Applied Statistical Methods Bootstrap

Xuemao Zhang East Stroudsburg University

April 14, 2023

Outline

- Introduction
- Bootstrap concept
- Bootstrap method
- Python: Bootstrap
- Python: Block bootstrap

Introduction

- In the section we discuss another resampling method: bootstrap.
- The bootstrap is a flexible and powerful statistical tool that can be used to quantify the uncertainty associated with a given estimator or statistical learning method.
- It can provide an estimate of the standard error of a coeffcient, or a confdence interval for that coeffcient.
 - ▶ For example, we used bootstrap to estimate a population mean (Lecture 15).

- Suppose that we wish to invest a fixed sum of money in two financial assets that yield returns of X and Y, respectively, where X and Y are random quantities.
- We will invest a fraction α of our money in X, and will invest the remaining $1-\alpha$ in Y.
- We wish to choose α to minimize the total risk, or variance, of our investment. In other words, we want to minimize $Var[\alpha X + (1 \alpha)Y]$.
- One can show that the value that minimizes the risk is given by

$$\alpha = \frac{\sigma_Y^2 - \sigma_{XY}}{\sigma_X^2 + \sigma_Y^2 - 2\sigma_{XY}},$$

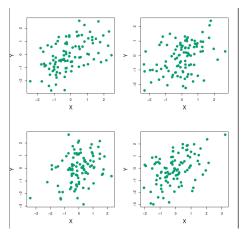
where $\sigma_X^2 = Var(X), \sigma_Y^2 = Var(Y)$, and $\sigma_{XY} = Cov(X, Y)$.

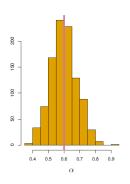
- But the values of σ_X^2, σ_Y^2 and σ_{XY} are unknown.
- We can compute estimates for these quantities, $\hat{\sigma}_X^2$, $\hat{\sigma}_Y^2$ and $\hat{\sigma}_{XY}$, using a data set that contains measurements for X and Y.
- \bullet We can then estimate the value of α that minimizes the variance of our investment using

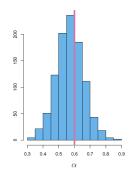
$$\hat{\alpha} = \frac{\hat{\sigma}_Y^2 - \hat{\sigma}_{XY}}{\hat{\sigma}_X^2 + \hat{\sigma}_Y^2 - 2\hat{\sigma}_{XY}},$$

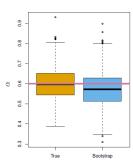
• We can estimate α using a single set of data. How do we estimate the standard deviation of $\hat{\alpha}$.

• Each panel displays 100 simulated returns for investments X and Y. From left to right and top to bottom, the resulting estimates for α are 0.576, 0.532, 0.657, and 0.651.





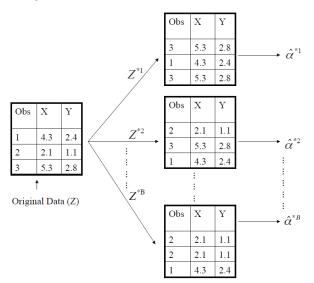




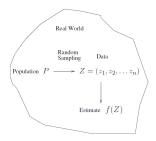
- To estimate the standard deviation of $\hat{\alpha}$, we repeated the process of simulating 1,000 paired observations of X and Y, and estimating α 1,000 times.
- We thereby obtained 1,000 estimates for α , which we can call $\hat{\alpha}_1, \hat{\alpha}_2, \dots, \hat{\alpha}_{1000}$.
- The left-hand panel of the Figure displays a histogram of the resulting estimates.
- For these simulations the parameters were set to $\sigma_X^2 = 1, \sigma_Y^2 = 1.25$ and $\sigma_{XY} = 0.5$, and so we know that the true value of α is 0.6 (indicated by the red line).
- ullet The standard deviation of the 1000 estimates gives us $SE(\hat{lpha}) pprox 0.0083$

