# All Cheat Sheets

Machine Learning, Deep Learning, Artificial Intelligence

BY
STANFORD UNIVERSITY
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**Technology** 

## Probability—the Science of Uncertainty and Data

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#### **PROBABILITY**

#### Probability models and axioms

Definition (Sample space) A sample space  $\Omega$  is the set of all possible outcomes. The set's elements must be mutually exclusive, collectively exhaustive and at the right granularity.

Definition (Event) An event is a subset of the sample space. Probability is assigned to events.

Definition (Probability axioms) A probability law  $\mathbb{P}$  assigns probabilities to events and satisfies the following axioms:

**Nonnegativity**  $\mathbb{P}(A) \geq 0$  for all events A.

**Normalization**  $\mathbb{P}(\Omega) = 1$ .

(Countable) additivity For every sequence of events  $A_1, A_2, \ldots$  such that  $A_i \cap A_j = \emptyset$ :  $\mathbb{P}\left(\bigcup_i A_i\right) = \sum_i \mathbb{P}(A_i)$ .

Corollaries (Consequences of the axioms)

- $\mathbb{P}(\emptyset) = 0$ .
- For any finite collection of disjoint events  $A_1, \ldots, A_n$ ,  $\mathbb{P}\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n \mathbb{P}(A_i).$
- $\mathbb{P}(A) + \mathbb{P}(A^c) = 1$ .
- $\mathbb{P}(A) \leq 1$ .
- If  $A \subset B$ , then  $\mathbb{P}(A) \leq \mathbb{P}(B)$ .
- $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) \mathbb{P}(A \cap B)$ .
- $\mathbb{P}(A \cup B) \leq \mathbb{P}(A) + \mathbb{P}(B)$ .

Example (Discrete uniform law) Assume  $\Omega$  is finite and consists of n equally likely elements. Also, assume that  $A \subset \Omega$  with k elements. Then  $\mathbb{P}(A) = \frac{k}{n}$ .

#### Conditioning and Bayes' rule

Definition (Conditional probability) Given that event B has occurred and that P(B) > 0, the probability that A occurs is

$$\mathbb{P}(A|B) \stackrel{\triangle}{=} \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}.$$

Remark (Conditional probabilities properties) They are the same as ordinary probabilities. Assuming  $\mathbb{P}(B) > 0$ :

- $\mathbb{P}(A|B) \geq 0$ .
- $\mathbb{P}(\Omega|B) = 1$
- $\mathbb{P}(B|B) = 1$ .
- If  $A \cap C = \emptyset$ ,  $\mathbb{P}(A \cup C|B) = \mathbb{P}(A|B) + \mathbb{P}(C|B)$ .

Proposition (Multiplication rule)

 $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_n) = \mathbb{P}(A_1) \cdot \mathbb{P}(A_2 | A_1) \cdots \mathbb{P}(A_n | A_1 \cap A_2 \cap \cdots \cap A_{n-1}).$ 

Theorem (Total probability theorem) Given a partition  $\{A_1, A_2, \ldots\}$  of the sample space, meaning that  $\bigcup_i A_i = \Omega$  and the events are disjoint, and for every event B, we have

$$\mathbb{P}(B) = \sum_{i} \mathbb{P}(A_i) \mathbb{P}(B|A_i)$$

Theorem (Bayes' rule) Given a partition  $\{A_1, A_2, \ldots\}$  of the sample space, meaning that  $\bigcup A_i = \Omega$  and the events are disjoint,

and if  $\mathbb{P}(A_i) > 0$  for all i, then for every event B, the conditional probabilities  $\mathbb{P}(A_i|B)$  can be obtained from the conditional probabilities  $\mathbb{P}(B|A_i)$  and the initial probabilities  $\mathbb{P}(A_i)$  as follows:

$$\mathbb{P}(A_i|B) = \frac{\mathbb{P}(A_i)\mathbb{P}(B|A_i)}{\sum_j \mathbb{P}(A_j)\mathbb{P}(B|A_j)}.$$

#### Independence

Definition (Independence of events) Two events are independent if occurrence of one provides no information about the other. We say that A and B are independent if

$$\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B).$$

Equivalently, as long as  $\mathbb{P}(A) > 0$  and  $\mathbb{P}(B) > 0$ ,

$$\mathbb{P}(B|A) = \mathbb{P}(B)$$
  $\mathbb{P}(A|B) = \mathbb{P}(A)$ 

#### Remarks

- The definition of independence is symmetric with respect to A and B
- The product definition applies even if  $\mathbb{P}(A) = 0$  or  $\mathbb{P}(B) = 0$ .

