

CS 480/680

Introduction to Machine Learning

Lecture 12

Expectation Maximization and Gaussian Mixture Models

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We know how to estimate parameters and make predictions

Problem Type 1:

Given: $\{x_1 = 1, x_2 = 2, x_3\}, x_i \sim \mathcal{N}(\mu = 1.0, \sigma^2 = 1.0)$

Task: Predict x_3

Problem Type 2:

Given: $\{x_1 = 1, x_2 = 2, x_3 = 0\}, x_i \sim \mathcal{N}(\mu, \sigma^2 = 1.0)$

Task: Estimate μ

Can we estimate parameters if data is missing?

Problem Type 3:

Given: $\{x_1 = 1, x_2 = 2, x_3\}, x_i \sim \mathcal{N}(\mu, \sigma^2 = 1.0)$

Task: Estimate (x_3, μ)

How could we solve it?

$\mu:$

$x_3:$

KEY IDEA BEHIND EM ALGORITHM

Lecture Outline

- I. How does the EM algorithm work in a special case?
- II. How does the EM algorithm work in general?

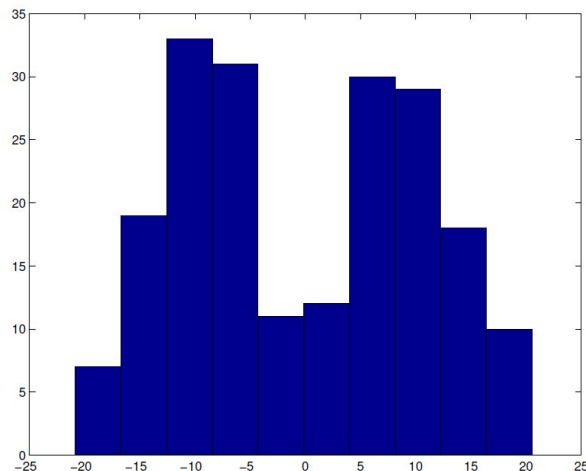


Lecture Outline

- I. How does the EM algorithm work in a special case?**
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Estimating the parameters of a *mixture* of Gaussians



$$X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$$

$$X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$$

$$X = (1 - \Delta) \cdot X_1 + \Delta \cdot X_2$$

Where Δ is a binary random variable:

$$\Delta \in \{0, 1\}$$

Let π denote the probability of Δ taking on the value of 1:

$$\Pr[\Delta = 1] = \pi$$

Let $\mathcal{N}_{\mu, \sigma^2}$ denote the normal density with mean μ and variance σ^2 . Then the density of x is

$$p(x) = (1 - \pi)\mathcal{N}_{\mu_1, \sigma_1^2}(x) + \pi\mathcal{N}_{\mu_2, \sigma_2^2}(x)$$

Can we find the parameters through direct maximization?

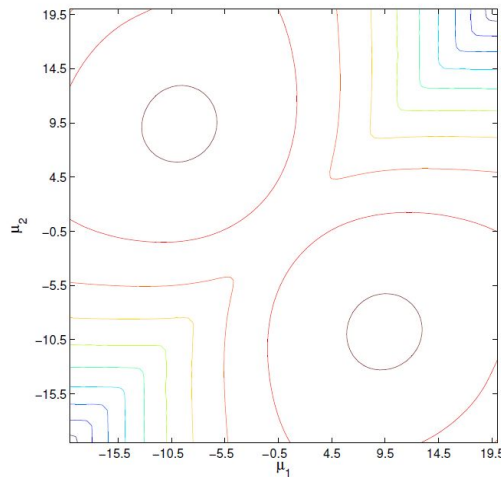
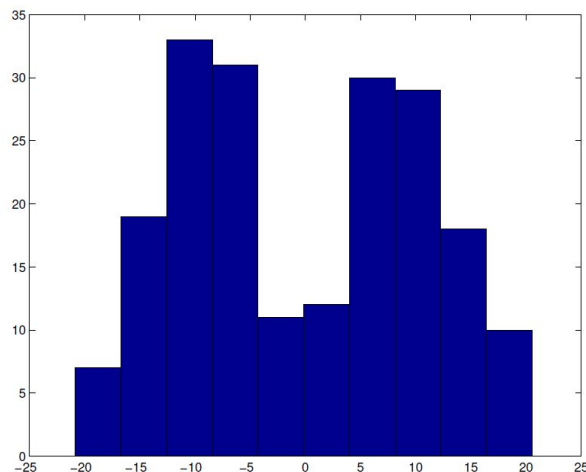
$$p(x) = (1 - \pi)\mathcal{N}_{\mu_1, \sigma_1^2}(x) + \pi\mathcal{N}_{\mu_2, \sigma_2^2}(x)$$

$$\mathcal{L}(\pi, \mu_1, \sigma_1^2, \mu_2, \sigma_2^2 \mid \mathbf{X}) = \prod_{i=1}^n \left[(1 - \pi)\mathcal{N}_{\mu_1, \sigma_1^2}(x_i) + \pi\mathcal{N}_{\mu_2, \sigma_2^2}(x_i) \right]$$

$$\log \mathcal{L}(\pi, \mu_1, \sigma_1^2, \mu_2, \sigma_2^2 \mid \mathbf{X}) = \sum_{i=1}^n \log \left[(1 - \pi)\mathcal{N}_{\mu_1, \sigma_1^2}(x_i) + \pi\mathcal{N}_{\mu_2, \sigma_2^2}(x_i) \right]$$

$$\frac{\partial \log \mathcal{L}}{\partial \sigma_2} = ?? \quad \frac{\partial \log \mathcal{L}}{\partial \mu_2} = ?? \quad \frac{\partial \log \mathcal{L}}{\partial \sigma_1} = ?? \quad \frac{\partial \log \mathcal{L}}{\partial \mu_1} = ?? \quad \frac{\partial \log \mathcal{L}}{\partial \pi} = ??$$

The likelihood function for a mixture model is nonconvex



Label-switching problem:

- Parameters are unidentifiable because likelihood surface has two symmetric modes
- Even with mixing weight π , and variances σ_1^2, σ_2^2 known!



thought experiments ...

