

1. Introduction (1-3 paragraphs)

According to a report presenting statistics on health insurance coverage in the United States by Current Population Survey Annual Social and Economic Supplements (CPS ASEC) and the American Community Survey (ACS), 91.2% of American have health insurance in 2017. One of the most frequently asked questions that people have is that how much they will be charged if they want to have a health insurance. Thus, we would like to use the technique of linear regression to deal with this issue. To be more specific, we would like to apply different linear regression skills with a health insurance dataset to build a regression model. From this model, we will be able to discover the relationships between different factors and charge.

We got the dataset from this link: <https://www.kaggle.com/mirichoi0218/insurance>

2. Method (2 pages)

(a) Brief Description of Data

This is a health insurance dataset containing totally 1338 rows with 6 predictors and 1 response variable. The following is the information of the variables.

Response variable:

charges: individual medical costs billed by health insurance; numerical

Input variables:

age: age of primary beneficiary; numerical

sex: insurance contractor gender, may be female or male; categorical

bmi: body mass index; numerical

children: number of children covered by health insurance; numerical

smoker: whether smoke or not, may be yes or no; categorical

region: the beneficiary's residential area in the US, may be northeast, southeast, southwest, northwest; categorical

(b) Description of Preliminary Exploratory Analysis

Firstly, we drew a histogram to show the distribution of response variable charges. We realized that most of the values are between 0 to 20000. To make it more spread out, we decided to do a logarithm transformation on the charges. Then, we drew multiple scatter plots between $\log(\text{charges})$ and predictors. From the plots, we may find out several relationships between predictors and response variable. Firstly, $\log(\text{charges})$ increase as the age grows. Also, it seems that $\log(\text{charges})$ increase as bmi gets higher, though the relationship is pretty weak. Also, $\log(\text{charges})$ and children have positive correlation. As for categorical variables, we can tell from box plots that female and male have similar medians with male has larger variability. Similarly, the four regions also have similar medians. However, smoker with value yes have much higher charges than those with no.

(c,d) Models Building Process

Firstly, we built a main effect model *m1* using all six predictor variables and $\log(\text{charge})$ as response variable. The model has the adjusted R-squared value of 0.7666, which means that this model can explain only about 77% of the response variable. However, since the p-value is less than $2.2\text{e-}16$, we can tell that most of the predictors in the model are significant.

Next, we tried to build a model *m2* with full interaction, which contains all possible interaction between predictor variables. Now the adjusted R-squared has increased to 0.8316 with smaller AIC and a small p-value less than $2.2\text{e-}16$. Though this model seems ideal, it contains many statistically insignificant predictors. Furthermore, it's too complex to interpret. Therefore, we would like to use this model to build a more precise model.

We then used stepwise variable selection method with both direction to build model *m3*. We used the main effect model *m1* as our base step and full interaction model *m2* as an upper limit, hoping to get significant interactions. It turns out that *m3* has similar adjusted R-squared as *m2*. However, this step was just needed for us to determine which interactions should be included in our model. Based on the result, the interaction *age:smoker*, *bmi:smoker*, *age:children*, *children:smoker*, *age:sex*, *age:region* are significant. This led us to think about the variable *bmi* in depth. Thus, we made an assumption that since *bmi* is considered 'normal' until 29.9, only *bmi* with higher than 30 might affect charges significantly.

To further explore the effect of *bmi*, we created a variable called *bmiOver30* which is a factor. Then, we added this variable to build the model *m4* with 6 original predictors and new interactions *age:smoker*, *bmiOver30:smoker*, *age:children*, *children:smoker*, *age:sex*, *age:region*, *bmiOver30:region*, *age:children:smoker*. This model has adjusted R-squared 0.8367. Besides, it has better AIC value. However, since there is still predictor that is not significant in the model, we would like to build a model *m5* based on *m4* by deleting those insignificant variables. Now all the predictors are significant and the adjusted R-squared is 0.8324.

(e) Model Checking

To make sure the final model is feasible, significance test, ANOVA test, residual analysis, normality test, error independence test, homoscedasticity test and multicollinearity test are applied. The final model passes all of above tests so that it is feasible and reasonable.

3. Results (2 pages)

(a) Model Selection

Name	Model	Adjusted R ²	AIC	p-value
m5	lm(log(charges)~age+sex+children+smoker+bmiOver30:smoker+age:children+age:smoker+smoker:children)	0.8324	1194.529	<2.2e-16
m4	lm(log(charges)~age+sex+bmi+children+smoker+age:smoker+bmiOver30:smoker+age:children+children:smoker+age:sex+age:region+bmiOver30:region+age:children:smoker)	0.8367	1168.562	<2.2e-16
m3	step(m1, scope=list(upper=m2), direction = "both")	0.8374	1169.492	<2.2e-16
m2	lm(log(charges) ~ age * sex * bmi * children * smoker * region)	0.8316	1312.767	<2.2e-16
m1	lm(log(charges) ~ age + sex + bmi + children + smoker + region)	0.7666	1637.033	<2.2e-16

(b) ANOVA table

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
age	1	314.96	314.96	2222.7288	< 2.2e-16 ***
sex	1	0.31	0.31	2.2101	0.1374
children	1	21.77	21.77	153.6260	< 2.2e-16 ***
smoker	1	520.59	520.59	3673.8955	< 2.2e-16 ***
smoker:bmiOver30	2	28.78	14.39	101.5526	< 2.2e-16 ***
age:children	1	6.02	6.02	42.4722	1.015e-10 ***
age:smoker	1	45.67	45.67	322.2676	< 2.2e-16 ***
children:smoker	1	4.20	4.20	29.6514	6.154e-08 ***
Residuals	1328	188.18	0.14	-	-

(c) Parameter estimates table for final model

Parameter	Point Estimate	Standard Error	t-statistic	p-value	Confidence Interval	
					2.5 %	97.5 %
(Intercept)	6.9340643	0.042773	162.113	< 2e-16	6.850154218	7.01797433

<i>age</i>	0.0447368	0.000965	46.373	< 2e-16	0.042844285	0.04662934
<i>sexmale</i>	-0.086777	0.020725	-4.187	3.01e-05	-0.12743272	-0.0461202
<i>children</i>	0.2881687	0.027424	10.508	< 2e-16	0.234370046	0.34196740
<i>smokeryes</i>	2.5666818	0.081120	31.641	< 2e-16	2.407544836	2.72581867
<i>smokerno:bmiOver30</i>	-0.005278	0.023250	-0.227	0.82	-0.05088825	0.04033223
<i>smokeryes:bmiOver30</i>	0.6855582	0.045708	14.999	< 2e-16	0.595891332	0.77522505
<i>age:children</i>	-0.004069	0.000653	-6.23	6.07e-10	-0.00534855	-0.0027883
<i>age:smokeryes</i>	-0.032201	0.001839	-17.508	< 2e-16	-0.03580850	-0.0285925
<i>children:smokeryes</i>	-0.119442	0.021935	-5.445	6.15e-08	-0.16247248	-0.0764112

4. Discussion (1 page)

After running five models and concise modeling selection process, we choose m5 as our final model. Based on several types of modeling checking methods, we could reach the conclusion that our model is reasonable and feasible.