Corollary If A and B are independent, then A and  $B^c$  are independent. Similarly for  $A^c$  and B, or for  $A^c$  and  $B^c$ .

Definition (Conditional independence) We say that A and B are independent conditioned on C, where  $\mathbb{P}(C) > 0$ , if

$$\mathbb{P}(A \cap B|C) = \mathbb{P}(A|C)\mathbb{P}(B|C).$$

Definition (Independence of a collection of events) We say that events  $A_1, A_2, \ldots, A_n$  are independent if for every collection of distinct indices  $i_1, i_2, \ldots, i_k$ , we have

$$\mathbb{P}(A_{i_1} \cap \ldots \cap A_{i_k}) = \mathbb{P}(A_{i_1}) \cdot \mathbb{P}(A_{i_2}) \cdots \mathbb{P}(A_{i_k}).$$

#### Counting

This section deals with finite sets with uniform probability law. In this case, to calculate  $\mathbb{P}(A)$ , we need to count the number of elements in A and in  $\Omega$ .

Remark (Basic counting principle) For a selection that can be done in r stages, with  $n_i$  choices at each stage i, the number of possible selections is  $n_1 \cdot n_2 \cdots n_r$ .

Definition (Permutations) The number of permutations (orderings) of n different elements is

$$n! = 1 \cdot 2 \cdot 3 \cdots n$$
.

Definition (Combinations) Given a set of n elements, the number of subsets with exactly k elements is

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}.$$

Definition (Partitions) We are given an n-element set and nonnegative integers  $n_1, n_2, \ldots, n_r$ , whose sum is equal to n. The number of partitions of the set into r disjoint subsets, with the i<sup>th</sup> subset containing exactly  $n_i$  elements, is equal to

$$\binom{n}{n_1,\ldots,n_r} = \frac{n!}{n_1!n_2!\cdots n_r!}.$$

Remark This is the same as counting how to assign n distinct elements to r people, giving each person i exactly  $n_i$  elements.

#### Discrete random variables

Probability mass function and expectation

Definition (Random variable) A random variable X is a function of the sample space  $\Omega$  into the real numbers (or  $\mathbb{R}^n$ ). Its range can be discrete or continuous.

Definition (Probability mass funtion (PMF)) The probability law of a discrete random variable X is called its PMF. It is defined as

$$p_X(x) = \mathbb{P}(X = x) = \mathbb{P}\left(\left\{\omega \in \Omega : X(\omega) = x\right\}\right).$$

Properties

 $p_X(x) \ge 0, \ \forall \ x.$ 

 $\sum_{x} p_X(x) = 1.$ 

Example (Bernoulli random variable) A Bernoulli random variable X with parameter  $0 \le p \le 1$  ( $X \sim \text{Ber}(p)$ ) takes the following values:

$$X = \begin{cases} 1 & \text{w.p. } p, \\ 0 & \text{w.p. } 1 - p. \end{cases}$$

An indicator random variable of an event ( $I_A$  = 1 if A occurs) is an example of a Bernoulli random variable.

Example (Discrete uniform random variable) A Discrete uniform random variable X between a and b with  $a \le b$  ( $X \sim \mathrm{Uni}[a,b]$ ) takes any of the values in  $\{a,a+1,\ldots,b\}$  with probability  $\frac{1}{b-a+1}$ .

Example (Binomial random variable) A Binomial random variable X with parameters n (natural number) and  $0 \le p \le 1$   $(X \sim \operatorname{Bin}(n,p))$  takes values in the set  $\{0,1,\ldots,n\}$  with probabilities  $p_X(i) = \binom{n}{i} p^i (1-p)^{n-i}$ .

It represents the number of successes in n independent trials where each trial has a probability of success p. Therefore, it can also be seen as the sum of n independent Bernoulli random variables, each with parameter p.

Example (Geometric random variable) A Geometric random variable X with parameter  $0 \le p \le 1$  ( $X \sim \text{Geo}(p)$ ) takes values in the set  $\{1,2,\ldots\}$  with probabilities  $p_X(i) = (1-p)^{i-1}p$ . It represents the number of independent trials until (and including)

the first success, when the probability of success in each trial is p.

