



b.

```
. estimates stats m1 m2 m3 m4
```

Akaike's information criterion and Bayesian information criterion

Model	N	ll(null)	ll(model)	df	AIC	BIC
m1	233	.	-194.9392	3	395.8784	406.2315
m2	233	.	-189.4545	4	386.9091	400.7132
m3	233	.	-189.435	5	388.8701	406.1253
m4	233	.	-190.5328	4	389.0657	402.8698

AIC selects Model 2 (386.9091). BIC selects Model 2 (400.7132).

c.

```
. estimates stats m1 m2 m3 m4
```

Akaike's information criterion and Bayesian information criterion

Model	N	ll(null)	ll(model)	df	AIC	BIC
m1	233	.	507.2147	3	-1008.429	-998.0763
m2	233	.	507.8574	4	-1007.715	-993.9107
m3	233	.	508.0733	5	-1006.147	-988.8914
m4	233	.	507.9372	4	-1007.874	-994.0702

AIC selects Model 1 (-1008.429). BIC selects Model 1 (-998.0763).

d.

```
. varsoc y1 y2, maxlag(6)
```

Lag-order selection criteria

Sample: 8 thru 237

Number of obs = 230

Lag	LL	LR	df	p	FPE	AIC	HQIC	SBIC
0	284.237				.000295	-2.45423	-2.44217	-2.42434
1	321.93	75.386	4	0.000	.00022	-2.74722	-2.71104*	-2.65753*
2	328.338	12.816*	4	0.012	.000215*	-2.76815*	-2.70786	-2.61867
3	329.794	2.9129	4	0.572	.00022	-2.74604	-2.66162	-2.53676
4	330.841	2.094	4	0.718	.000226	-2.72036	-2.61182	-2.45129
5	333.997	6.3115	4	0.177	.000227	-2.71302	-2.58036	-2.38416
6	335.467	2.9398	4	0.568	.000233	-2.69102	-2.53424	-2.30236

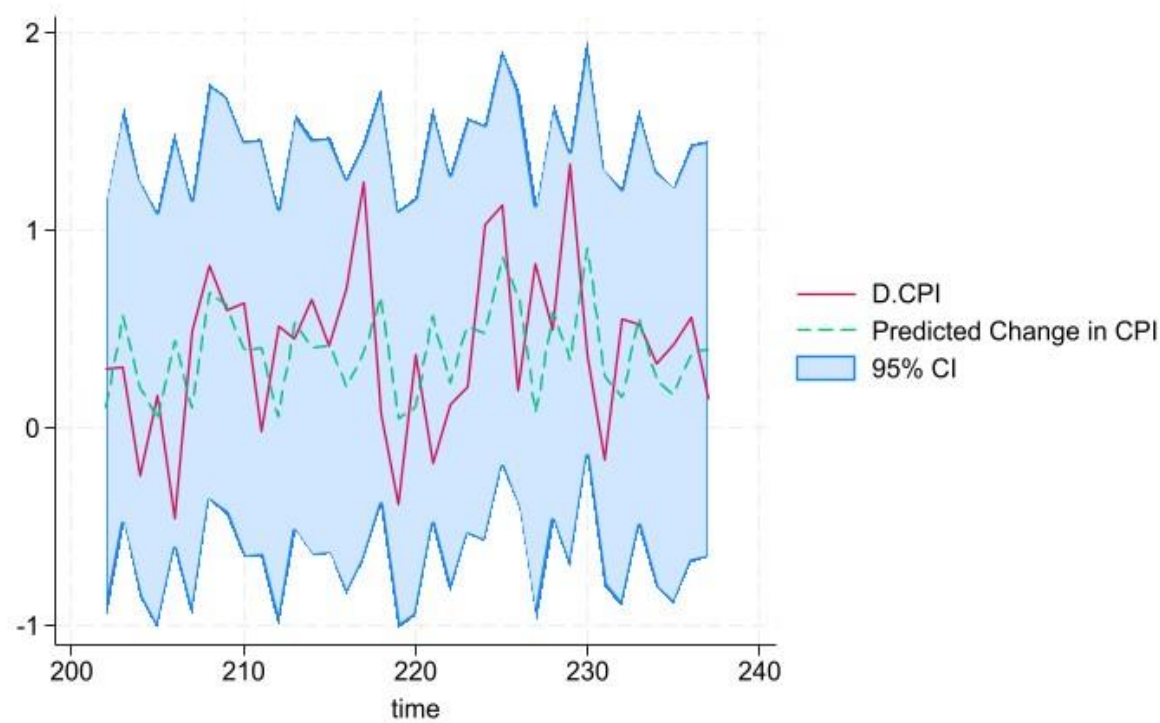
\* optimal lag

Endogenous: y1 y2

Exogenous: \_cons

AIC selects a VAR(2)

e.