According to our final model, the variable, age, is significant. Older people may have higher insurance charges. Female would pay less for their insurance. The number of children, is relevant to charges. When people have more children, they need to pay more for insurance. The charges for smokers are higher than those for non-smokers. In terms of interactions, the interactions between smokers with bmi over 30 and non-smokers with bmi over 30 are both significant for the insurance charges. Among people whose bmi is over 30, smokers tend to pay more than non-smokers.

However, the model and results given by this project have some limitations as well. On the one hand, we only collected about 1400 set of data in the dataset because of the limited resource. Though it is not a small number, it could help get a more accurate model if we could have more data. On the other hand, actually, there should be more variables which could influence the insurance charges, like the customers' job, salaries and so on. Here,

only six variables are considered. It is necessary to discuss how other variables can affect the insurance charges. A more comprehensive dataset can help solve this problem. The interactions between age and children, age and smoker, children and smoker are also significant.

5. References (1 page)

[1] <https://www.kaggle.com/mirichoi0218/insurance>

[2] Grumbach, Kevin, et al. "Charges for obstetric liability insurance and discontinuation of obstetric practice in New York." *Journal of Family Practice*, Jan. 1997, p. 61+. *Academic OneFile*,
<https://link.galegroup.com/apps/doc/A19122984/AONE?u=googlescholar&sid=AONE&xid=1a611fbd>. Accessed 26 Nov. 2018.

[3] <https://www.census.gov/library/publications/2018/demo/p60-264.html>

6. Appendix A: Code

```
# transform some variables into factor
insurance$sex = factor(insurance$sex)
insurance$smoker = factor(insurance$smoker)
insurance$region = factor(insurance$region)

# draw plots to see relationships
plot(age, charges)
plot(bmi, charges)
plot(children, charges)
boxplot(charges ~ sex, xlab='sex', ylab='charges')
boxplot(charges ~ smoker, xlab='smoker', ylab='charges')
boxplot(charges ~ region, xlab='region', ylab='charges')

plot(age, bmi)
plot(age, children)
plot(bmi, children)

plot(insurance)

# main effects model
m1 = lm(log(charges) ~ age + sex + bmi + children + smoker +
region, data = insurance)
summary(m1)
anova(m1)
AIC(m1)

layout(matrix(c(1,2,3,4), 2,2))
```

```

plot(m1)

# all possible interaction
m2 = lm(log(charges) ~ age * sex * bmi * children * smoker *
region, data = insurance)
summary(m2)
anova(m2)
AIC(m2)

# Half-way model
m3 <- step(m1,scope=list(upper=m2),direction = "both")
summary(m3)
AIC(m3)

# Considered model
bmiOver30 <- ifelse(insurance$bmi >= 30, 1, 0)
m4 <-
lm(log(charges)~age+sex+bmi+children+smoker+age:smoker+bmiOver30:s
moker+age:children+children:smoker+age:sex+age:region+bmiOver30:re
gion+age:children:smoker,data=insurance)
summary(m4)
AIC(m4)

# Final Model
m5 <-
lm(log(charges)~age+sex+children+smoker+bmiOver30:smoker+age:child
ren+age:smoker+smoker:children,data=insurance)
summary(m5)
AIC(m5)
confint(m5)

#Compare m4 and m5
anova(m4,m5)

# visualize model checking
library(ggfortify)
autoplot(m5)

# significance test # comprehensive test
summary(m5)

# anova test

```

```
anova(m5)
plot(m5)

# residual analysis:delete outliers
plot(m5,which=1:4)
qqPlot(m5,id.method='identify',simulate=T)
outlierTest(m5)

# normality test
res <- residuals(m5)
plot(density(res))
shapiro.test(res)
# => pass!

# homoscedasticity test
ncvTest(m5)
# => pass!
spreadLevelPlot(m5)

# multicollinearity test
vif(m5)
sqrt(vif(m5))>10
# => pass!

# comprehensive test
library(gvlma)
gvmodel <- gvlma(m5)
gvmodel
```