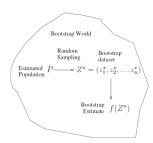
- In practice, however, the procedure for estimating $SE(\hat{\alpha})$ outlined above cannot be applied, because for real data we cannot generate new samples from the original population.
- However, the bootstrap approach allows us to use a computer to mimic the process of obtaining new data sets, so that we can estimate the variability of our estimate without generating additional samples.
- Rather than repeatedly obtaining independent data sets from the population, we instead obtain distinct data sets by repeatedly sampling observations from the original data set with replacement.
- Each of these "bootstrap data sets" is created by sampling with replacement, and is the same size as our original dataset. As a result some observations may appear morethan once in a given bootstrap data set and some not at all.

• Each bootstrap data set contains *n* observations, sampled with replacement from the original data set.



• A general picture for the bootstrap





- Primarily used to obtain standard errors of an estimate.
- A Bootstrap Percentile confidence interval can be obtained by the distribution of the bootstrap estimates.

- In more complex data situations, figuring out the appropriate way to generate bootstrap samples can require some thought.
- For example, if the data is a time series, we can't simply sample the observations with replacement (why not?).
- We can instead create blocks of consecutive observations, and sample those with replacements. Then we paste together sampled blocks to obtain a bootstrap dataset.

Can the bootstrap estimate prediction error?

- To estimate prediction error using the bootstrap, we could think about using each bootstrap dataset as our training sample, and the original sample as our validation sample.
- But each bootstrap sample has significant overlap with the original data. About two-thirds of the original data points appear in each bootstrap sample
- This will cause the bootstrap to seriously underestimate the true prediction error.
- The other way around with original sample = training sample, bootstrap dataset = validation sample - is worse!
- Can partly fix this problem by only using predictions for those observations that did not (by chance) occur in the current bootstrap sample.
- But the method gets complicated, and in the end, cross-validation provides a simpler, more attractive approach for estimating prediction error.

- Performing a bootstrap analysis in Python entails only two steps.
 - ▶ First, we must create a function that computes the statistic of interest.
 - Second, we use the bootstrap() function, which is from scipy.stats, to perform the bootstrap by repeatedly sampling observations from the data set with replacement.
- We illustrate the use of the bootstrap in the simple example on the Portfolio data set.
- We use bootstrap to estimate the accuracy of the linear regression model on the Auto data set.

Consider the Portfolio data set in the ISLR package in R.

```
## 0 -0.895251 -0.234924

## 1 -1.562454 -0.885176

## 2 -0.417090 0.271888

## 3 1.044356 -0.734198

## 4 -0.315568 0.841983
```

• We first create a function, alpha_fn(), which takes as input the (X, Y) data as well as a vector indicating which observations should be used to estimate α .

```
import numpy as np
def alpha_fn(X, Y):
    return ((np.var(Y)-np.cov(X,Y)[0,1])/
    (np.var(X)+np.var(Y)-2*np.cov(X,Y)[0,1]))
```

ullet This function returns an estimate for lpha from the solution we just saw. For instance, the following command tells python to estimate lpha using all 100 observations.

```
x=Portfolio.X
y=Portfolio.Y
print(alpha_fn(x,y))
```

0.5766511516664772

- Next, We use the sample() function to randomly select 100 observations from the range 1 to 100, with replacement.
 - https: //pandas.pydata.org/docs/reference/api/pandas.DataFrame.sample.html

```
import random
random.seed(1)
dfsample = Portfolio.sample(frac=1, replace=True)
x=dfsample.X
y=dfsample.Y
print(alpha_fn(x,y))
```

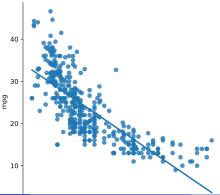
0.6414204016475976

• We can implement a bootstrap analysis by performing this command many times, recording all of the corresponding estimates for α , and computing the resulting standard deviation.

```
def bstrap(df, R):
  import numpy as np
 estimate_boot=np.zeros(R)
  for i in range(0,R):
      dfsample = df.sample(frac=1, replace=True)
      x = dfsample.X
      y = dfsample.Y
      estimate_boot[i] = alpha_fn(x,y)
  return(np.mean(estimate_boot), np.std(estimate_boot))
bstrap(Portfolio, 1000)
## (0.573447645233924, 0.0912081001073443)
```

