Applied Statistical Methods

Simple Linear Regression Models

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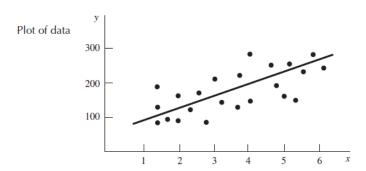
Outline

- Introduction to Simple Linear Regression Models
- Probabilistic Model
- Point Estimations
- Statistical Inferences
 - Hypothesis testing
 - Interval estimations
 - Overall accuracy of the model
 - ANOVA (Analysis of Variance)
 - An example
- Gradient Descent Algorithm

When two variables are measured (not always but usually on a single experimental unit), the resulting data are called bivariate data (or Paired data). When both of the variables (X,Y) are quantitative, call the variable X - the *independent variable*, and Y - the *dependent variable*. A random sample is of the form

$$(X_1, Y_1), (X_2, Y_2), \ldots, (X_n, Y_n).$$

Scatter plot can be used to check the relationship between X and Y. A typical scatter plot is like



Assume that visual examination of the scatter plot confirms that the points approximate a straight-line pattern

$$y=\beta_0+\beta_1x.$$

This model is called a **deterministic** mathematical model because it does not allow for any error in predicting y as a function of x.

However, the bivariate measurements that we observe do not generally fall exactly on a straight line, we choose to use a **probabilistic** model: for any fixed value of x,

$$E(Y|X=x) = \beta_0 + \beta_1 x.$$

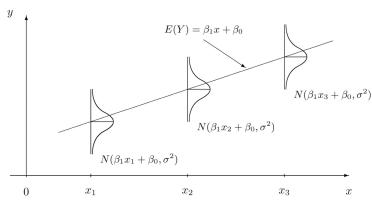
or, equivalently,

$$Y|_{X=x} = \beta_0 + \beta_1 x + \varepsilon,$$

where ε is a random variable possessing a specified probability distribution with mean 0.

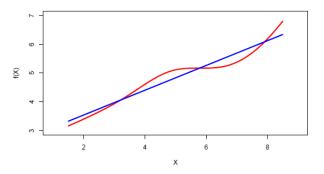
For example, assume that ε 's are independent normal random variables with mean 0 and common variance σ^2 .

Simple Linear Regression Model Introductory Statistics (Shafer and Zhang), UC Davis Stat Wiki



We estimate the population parameters β_0 and β_1 using sample information.

- Linear regression is a simple approach to supervised learning. It assumes that the dependence of Y on the predictors X_1, X_2, \ldots, X_p is linear.
- True regression functions are never linear!



 Although it may seem overly simplistic, linear regression is extremely useful both conceptually and practically.

- Consider the advertising data shown.
- Questions we might ask:
 - Is there a relationship between advertising budget and sales?
 - ▶ How strong is the relationship between advertising budget and sales?
 - Which media contribute to sales?
 - How accurately can we predict future sales?
 - ▶ Is the relationship linear?
 - Is there synergy among the advertising media?

SLR Probabilistic Model

Consider one dependent variable Y and 1 independent variables, X. The data will be in the form of

$$(x_1, y_1), \ldots, (x_n, y_n).$$

Or

Y	X
<i>y</i> ₁	<i>x</i> ₁
:	:
Уn	Xn

Our objective is to use the information provided by the X to predict the value of Y.

SLR Probabilistic Model

Definition. A simple linear statistical model relating a random response Y to an independent variables X is of the form

$$Y|_{X=x_i}=\beta_0+\beta_1x_i+\varepsilon_i,$$

where β_0 and β_1 are unknown parameters, ε is a random variable, and the variables x_1, x_2, \ldots, x_n assume known values. We will assume that $E(\varepsilon_i) = 0, i = 1, \ldots, n$, and hence that

$$E(Y|_{X=x_i}) = \beta_0 + \beta_1 x_i, i = 1, ..., n.$$

• Method of Least Squares for simple linear regression models: The line of means

$$E(Y_i) = \beta_0 + \beta_1 x_i, \quad i = 1, 2, ..., n$$

describes average value of Y_i for any fixed value of x_i , i = 1, 2, ..., n.