7. Appendix B: Output

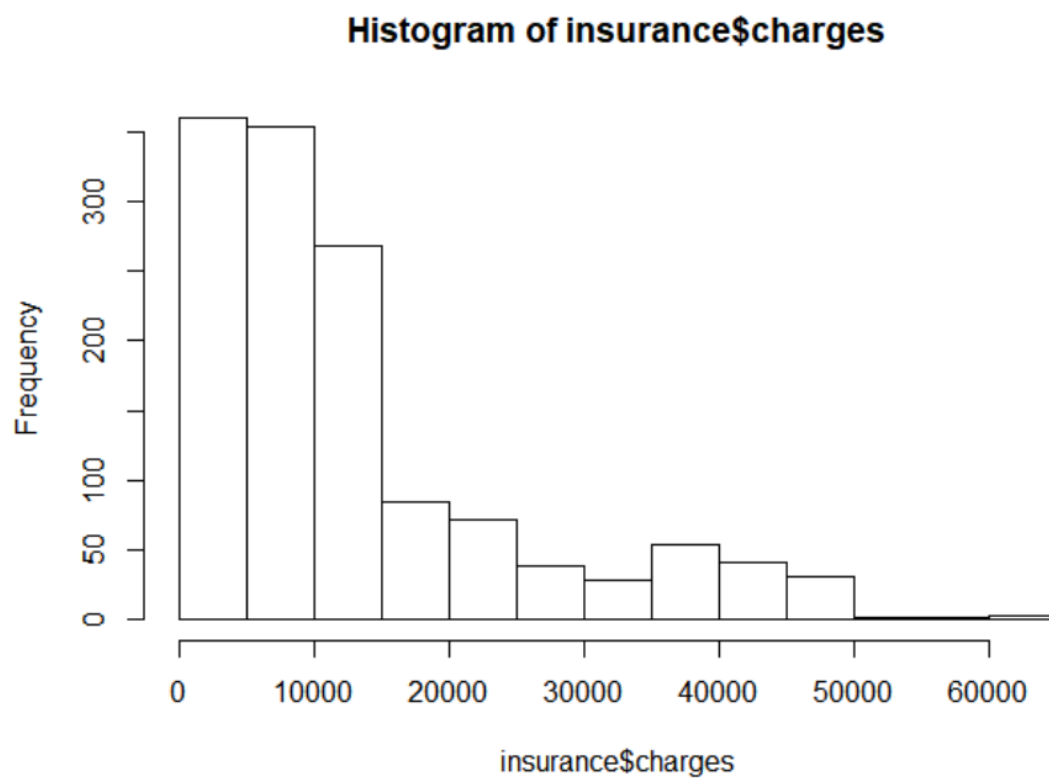


Fig. 7.1 Histogram of charges

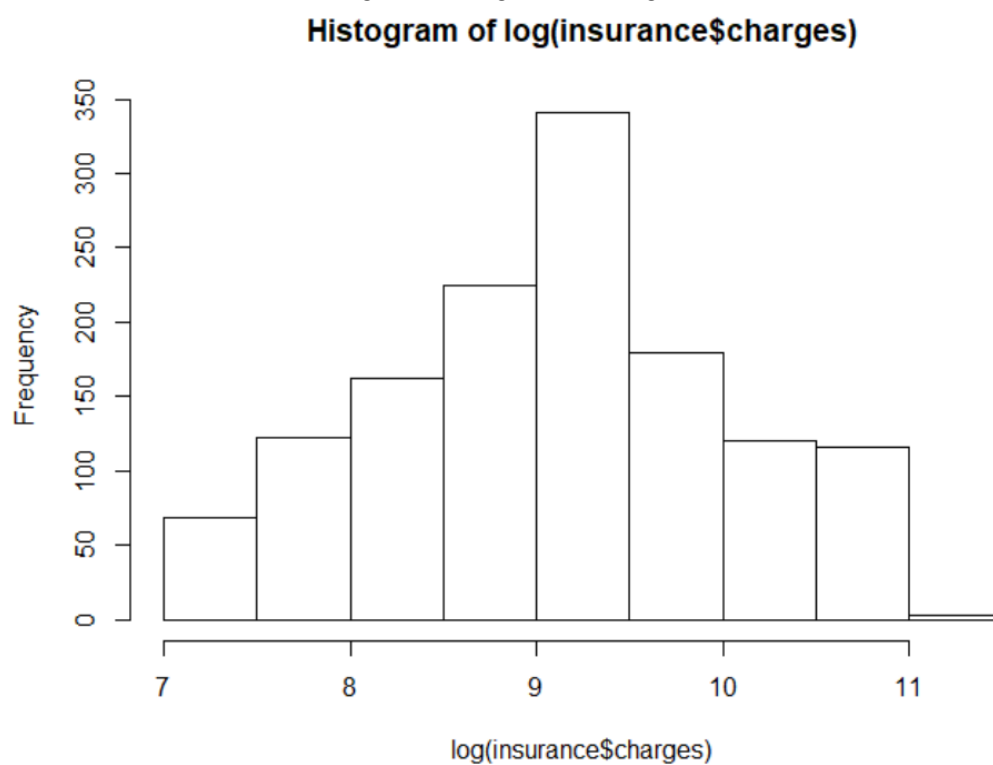


Fig. 7.2 Histogram of log(charges)