- We use the bootstrap approach to assess the variability of the estimates of the regression coefficients.
- Consider the Auto data set and use horsepower to predict mpg.

```
import pandas as pd
Auto= pd.read_csv('.../data/Auto.csv')
import matplotlib.pyplot as plt
import seaborn as sns
sns.lmplot(data=Auto, x='horsepower', y='mpg',ci=None, fit_reg=True)
```



 We first create a simple function, boot_fn(), which takes in the Auto data set and the number of samples, and returns the standard errors of the bootstrap estimates for the intercept and slope terms.

```
import statsmodels.api as sm
import statsmodels.formula.api as smf
import numpy as np
def boot fn(df, R):
 intercept_boot=np.zeros(R)
 slope_boot=np.zeros(R)
 for i in range(0,R):
     dfsample = df.sample(frac=1, replace=True)
     model=smf.ols(formula='mpg~horsepower', data=dfsample)
     fit=model.fit()
     intercept_boot[i]=fit.params[0]
      slope_boot[i]=fit.params[1]
 return( np.mean(intercept_boot), np.std(intercept_boot),
np.mean(slope_boot), np.std(slope_boot) )
boot fn(Auto, 1000)
## (39.994132348560456, 0.8398902170755723, -0.1583891463241258, 0.0073347172520215
```

• Compare the bootstrap estimates with the general linear regression model fit.

```
model=smf.ols(formula='mpg~horsepower', data=Auto)
fit=model.fit()
print(fit.summary())
```

##	6								
##									
##	Dep. Variable:			mpg	R-sqı	ıared:			0.606
##	Model:			OLS	Adj.	R-squar	red:		0.605
##	Method:		Least Squ	ares	F-sta	atistic:			599.7
##	Date:	7	Thu, 13 Apr	2023	Prob	(F-stat	istic):	:	7.03e-81
##	Time:		14:2	5:16	Log-I	Likeliho	ood:		-1178.7
##	No. Observation	ns:		392	AIC:				2361.
##	Df Residuals:			390	BIC:				2369.
##	Df Model:			1					
##	Covariance Type	e:	nonro	bust					
##	==========								
##		coef	std err		t	P>	t	[0.025	0.975]
##									
##	Intercept	39.9359	0.717	55	660	0.0	000	38.525	41.347
##	horsepower	-0.1578	0.006	-24	489	0.0	000	-0.171	-0.145
##	==========								=======
##	Omnibus:		16	.432	Durb	in-Watso	n:		0.920
##	<pre>Prob(Omnibus):</pre>		0	.000		ıe-Bera	(JB):		17.305
Xue	mao Zhang East Stroudsbur	g University	Applie	d Statistica	l Method			April 14, 2023	21/32

We also can consider the quadratic regression model

```
import statsmodels.api as sm
import statsmodels.formula.api as smf
import numpy as np
def boot fn2(df, R):
  beta0_boot=np.zeros(R)
 beta1_boot=np.zeros(R)
 beta2 boot=np.zeros(R)
 for i in range(0,R):
     dfsample = df.sample(frac=1, replace=True)
     model=smf.ols(formula='mpg~horsepower+I(horsepower^2)', data=dfsample)
     fit=model.fit()
     beta0 boot[i]=fit.params[0]
     beta1 boot[i]=fit.params[1]
     beta2_boot[i]=fit.params[2]
 from tabulate import tabulate
  col names = ["", "Estimate", "Std. Error"]
  output = [["Intercept", np.mean(beta0_boot), np.std(beta0_boot)],
        ["horsepower", np.mean(beta1 boot), np.std(beta1 boot)],
        ["I(horsepower^2)", np.mean(beta2_boot), np.std(beta2_boot)]]
 print(tabulate(output, headers=col names) )
```

boot fn2(Auto, 1000)

```
##
                 Estimate
                        Std. Error
               39.9642 0.874706
## Intercept
## horsepower -0.136747 0.125158
## I(horsepower^2) -0.021405
                            0.124721
```