Let $\widehat{\beta}_0$ and $\widehat{\beta}_1$ be the estimator of β_0 and β_1 respectively. Then

$$\widehat{Y} = \widehat{\beta}_0 + \widehat{\beta}_1 x$$

is clearly an estimator of E(Y) when X = x. Let

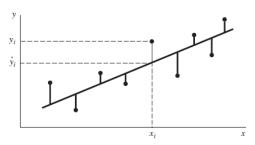
$$\widehat{y}_i = \widehat{\beta}_0 + \widehat{\beta}_1 x_i$$

be the predicted value of the *i*th *y* value (when $X = x_i$).

We choose our estimates $\widehat{\beta}_0$ and $\widehat{\beta}_1$ to estimate β_0 and β_1 so that the vertical distances of the points y_i from the line, are minimized. That is, $\widehat{\beta}_0$ and $\widehat{\beta}_1$ are chosen to minimize the sum of squares of deviations

$$SSE = \sum_{i=1}^{n} (y_i - \widehat{y}_i)^2 = \sum_{i=1}^{n} [y_i - (\widehat{\beta}_0 + \widehat{\beta}_1 x_i)]^2.$$

Fitting a straight line through a set of data points



Least-Squares Estimators for the Simple Linear Regression Model.

1.
$$\widehat{\beta}_1 = \frac{S_{xy}}{S_{xx}}$$
 where $S_{xy} = \sum_{i=1}^n (x_i - \overline{x})(y_i - \overline{y})$ and $S_{xx} = \sum_{i=1}^n (x_i - \overline{x})^2$.
2. $\widehat{\beta}_0 = \overline{y} - \widehat{\beta}_1 \overline{x}$.

$$2. \ \widehat{\beta}_0 = \overline{y} - \widehat{\beta}_1 \overline{x}$$

In the SLR model

$$Y_i|_{X=x_i}=\beta_0+\beta_1x_i+\varepsilon_i,$$

where $E(\varepsilon_i)=0$. Furthermore, assume $V(\varepsilon_i)=\sigma^2$, $i=1,2,\ldots,n$.

Properties of the Least-Squares Estimators:

Summary of properties:

- 1. $\widehat{\beta}_0$ and $\widehat{\beta}_1$ are unbiased. That is, $E(\widehat{\beta}_i) = \beta_i$, for i = 0, 1.
- **2.** $V(\widehat{\beta}_0) = c_{00}\sigma^2$, where $c_{00} = \sum x_i^2/(nS_{xx})$.
- 3. $V(\widehat{\beta}_1) = c_{11}\sigma^2$, where $c_{11} = 1/S_{xx}$.
- **4.** $Cov(\widehat{\beta}_0, \widehat{\beta}_1) = c_{01}\sigma^2$, where $c_{01} = -\overline{x}/S_{xx}$.
- **5.** An unbiased estimator of σ^2 is MSE = SSE/(n-2), and $SSE = S_{yy} \widehat{\beta}_1 S_{xy}$.

If, in addition, the ε_i , for $i=1,2,\ldots,n$ are normal $N(0,\sigma^2)$,

- **6.** Both $\widehat{\beta}_0$ and $\widehat{\beta}_1$ are normally distributed.
- 7. The random variable $\frac{(n-2)MSE}{\sigma^2}$ has a χ^2 distribution with n-2 df.
- **8.** $Cov(\overline{Y}, \widehat{\beta}_1) = 0.$
- **9.** The statistic S^2 is independent of both $\widehat{\beta}_0$ and $\widehat{\beta}_1$.