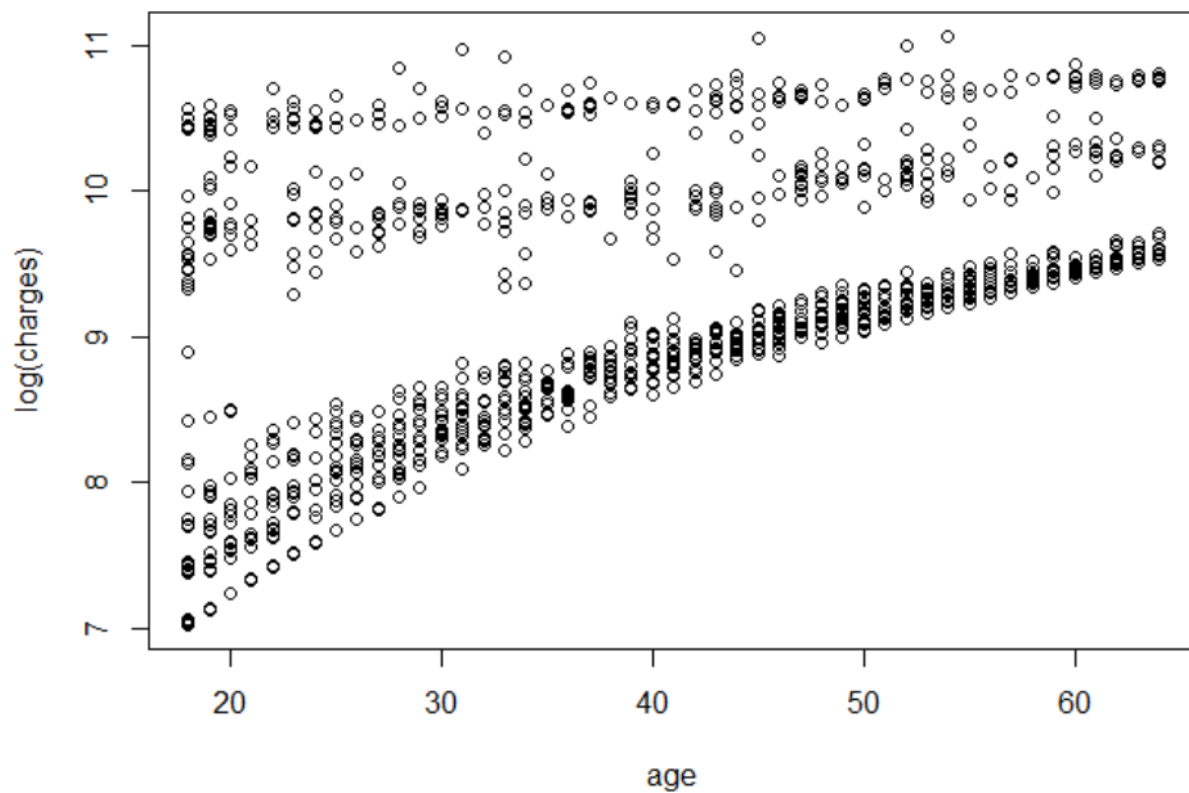


Fig. 7.3 Scatter plot of $\log(\text{charges}) \sim \text{age}$

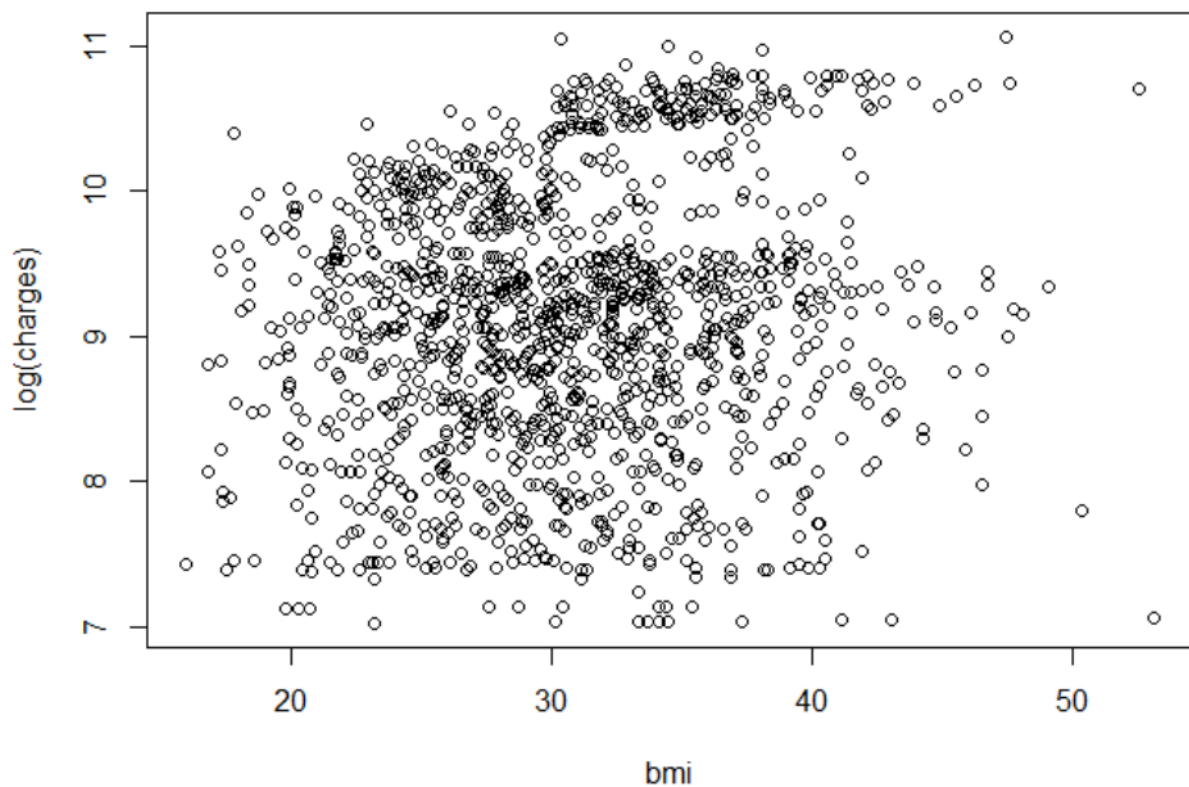


Fig. 7.4 Scatter plot of $\log(\text{charges}) \sim \text{bmi}$

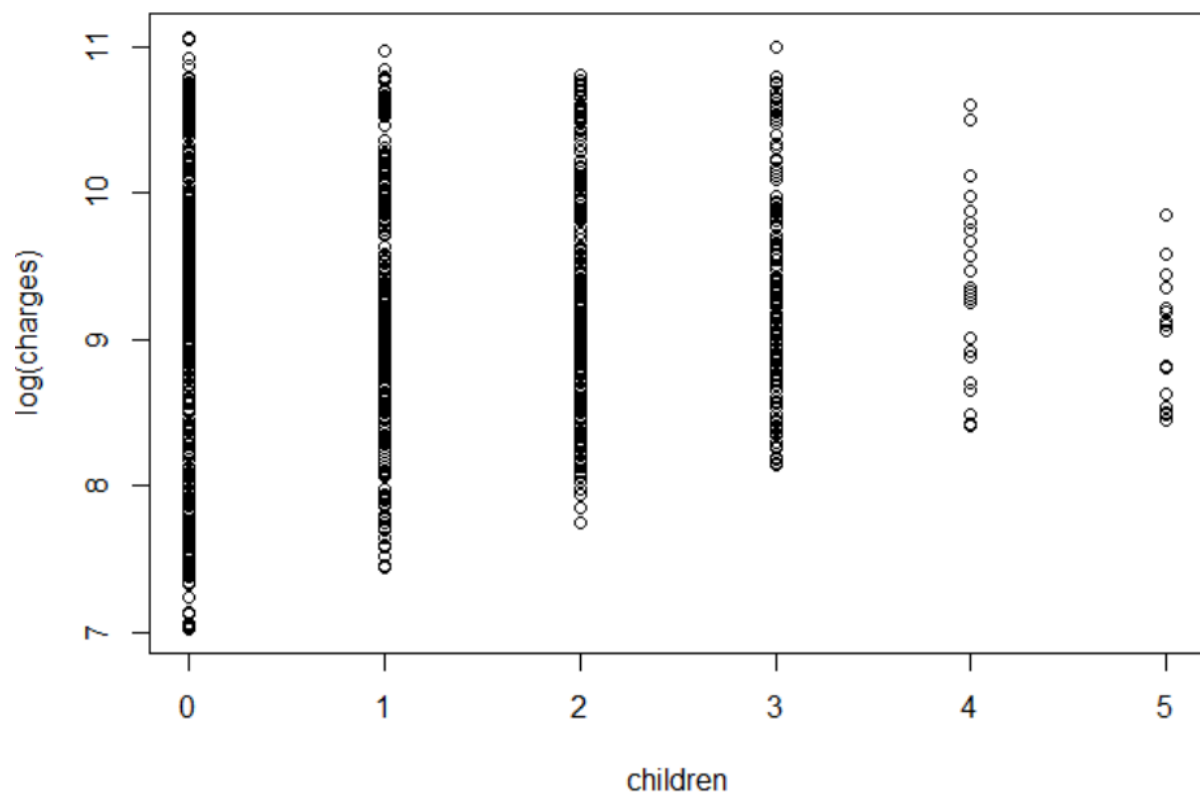


Fig. 7.5 Scatter plot of $\log(\text{charges})$ versus children