Thought experiment 1: If we knew the parameters...

$$p(x) = (1 - \pi)\mathcal{N}_{\mu_1, \sigma_1^2}(x) + \pi\mathcal{N}_{\mu_2, \sigma_2^2}(x)$$

$$\Pr[\Delta_i = 1 \mid \pi, \mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \mathbf{X}]$$

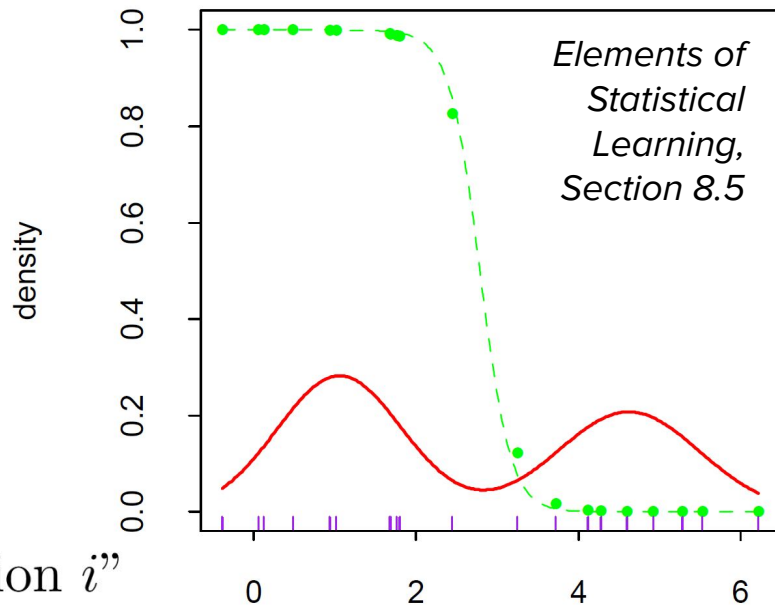
$$= \frac{\pi\mathcal{N}_{\mu_2, \sigma_2^2}(x_i)}{(1 - \pi)\mathcal{N}_{\mu_1, \sigma_1^2}(x_i) + \pi\mathcal{N}_{\mu_2, \sigma_2^2}(x_i)}$$

$$= \gamma_i : \text{“responsibility of mode 2 for observation } i\text{”}$$

$$= \mathbb{E}[\Delta_i \mid \pi, \mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \mathbf{X}] :$$

“expectation of Δ_i given parameters and data”

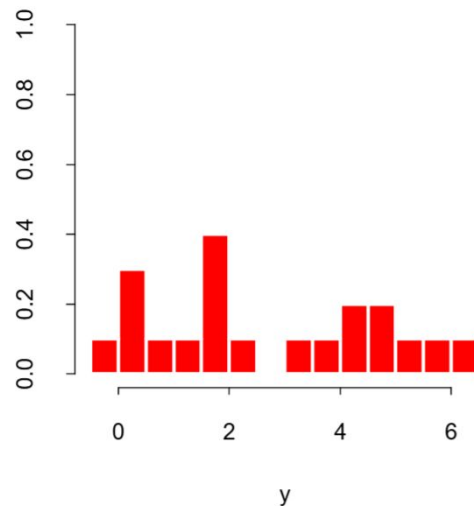
...we could compute the probability of a sample assignment



Thought experiment 2: If we knew the sample assignments...

$$\log \mathcal{L}(\pi, \mu_1, \sigma^2, \mu_2, \sigma_2 \mid \mathbf{X})$$

$$\begin{aligned} &= \sum_{i=1}^n \log \left[(1 - \pi) \mathcal{N}_{\mu_1, \sigma_1^2}(x) + \pi \mathcal{N}_{\mu_2, \sigma_2^2}(x) \right] \\ &= \sum_{i=1}^n \left[(1 - \Delta_i) \log \mathcal{N}_{\mu_1, \sigma_1^2}(x) + \Delta_i \log \mathcal{N}_{\mu_2, \sigma_2^2}(x) \right] \\ &\quad + \sum_{i=1}^n \left[(1 - \Delta_i) \log(1 - \pi) + \Delta_i \log \pi \right] \\ &= \begin{cases} \sum_{i=1}^n \log \mathcal{N}_{\mu_1, \sigma_1^2}(x) + \sum_{i=1}^n \log(1 - \pi) & \text{if } \Delta = 0 \\ \sum_{i=1}^n \log \mathcal{N}_{\mu_2, \sigma_2^2}(x) + \sum_{i=1}^n \log \pi & \text{if } \Delta = 1 \end{cases} \end{aligned}$$



Thought experiment 2: If we knew the sample assignments...

$$= \begin{cases} \sum_{i=1}^n \log \mathcal{N}_{\mu_1, \sigma_1^2}(x) + \sum_{i=1}^n \log(1 - \pi) & \text{if } \Delta = 0 \\ \sum_{i=1}^n \log \mathcal{N}_{\mu_2, \sigma_2^2}(x) + \sum_{i=1}^n \log \pi & \text{if } \Delta = 1 \end{cases}$$

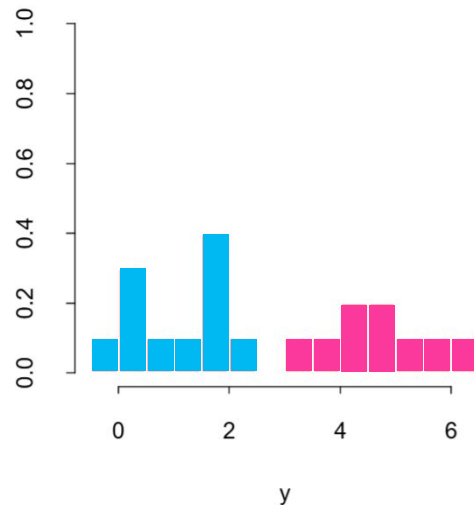
$$\hat{\mu}_1 = \frac{1}{|\Delta_0|} \sum_{i \in \Delta_0} x_i$$

$$\hat{\mu}_2 = \frac{1}{|\Delta_1|} \sum_{i \in \Delta_1} x_i$$

$$\hat{\sigma}_1^2 = \frac{1}{|\Delta_0|} \sum_{i \in \Delta_0} (x_i - \hat{\mu}_1)^2$$