SLR Hypothesis testing

• Let $S = \sqrt{MSE}$. Then under $H_0: \beta_i = \beta_{i0}, i = 0, 1$,

$$T = \frac{\widehat{\beta}_i - \beta_{i0}}{S\sqrt{c_{ii}}}, \quad i = 0, 1$$

possess a Student's t distribution with n-2 df, where $c_{00} = \sum x_i^2/(nS_{xx})$ and $c_{11} = 1/S_{xx}$.

ullet Note that the testing of eta_1 is actually to test

 H_0 : There is no relationship between X and Y versus

 H_a : There is some relationship between X and Y

SLR Hypothesis testing

• Test of Hypothesis for β_i :

```
H_0: \beta_i = \beta_{i0}
H_a: \begin{cases} \beta_i > \beta_{i0}, & \text{(upper-tail alternative);} \\ \beta_i < \beta_{i0}, & \text{(lower-tail alternative);} \\ \beta_i \neq \beta_{i0}, & \text{(two-tailed alternative).} \end{cases}
\text{Test statistic: } T = \frac{\widehat{\beta}_i - \beta_{i0}}{S\sqrt{c_{ii}}}
\text{Rejection region: } \begin{cases} t > t_{\alpha}, & \text{(upper-tail rejection region);} \\ t < -t_{\alpha}, & \text{(lower-tail rejection region);} \\ |t| > t_{\alpha/2}, & \text{(two-tailed rejection region).} \end{cases}
         where
                             c_{00} = \sum x_i^2/(nS_{xx}), c_{11} = 1/S_{xx}.
             Notice that the t-distribution is based on (n-2) df.
```

SLR Interval Estimations

• A $100(1-\alpha)\%$ Confidence Interval for β_i

$$\widehat{eta}_i \pm t_{lpha/2,n-2} S \sqrt{c_{ii}}$$
 where $c_{00} = \sum x_i^2/(n S_{ ext{xx}}), c_{11} = 1/S_{ ext{xx}}.$

• One useful application of the hypothesis-testing and confidence interval techniques just presented is to the problem of estimating E(Y), the mean value of Y, for a fixed value of the independent variable $x=x^*$.

A 100(1
$$-\alpha$$
)% Confidence Interval for $E(Y) = \beta_0 + \beta_1 x^*$:
$$\widehat{\beta}_0 + \widehat{\beta}_1 x^* \pm t_{\alpha/2, n-2} S \sqrt{\tfrac{1}{n} + \tfrac{(x^* - \overline{x})^2}{S_{xx}}}$$

SLR Interval Estimations

• Predicting a particular value of Y:

Let $x=x^*$ be a fixed value of the independent variable. Instead of estimating the mean Y value at $x=x^*$, we wish to predict the particular (individual) response Y that we will observe if the experiment is run at some time in the future (such as next Monday), denoted by Y^* . Then

$$Y^* = \beta_0 + \beta_1 x^* + \varepsilon.$$

It is natural to estimate Y^* by $\widehat{Y^*} = \widehat{\beta}_0 + \widehat{\beta}_1 x^*$.

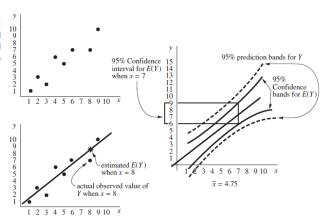
• A $100(1-\alpha)\%$ Prediction Confidence Interval for Y when $x=x^*$

$$\widehat{\beta}_0 + \widehat{\beta}_1 x^* \pm t_{\alpha/2, n-2} S \sqrt{1 + \frac{1}{n} + \frac{(x^* - \overline{x})^2}{S_{xx}}}.$$

SLR Interval Estimations

Remark. Prediction intervals for the actual value of Y are longer than confidence intervals for E(Y) if both confidence levels are the same and both are determined for the same value of $x = x^*$.

Some hypothetical data and associated confidence and prediction bands



The total variation is measured by the Total Sum of Squares (SS_{total}), a measure of the variation in the response values ignoring the regression model:

$$SS_{total} = \sum_{i=1}^{n} (y_i - \overline{y})^2.$$

Now

$$SSE = \sum_{i=1}^{n} (y_i - \widehat{y}_i)^2$$

is a measure of the variation remaining in the response values after predicting them using the fitted regression equation and

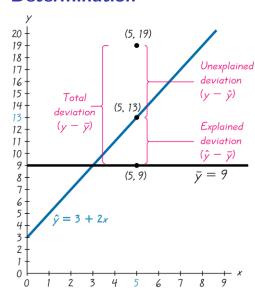
$$SSR = \sum_{i=1}^{n} (\widehat{y}_i - \overline{y})^2$$

is a measure of the variation explained by using x in the SLR model.