Definition (Expectation/mean of a random variable) The expectation of a discrete random variable is defined as

$$\mathbb{E}[X] \stackrel{\triangle}{=} \sum_{x} x p_X(x).$$

assuming  $\sum_{x} |x| p_X(x) < \infty$ .

Properties (Properties of expectation)

- If  $X \ge 0$  then  $\mathbb{E}[X] \ge 0$ .
- If a < X < b then  $a < \mathbb{E}[X] < b$ .
- If X = c then  $\mathbb{E}[X] = c$ .

Example Expected value of know r.v.

- If  $X \sim \text{Ber}(p)$  then  $\mathbb{E}[X] = p$ .
- If  $X = I_A$  then  $\mathbb{E}[X] = \mathbb{P}(A)$ .
- If  $X \sim \text{Uni}[a, b]$  then  $\mathbb{E}[X] = \frac{a+b}{2}$ .
- If  $X \sim \text{Bin}(n, p)$  then  $\mathbb{E}[X] = np$ .
- If  $X \sim \text{Geo}(p)$  then  $\mathbb{E}[X] = \frac{1}{p}$ .

Theorem (Expected value rule) Given a random variable X and a Properties (Properties of joint PMF) function  $q: \mathbb{R} \to \mathbb{R}$ , we construct the random variable Y = q(X). Then

$$\sum_{y} y p_{Y}(y) = \mathbb{E}[Y] = \mathbb{E}[g(X)] = \sum_{x} g(x) p_{X}(x).$$

Remark (PMF of Y = q(X)) The PMF of Y = q(X) is  $p_Y(y) = \sum_{x:g(x)=y} p_X(x).$ 

Remark In general  $g(\mathbb{E}[X]) \neq \mathbb{E}[g(X)]$ . They are equal if a(x) = ax + b.

Variance, conditioning on an event, multiple r.v.

Definition (Variance of a random variable) Given a random variable X with  $\mu = \mathbb{E}[X]$ , its variance is a measure of the spread of the random variable and is defined as

$$\operatorname{Var}(X) \stackrel{\triangle}{=} \mathbb{E}\left[(X - \mu)^2\right] = \sum_{x} (x - \mu)^2 p_X(x).$$

$$\sigma_X = \sqrt{\operatorname{Var}(X)}$$
.

Properties (Properties of the variance)

- $Var(aX) = a^2 Var(X)$ , for all  $a \in \mathbb{R}$ .
- Var(X + b) = Var(X), for all  $b \in \mathbb{R}$ .
- $\operatorname{Var}(aX + b) = a^2 \operatorname{Var}(X)$ .
- $\operatorname{Var}(X) = \mathbb{E}[X^2] (\mathbb{E}[X])^2$ .

Example (Variance of known r.v.)

- If  $X \sim \text{Ber}(p)$ , then Var(X) = p(1-p).
- If  $X \sim \text{Uni}[a,b]$ , then  $\text{Var}(X) = \frac{(b-a)(b-a+2)}{12}$ .
- If  $X \sim \text{Bin}(n, p)$ , then Var(X) = np(1 p).
- If  $X \sim \text{Geo}(p)$ , then  $\text{Var}(X) = \frac{1-p}{2}$

Proposition (Conditional PMF and expectation, given an event) Given the event A, with  $\mathbb{P}(A) > 0$ , we have the following

- $p_{X|A}(x) = \mathbb{P}(X = x|A)$ .
- If A is a subset of the range of X, then:  $p_{X|A}(x) \stackrel{\triangle}{=} p_{X|\{X \in A\}}(x) = \begin{cases} \frac{1}{P(A)} p_X(x), & \text{if } x \in A, \\ 0, & \text{otherwise.} \end{cases}$
- $\sum_{x} p_{X|A}(x) = 1$ .
- $\mathbb{E}[X|A] = \sum_{x} x p_{X|A}(x)$ .
- $\mathbb{E}[g(X)|A] = \sum_{x} g(x) p_{X|A}(x)$ .