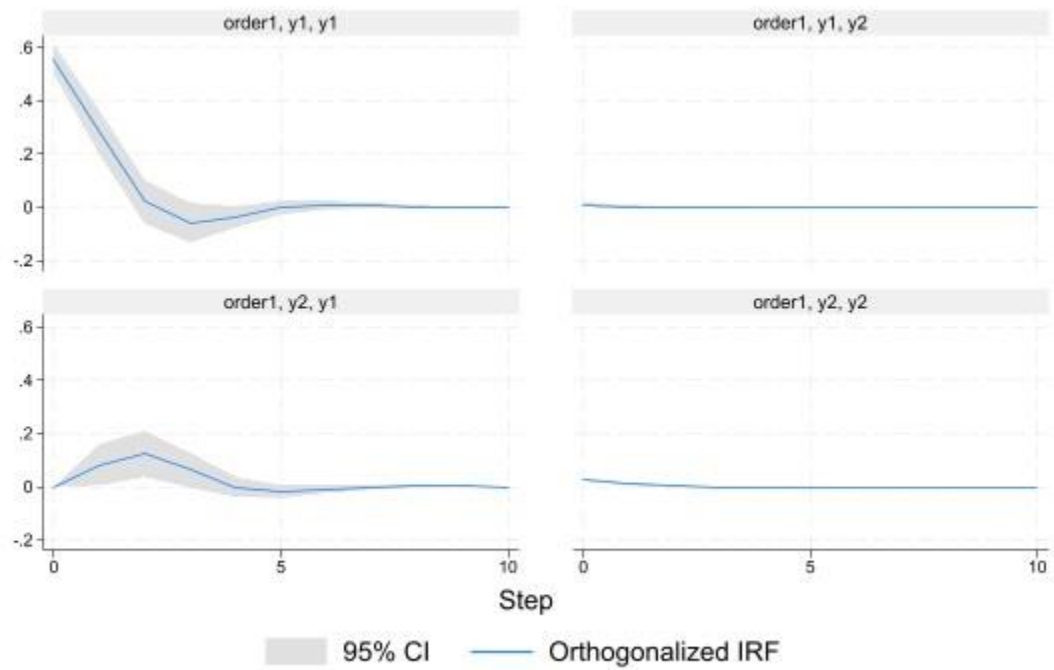
f.