The coefficient of determination, denoted by R^2 , is the proportion of the variation in y that is explained by the regression line

$$R^2 = \frac{\text{explained variation}}{\text{total variation}}.$$

That is, it is a measure of: How much of the variation in the response is "explained" by the regression (the linear relationship between X and Y).



Example:

- There is sufficient evidence of a linear correlation
- The equation of the line is

$$\hat{y} = 3 + 2x$$

- The sample mean of the *y*-values is 9.
- One of the pairs of sample data is
 x = 5 and y = 19.
- The point (5,13) is on the fitted regression line.

It can be shown that

$$\sum_{i=1}^{n} (y_i - \overline{y})^2 = \sum_{i=1}^{n} (\widehat{y}_i - \overline{y})^2 + \sum_{i=1}^{n} (y_i - \widehat{y}_i)^2.$$

That is,

$$SS_{total} = SSR + SSE$$
.

Therefore.

$$R^2 = \frac{SSR}{SS_{total}} = 1 - \frac{SSE}{SS_{total}}.$$

Analysis of Variance

The procedure of Analysis of Variance for SLR models can be summarized in the following table.

Source	df	SS	MS	F
Regression	1	SSR	MSR = SSR/1	MSR/MSE
Error	n-2	SSE	MSE = SSE/(n-2)	
Total	n-1	SS_{total}		

Note. The F-test for H_0 : $\beta_1 = 0$ versus H_a : $\beta_1 \neq 0$ is exactly equivalent to the t-test, with

$$t^2 = F$$
.

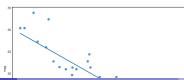
And the F-test statistic has an F distribution under H_0 with $df_1 = 1$, $df_2 = n - 2$.

- Scatter plot with SLR model fit
 - https://seaborn.pydata.org/generated/seaborn.regplot.html

```
import pandas as pd
import matplotlib.pyplot as plt
import seaborn as sns
import statsmodels.api as sm

plt.style.use('seaborn-white')

## <string>:1: MatplotlibDeprecationWarning: The seaborn styles shipped by
mtcars = sm.datasets.get_rdataset('mtcars', "datasets")
mtcars_data = pd.DataFrame(mtcars.data)
sns.regplot(data=mtcars_data, x='wt', y='mpg', ci=None)
plt.show()
```



R code

```
library(ggplot2)
mtcars$cyl <- as.factor(mtcars$cyl)</pre>
ggplot(mtcars, aes(x=wt, y=mpg)) +
 geom_point(size=2);
# Add the regression line:
ggplot(mtcars, aes(x=wt, y=mpg)) +
 geom_point()+
  geom_smooth(method=lm, se=FALSE);
# Add confidence band:
ggplot(mtcars, aes(x=wt, y=mpg)) +
 geom point()+
 geom_smooth(method=lm);
# Scatter plot with fitted regression line and equation:
library(ggpubr);
ggplot(mtcars, aes(x=wt, y=mpg)) +
 geom_point()+
  geom_smooth(method=lm)+
  stat_regline_equation(formula=y ~ x, label.x = 3, label.y = 32);
# label.x and label.y specifies the absolute positioning of the label
```

- There are two main ways to perform linear regression in Python with statsmodels and scikit-learn.
 - statsmodels is "a Python module that provides classes and functions for the estimation of many different statistical models, as well as for conducting statistical tests, and statistical data exploration." (from the documentation)
 - See https://www.statsmodels.org/stable/generated/statsmodels.regression.line ar_model.OLS.html

```
#import statsmodels.formula.api as smf
from statsmodels.formula.api import ols
model1=ols('mpg~wt', mtcars_data).fit()
print(model1.params)
```

```
## Intercept 37.285126
## wt -5.344472
## dtype: float64
```