Proposition (Total expectation rule) Given a partition of disjoint events  $A_1, \ldots, A_n$  such that  $\sum_i \mathbb{P}(A_i) = 1$ , and  $\mathbb{P}(A_i) > 0$ ,

$$\mathbb{E}[X] = \mathbb{P}(A_1)\mathbb{E}[X|A_1] + \dots + \mathbb{P}(A_n)\mathbb{E}[X|A_n].$$

Definition (Memorylessness of the geometric random variable)

When we condition a geometric random variable X on the event X > n we have memorylessness, meaning that the "remaining time" X-n, given that X>n, is also geometric with the same parameter. Formally,

$$p_{X-n|X>n}(i) = p_X(i).$$

Definition (Joint PMF) The joint PMF of random variables  $X_1, X_2, ..., X_n$  is  $p_{X_1,X_2,...,X_n}(x_1,...,x_n) = \mathbb{P}(X_1 = x_1,...,X_n = x_n).$ 

- $\bullet \sum_{x_1} \cdots \sum_{x_n} p_{X_1, \dots, X_n} (x_1, \dots, x_n) = 1.$
- $p_{X_1}(x_1) = \sum_{x_2} \cdots \sum_{x_n} p_{X_1,...,X_n}(x_1,x_2,...,x_n).$
- $p_{X_2,...,X_n}(x_2,...,x_n) = \sum p_{X_1,X_2,...,X_n}(x_1,x_2,...,x_n)$ .

Definition (Functions of multiple r.v.) If  $Z = g(X_1, \ldots, X_n)$ , where  $g: \mathbb{R}^n \to \mathbb{R}$ , then  $p_Z(z) = \mathbb{P}(g(X_1, \dots, X_n) = z)$ .

Proposition (Expected value rule for multiple r.v.) Given

$$\mathbb{E}\left[g(X_1,\ldots,X_n)\right] = \sum_{x_1,\ldots,x_n} g(x_1,\ldots,x_n) p_{X_1,\ldots,X_n}(x_1,\ldots,x_n).$$

Properties (Linearity of expectations)

- $\mathbb{E}[aX + b] = a\mathbb{E}[X] + b$ .
- $\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$

Conditioning on a random variable, independence

Definition (Conditional PMF given another random variable)

Given discrete random variables X, Y and y such that  $p_Y(y) > 0$ we define

$$p_{X|Y}(x|y) \stackrel{\triangle}{=} \frac{p_{X,Y}(x,y)}{p_Y(y)}.$$

Proposition (Multiplication rule) Given jointly discrete random variables X, Y, and whenever the conditional probabilities are defined,

$$p_{X,Y}(x,y) = p_X(x)p_{Y|X}(y|x) = p_Y(y)p_{X|Y}(x|y).$$

Definition (Conditional expectation) Given discrete random variables X, Y and y such that  $p_Y(y) > 0$  we define

$$\mathbb{E}[X|Y = y] = \sum_{x} x p_{X|Y}(x|y)$$

Additionally we have

$$\mathbb{E}\left[g(X)|Y=y\right] = \sum_{x} g(x) p_{X|Y}(x|y).$$

Theorem (Total probability and expectation theorems) If  $p_Y(y) > 0$ , then

$$p_X(x) = \sum_{y} p_Y(y) p_{X|Y}(x|y),$$

$$\mathbb{E}[X] = \sum_{y} p_{Y}(y) \mathbb{E}[X|Y = y].$$

Definition (Independence of a random variable and an event) A discrete random variable X and an event A are independent if  $\mathbb{P}(X = x \text{ and } A) = p_X(x)\mathbb{P}(A), \text{ for all } x.$ 

Definition (Independence of two random variables) Two discrete random variables X and Y are independent if  $p_{X,Y}(x,y) = p_X(x)p_Y(y)$  for all x,y.

Remark (Independence of a collection of random variables) A collection  $X_1, X_2, \dots, X_n$  of random variables are independent if

$$p_{X_1,...,X_n}(x_1,...,x_n) = p_{X_1}(x_1) \cdots p_{X_n}(x_n), \forall x_1,...,x_n.$$

Remark (Independence and expectation) In general,  $\mathbb{E}[g(X,Y)] \neq g(\mathbb{E}[X],\mathbb{E}[Y])$ . An exception is for linear functions:  $\mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y].$ 

Proposition (Expectation of product of independent r.v.) If X and Y are discrete independent random variables,

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

Remark If X and Y are independent,  $\mathbb{E}[g(X)h(Y)] = \mathbb{E}[g(X)]\mathbb{E}[h(Y)].$ 

Proposition (Variance of sum of independent random variables) IF X and Y are discrete independent random variables.

$$Var(X + Y) = Var(X) + Var(Y).$$

#### Continuous random variables

PDF. Expectation. Variance. CDF

Definition (Probability density function (PDF)) A probability density function of a r.v. X is a non-negative real valued function  $f_X$  that satisfies the following

- $\bullet \int_{-\infty}^{\infty} f_X(x) dx = 1.$
- $\mathbb{P}(a \le X \le b) = \int_{a}^{b} f_X(x) dx$  for some random variable X.