$$\hat{\sigma}_2^2 = \frac{1}{|\Delta_1|} \sum_{i \in \Delta_1} (x_i - \hat{\mu}_2)^2$$

$$\hat{\pi} = \frac{1}{N} \sum_{i=1}^n \Delta_i$$

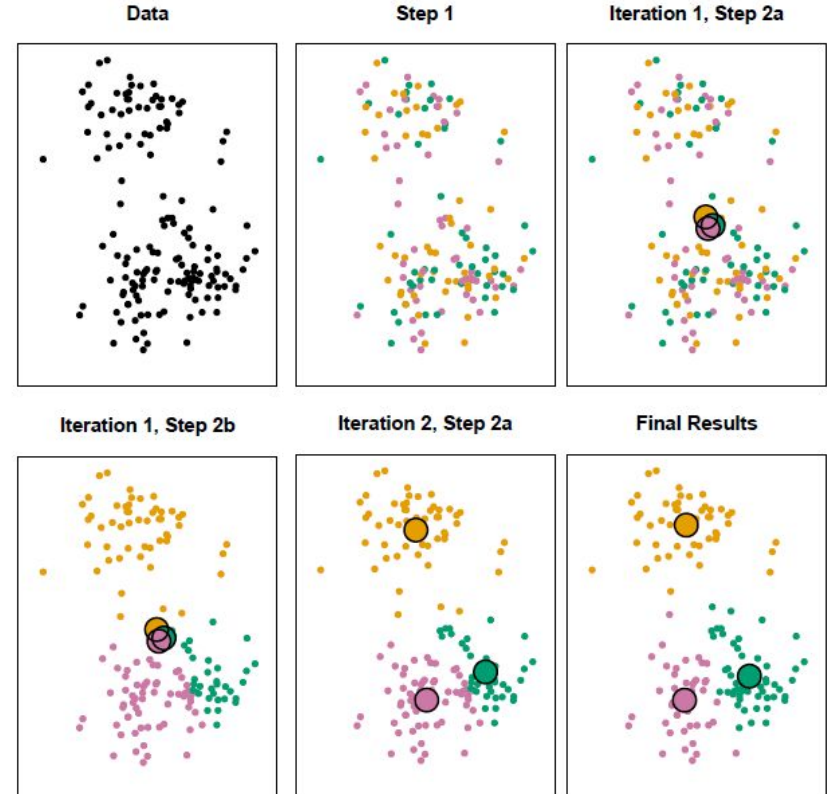
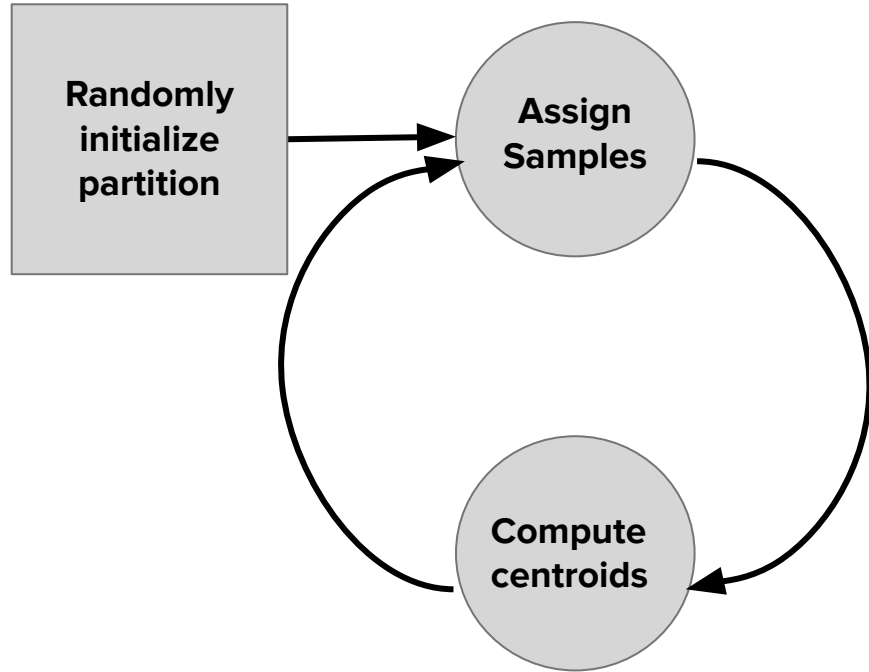


... we could compute the parameters empirically

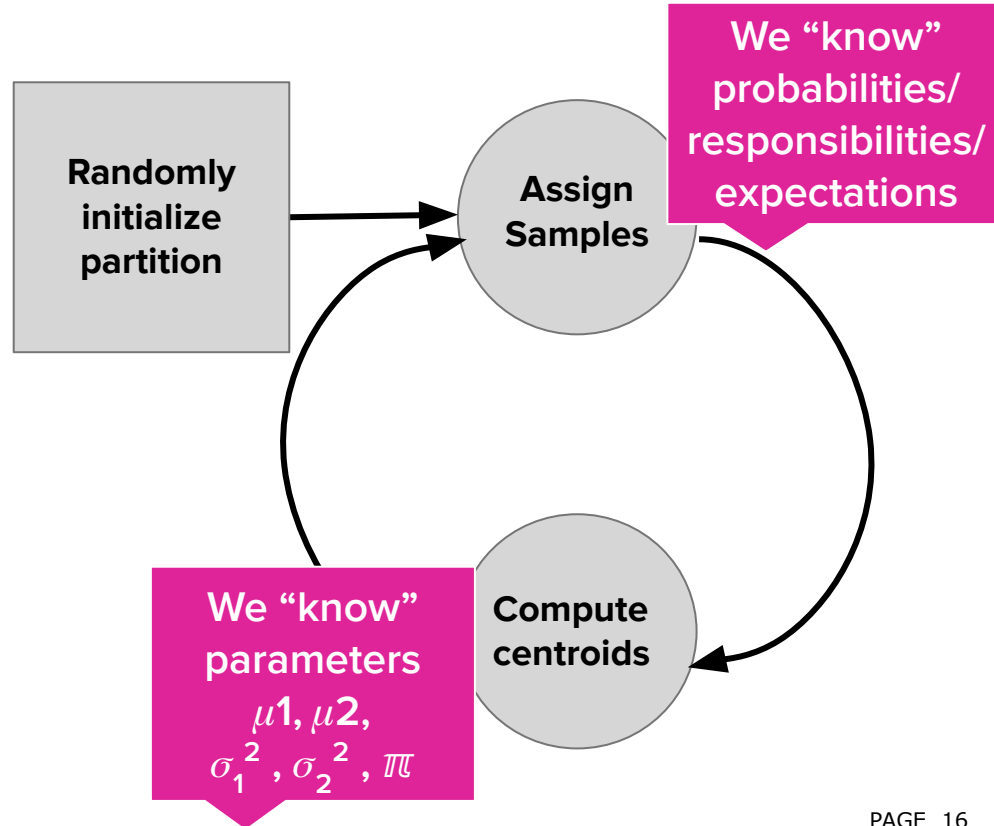
Could we combine these two somehow?