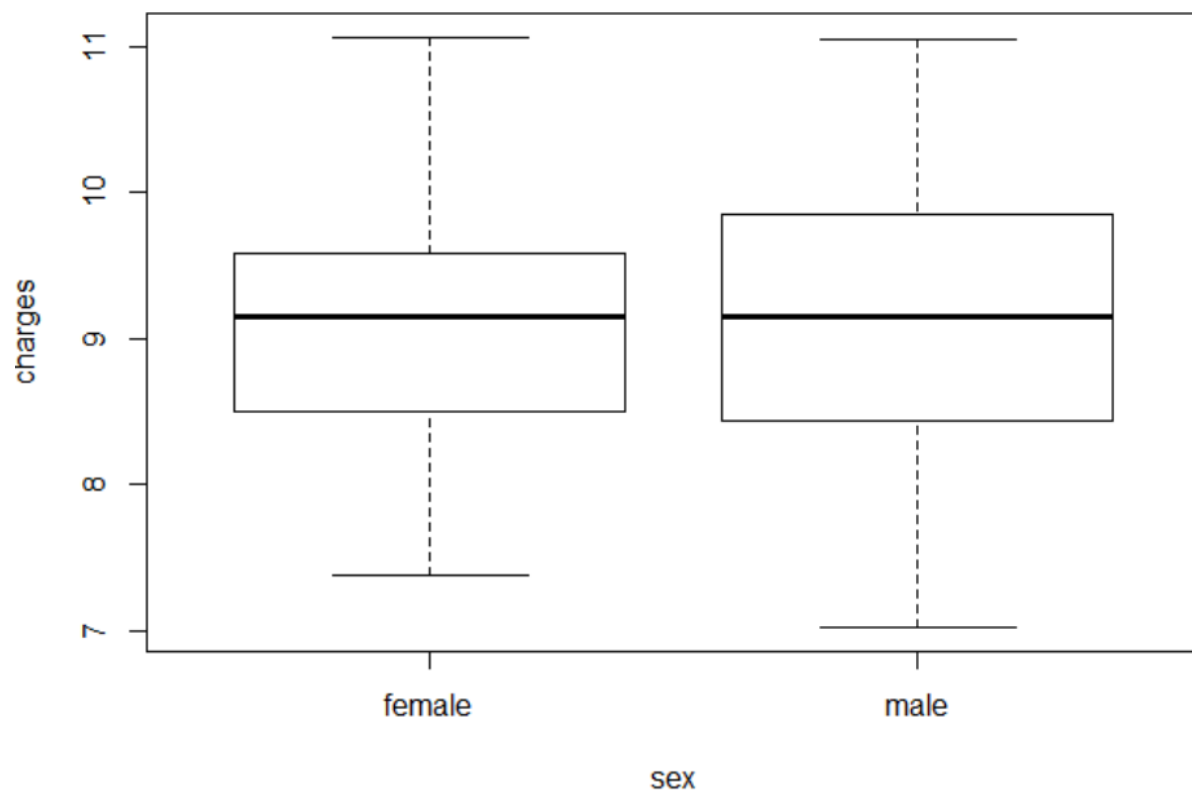


Fig. 7.6 Box plot of $\log(\text{charges})$ versus sex

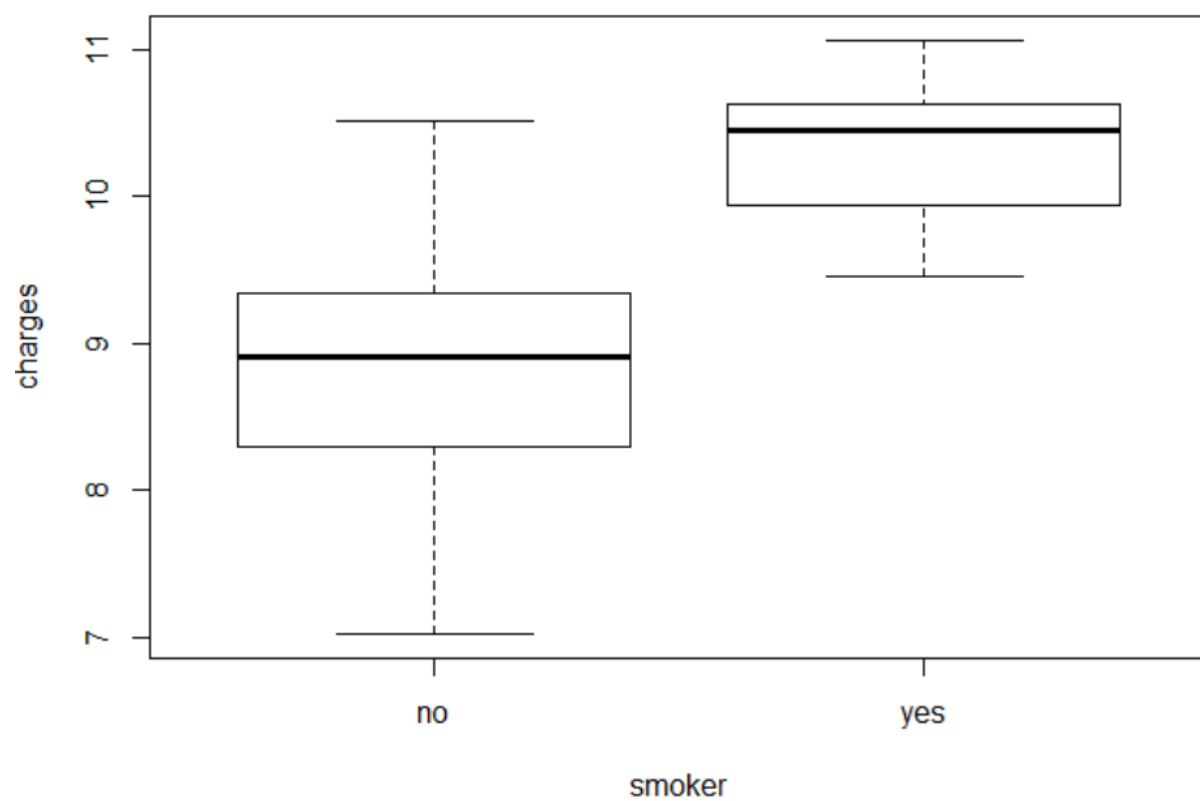


Fig. 7.7 Box plot of log(charges) versus smoker

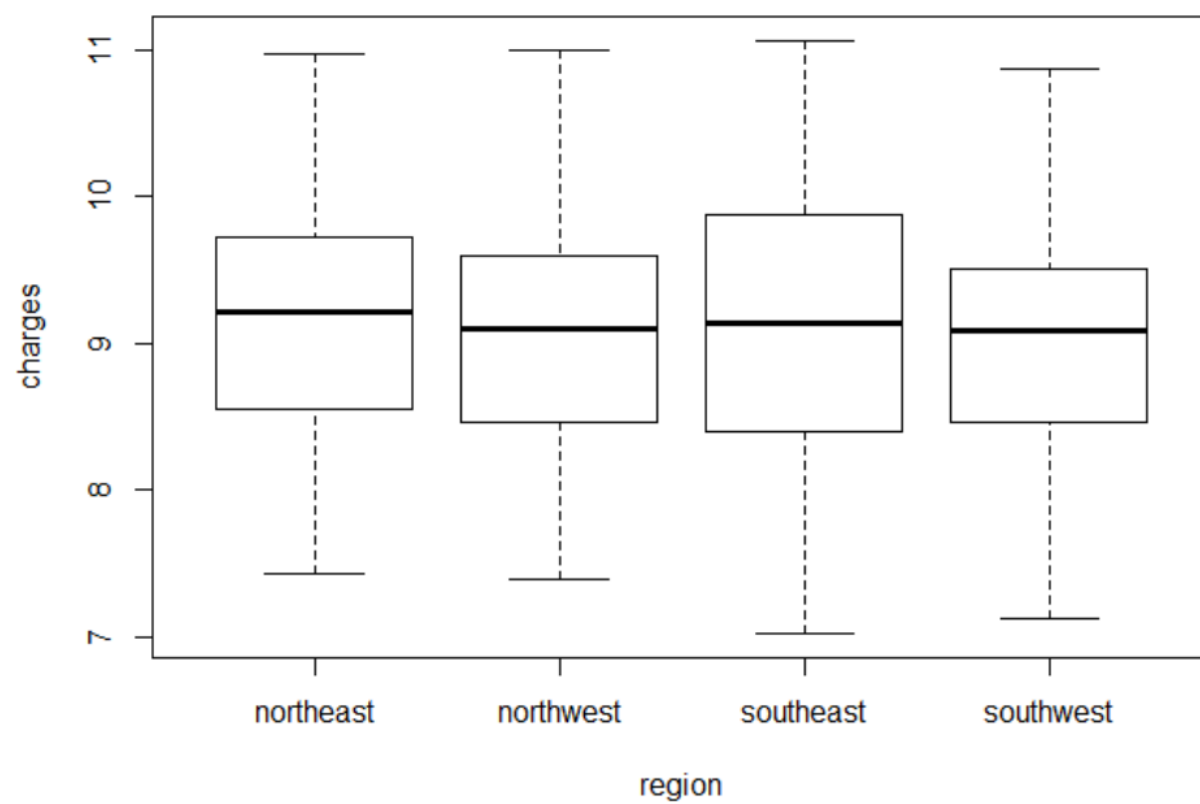


Fig. 7.8 Box plot of log(charges) versus region

```
# main effects model
```

```
> summary(m1)
```