Definition (Continuous random variable) A random variable X is continuous if its probability law can be described by a PDF  $f_X$ . Remark Continuous random variables satisfy:

- For small  $\delta > 0$ ,  $\mathbb{P}(a \le X \le a + \delta) \approx f_X(a)\delta$ .
- $\mathbb{P}(X = a) = 0, \forall a \in \mathbb{R}.$

Definition (Expectation of a continuous random variable) The expectation of a continuous random variable is

$$\mathbb{E}[X] \stackrel{\triangle}{=} \int_{-\infty}^{\infty} x f_X(x) \mathrm{d}x.$$

assuming  $\int_{0}^{\infty} |x| f_X(x) dx < \infty$ .

Properties (Properties of expectation)

- If  $X \ge 0$  then  $\mathbb{E}[X] \ge 0$ .
- If  $a \le X \le b$  then  $a \le \mathbb{E}[X] \le b$ .
- $\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx$ .
- $\mathbb{E}[aX + b] = a\mathbb{E}[X] + b$ .

Definition (Variance of a continuous random variable) Given a continuous random variable X with  $\mu = \mathbb{E}[X]$ , its variance is

$$\operatorname{Var}(X) = \mathbb{E}\left[(X - \mu)^2\right] = \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) dx.$$

It has the same properties as the variance of a discrete random variable.

Example (Uniform continuous random variable) A Uniform continuous random variable X between a and b, with a < b,  $(X \sim \text{Uni}(a,b))$  has PDF

$$f_X(x) = \begin{cases} \frac{1}{b-a}, & \text{if } a < x < b, \\ 0, & \text{otherwise.} \end{cases}$$

We have  $\mathbb{E}[X] = \frac{a+b}{2}$  and  $\operatorname{Var}(X) = \frac{(b-a)^2}{12}$ .

Example (Exponential random variable) An Exponential random variable X with parameter  $\lambda > 0$  ( $X \sim Exp(\lambda)$ ) has PDF

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \ge 0, \\ 0, & \text{otherwise.} \end{cases}$$

We have  $E[X] = \frac{1}{\lambda}$  and  $Var(X) = \frac{1}{\lambda^2}$ .

Definition (Cumulative Distribution Function (CDF)) The CDF of a random variable X is  $F_X(x) = \mathbb{P}(X \le x)$ .

In particular, for a continuous random variable, we have

$$F_X(x) = \int_{-\infty}^{x} f_X(x) dx,$$
$$f_X(x) = \frac{dF_X(x)}{dx}.$$

Properties (Properties of CDF)

- If  $y \ge x$ , then  $F_X(y) \ge F_X(x)$ .
- $\bullet \lim_{x \to -\infty} F_X(x) = 0.$
- $\lim_{x \to \infty} F_X(x) = 1$ .

Definition (Normal/Gaussian random variable) A Normal random variable X with mean  $\mu$  and variance  $\sigma^2 > 0$  ( $X \sim \mathcal{N}(\mu, \sigma^2)$ ) has PDF

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}.$$

We have  $E[X] = \mu$  and  $Var(X) = \sigma^2$ .

Remark (Standard Normal) The standard Normal is  $\mathcal{N}(0,1)$ .

Proposition (Linearity of Gaussians) Given  $X \sim \mathcal{N}(\mu, \sigma^2)$ , and if  $a \neq 0$ , then  $aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$ .

Using this  $Y = \frac{X - \mu}{\sigma}$  is a standard gaussian.

Conditioning on an event, and multiple continuous r.v.