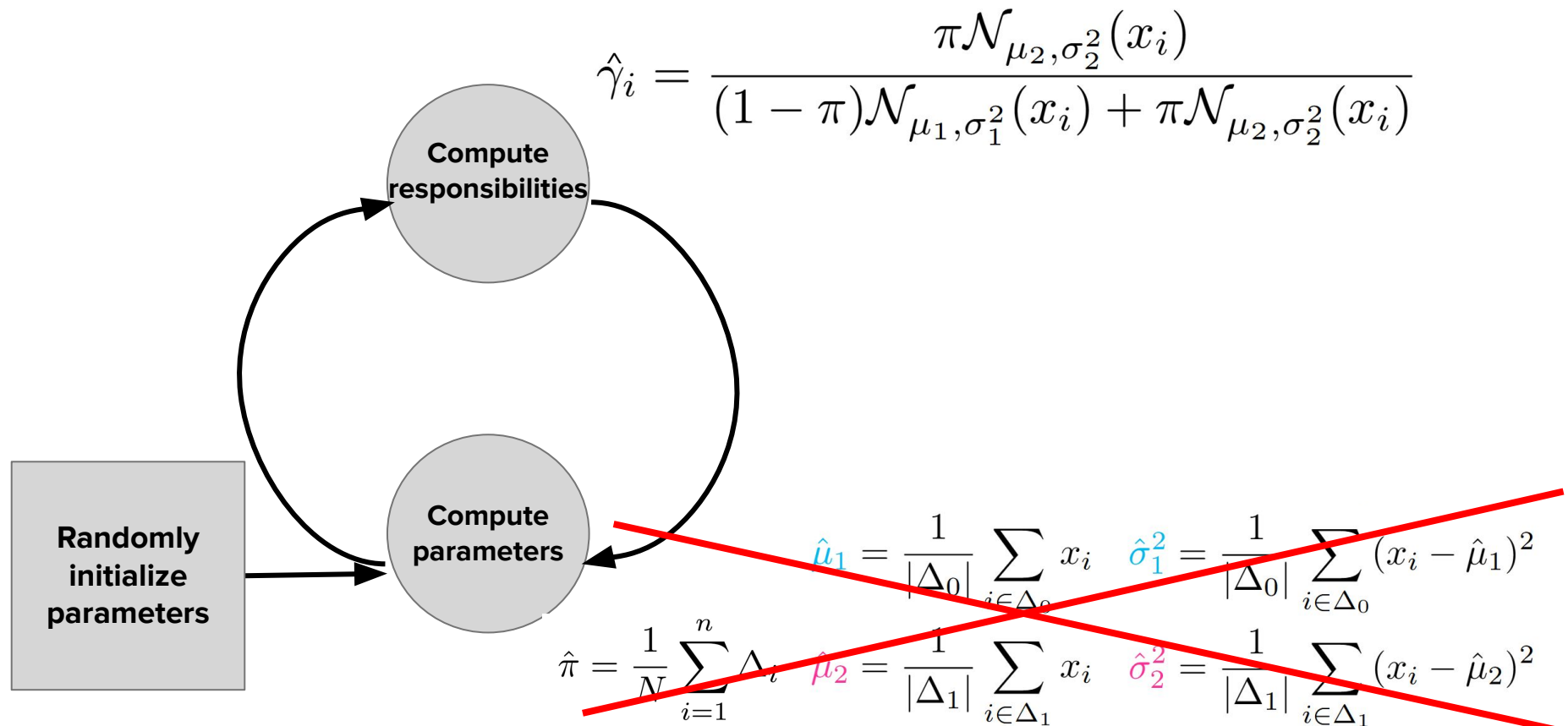
Recall Lloyd's algorithm for K-Means clustering



