# **Applied Statistics** for Computer Science BSc, Exam

Probability Theory and Mathematical Statistics for Computer Science Engineering BSc, Term grade

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# Main topics

- 1. Probability theory
- 2. Statistics

Mathematical tools: combinatorics, calculus

Computer tool: Matlab

Book:

Yates, Goodman:

Probability and Stochastic Processes: A Friendly Introduction for

Electrical and Computer Engineers

## Lecture 11

**Testing statistical hypotheses** 

**Example.** We want to check if the weight of a bar of chocolate is equal to 100 grams. We weight n = 16 bars. The results are

$$x_1 = 100.3, x_2 = 99.8, \dots, x_{16} = 99.9$$

So it is our sample realization.

#### Notation.

Let m denote the theoretical value of the weight. It is the expectation of the generic random variable X, i.e.  $m=\mathbb{E}X$ . X is the generic random variable of the underlying population.

$$X_1, X_2, \ldots, X_n$$

is the sample from this population.

We denote the prescribed value of m by  $m_0$ . In our example  $m_0 = 100$ .

$$H_0 : m = m_0$$

is called the null hypothesis.

The hypothesis

$$H_1: m \neq m_0$$

is called the alternative hypothesis. It is a two-sided alternative hypothesis.

We should decide if  $H_0$  or  $H_1$  is true.

In the case of z-test it is assumed that the sample comes from a normally distributed population with known variance. That is  $X \sim \mathcal{N}(m, \sigma^2)$ , where m is unknown, but  $\sigma$  is known.



Calculate the empirical mean. We know, that

$$\bar{X} \sim \mathcal{N}(m, \sigma^2/n).$$

So we standardize it, then we obtain a standard normal random variable

$$rac{ar{X}-m}{rac{\sigma}{\sqrt{n}}}\sim \mathcal{N}(0,1).$$

So the test statistic in z-test is

$$z = \frac{\bar{X} - m_0}{\frac{\sigma}{\sqrt{n}}}$$

We emphasize that here we use  $m_0$ . So z is standard normal if and only if  $m=m_0$ , i.e. the the null hypothesis  $H_0$  is true.



A realistic decision is the following. We reject  $H_0$  if the value of |z| is too large, larger than a critical value.

How to find the critical value?

First assume that a small number is preliminary given:

 $\alpha = 0.1, 0.05, 0.01, \dots$  This  $\alpha$  is the error of first kind.

It means that we reject  $H_0$  with probability  $\alpha$  when  $H_0$  is true. And we accept  $H_0$  with probability  $1 - \alpha$ . So we have

$$1 - \alpha = P\left(-z_{\alpha/2} < \frac{\overline{X} - m_0}{\sigma} \sqrt{n} < z_{\alpha/2} \mid H_0\right)$$
  
=  $\Phi(z_{\alpha/2}) - \Phi(-z_{\alpha/2}) = 2\Phi(z_{\alpha/2}) - 1$ ,

where  $\Phi(x)$  is the standard normal CDF . That is  $z_{\alpha/2}$  is the number for which

$$1 - \frac{\alpha}{2} = \Phi(z_{\alpha/2}) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{z_{\alpha/2}} e^{-\frac{x^2}{2}} dx.$$

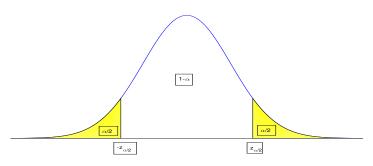


Figure: How to find the critical value for two-sided *z*-test? We cut both tails of the standard normal PDF

**Example.** In the previous example we had  $m_0 = 100$  and n = 16.

Assume that  $\sigma = 0.1$ .

We calculated  $\bar{x} = 99.9$ .

Therefore

$$z = \frac{\bar{X} - m_0}{\frac{\sigma}{\sqrt{n}}} = \frac{99.9 - 100}{\frac{0.1}{\sqrt{16}}} = -4.$$

Choose the significance level as 99% that is  $\alpha = 0.01$ 

As 
$$\Phi(2.58) = 0.995$$
, so  $z_{0.005} = 2.58$ .

As |-4| > 2.58, we reject the null-hypothesis.

The region of acceptance is

$$C_0 = \left\{ (x_1, \ldots, x_n) : \left| \frac{\overline{X} - m_0}{\sigma} \sqrt{n} \right| < z_{\alpha/2} \right\},$$

If the sample falls into  $C_0$ , then we accept  $H_0$ . The critical region is

$$C_1 = \left\{ (x_1, \ldots, x_n) : \left| \frac{\overline{X} - m_0}{\sigma} \sqrt{n} \right| \geq z_{\alpha/2} \right\}.$$

If the sample falls into  $C_1$ , then we reject  $H_0$ .