. vargranger

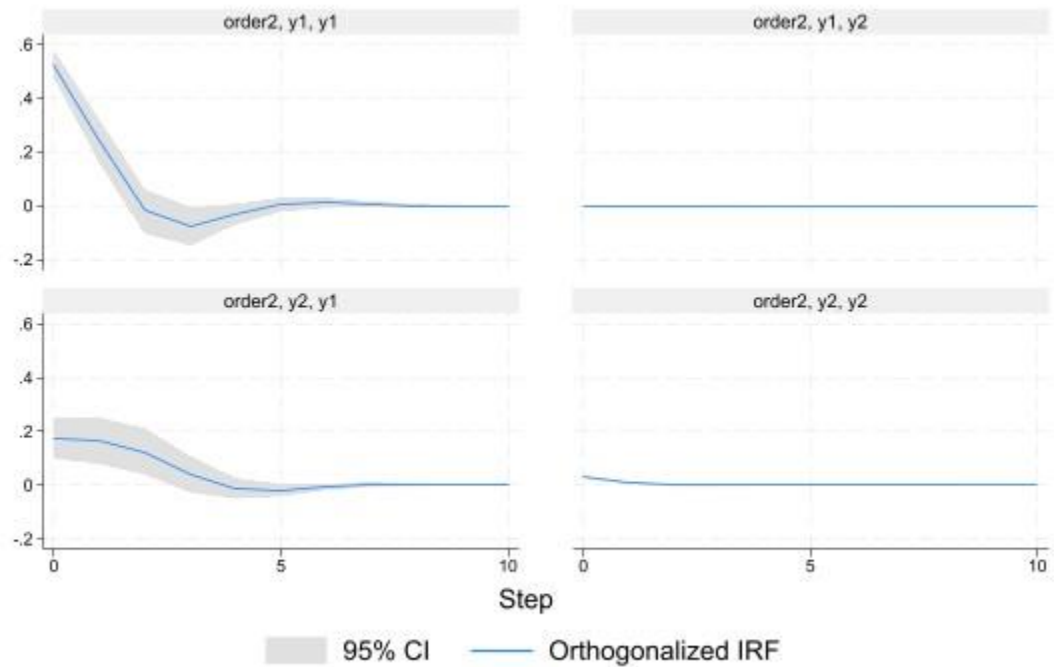
Granger causality Wald tests

Equation	Excluded	chi2	df	Prob > chi2
y1	y2	8.9601	2	0.011
y1	ALL	8.9601	2	0.011
y2	y1	.88437	2	0.643
y2	ALL	.88437	2	0.643

g.



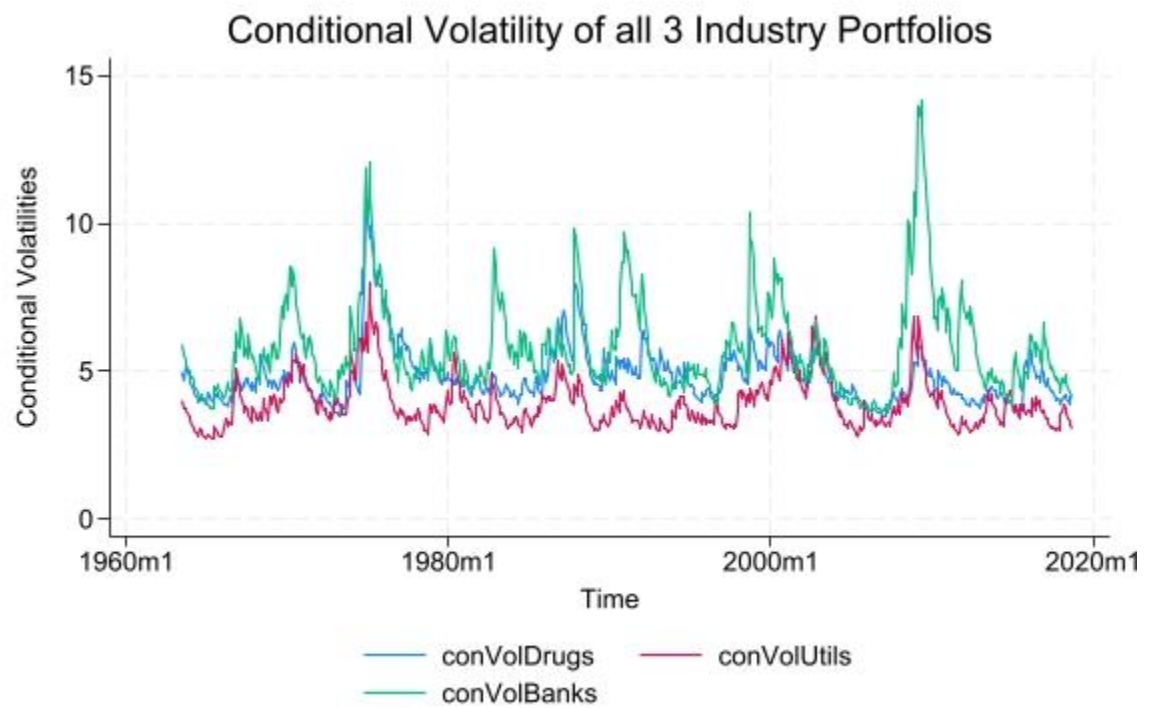
Graphs by irfname, impulse variable, and response variable



Graphs by irfname, impulse variable, and response variable

Q2.

a.