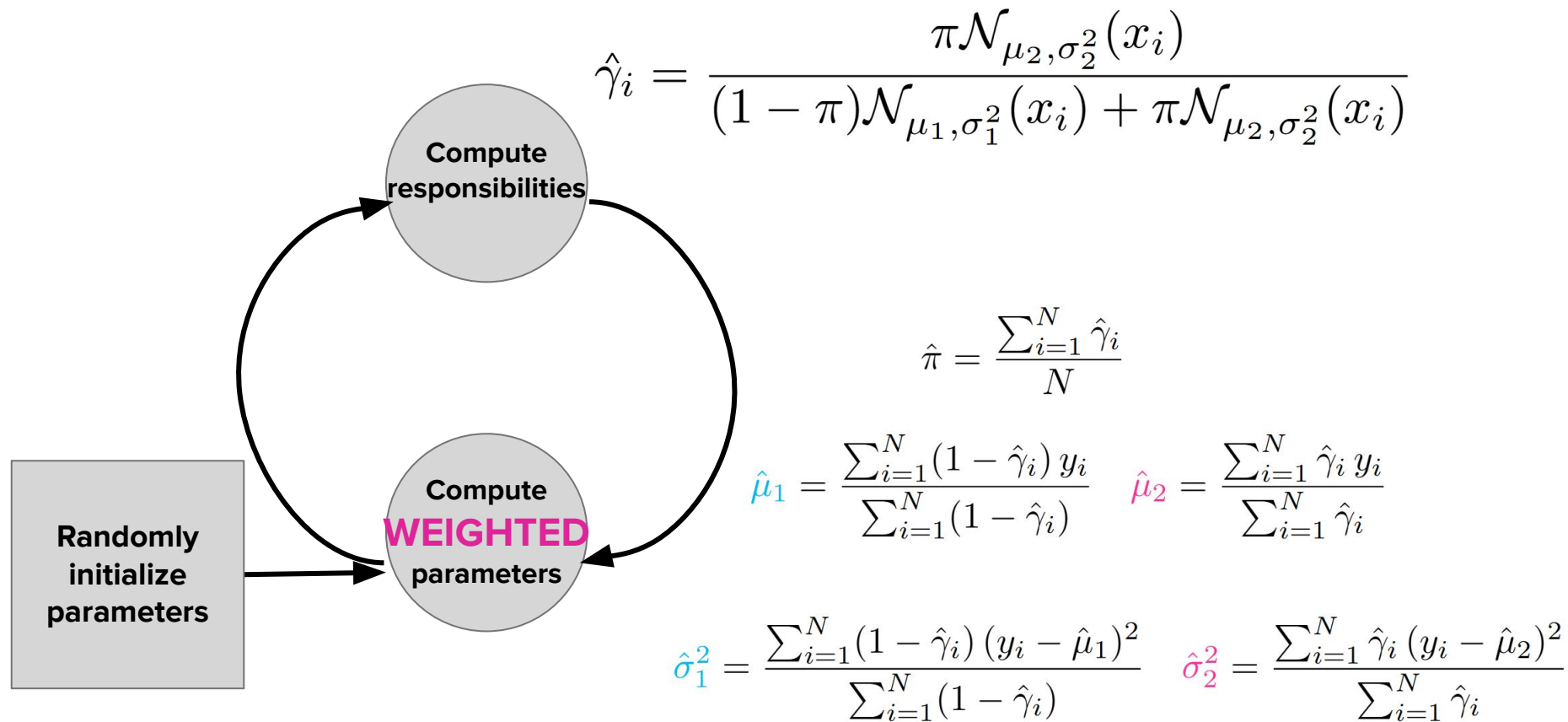
Comparing Lloyd's algorithm to GMM parameter estimation



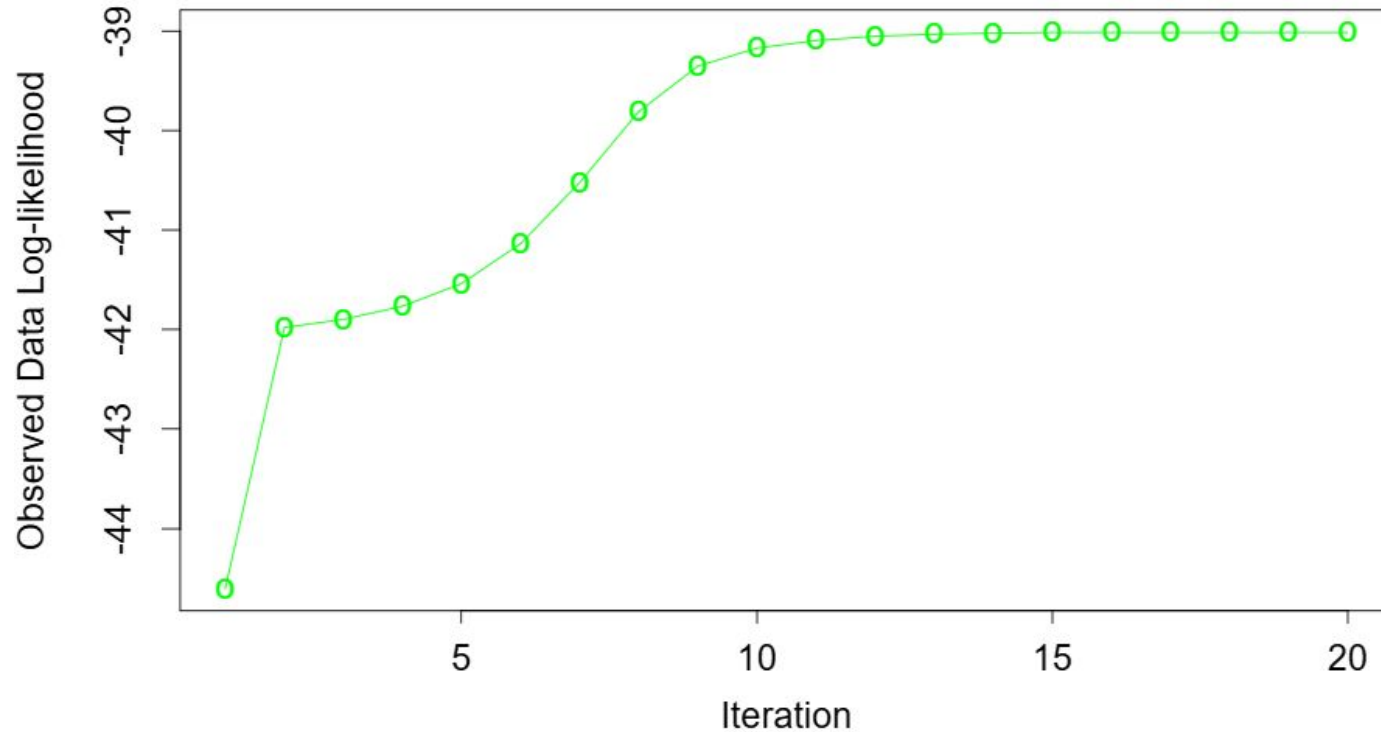
Adapting Lloyd's algorithm for GMM parameter estimation?



Adapting Lloyd's algorithm for GMM parameter estimation?