Model fit summary

```
print(model1.summarv())
##
                      OLS Regression Results
## Dep. Variable:
                               R-squared:
                                                       0.753
                          mpg
                          OLS Adj. R-squared:
## Model:
                                                       0.745
                  Least Squares F-statistic:
## Method:
                                                       91.38
                 Thu, 16 Mar 2023 Prob (F-statistic):
## Date:
                                                   1.29e-10
                       09:09:59 Log-Likelihood:
## Time:
                                                    -80.015
## No. Observations:
                           32 ATC:
                                                       164.0
## Df Residuals:
                           30
                              BIC:
                                                       167.0
## Df Model:
## Covariance Type:
                     nonrobust
## -----
            coef
                  std err
                          t P>|t|
                                             [0.025 0.975]
##
## -----
## Omnibus:
                       2.988 Durbin-Watson:
                                                       1.252
                        0.225 Jarque-Bera (JB):
## Prob(Omnibus):
                                                       2.399
                         0.668 Prob(JB):
## Skew:
                                                       0.301
## Kurtosis:
                         2.877 Cond. No.
                                                        12.7
##
## Notes:
## [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
```

 Confidence interval of the parameters. For example, to get the 95% confidence intervals of the parameters,

```
## 0 1
## Intercept 33.450500 41.119753
## wt -6.486308 -4.202635
```

model1.conf int(alpha=0.05)

Fitted values

```
model1.predict()

## array([23.28261065, 21.9197704 , 24.88595212, 20.10265006, 18.90014396,
## 18.79325453, 18.20536265, 20.23626185, 20.45004071, 18.90014396,
## 18.90014396, 15.53312687, 17.3502472 , 17.08302362, 9.22665041,
## 8.29671236, 8.71892561, 25.52728871, 28.65380458, 27.47802083,
## 24.11100374, 18.47258623, 18.92686632, 16.76235533, 16.73563297,
## 26.94357367, 25.847957 , 29.19894068, 20.34315128, 22.48093991,
## 18.20536265, 22.4274952 ])
model1.fittedvalues
```

```
## Mazda RX4
                         23.282611
## Mazda RX4 Wag
                         21.919770
## Datsun 710
                         24.885952
## Hornet 4 Drive
                       20.102650
## Hornet Sportabout 18.900144
## Valiant
                         18.793255
                         18.205363
## Duster 360
## Merc 240D
                         20.236262
## Merc 230
                         20.450041
```

Residuals

model1.resid

```
## Mazda RX4
                         -2.282611
## Mazda RX4 Wag
                         -0.919770
## Datsun 710
                         -2.085952
## Hornet 4 Drive
                         1.297350
                         -0.200144
## Hornet Sportabout
## Valiant
                         -0.693255
## Duster 360
                         -3.905363
## Merc 240D
                          4.163738
## Merc 230
                          2.349959
## Merc 280
                          0.299856
## Merc 280C
                         -1.100144
## Merc 450SE
                          0.866873
## Merc 450SL
                         -0.050247
## Merc 450SLC
                         -1.883024
## Cadillac Fleetwood 1.173350
## Lincoln Continental
                          2.103288
## Chrysler Imperial
                          5.981074
## Fiat 128
                          6.872711
```

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 ANOVA table: https: //www.statsmodels.org/stable/generated/statsmodels.stats.anova.anova_lm.html

```
anova = sm.stats.anova_lm(model1)
print(anova)
```

```
## df sum_sq mean_sq F PR(>F)
## wt 1.0 847.725250 847.725250 91.375325 1.293959e-10
## Residual 30.0 278.321938 9.277398 NaN NaN
```

- Model fit using package sklearn: https://scikit-learn.org/stable/supervised_learning.html#supervised-learning
 - Linear Regression: https://scikit-learn.org/stable/modules/generated/sklearn.linear model.LinearRegression.html

```
import numpy as np
from sklearn.linear_model import LinearRegression
regr = LinearRegression()
y=mtcars data['mpg']
X=np.array(mtcars_data['wt']).reshape(-1, 1)
#reshape as 1-column matrix since there is only one single feature
regr.fit(X,v)
## LinearRegression()
print(regr.intercept_)
## 37.28512616734204
print(regr.coef_)
## [-5.34447157]
```

Fitted values

```
regr.predict(X)

## array([23.28261065, 21.9197704 , 24.88595212, 20.10265006, 18.90014396, ## 18.79325453, 18.20536265, 20.23626185, 20.45004071, 18.90014396, ## 18.90014396, 15.53312687, 17.3502472 , 17.08302362, 9.22665041, ## 8.29671236, 8.71892561, 25.52728871, 28.65380458, 27.47802083, ## 24.11100374, 18.47258623, 18.92686632, 16.76235533, 16.73563297, ## 26.94357367, 25.847957 , 29.19894068, 20.34315128, 22.48093991, ## 18.20536265, 22.4274952 ])
```