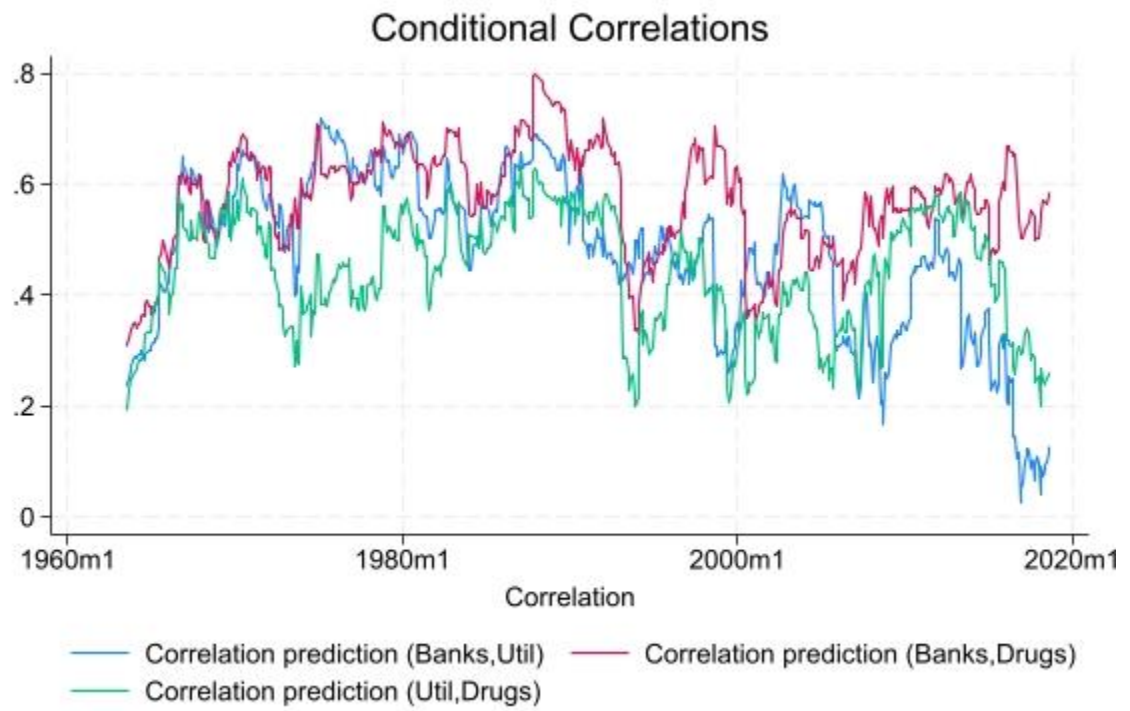
b.

Sample: 1 thru 663  
 Distribution: Gaussian  
 Log likelihood = -5615.428

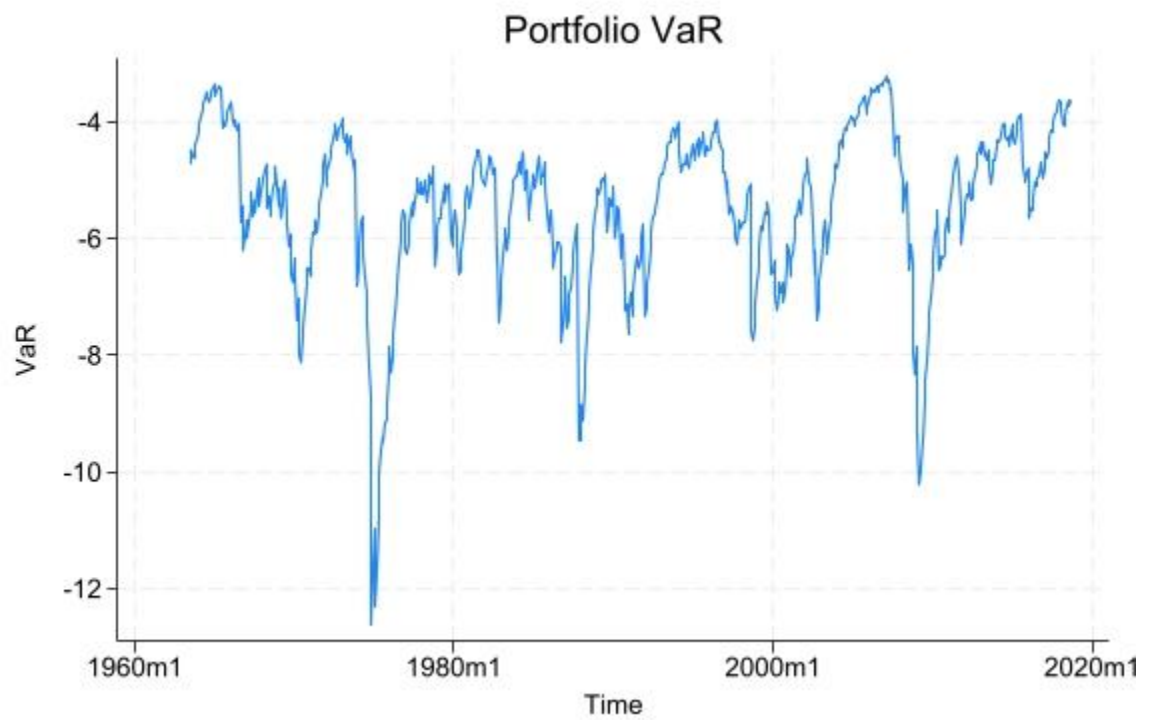
Number of obs = 663  
 Wald chi2(.) = .  
 Prob > chi2 = .