Call:

```
lm(formula = log(charges) ~ age + sex + bmi + children + smoker +  
    region, data = insurance)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.07186	-0.19835	-0.04917	0.06598	2.16636

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	7.0305581	0.0723960	97.112	< 2e-16	***
age	0.0345816	0.0008721	39.655	< 2e-16	***
sexmale	-0.0754164	0.0244012	-3.091	0.002038	**
bmi	0.0133748	0.0020960	6.381	2.42e-10	***
children	0.1018568	0.0100995	10.085	< 2e-16	***
smokeryes	1.5543228	0.0302795	51.333	< 2e-16	***
regionnorthwest	-0.0637876	0.0349057	-1.827	0.067860	.
regionsoutheast	-0.1571967	0.0350828	-4.481	8.08e-06	***
regionsouthwest	-0.1289522	0.0350271	-3.681	0.000241	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.4443 on 1329 degrees of freedom

Multiple R-squared: 0.7679, Adjusted R-squared: 0.7666

F-statistic: 549.8 on 8 and 1329 DF, p-value: < 2.2e-16

```
> AIC(m1)
```

```
[1] 1637.033
```

```
# all possible interaction
```

```
> m2 = lm(log(charges) ~ age * sex * bmi * children * smoker * region,  
data = insurance)
```

Call:

```
lm(formula = log(charges) ~ age * sex * bmi * children * smoker *  
    region, data = insurance)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.68444	-0.15029	-0.06529	0.00265	2.46927

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3774 on 1210 degrees of freedom

Multiple R-squared: 0.8476, Adjusted R-squared: 0.8316

F-statistic: 52.98 on 127 and 1210 DF, p-value: < 2.2e-16

```
> AIC(m2)
```

```
[1] 1312.767
```

```
> anova(m1,m2)
```

Analysis of Variance Table

Model 1: log(charges) ~ age + sex + bmi + children + smoker + region

Model 2: log(charges) ~ age * sex * bmi * children * smoker * region

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	1329	262.33				
2	1210	172.32	119	90.006	5.3109	< 2.2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
> # Half-way model
```

```
> summary(m3)
```

Call:

```
lm(formula = log(charges) ~ age + sex + bmi + children + smoker +  
    region + age:smoker + bmi:smoker + age:children + children:smoker +  
    age:sex + age:region + bmi:region + sex:smoker + age:bmi +  
    smoker:region + age:children:smoker, data = insurance)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.52439	-0.15504	-0.07881	-0.01138	2.55417

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	6.7938672	0.1871676	36.298	< 2e-16
age	0.0442536	0.0040504	10.926	< 2e-16
sexmale	-0.2956321	0.0616672	-4.794	1.82e-06
bmi	0.0169406	0.0060567	2.797	0.005234

children	0.3243387	0.0298037	10.882	< 2e-16
smokeryes	1.4536886	0.1585072	9.171	< 2e-16
regionnorthwest	-0.0066814	0.1711631	-0.039	0.968868
regionsoutheast	-0.0027048	0.1634594	-0.017	0.986800
regionsouthwest	-0.2640395	0.1645770	-1.604	0.108877
age:smokeryes	-0.0368050	0.0022844	-16.111	< 2e-16
bmi:smokeryes	0.0497069	0.0043119	11.528	< 2e-16
age:children	-0.0049765	0.0007179	-6.932	6.50e-12
children:smokeryes	-0.3044677	0.0674578	-4.513	6.95e-06
age:sexmale	0.0048209	0.0014553	3.313	0.000950
age:regionnorthwest	0.0026388	0.0020957	1.259	0.208211
age:regionsoutheast	0.0080687	0.0021003	3.842	0.000128
age:regionsouthwest	0.0078965	0.0021344	3.700	0.000225
bmi:regionnorthwest	-0.0056689	0.0053729	-1.055	0.291579
bmi:regionsoutheast	-0.0149898	0.0046520	-3.222	0.001303
bmi:regionsouthwest	-0.0078903	0.0051422	-1.534	0.125170
sexmale:smokeryes	0.0960914	0.0516605	1.860	0.063101
age:bmi	-0.0001964	0.0001249	-1.573	0.116030
smokeryes:regionnorthwest	0.0504945	0.0743255	0.679	0.497023
smokeryes:regionsoutheast	0.0798718	0.0707557	1.129	0.259173
smokeryes:regionsouthwest	0.1828824	0.0751375	2.434	0.015067
age:children:smokeryes	0.0046933	0.0016548	2.836	0.004636

(Intercept)	***
age	***
sexmale	***
bmi	**
children	***
smokeryes	***
regionnorthwest	
regionsoutheast	
regionsouthwest	
age:smokeryes	***
bmi:smokeryes	***
age:children	***
children:smokeryes	***
age:sexmale	***
age:regionnorthwest	
age:regionsoutheast	***
age:regionsouthwest	***
bmi:regionnorthwest	
bmi:regionsoutheast	**
bmi:regionsouthwest	
sexmale:smokeryes	.
age:bmi	
smokeryes:regionnorthwest	

```
smokeryes:regionsoutheast
smokeryes:regionsouthwest *
age:children:smokeryes      **
---
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3707 on 1312 degrees of freedom
Multiple R-squared: 0.8405, Adjusted R-squared: 0.8374
F-statistic: 276.5 on 25 and 1312 DF, p-value: < 2.2e-16