### The one sided z-test

#### Example.

Assume that a student is on slimming diet, so he/she rejects a bar of chocolate if it is too large. Find the appropriate version of the z-test!

Now

$$H_0 : m = m_0$$

is the null hypothesis.

$$H_1 : m > m_0$$

is the alternative hypothesis. It is a one-sided alternative hypothesis.

We should decide if  $H_0$  or  $H_1$  is true.

Again, it is assumed that the sample comes from a normally distributed population with known variance. That is  $X \sim \mathcal{N}(m, \sigma^2)$ , where m is unknown, but  $\sigma$  is known.

## The one sided z-test

We reject  $H_0$  if the test statistics

$$z = \frac{\bar{X} - m_0}{\frac{\sigma}{\sqrt{n}}}$$

is too large, i.e.  $z > z_{\alpha}$ .

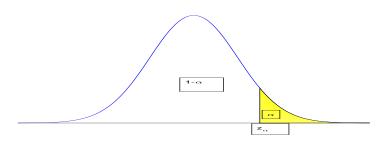


Figure: We cut the right tail of the standard normal PDF to find the critical value for this one-sided z-test



## The one sided z-test

**Example.** In the previous example let  $m_0 = 100$  and n = 25.

Assume that  $\sigma = 0.4$ .

We calculated  $\bar{x} = 100.1$ .

Therefore

$$z = \frac{\bar{X} - m_0}{\frac{\sigma}{\sqrt{n}}} = \frac{100.1 - 100}{\frac{0.4}{\sqrt{25}}} = 1.25.$$

Choose the significance level as 99% that is  $\alpha = 0.01$ 

As 
$$\Phi(2.32) = 0.99$$
, so  $z_{0.01} = 2.32$ .

As 1.25 < 2.32, we accept the null-hypothesis.

## The one sided z-test (the other side)

### Example.

Assume that a student likes chocolate very much, so he/she rejects a bar of chocolate if it is too small. Find the appropriate version of the z-test!

Now

$$H_0 : m = m_0$$

is the null hypothesis.

$$H_1 : m < m_0$$

is the alternative hypothesis. It is a one-sided alternative hypothesis.

We should decide if  $H_0$  or  $H_1$  is true.

Again, it is assumed that the sample comes from a normally distributed population with known variance. That is  $X \sim \mathcal{N}(m, \sigma^2)$ , where m is unknown, but  $\sigma$  is known.

# The one sided z-test (the other side)

We reject  $H_0$  if the test statistics  $z=\frac{\bar{X}-m_0}{\frac{\sigma}{\sqrt{n}}}$  is too small, i.e.  $z<-z_{\alpha}$ .

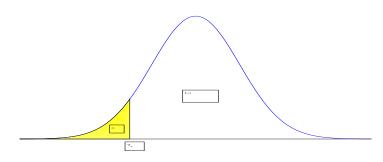


Figure: How to find the critical value for one-sided *z*-test? We should cut the left tail of the standard normal PDF

# The one sided z-test (the other side)

**Example.** In the previous example let  $m_0 = 100$  and n = 36.

Assume that  $\sigma = 0.8$ .

We calculated  $\bar{x} = 99.7$ .

Therefore

$$z = \frac{\bar{X} - m_0}{\frac{\sigma}{\sqrt{n}}} = \frac{99.7 - 100}{\frac{0.8}{\sqrt{36}}} = -2.25.$$

Choose the significance level as 99% that is  $\alpha = 0.01$ 

As 
$$\Phi(2.32) = 0.99$$
, so  $z_{0.01} = 2.32$ .

As -2.32 < -2.25, we accept the null-hypothesis.

## The two-sample z-test

We compare the expectations of two samples.

Let  $X \sim \mathcal{N}(m_1, \sigma_1^2)$ , and  $Y \sim \mathcal{N}(m_2, \sigma_2^2)$ , where  $\sigma_1$  and  $\sigma_2$  are known. Let

$$X_1, X_2, \dots, X_{n_1}$$
; and  $Y_1, Y_2, \dots, Y_{n_2}$ 

be independent samples from populations  $\mathcal{N}(m_1, \sigma_1^2)$ , resp.  $\mathcal{N}(m_2, \sigma_2^2)$ . Let

$$H_0: m_1 = m_2$$

be the null hypothesis. Then the test statistic

$$z = \frac{\overline{X} - \overline{Y}}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

has standard normal distribution if and only if  $H_0$  is true.

The possible alternative hypotheses are  $H_1: m_1 
eq m_2$  or

$$H_1: m_1 > m_2 \text{ or } H_1: m_1 < m_2.$$

The decision is the same as in the one-sample case?