Definition (Conditional PDF given an event) Given a continuous random variable X and event A with P(A) > 0, we define the conditional PDF as the function that satisfies

$$\mathbb{P}(X \in B|A) = \int_B f_{X|A}(x) dx.$$

Definition (Conditional PDF given  $X \in A$ ) Given a continuous random variable X and an  $A \subset \mathbb{R}$ , with P(A) > 0:

$$f_{X|X\in A}(x) = \begin{cases} \frac{1}{\mathbb{P}(A)} f_X(x), & x \in A, \\ 0, & x \notin A. \end{cases}$$

Definition (Conditional expectation) Given a continuous random variable X and an event A, with P(A) > 0:

$$\mathbb{E}[X|A] = \int_{-\infty}^{\infty} f_{X|A}(x) dx.$$

#### Definition (Memorylessness of the exponential random variable)

When we condition an exponential random variable X on the event X > t we have memorylessness, meaning that the "remaining time" X - t given that X > t is also geometric with the same parameter i.e.,

$$\mathbb{P}(X-t>x|X>t)=\mathbb{P}(X>x).$$

Theorem (Total probability and expectation theorems) Given a partition of the space into disjoint events  $A_1, A_2, \ldots, A_n$  such that  $\sum_i \mathbb{P}(A_i) = 1$  we have the following:

$$\begin{split} F_X(x) &= \mathbb{P}(A_1) F_{X|A_1}(x) + \dots + \mathbb{P}(A_n) F_{X|A_n}(x), \\ f_X(x) &= \mathbb{P}(A_1) f_{X|A_1}(x) + \dots + \mathbb{P}(A_n) f_{X|A_n}(x), \\ \mathbb{E}[X] &= \mathbb{P}(A_1) \mathbb{E}[X|A_1] + \dots + \mathbb{P}(A_n) \mathbb{E}[X|A_n]. \end{split}$$

Definition (Jointly continuous random variables) A pair (collection) of random variables is jointly continuous if there exists a joint PDF  $f_{X,Y}$  that describes them, that is, for every set  $B \subset \mathbb{R}^n$ 

$$\mathbb{P}\left((X,Y)\in B\right) = \iint_B f_{X,Y}(x,y) \mathrm{d}x\mathrm{d}y.$$

Properties (Properties of joint PDFs)

• 
$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$
.

• 
$$F_{X,Y}(x,y) = \mathbb{P}(X \le x, Y \le y) = \int_{-\infty}^{x} \left[ \int_{-\infty}^{y} f_{X,Y}(u,v) dv \right] du.$$

• 
$$f_{X,Y}(x) = \frac{\partial^2 F_{X,Y}(x,y)}{\partial x \, \partial y}$$

Example (Uniform joint PDF on a set S) Let  $S \subset \mathbb{R}^2$  with area s > 0, then the random variable (X, Y) is uniform over S if it has PDF

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{s}, & (x,y) \in S, \\ 0, & (x,y) \notin S. \end{cases}$$

Conditioning on a random variable, independence, Bayes' rule

#### Definition (Conditional PDF given another random variable)

Given jointly continuous random variables X, Y and a value y such that  $f_Y(y) > 0$ , we define the conditional PDF as

$$f_{X|Y}(x|y) \stackrel{\triangle}{=} \frac{f_{X,Y}(x,y)}{f_Y(y)}.$$

Additionally we define  $\mathbb{P}(X \in A|Y = y) \int_A f_{X|Y}(x|y) dx$ . Proposition (Multiplication rule) Given jointly continuous random variables X, Y, whenever possible we have

$$f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x) = f_Y(y)f_{X|Y}(x|y).$$

Definition (Conditional expectation) Given jointly continuous random variables X, Y, and y such that  $f_Y(y) > 0$ , we define the conditional expected value as

$$\mathbb{E}[X|Y=y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx.$$

Additionally we have

$$\mathbb{E}[g(X)|Y=y] = \int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx.$$

Theorem (Total probability and total expectation theorems)

$$f_X(x) = \int_{-\infty}^{\infty} f_Y(y) f_{X|Y}(x|y) dy,$$
$$\mathbb{E}[X] = \int_{-\infty}^{\infty} f_Y(y) \mathbb{E}[X|Y = y] dy.$$

Definition (Independence) Jointly continuous random variables X, Y are independent if  $f_{X,Y}(x,y) = f_X(x)f_Y(y)$  for all x,y.

Proposition (Expectation of product of independent r.v.) If X and Y are independent continuous random variables,

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

Remark If X and Y are independent,  $\mathbb{E}[g(X)h(Y)] = \mathbb{E}[g(X)]\mathbb{E}[h(Y)].$ 

Proposition (Variance of sum of independent random variables) If X and Y are independent continuous random variables,

$$Var(X + Y) = Var(X) + Var(Y)$$
.