```
> AIC(m3)
[1] 1169.492
```

```
> # Considered model
> summary(m4)
```

Call:
lm(formula = log(charges) ~ age + sex + bmi + children + smoker +
 age:smoker + bmiOver30:smoker + age:children + children:smoker +
 age:sex + age:region + bmiOver30:region + age:children:smoker,
 data = insurance)

Residuals:

Min	1Q	Median	3Q	Max
-0.58866	-0.14054	-0.07513	-0.00679	2.47407

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	6.8558714	0.0883058	77.638	< 2e-16	***
age	0.0436244	0.0013549	32.197	< 2e-16	***
sexmale	-0.2418968	0.0606713	-3.987	7.06e-05	***
bmi	0.0052866	0.0028828	1.834	0.06690	.
children	0.3216612	0.0297813	10.801	< 2e-16	***
smokeryes	2.7243129	0.0955219	28.520	< 2e-16	***
age:smokeryes	-0.0363313	0.0022811	-15.927	< 2e-16	***
smokerno:bmiOver30	0.0380279	0.0509148	0.747	0.45526	
smokeryes:bmiOver30	0.7293499	0.0643959	11.326	< 2e-16	***
age:children	-0.0048794	0.0007176	-6.799	1.59e-11	***
children:smokeryes	-0.2994366	0.0674406	-4.440	9.75e-06	***
age:sexmale	0.0040394	0.0014558	2.775	0.00560	**
age:regionnorthwest	-0.0001738	0.0009301	-0.187	0.85182	
age:regionsoutheast	-0.0003806	0.0010014	-0.380	0.70393	
age:regionsouthwest	-0.0010771	0.0009761	-1.103	0.27001	
bmiOver30:regionnorthwest	-0.0624980	0.0581025	-1.076	0.28228	

```
bmiOver30:regionsoutheast -0.1361595 0.0564043 -2.414 0.01591 *
bmiOver30:regionsouthwest -0.1393545 0.0584029 -2.386 0.01717 *
age:children:smokeryes 0.0045806 0.0016527 2.772 0.00566 **
---
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3716 on 1319 degrees of freedom
Multiple R-squared: 0.8389, Adjusted R-squared: 0.8367
F-statistic: 381.6 on 18 and 1319 DF, p-value: < 2.2e-16

```
> AIC(m4)
[1] 1168.562
```

```
> # Final Model
> summary(m5)
```

Call:
lm(formula = log(charges) ~ age + sex + children + smoker +
bmiOver30:smoker +
age:children + age:smoker + smoker:children, data = insurance)

Residuals:

Min	1Q	Median	3Q	Max
-0.62979	-0.13713	-0.07356	-0.00603	2.35484

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	6.9340643	0.0427730	162.113	< 2e-16	***
age	0.0447368	0.0009647	46.373	< 2e-16	***
sexmale	-0.0867765	0.0207245	-4.187	3.01e-05	***
children	0.2881687	0.0274238	10.508	< 2e-16	***
smokeryes	2.5666818	0.0811198	31.641	< 2e-16	***
smokerno:bmiOver30	-0.0052780	0.0232497	-0.227	0.82	
smokeryes:bmiOver30	0.6855582	0.0457075	14.999	< 2e-16	***
age:children	-0.0040684	0.0006525	-6.235	6.07e-10	***
age:smokeryes	-0.0322005	0.0018392	-17.508	< 2e-16	***
children:smokeryes	-0.1194418	0.0219348	-5.445	6.15e-08	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3764 on 1328 degrees of freedom
Multiple R-squared: 0.8335, Adjusted R-squared: 0.8324
F-statistic: 738.9 on 9 and 1328 DF, p-value: < 2.2e-16


```
> AIC(m5)
[1] 1194.529
```

```
> #Compare m4 and m5
```

```
> anova(m4,m5)
```

Analysis of Variance Table

Model 1: log(charges) ~ age + sex + bmi + children + smoker + age:smoker +

bmiOver30:smoker + age:children + children:smoker + age:sex +
age:region + bmiOver30:region + age:children:smoker

Model 2: log(charges) ~ age + sex + children + smoker + bmiOver30:smoker +

age:children + age:smoker + smoker:children

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	1319	182.09				
2	1328	188.18	-9	-6.083	4.8958	1.777e-06 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
> confint(m5)
```

	2.5 %	97.5 %
(Intercept)	6.850154218	7.017974330
age	0.042844285	0.046629335
sexmale	-0.127432720	-0.046120240
children	0.234370046	0.341967399
smokeryes	2.407544836	2.725818673
smokerno:bmiOver30	-0.050888250	0.040332226
smokeryes:bmiOver30	0.595891332	0.775225046
age:children	-0.005348547	-0.002788304
age:smokeryes	-0.035808503	-0.028592502
children:smokeryes	-0.162472480	-0.076411186

```
> # anova test
```

```
> anova(m5)
```

Analysis of Variance Table

Response: log(charges)

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
age	1	314.96	314.96	2222.7288	< 2.2e-16 ***
sex	1	0.31	0.31	2.2101	0.1374
children	1	21.77	21.77	153.6260	< 2.2e-16 ***
smoker	1	520.59	520.59	3673.8955	< 2.2e-16 ***
smoker:bmiOver30	2	28.78	14.39	101.5526	< 2.2e-16 ***

```

age:children      1    6.02    6.02   42.4722 1.015e-10 ***
age:smoker        1   45.67   45.67  322.2676 < 2.2e-16 ***
children:smoker   1    4.20    4.20   29.6514 6.154e-08 ***
Residuals        1328 188.18    0.14

---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```
# Modeling Checking for Final Model
```

```
# residual analysis
```

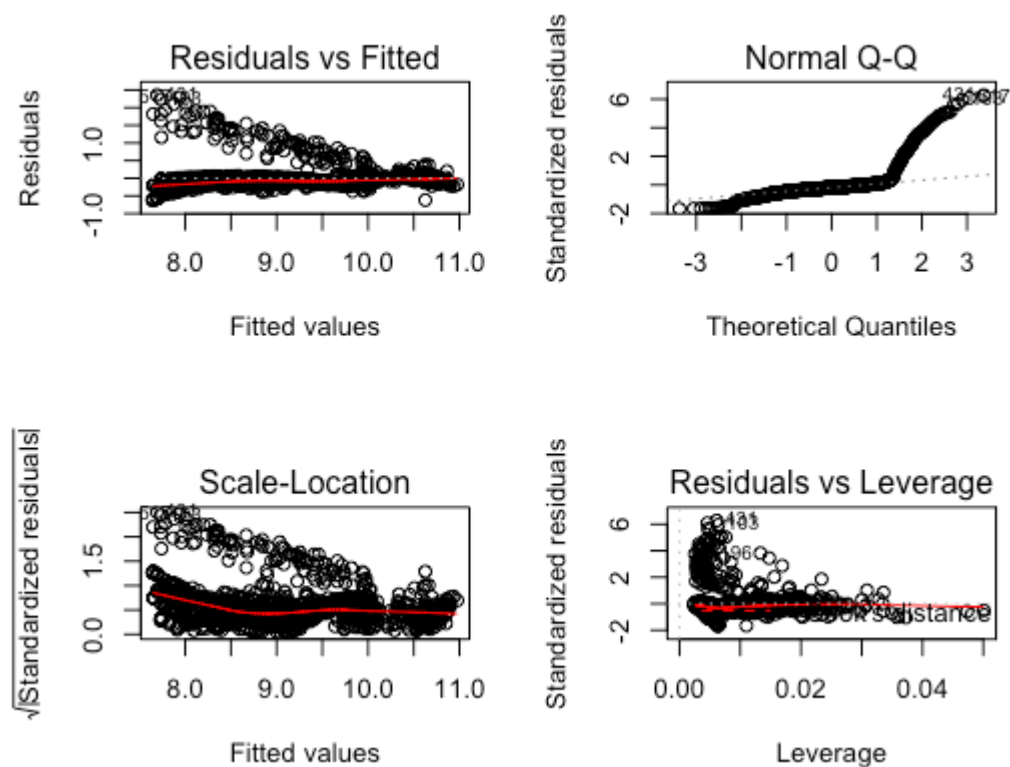


Figure 7.9 Residual Plot