	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
Drugs						
_cons	1.231964	.1725749	7.14	0.000	.8937234	1.570205
ARCH_Drugs						
arch L1.	.0799133	.0172312	4.64	0.000	.0461409	.1136858
garch L1.	.8780659	.0263463	33.33	0.000	.826428	.9297037
_cons	1.054416	.4396656	2.40	0.016	.1926873	1.916145
Util						
_cons	.945712	.1344151	7.04	0.000	.6822633	1.209161
ARCH_Util						
arch L1.	.1047558	.0256416	4.09	0.000	.0544991	.1550125
garch L1.	.8170652	.0392431	20.82	0.000	.7401502	.8939802
_cons	1.163961	.3771213	3.09	0.002	.4248172	1.903106
Banks						
_cons	1.008646	.1952552	5.17	0.000	.6259533	1.39134
ARCH_Banks						
arch L1.	.106573	.0194248	5.49	0.000	.0685011	.1446449
garch L1.	.8363687	.0287276	29.11	0.000	.7800636	.8926738
_cons	1.895741	.6609275	2.87	0.004	.6003469	3.191135
corr(Drugs,Util)	.4154742	.0800977	5.19	0.000	.2584855	.5724629
corr(Drugs,Banks)	.5765296	.0646339	8.92	0.000	.4498495	.7032097
corr(Util,Banks)	.4340148	.0818921	5.30	0.000	.2735093	.5945204
/Adjustment						
lambda1	.0358398	.0078648	4.56	0.000	.0204249	.0512546
lambda2	.9390758	.0134624	69.76	0.000	.91269	.9654615

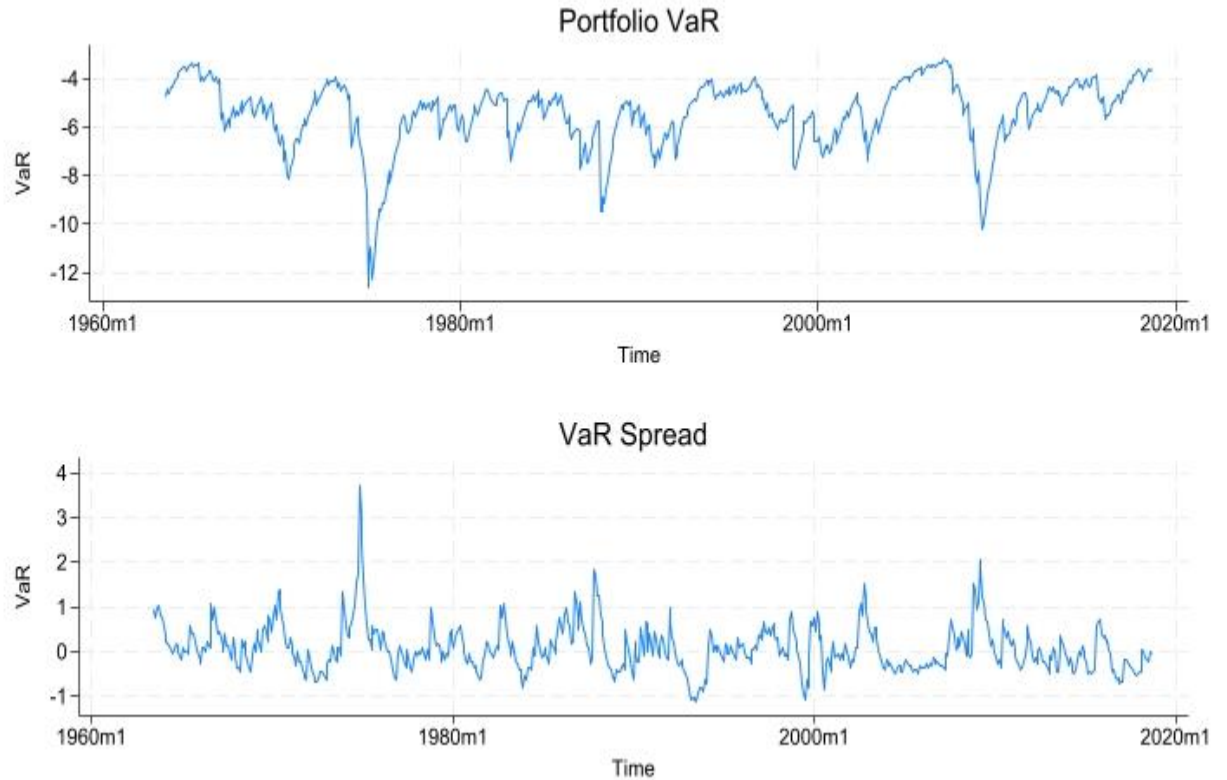
c.