R code

```
model1 = lm(mpg ~ wt, data=mtcars);
coef(model1);
model1summary=summary(model1);
model1summary;
model1summary$fstatistic;
confint(model1, level = 0.95);
anova(model1);
```

- Confidence intervals for future observations: https://www.statsmodels.org/stable/g enerated/statsmodels.regression.linear_model.OLSResults.get_prediction.html
- Suppose Suppose there are two future observation $x_1=3.3$ and $x_2=3.5$.
- Confidence intervals and Prediction confidence intervals

```
import pandas as pd
X1=pd.DataFrame({'wt': [3.3, 3.5]})
predictions=model1.get_prediction(X1)
print(round(predictions.summary_frame(alpha=0.05),3))
```

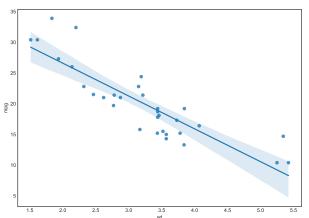
```
##
                      mean_ci_lower mean_ci_upper
                                                    obs ci lower
                                                                  obs_ci_u
       mean
             mean se
## 0
      19.648
               0.540
                              18.545
                                            20.752
                                                          13.331
                                                                        25
## 1
      18.579
               0.561
                              17.433
                                            19.726
                                                          12,254
                                                                        24
```

R code

```
new =data.frame( wt=c(3.3, 3.5) )
predict(model1, newdata=new, interval="confidence", level=0.95)
predict(model1, newdata=new, interval="prediction",level=0.95)
```

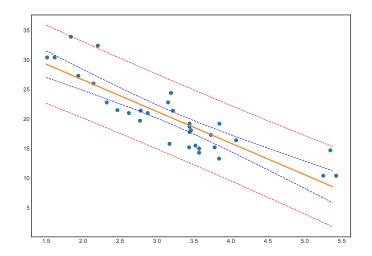
• Plot both confidence band and the scatter plot

```
sns.regplot(data=mtcars_data,x='wt', y='mpg')
#sns.lmplot(data=mtcars_data,x='wt', y='mpg')
plt.show()
```



Plot both confidence band and prediction confidence band to the scatter plot

```
import matplotlib.pyplot as plt
X=pd.DataFrame(np.arange(1.5, 5.5, 0.125),columns=['wt'])
#make sure the number of rows is 32
x=mtcars_data['wt']
predictions=model1.get_prediction(X)
pred_data=predictions.summary_frame(alpha=0.05)
fittedvalues = pred_data.iloc[:,0]
predict_mean_se = pred_data.iloc[:, 1]
predict_mean_ci_low = pred_data.iloc[:, 2]
predict_mean_ci_upp = pred_data.iloc[:, 3]
predict_ci_low = pred_data.iloc[:, 4]
predict_ci_upp = pred_data.iloc[:, 5]
y=mtcars_data['mpg']
plt.plot(x, y, 'o')
plt.plot(X, fittedvalues, '-', lw=2)
plt.plot(X, predict ci low, 'r--', lw=1)
plt.plot(X, predict_ci_upp, 'r--', lw=1)
plt.plot(X, predict_mean_ci_low, 'b--', lw=1)
plt.plot(X, predict_mean_ci_upp, 'b--', lw=1)
plt.show()
```

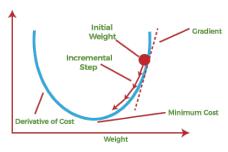


R code

```
library(ggplot2)
temp_var=predict(model1, interval="prediction");
new_df = cbind(mtcars, temp_var);
ggplot(new_df, aes(wt, mpg))+
    geom_point() +
    geom_line(aes(y=lwr), color = "red", linetype = "dashed")+
    geom_line(aes(y=upr), color = "red", linetype = "dashed")+
    geom_smooth(method=lm, se=TRUE);
```