# Type 1 and type 2 errors

|                | We accept $H_0$              | We reject $H_0$                     |
|----------------|------------------------------|-------------------------------------|
| $H_0$ is true  | Good decision                | Bad decision Type 1 error, $\alpha$ |
| $H_0$ is false | Bad decision<br>Type 2 error | Good decision                       |

It is not possible to minimize both type 1 error and type 2 error at the same time.

# The *p*-value

- 1. The significance level  $\alpha$  for a study is chosen before data collection, and it is typically set to  $\alpha=0.1$  (i.e. 10%) or  $\alpha=0.05$  (i.e. 5%) or much lower, depending on the field of study. Sometimes  $1-\alpha$  is given, that is significance level 90% means that  $\alpha=0.1$ . The preliminarily given  $\alpha$  means, that we want to decide as follows. We reject the null hypotheses with probability  $\alpha$ , given that the null hypothesis is true.
- 2. Statistical computer programs often do not require the preliminarily given  $\alpha$ , but they offer p-value. They calculate the value of the test statistic, z=1.85, say. The p-value is the probability of obtaining a result at least as extreme as z=1.85, given that the null hypothesis is true.
- 3. Of course, the above two points of view are comparable. We should reject  $H_0$ , when  $p \le \alpha$ .

# The *p*-value

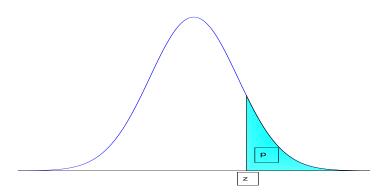


Figure: For one sided *z*-test we obtained z = 1.2. Then the *p*-value is the area on the right tail of the standard normal PDF

# The chi-square distribution

Let  $X_1, \ldots, X_n$  be independent with standard normal distribution. Then the distribution of

$$Y_n = X_1^2 + \cdots + X_n^2$$

is called chi-square distribution with degree of freedom n.

Notation:  $Y_n \sim \chi_n^2$ .

If  $X_1 \sim \mathcal{N}(0,1)$ , then  $\mathbb{E} X_1^2 = 1$ , and  $\mathsf{Var} X_1^2 = 2$ , so

$$\mathbb{E}Y_n = n$$
,  $\operatorname{Var}Y_n = 2n$ .

If  $Y_m$  and  $Y_n$  are independent with  $Y_m \sim \chi_m^2$  and  $Y_n \sim \chi_n^2$ , then  $Y_m + Y_n \sim \chi_{m+n}^2$ .

# The PDF of the chi-square distribution

The PDF of the  $\chi_n^2$  distribution is

$$f(x) = \begin{cases} \frac{x^{\frac{n}{2} - 1} e^{-\frac{x}{2}} 2^{-\frac{n}{2}}}{\Gamma(\frac{n}{2})}, & x > 0, \\ 0, & x \le 0. \end{cases}$$

Here

$$\Gamma(\alpha) = \int_0^\infty u^{\alpha - 1} e^{-u} du$$

is the Gamma function ( $\alpha > 0$ ).

# The PDF of the chi-square distribution

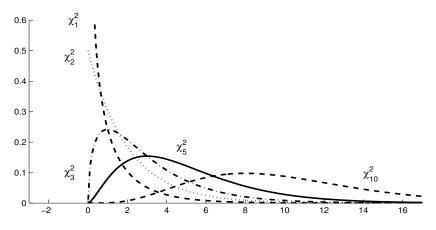


Figure: The PDF of the chi-square distribution for degrees of freedom 1, 2, 3, 5, 10.

# The chi-square distribution converges to the normal distribution

distribution If  $Y_n \sim \chi_n^2$ , then its standardized version converges to standard normal distribution

$$\frac{Y_n-n}{\sqrt{2n}} \implies \mathcal{N}(0,1) \quad \text{as} \ n\to\infty.$$

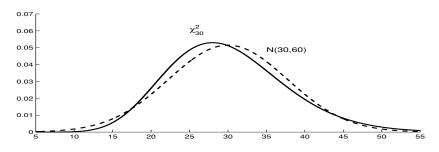


Figure: The PDF's of  $\chi^2_{30}$  and  $\mathcal{N}(30,60)$ .

## The Student distribution

Let X and Y be independent with  $X \sim \mathcal{N}(0,1)$  and  $Y \sim \chi_n^2$ . Then the distribution of

$$\frac{X}{\sqrt{Y/n}}$$

is called Student's t-distribution with degree of freedom n. The PDF of the t-distribution with degree of freedom n is

$$f(x) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{\pi}\sqrt{n}\Gamma\left(\frac{n}{2}\right)\left(1+\frac{x^2}{n}\right)^{\frac{n+1}{2}}},$$

for  $x \in \mathbb{R}$ .