d.



e.



It appears that GARCH(1,1) volatility model is more volatile than the dynamic conditional correlation model. The DCC model is a multivariate model that possesses the characteristics of GARCH models while also having a parametric model for correlations. On the other hand, GARCH is a univariate model and does not account for correlations between holdings in the portfolio. The difference in structure between the models should explain this spread.



f.

```
. summarize hit1
```

Variable	Obs	Mean	Std. dev.	Min	Max
hit1	663	.0558069	.2297218	0	1

```

-
. egen aHat1=mean(hit1)

. egen tOne1=sum(portfolioReturns-VaR < 0)

-
. gen logNum1 = (tOne1*log(0.05))+((_N-tOne1)*log(0.95))

. gen logDenom1 = (tOne1*log(aHat1))+((_N-tOne1)*log(1-aHat1))

. gen UC1=-2*(logNum1-logDenom1)

-
. gen pval1=1-chi2(1,UC1)

-
. disp UC1[_N]
.4543457

. disp pval1[_N]
.50027841

```

```
. summarize hit2
```

Variable	Obs	Mean	Std. dev.	Min	Max
hit2	663	.0603318	.2382804	0	1

```

-
. egen aHat2=mean(hit2)

. egen tOne2=sum(portfolioReturns-VaR2 < 0)

-
. gen logNum2 = (tOne2*log(0.05))+((_N-tOne2)*log(0.95))

. gen logDenom2 = (tOne2*log(aHat2))+((_N-tOne2)*log(1-aHat2))

. gen UC2=-2*(logNum2-logDenom2)

-
. gen pval2=1-chi2(1,UC2)

-
. disp UC2[_N]
1.4017334

. disp pval2[_N]
.23643357

```

There is no significant difference between both tests.

## STATA CODE

Q1.

- a. 

```
gen time=_n
tsset time

dfuller cpi, lags(2) trend regress
dfuller usdeuro, lags(2) regress
```
- b. 

```
arima cpi if time >=5, arima(1,1,0)
estimates store m1
arima cpi if time >=5, arima(2,1,0)
estimates store m2
arima cpi if time >=5, arima(3,1,0)
estimates store m3
arima cpi if time >=5, arima(1,1,1)
estimates store m4

estimates stats m1 m2 m3 m4
```
- c. 

```
arima usdeuro if time >=5, arima(1,1,0)
estimates store m1
arima usdeuro if time >=5, arima(2,1,0)
estimates store m2
arima usdeuro if time >=5, arima(3,1,0)
estimates store m3
arima usdeuro if time >=5, arima(1,1,1)
estimates store m4

estimates stats m1 m2 m3 m4
```
- d. 

```
gen y1=D.cpi
gen y2=D.usdeuro
label variable y1 "D.CPI"
label variable y2 "D.usdeuro"

varsoc y1 y2, maxlag(6)
```
- e. 

```
quietly var y1 y2 if time<=201, lags(1/2)
predict y1for, equation(#1)
label variable y1for "Predicted Change in CPI"
predict res, equation(#1) residuals
egen sdf=sd(res)
gen lcl = y1for - 1.96*sdf
gen hcl = y1for + 1.96*sdf
twoway (rarea lcl hcl time, fintensity(inten20)) ///
(line y1 y1for time, lpattern(solid dash)) in -36/1, ///
```

```
legend(order(2 3 1) label(1 "95% CI"))
```