Prob(Omnibus):

Xuemao Zhang East Stroudsburg University

Skew:

Kurtosis.

model2=smf.ols(formula='mpg~horsepower+I(horsepower^2)', data=Auto) fit2=model2.fit()

```
print(fit2.summary())
##
                                OLS Regression Results
                                            R-squared:
                                                                              0.606
## Dep. Variable:
                                      mpg
## Model:
                                      OLS Adj. R-squared:
                                                                              0.604
## Method:
                           Least Squares F-statistic:
                                                                              299.1
                                                                           2.12e-79
## Date:
```

Thu, 13 Apr 2023 Prob (F-statistic):

Time: 14:25:24 Log-Likelihood: -1178.6392 ATC:

No. Observations: 2363. ## Df Residuals: 389 BIC: 2375. ## Df Model: nonrobust

Covariance Type:

##		coef	std err	t	P> t	[0.025	0
	Intercept	39.9433	0.719	55.534	0.000	38.529	4:
	horsepower	-0.1321	0.124	-1.065	0.288	-0.376	(
444	T(b ^ 0)	0.0050	0 101	0 000	0.025	0 070	,

I(horsepower

Applied Statistical Methods

0.000

0.499

3 307

Jarque-Bera (JB):

Prob(JB):

Cond No.

0.921

17.833

456

23 / 32

0.000134

April 14, 2023

Omnibus: 16.878 Durbin-Watson:

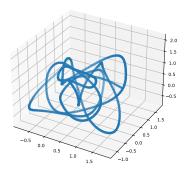
Suppose there is a time series data set.

```
Xy=pd.read_csv("../data/Xy.csv")
Xy.head(10)
```

```
##
       t
                 X 1
                           X2
                                       У
                     0.805921
## 0
          1.297720
                                0.298968
          1.267323
                     0.799034
                                0.318134
## 1
## 2
          1.236882
                     0.792169
                                0.337201
## 3
          1.206317
                     0.785296
                                0.356121
          1.175553
                     0.778385
                                0.374842
## 4
## 5
          1.144513
                     0.771404
                                0.393312
## 6
          1.113118
                     0.764324
                                0.411482
## 7
          1.081305
                     0.757109
                                0.429304
          1.049059
                     0.749703
                                0.446740
## 8
       9
  9
          1.016378
                     0.742045
                                0.463757
##
      10
```

• 3-d plot: https://matplotlib.org/stable/gallery/mplot3d/scatter3d.html

```
import matplotlib.pyplot as plt
fig = plt.figure()
ax = plt.axes(projection='3d')
ax.scatter3D(Xy['X1'], Xy['X2'], Xy['y'])
plt.show()
```



```
import statsmodels.api as sm
import statsmodels.formula.api as smf
model3=smf.ols(formula='y~X1+X2', data=Xy)
```

```
fit3=model3.fit()
print(fit3.summary())
##
                               OLS Regression Results
## Dep. Variable:
                                                                             0.117
                                           R-squared:
## Model:
                                     OLS Adj. R-squared:
                                                                             0.115
                           Least Squares F-statistic:
## Method:
                                                                             66.14
                        Thu, 13 Apr 2023 Prob (F-statistic):
                                                                          1.06e-27
## Date:
## Time:
                                14:25:30 Log-Likelihood:
                                                                           -810.66
## No. Observations:
                                    1000 AIC:
                                                                             1627.
## Df Residuals:
                                     997
                                           BTC:
                                                                             1642.
## Df Model:
                                       2
```

## ##	Covariance Type:		nonrobust				
##		coef	std err	t	P> t	[0.025	0.975]
##	Intercept	0.2658	0.020	13.372	0.000	0.227	0.305
##	X1	0.1453	0.026	5.604	0.000	0.094	0.196
##	X2	0.3134	0.029	10.722	0.000	0.256	0.371
##	========		========				======
##	Omnibus:		9.7	703 Durbin	-Watson:		0.002
	Prob(Omnibi mao Zhang East Stro			008 Jarque Statistical Methods	-Rera (.IR):	April 14, 2023	13.694 26/32