Iterative procedure convergences on the given dataset



Algorithm 8.1 *EM Algorithm for Two-component Gaussian Mixture.*

1. Take initial guesses for the parameters $\hat{\mu}_1, \hat{\sigma}_1^2, \hat{\mu}_2, \hat{\sigma}_2^2, \hat{\pi}$ (see text).
2. *Expectation Step*: compute the responsibilities

$$\hat{\gamma}_i = \frac{\hat{\pi} \phi_{\hat{\theta}_2}(y_i)}{(1 - \hat{\pi}) \phi_{\hat{\theta}_1}(y_i) + \hat{\pi} \phi_{\hat{\theta}_2}(y_i)}, \quad i = 1, 2, \dots, N. \quad (8.42)$$

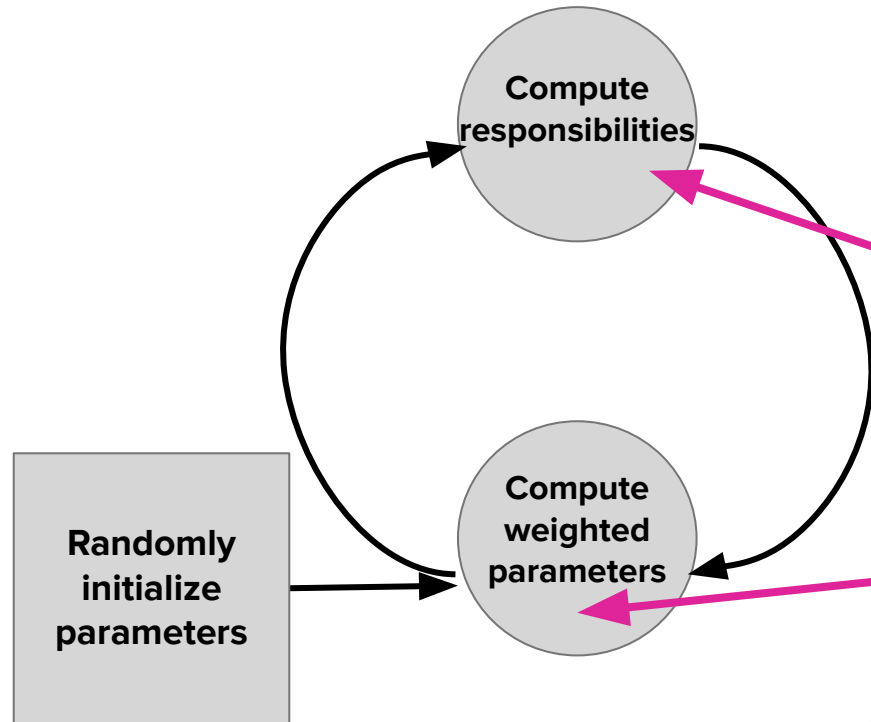
3. *Maximization Step*: compute the weighted means and variances:

$$\begin{aligned} \hat{\mu}_1 &= \frac{\sum_{i=1}^N (1 - \hat{\gamma}_i) y_i}{\sum_{i=1}^N (1 - \hat{\gamma}_i)}, & \hat{\sigma}_1^2 &= \frac{\sum_{i=1}^N (1 - \hat{\gamma}_i) (y_i - \hat{\mu}_1)^2}{\sum_{i=1}^N (1 - \hat{\gamma}_i)}, \\ \hat{\mu}_2 &= \frac{\sum_{i=1}^N \hat{\gamma}_i y_i}{\sum_{i=1}^N \hat{\gamma}_i}, & \hat{\sigma}_2^2 &= \frac{\sum_{i=1}^N \hat{\gamma}_i (y_i - \hat{\mu}_2)^2}{\sum_{i=1}^N \hat{\gamma}_i}, \end{aligned}$$

and the mixing probability $\hat{\pi} = \sum_{i=1}^N \hat{\gamma}_i / N$.

4. Iterate steps 2 and 3 until convergence.

Why is it called Expectation-Maximization (EM)?



$$\hat{\gamma}_i = \frac{\pi \mathcal{N}_{\mu_2, \sigma_2^2}(x_i)}{(1 - \pi) \mathcal{N}_{\mu_1, \sigma_1^2}(x_i) + \pi \mathcal{N}_{\mu_2, \sigma_2^2}(x_i)}$$
$$= \mathbb{E}[\Delta_i \mid \pi, \mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \mathbf{X}]$$

Compute **expectation** of latent variables
“E” step

Compute parameters that
maximize the weighted log likelihood
“M” step

Gaussian Mixture Models

The probability density for a point x is determined by the sum of densities of independent Gaussian distributions

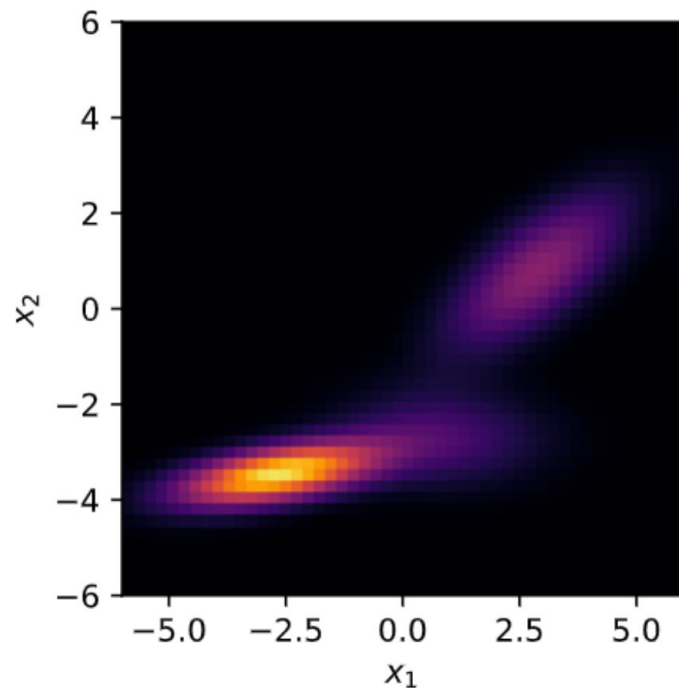
$$p(x) = \sum_{j=1}^k \pi_j \mathcal{N}(\mu_j, \Sigma_j, x)$$

Where:

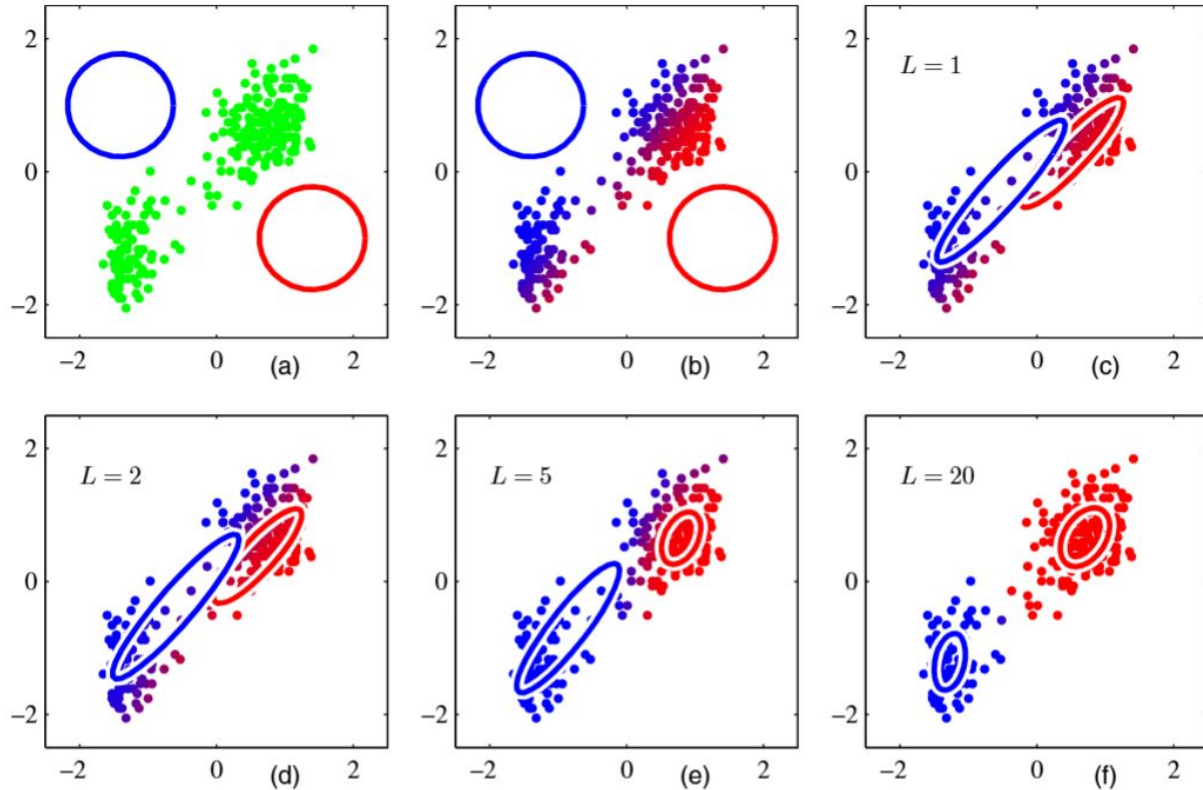
μ_j, Σ_j : mean vector and covariance matrix of j^{th} Gaussian,
for $x \in \mathbb{R}^d, d > 1$ each Gaussian is multivariate

k : number of Gaussians in the model,

π_j : mixing weight associated with the j^{th} Gaussian;
 $\pi_j \in [0, 1]$ and $\sum_{j=1}^k \pi_j = 1$



EM for mixtures of multivariate Gaussians



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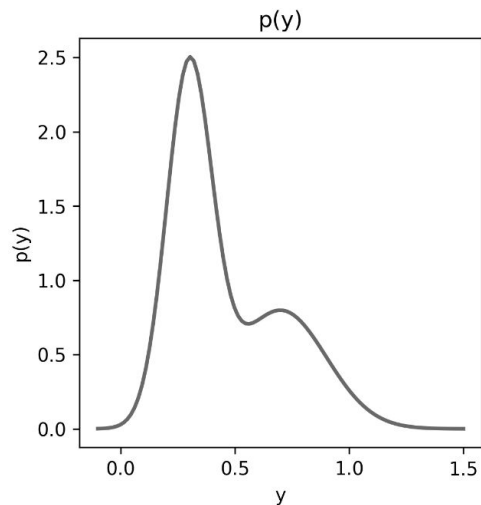
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General Expectation Maximization

$$\ell(\theta) = \sum_{n=1}^N \log p(y_n \mid \theta)$$

y_n : observed data

θ : parameters to estimate



General Expectation Maximization

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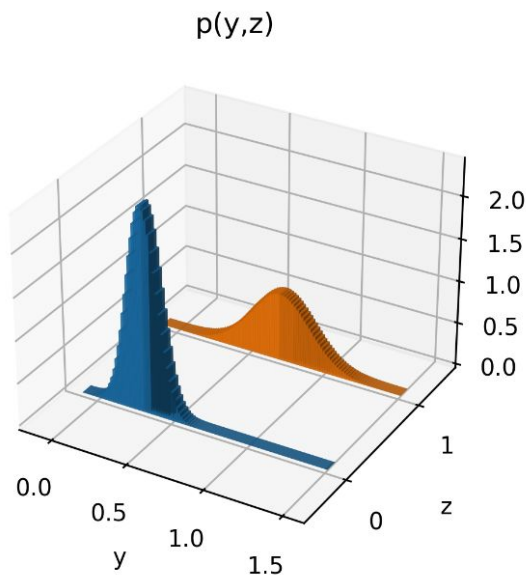
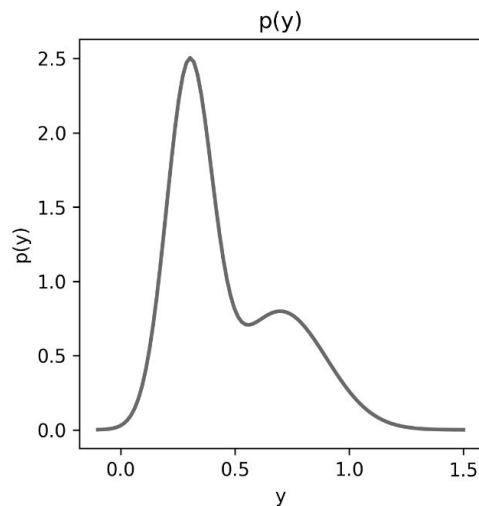
$$\ell(\theta) = \sum_{n=1}^N \log \left[\sum_{z_n} p(y_n, z_n \mid \theta) \right]$$

y_n : observed data

θ : parameters to estimate

z_n : hidden variables

$p(y_n, z_n \mid \theta)$: joint distribution of y_n and z_n



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General Expectation Maximization