f. quietly var y1 y2 if time<=201, lags(1/2)  
vargranger

g. irf create order1, set(myExamp) step(10) replace  
irf graph oirf

```
erase myExamp.irf
```

```
quietly var y1 y2 if time<=201, lags(1/2)
```

```
irf create order2, set(myExamp) step(10) replace  
irf graph oirf
```

```
erase myExamp.irf
```

Q2.

- a. 

```
clear all
ssc install tsmktime
import excel "\\apporto.com\dfs\CLT\Users\rkamath1_clt\Documents\industry.xlsx",
sheet("49_Industry Portfolios") firstrow
gen time=_n
tsset time

arch Drugs, garch(1) arch(1)
predict conVar1, variance
gen conVolDrugs = sqrt(conVar1)

arch Util, garch(1) arch(1)
predict conVar2, variance
gen conVolUtils = sqrt(conVar2)

arch Banks, garch(1) arch(1)
predict conVar3, variance
gen conVolBanks = sqrt(conVar3)

tsline conVolDrugs conVolUtils conVolBanks, lpattern(solid) xtitle("Time") ytitle("Conditional
Volatilities") title("Conditional Volatility of all 3 Industry Portfolios")
```
- b. 

```
mgarch dcc (Drug Util Banks), arch(1) garch(1)
```
- c. 

```
predict rho*, correlation
tsline rho_Banks_Util rho_Banks_Drugs rho_Util_Drugs, lpattern(solid) xtitle("Correlation")
title("Conditional Correlations")
```
- d. 

```
predict m*, xb
predict s*, variance
gen m_Hat = (m_Drugs+m_Banks+m_Util)/3
gen v_Hat = (s_Drugs_Drugs+s_Util_Util+s_Banks_Banks +
2*(s_Util_Drugs+s_Banks_Drugs+s_Banks_Util))/9
gen VaR = (m_Hat - (1.645 * sqrt(v_Hat)))
tsline VaR, xtitle("Time") ytitle("VaR") title("Portfolio VaR") name(VAR1)
```
- e. 

```
gen portfolioReturns = (Drugs + Util + Banks)/3
arch portfolioReturns, garch(1) arch(1)
predict me, xb
predict conVar, variance
gen VaR2 = (me - (1.645 * sqrt(conVar)))
tsline VaR2, pattern(solid) xtitle("Time") ytitle("Conditional Volatility") title("VaR from GARCH")
name(VAR2)
graph combine VAR1 VAR2, row(2) col(1)
```

```

gen VaR_spread = VaR - VaR2
tsline VaR_spread, lpattern(solid) xtitle ("Time") ytitle("VaR") title("VaR Spread") name
(VARspread)
graph combine VAR1 VARspread, row(2) col(1)

```

- f. generate hit1 = (portfolioReturns-VaR < 0)  
summarize hit1

```

egen aHat1=mean (hit1)
egen tOne1=sum(portfolioReturns-VaR < 0)

gen LogNum1 = (tOne1*log (0.05))+((_N-tOne1)*log(0.95))
gen logDenom1 = (tOne1*log(aHat1))+((_N-tOne1)*log(1-aHat1))
gen UC1=-2* (logNum1-logDenom1)

gen pval1=1-chi2(1, UC1)

disp UC1[_N]
disp pval1[_N]

```

```

generate hit2 = (portfolioReturns-VaR2 < 0)
summarize hit2

```

```

egen aHat2=mean (hit2)
egen tOne2=sum(portfolioReturns-VaR2 < 0)

gen logNum2 = (tOne2*log(0.05))+((_N-tOne2)*log(0.95))
gen logDenom2 = (tOne2*log(aHat2))+((_N-tOne2)*log(1-aHat2))
gen UC2=-2* (logNum2-logDenom2)

gen pval2=1-chi2(1,UC2)

disp UC2[_N]
disp pval2[_N]

```