- In mathematics, gradient descent (also often called steepest descent) is a first-order iterative optimization algorithm for finding a local minimum of an objective function, a differentiable function.
- It is a greedy technique that finds the optimal solution by taking a step in the direction of the maximum rate of decrease of the function.
- Gradient descent is by far the most popular optimization strategy used in deep learning.
- We introduce the algorithm for simple linear regression model fit.
 - ▶ Watch Lecture 2.6 Linear Regression With One Variable | Gradient Descent Intuition and Lecture 2.7 Linear Regression With One Variable | Gradient Descent For Linear Regression

- Idea: suppose the objective function y = f(x) has one parameter x only and we try to find a local minimizer or **local minimum** of f(x).
 - If we move towards a negative gradient or away from the gradient of the function at the current point, it will give the **local minimum** of that function.

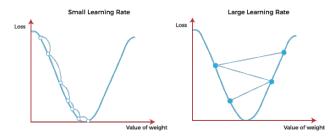


 The main objective of using a gradient descent algorithm is to minimize the cost function using iteration.



- To achieve this goal, it performs two steps iteratively:
 - Calculates the first-order derivative of the function to compute the gradient or slope of that function.
 - Move away from the direction of the gradient, which means slope increased from the current point by α times, where α is defined as **Learning Rate**.
 - It is a tuning parameter in the optimization process which helps to decide the length of the steps.

 Learning Rate: It is defined as the step size taken to reach the minimum or lowest point. This is typically a small value that is evaluated and updated based on the behavior of the cost function.



- The cost/loss function is defined as the measurement of difference or error between actual values and expected values at the current position and present in the form of a single real number.
- Linear Regression Cost/Loss Function:

$$J = J(\beta_0, \beta_1) = \sum_{i=1}^n [y_i - (\beta_0 + \beta_1 x_i)]^2.$$

Gradient:

$$\nabla J = \left(\frac{\partial J}{\partial \beta_0}, \frac{\partial J}{\partial \beta_1}\right)$$

• Gradient Descent:

$$\Delta J = -\alpha \nabla J = -\alpha \left(\frac{\partial J}{\partial \beta_0}, \frac{\partial J}{\partial \beta_1} \right)$$

• Update $\beta = (\beta_0, \beta_1)$ (so we need to choose an initial value) as

$$\beta := \beta + \Delta J$$
.

- If a data set is large, it may take very long time for GD to converge. The solution is SGD.
- SGD (Stochastic Gradient Descent) method uses a random sample of data to calculate the cost/loss function and its derivative in each iteration during the iterating procedure.

 Let's use Gradient Descent Algorithm to estimate parameters in simple linear regression models

```
import numpy as np
def gradient_descent(X, y, theta0, theta1, learning_rate, tol, max_iter):
   error=1
   m = len(X)
   iter count =0
   while error>tol and iter count<max iter:
        # Calculate the predicted values
        v pred = theta0 + theta1*X
        cost = (1/(2*m))*sum((y-y_pred)**2) #cost function
        d theta0 = -(1/m)*sum(y-y pred)
        d_theta1 = -(1/m)*sum((y-y_pred)*X) # Calculate the gradients
        theta0 prev=theta0
        theta1 prev=theta1
        theta0=theta0-learning rate*d theta0
        theta1=theta1-learning rate*d theta1 #Update the parameters
        error=max(abs(theta0 - theta0 prev).abs(theta1 - theta1 prev))
        iter count+=1
   return theta0, theta1, cost, iter count
```

Apply the function above to mtcars data

```
v=np.arrav(mtcars data.mpg)
X=np.array(mtcars data.wt)
theta0 = 0
theta1 = 0 #initial values of beta0 and beta1
learning rate = 0.1
tol=0.0001
max iter=2000
theta0, theta1, cost, iter_count = gradient_descent(X,y,
theta0, theta1, learning rate, tol, max iter)
print("theta0 = ", theta0)
## theta0 = 37.27209076474596
print("theta1 = ", theta1)
## theta1 = -5.340727813952249
print("Cost = ", cost)
## Cost = 4.348787371960817
print("Iterations = ", iter_count)
## Iterations = 1038
```

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