Proposition (Bayes' rule summary)

- For X, Y discrete:  $p_{X|Y}(x|y) = \frac{p_X(x)p_{Y|X}(y|x)}{p_Y(y)}$ .
- For X, Y continuous:  $f_{X|Y}(x|y) = \frac{f_X(x)f_{Y|X}(y|x)}{f_Y(y)}$
- For X discrete, Y continuous:  $p_{X|Y}(x|y) = \frac{p_X(x)f_{Y|X}(y|x)}{f_Y(y)}$ .
- For X continuous, Y discrete:  $f_{X|Y}(x|y) = \frac{f_X(x)p_{Y|X}(y|x)}{p_Y(y)}$ .

#### Derived distributions

Proposition (Discrete case) Given a discrete random variable X and a function g, the r.v. Y = g(X) has PMF

$$p_Y(y) = \sum_{x:g(x)=y} p_X(x).$$

Remark (Linear function of discrete random variable) If g(x) = ax + b, then  $p_Y(y) = p_X\left(\frac{y-b}{a}\right)$ .

Proposition (Linear function of continuous r.v.) Given a continuous random variable X and Y = aX + b, with  $a \neq 0$ , we have

$$f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y-b}{a}\right).$$

Corollary (Linear function of normal r.v.) If  $X \sim \mathcal{N}(\mu, \sigma^2)$  and Y = aX + b, with  $a \neq 0$ , then  $Y \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$ .

Example (General function of a continuous r.v.) If X is a continuous random variable and g is any function, to obtain the pdf of Y = g(X) we follow the two-step procedure:

- 1. Find the CDF of Y:  $F_Y(y) = \mathbb{P}(Y \le y) = \mathbb{P}(q(X) \le y)$ .
- 2. Differentiate the CDF of Y to obtain the PDF:  $f_Y(y) = \frac{dF_Y(y)}{dy}$ .

Proposition (General formula for monotonic g) Let X be a continuous random variable and g a function that is monotonic wherever  $f_X(x) > 0$ . The PDF of Y = g(X) is given by

$$f_Y(y) = f_X(h(y)) \left| \frac{\mathrm{d}h}{\mathrm{d}y}(y) \right|.$$

where  $h = g^{-1}$  in the interval where g is monotonic.

#### Sums of independent r.v., covariance and correlation

Proposition (Discrete case) Let X, Y be discrete independent random variables and Z = X + Y, then the PMF of Z is

$$p_Z(z) = \sum_x p_X(x) p_Y(z-x).$$

Proposition (Continuous case) Let X, Y be continuous independent random variables and Z = X + Y, then the PDF of Z is

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx.$$

Proposition (Sum of independent normal r.v.) Let  $X \sim \mathcal{N}(\mu_x, \sigma_x^2)$  and  $Y \sim \mathcal{N}(\mu_y, \sigma_y^2)$  independent. Then  $Z = X + Y \sim \mathcal{N}(\mu_x + \mu_y, \sigma_x^2 + \sigma_y^2)$ .

Definition (Covariance) We define the covariance of random variables X, Y as

$$Cov(X,Y) \stackrel{\triangle}{=} \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])].$$

Properties (Properties of covariance)

- If X, Y are independent, then Cov(X, Y) = 0.
- Cov(X, X) = Var(X)
- Cov(aX + b, Y) = a Cov(X, Y).
- Cov(X, Y + Z) = Cov(X, Y) + Cov(X, Z).
- $Cov(X, Y) = \mathbb{E}[XY] \mathbb{E}[X]\mathbb{E}[Y].$

Proposition (Variance of a sum of r.v.)

$$\operatorname{Var}(X_1+\cdots+X_n)=\sum_i\operatorname{Var}(X_i)+\sum_{i\neq j}\operatorname{Cov}(X_i,X_j).$$

Definition (Correlation coefficient) We define the correlation coefficient of random variables X, Y, with  $\sigma_X, \sigma_Y > 0$ , as

$$\rho(X,Y) \stackrel{\triangle}{=} \frac{\mathrm{Cov}(X,Y)}{\sigma_X \sigma_Y}.$$

Properties (Properties of the correlation coefficient)

- $-1 \le \rho \le 1$ .
- If X, Y are independent, then  $\rho = 0$ .
- $|\rho| = 1$  if and only if  $X \mathbb{E}[X] = c(Y \mathbb{E}[Y])$ .
- $\rho(aX + b, Y) = \operatorname{sign}(a)\rho(X, Y)$ .

