210	1574	1.0374
	4.00	1.00	51972	.22647	.122	-1.1355	.0961
		2.00	75472*	.22647	.012	-1.3705	1389
		3.00	44000	.21970	.210	-1.0374	.1574

#### **Homogeneous Subsets**

#### diversity

			Subset for a	lpha = 0.05
	zinc_factor	И	1	2
Tukey HSD*,b	4.00	9	1.2778	
	3.00	9	1.7178	1.7178
	1.00	8	1.7975	1.7975
	2.00	8		2.0325
	Sig.		.122	.515
Scheffe <sup>a,b</sup>	4.00	9	1.2778	
	3.00	9	1.7178	1.7178
	1.00	8	1.7975	1.7975
	2.00	8		2.0325
	Sig.		.177	.593

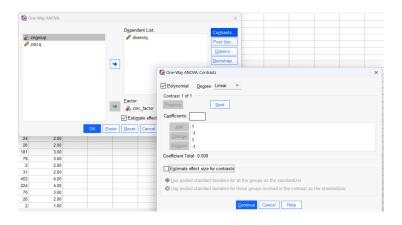
Means for groups in homogeneous subsets are displayed.

b. The group sizes are unequal. The harmonic mean of the group sizes is used. Type I error levels are not guaranteed.

a. Uses Harmonic Mean Sample Size = \$.471.

- The Homogeneous Subsets table in SPSS output for post hoc tests helps to identify groups of means that are not significantly different from each other.
  - ► The table lists the means for each group (identified by the zinc\_factor values 4.00, 3.00, 1.00, and 2.00).
  - Groups that appear in the same column subset (e.g., 1 or 2) are considered homogeneous, meaning their means are not significantly different from each other.
  - ▶ The Sig. row indicates the p-value for the comparisons within each subset.
- For both Tukey HSD and Scheffe tests, the means of groups 4.00, 3.00, and 1.00 are not significantly different from each other (Subset 1).
- Additionally, groups 3.00, 1.00, and 2.00 are also not significantly different from each other (Subset 2).

• Let's check one contrast  $H_0: \mu_1 - \mu_2 + \mu_3 - \mu_4 = 0$  (equivalent to  $H_0: \mu_1 + \mu_3 = \mu_2 + \mu_4$ )



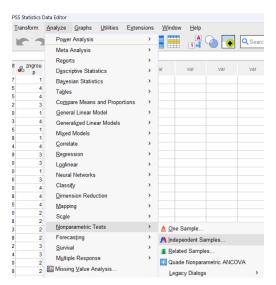
#### Contrast Coefficients

		zinc	factor	
Contrast	1.00	2.00	3.00	4.00
1	1	-1	1	-1

#### Contrast Tests

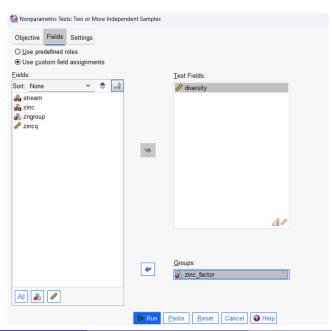
			Value of					95% Confid	ence Interval
		Contrast	Contrast	Std. Error	t	df	Sig. (2-tailed)	Lower	Upper
diversity	Assumes equal variances	1	.2050	.32027	.640	30	.527	4491	.8591
	Does not assume equal variances	1	.2050	.32023	.640	29.103	.527	4498	.8598

 $\bullet$  Kruskal-Wallis Test: Analyze  $\to$  Nonparametric Tests  $\to$  Independent Samples ....

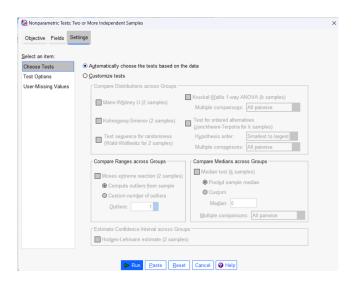


#### I ah





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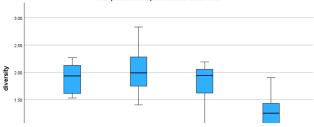


#### Independent-Samples Kruskal-Wallis Test

#### diversity across zinc\_factor

Independent-Samples Kruskal-Walii	s 1est summary
Total N	34
Test Statistic	8.737*
Degree Of Freedom	3
Asymptotic Sig.(2-sided test)	.033

#### Independent-Samples Kruskal-Wallis Test



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