## The PDF of the t-distribution

If  $n \to \infty$ , then the t-distribution with degree of freedom n converges to the standard normal distribution.

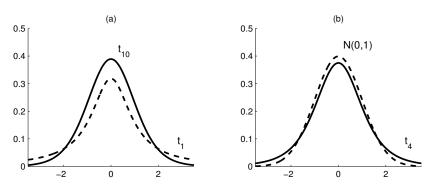


Figure: The PDF's of  $t_1$  and  $t_{10}$  (left) and the PDF's of  $t_4$  and the standard normal PDF (right).

# Sample from normally distributed population

#### Theorem. Let

$$X_1, X_2, \ldots, X_n$$

be a sample from  $\mathcal{N}(m, \sigma^2)$ .

Then  $\bar{X}$  and  $s_n^{*2}$  are independent with

$$\bar{X} \sim \mathcal{N}\left(m, \frac{\sigma^2}{n}\right), \qquad \frac{(n-1)s_n^{*2}}{\sigma^2} \sim \chi_{n-1}^2.$$

Therefore

$$\frac{\bar{X}-m}{\frac{s_n^*}{\sqrt{n}}}$$

is of Student distribution with degree of freedom n-1.

## Student's t-test

The *t*-test is similar to the *z*-test. The major difference of the two tests is that in the case of *t*-test **the variance is unknown**. So let

$$X_1, X_2, \ldots, X_n$$

be a sample from population  $\mathcal{N}(m, \sigma^2)$ , where m and  $\sigma$  are unknown.

The null hypothesis is

$$H_0: m=m_0,$$

where  $m_0$  is a fixed number.

We start with the two-sided alternative hypothesis

$$H_1: m \neq m_0$$



### The t-test

Calculate the empirical mean and the corrected empirical variance from the sample. From our previous theorem we know, that

$$\frac{\bar{X} - m}{\frac{s_n^*}{\sqrt{n}}}$$

has t-distribution with degree of freedom n-1. So the test statistic of t-test is

$$t = \frac{\bar{X} - m_0}{\frac{s_n^*}{\sqrt{n}}}$$

We emphasize that here we use  $m_0$ .

So t has  $t_{n-1}$  distribution if and only if  $m = m_0$ , i.e. the the null hypothesis  $H_0$  is true.

### The two-sided t-test

Using significance level  $\alpha$ , we reject  $H_0$  if  $|t| \geq t_{\alpha/2}$ , where the critical value  $t_{\alpha/2}$  satisfies  $F(t_{\alpha/2}) = 1 - \frac{\alpha}{2}$  and F is the CDF of t-distribution with degree of freedom n-1.

## The two-sided t-test

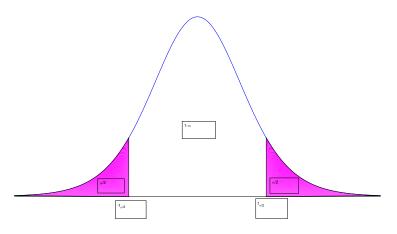


Figure: How to find the critical value for two-sided *t*-test? We cut both tails of the  $t_{n-1}$  PDF

## The two-sided t-test

**Example.** We want to test if the weight of a certain kind of apple is 150 gram. So

$$H_0 : m = 150, H_1 : m \neq 150$$

We collect 100 apples and weight them. Then calculate the empirical mean and the corrected empirical variance from the sample. We obtain  $\bar{X} = 146$ , and  $s_n^* = 25$ . Therefore

$$t = \frac{146 - 150}{\frac{25}{\sqrt{100}}} = -1.6$$

Use significance level 95%. From the table of  $t_{99}$  distribution, we obtain that  $t_{\alpha/2}=1.98$ .

As  $|t| < t_{\alpha/2}$ , so we accept  $H_0$ .



#### t-tests

#### One-sided t-tests

There are two versions of the one sided t-tests: the alternative hypothesis can be either  $H_1: m < m_0$  or  $H_1: m > m_0$ . To handle these cases we should modify the two-sided t-test like we done in the case of z-test.

Two-sample *t*-tests Let

$$X_1, X_2, \ldots, X_{n_1},$$
 and  $Y_1, Y_2, \ldots, Y_{n_2},$ 

be two independent samples from population  $\mathcal{N}(m_1, \sigma_1^2)$ , resp.  $\mathcal{N}(m_2, \sigma_2^2)$ , where  $\sigma_1$  and  $\sigma_2$  are unknown. The null hypothesis is

$$H_0: m_1 = m_2.$$