### Conditional expectation and variance, sum of random number of r.v.

Definition (Conditional expectation as a random variable) Given random variables X, Y the conditional expectation  $\mathbb{E}[X|Y]$  is the random variable that takes the value  $\mathbb{E}[X|Y=y]$  whenever Y=y. Theorem (Law of iterated expectations)

$$\mathbb{E}\left[\mathbb{E}[X|Y]\right] = \mathbb{E}[X].$$

Definition (Conditional variance as a random variable) Given random variables X, Y the conditional variance Var(X|Y) is the random variable that takes the value Var(X|Y=y) whenever Y=y.

Theorem (Law of total variance)

$$Var(X) = \mathbb{E}\left[Var(X|Y)\right] + Var\left(\mathbb{E}[X|Y]\right).$$

Proposition (Sum of a random number of independent r.v.)

Let N be a nonnegative integer random variable. Let  $X, X_1, X_2, \ldots, X_N$  be i.i.d. random variables. Let  $Y = \sum_i X_i$ . Then

$$\mathbb{E}[Y] = \mathbb{E}[N]\mathbb{E}[X],$$

$$Var(Y) = \mathbb{E}[N]Var(X) + (\mathbb{E}[X])^{2}Var(N).$$

#### Convergence of random variables

## Inequalities, convergence, and the Weak Law of Large Numbers

Theorem (Markov inequality) Given a random variable  $X \ge 0$  and, for every a > 0 we have

$$\mathbb{P}(X \ge a) \le \frac{\mathbb{E}[X]}{a}.$$

Theorem (Chebyshev inequality) Given a random variable X with  $\mathbb{E}[X] = \mu$  and  $\text{Var}(X) = \sigma^2$ , for every  $\epsilon > 0$  we have

$$\mathbb{P}\left(\left|X-\mu\right| \geq \epsilon\right) \leq \frac{\sigma^2}{\epsilon^2}.$$

Theorem (Weak Law of Large Number (WLLN)) Given a sequence of i.i.d. random variables  $\{X_1, X_2, ...\}$  with  $\mathbb{E}[X_i] = \mu$  and  $\text{Var}(X_i) = \sigma^2$ , we define

$$M_n = \frac{1}{n} \sum_{i=1}^n X_i,$$

for every  $\epsilon > 0$  we have

$$\lim_{n\to\infty} \mathbb{P}\left(|M_n - \mu| \ge \epsilon\right) = 0.$$

Definition (Convergence in probability) A sequence of random variables  $\{Y_i\}$  converges in probability to the random variable Y if

$$\lim_{n\to\infty} \mathbb{P}\left(|Y_i - Y| \ge \epsilon\right) = 0,$$

for every  $\epsilon > 0$ .

Properties (Properties of convergence in probability) If  $X_n \to a$  and  $Y_n \to b$  in probability, then

- $\bullet \ \ X_n + Y_n \to a + b.$
- If q is a continuous function, then  $q(X_n) \to q(a)$ .
- $\mathbb{E}[X_n]$  does not always converge to a.

#### The Central Limit Theorem

Theorem (Central Limit Theorem (CLT)) Given a sequence of independent random variables  $\{X_1, X_2, ...\}$  with  $\mathbb{E}[X_i] = \mu$  and  $\text{Var}(X_i) = \sigma^2$ , we define

$$Z_n = \frac{1}{\sigma\sqrt{n}} \sum_{i=1}^n (X_i - \mu).$$

Then, for every z, we have

$$\lim_{n\to\infty} \mathbb{P}(Z_n \le z) = \mathbb{P}(Z \le z),$$

where  $Z \sim \mathcal{N}(0, 1)$ .

Corollary (Normal approximation of a binomial) Let  $X \sim Bin(n, p)$  with n large. Then  $S_n$  can be approximated by  $Z \sim \mathcal{N}(np, np(1-p))$ .

Remark (De Moivre-Laplace 1/2 approximation) Let  $X \sim Bin$ , then  $\mathbb{P}(X=i) = \mathbb{P}\left(i-\frac{1}{2} \leq X \leq i+\frac{1}{2}\right)$  and we can use the CLT to approximate the PMF of X.