• First consider the general bootstrap method for the simple linear regression model.

```
import statsmodels.api as sm
import statsmodels.formula.api as smf
import numpy as np
from tabulate import tabulate
def boot_fn3(df, R):
 beta0 boot=np.zeros(R)
 beta1 boot=np.zeros(R)
 beta2_boot=np.zeros(R)
 for i in range(0,R):
     dfsample = df.sample(frac=1, replace=True)
     model=smf.ols(formula='v~X1+X2', data=dfsample)
     fit=model.fit()
     beta0_boot[i]=fit.params[0]
     beta1_boot[i]=fit.params[1]
      beta2_boot[i]=fit.params[2]
 col_names = ["", "Estimate", "Std. Error"]
 output = [["Intercept", np.mean(beta0_boot), np.std(beta0_boot)],
        ["X1", np.mean(beta1_boot), np.std(beta1_boot)],
        ["X2", np.mean(beta2_boot), np.std(beta2_boot)]]
 print(tabulate(output, headers=col names) )
```

boot_fn3(Xy, 1000)

##		Estimate	Std. Error
##			
##	Intercept	0.266028	0.0148548
##	X1	0.144487	0.0284277
##	X2	0.311909	0.0354403

• Finally, use the **block bootstrap** to estimate $s.e.(\hat{\beta}_1)$. Use blocks of 100 contiguous observations, and resample ten whole blocks with replacement then paste them together to construct each bootstrap time series. For example, one of bootstrap resamples could be:

```
#Concatenating the ranges
new_rows=[*range(100,200), *range(400,500), *range(100,200),
*range(900,1000), *range(300,400), *range(0,100), *range(0,100),
*range(800,900),*range(200,300), *range(700,800)]
new_Xy=Xy.iloc[new_rows, ]
new_Xy.shape
## (1000, 4)
new_Xy.head()
                   X 1
##
                             X2
             0.327028
                       0.054913 1.135742
## 100
        101
```

```
## t X1 X2 y
## 100 101 0.327028 0.054913 1.135742
## 101 102 0.358675 0.053752 1.134353
## 102 103 0.389793 0.053699 1.132062
## 103 104 0.420223 0.054714 1.128780
## 104 105 0.449808 0.056759 1.124419
```

Now let's write the new bootstrap function.

```
import numpy as np
from random import choices
from tabulate import tabulate
def newboot fn(df. R):
  beta0 boot=np.zeros(R)
  beta1_boot=np.zeros(R)
  beta2 boot=np.zeros(R)
 for i in range(0,R):
   seq=choices(list(range(10)), k=10)
   newrows=[]
   for i in range(10):
      row boot=choices( range(seq[i]*100, (seq[j]+1)*100), k=100) #numpy array of length 100
      newrows.extend(row boot)
   new df=df.iloc[newrows, ] #a block bootstrap sample
   model=smf.ols(formula='v~X1+X2', data=new df)
   fit=model.fit()
   beta0 boot[i]=fit.params[0]
   beta1 boot[i]=fit.params[1]
   beta2 boot[i]=fit.params[2]
  col names = ["", "Estimate", "Std. Error"]
  output = [["Intercept", np.mean(beta0_boot), np.std(beta0_boot)],
  ["X1", np.mean(beta1 boot), np.std(beta1 boot)].
  ["X2", np.mean(beta2 boot), np.std(beta2 boot)]]
  print(tabulate(output, headers=col names)
```

newboot_fn(Xy, 1000)

##		Estimate	Std. Error
##			
##	Intercept	0.278887	0.0919399
##	X1	0.137507	0.201131
##	X2	0.393714	0.324099

License



This work is licensed under a Creative Commons Attribution-NonCommercial-ShareAlike 4.0 International License.