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$p(y_n, z_n \mid \theta)$: joint distribution of y_n and z_n

$$\ell(\theta) = \sum_{n=1}^N \log \left[\sum_{z_n} p(y_n, z_n \mid \theta) \frac{q_n(z_n)}{q_n(z_n)} \right]$$

$$\ell(\theta) = \sum_{n=1}^N \log \left[\sum_{z_n} q_n(z_n) \frac{p(y_n, z_n \mid \theta)}{q_n(z_n)} \right]$$



General Expectation Maximization

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$$\ell(\theta) = \sum_{n=1}^N \log \left[\sum_{z_n} q_n(z_n) \frac{p(y_n, z_n | \theta)}{q_n(z_n)} \right]$$

Jensen's Inequality:

$$\log \mathbb{E}_{q_n} [Z] \geq \mathbb{E}_{q_n} [\log Z]$$

$$\ell(\theta) \geq \sum_n \sum_{z_n} q_n(z_n) \log \frac{p(y_n, z_n | \theta)}{q_n(z_n)}$$

$$\log \sum_{z_n} q_n(z_n) \frac{p(y_n, z_n | \theta)}{q_n(z_n)} \geq \sum_{z_n} q_n(z_n) \log \frac{p(y_n, z_n | \theta)}{q_n(z_n)}$$



How can we maximize $\ell(\theta)$?

$$\begin{aligned}\ell(\theta) &\geq \sum_{z_n} q_n(z_n) \log \frac{p(z_n | y_n, \theta) p(y_n | \theta)}{q_n(z_n)} \\ &\geq \sum_{z_n} q_n(z_n) \log \frac{p(z_n | y_n, \theta)}{q_n(z_n)} + \sum_{z_n} q_n(z_n) \log p(y_n | \theta) \\ &\geq -D_{\text{KL}}(q_n(z_n) \| p(z_n | y_n, \theta)) + \log p(y_n | \theta)\end{aligned}$$

Kullback-Leibler divergence

$$D_{\text{KL}}(q \| p) \triangleq \sum_z q(z) \log \frac{q(z)}{p(z)}$$

$$D_{\text{KL}}(q \| p) \geq 0$$

$$D_{\text{KL}}(q \| p) = 0 \quad \text{iff} \quad q = p$$

Select: $q_n^* = p(z_n | y_n, \theta)$

$$\implies \ell(\theta) = \sum_n \log p(y_n | \theta)$$

How can we maximize $\ell(\theta)$?

$$\ell^t(\theta) = \sum_n \log p(y_n \mid \theta)$$

$$\theta^{t+1} = \arg \max_{\theta} \sum_n \log p(y_n \mid \theta)$$

Select: $q_n^* = p(z_n | y_n, \theta)$

Expectation

$$\ell^t(\theta) \geq -D_{\text{KL}}(q_n(z_n) \| p(z_n | y_n, \theta)) + \log p(y_n | \theta)$$

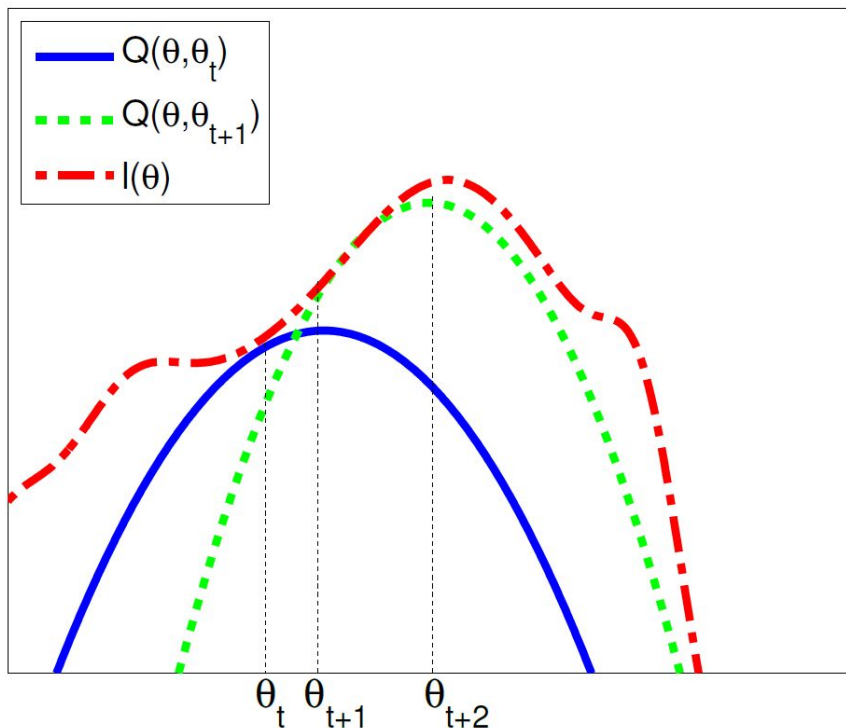
$$\ell^t(\theta) = \sum_n \log p(y_n | \theta)$$

Maximization

$$\theta^{t+1} = \arg \max_{\theta} \sum_n \log p(y_n | \theta)$$

**Initialize
parameters**

EM as bound optimization



$$\ell(\theta) \geq -D_{\text{KL}}(q_n(z_n) \parallel p(z_n \mid y_n, \theta)) + \log p(y_n \mid \theta)$$

$$\ell(\theta) \geq Q(\theta, \theta^t)$$

$$\ell(\theta^t) = Q(\theta^t, \theta^t)$$

EM as Maximization-Maximization