```

> outlierTest(m5)
      rstudent unadjusted p-value Bonferonni p
431  6.367722      2.6381e-10    3.5297e-07
517  6.172712      8.9032e-10    1.1912e-06
103  6.033966      2.0721e-09    2.7724e-06
220  5.717868      1.3310e-08    1.7808e-05
1028 5.668935      1.7608e-08    2.3560e-05
398  5.204898      2.2470e-07    3.0065e-04

```

1040	5.117985	3.5424e-07	4.7397e-04
1020	5.092837	4.0359e-07	5.4000e-04
527	5.080943	4.2918e-07	5.7424e-04
341	4.948271	8.4471e-07	1.1302e-03

```
# normality test
> shapiro.test(res)
```

Shapiro-Wilk normality test

```
data:  res
W = 0.59439, p-value < 2.2e-16
```

```
# => pass!
```

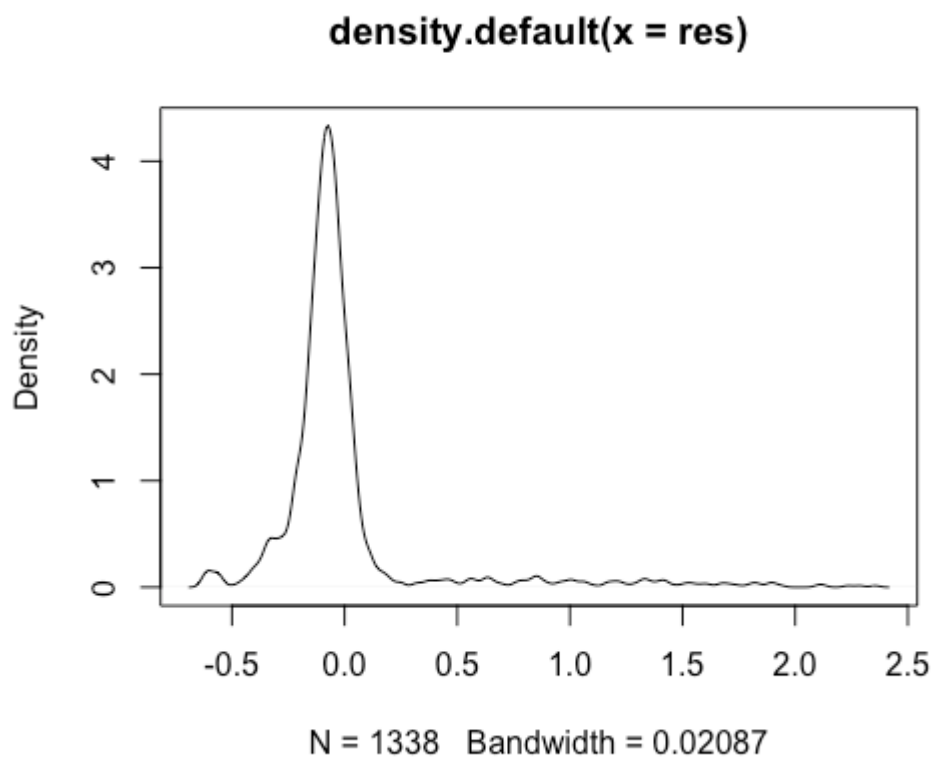


Figure 7.10 Normality Test

```
> # homoscedasticity test
> ncvTest(m5)
Non-constant Variance Score Test
Variance formula: ~ fitted.values
Chisquare = 422.9763, Df = 1, p = < 2.22e-16
```

=> pass!

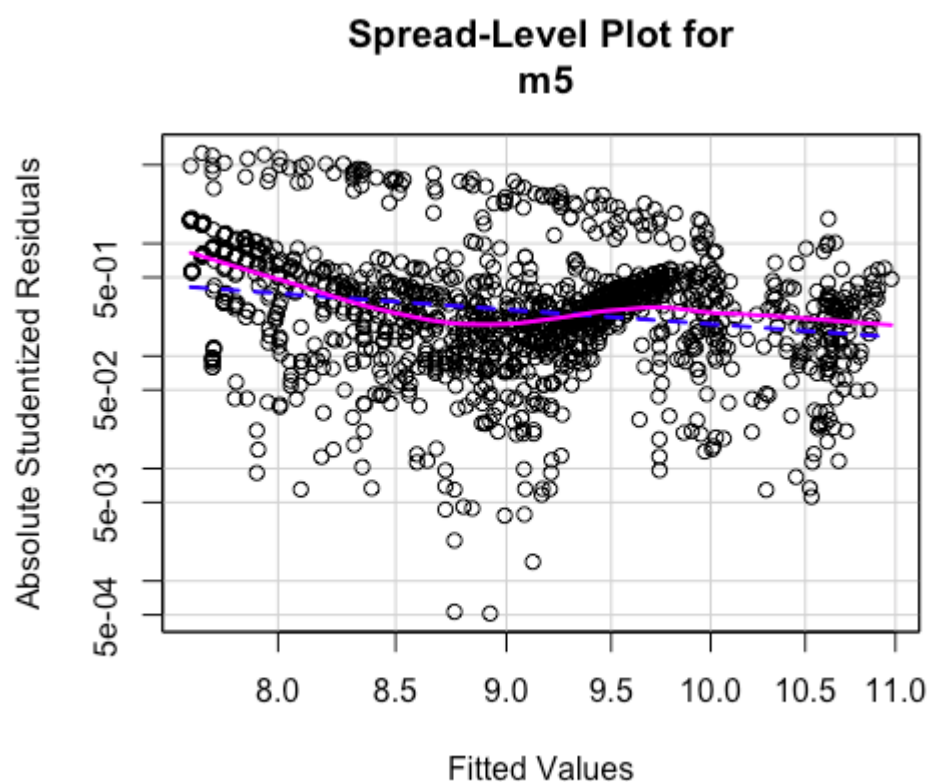


Figure 7.11 Homoscedasticity Test

```
> # multicollinearity test
> vif(m5)
```

	GVIF	Df	GVIF^(1/(2*Df))
age	1.733436	1	1.316600
sex	1.013785	1	1.006869
children	10.312093	1	3.211245
smoker	10.118613	1	3.180977
smoker:bmiOver30	2.161474	2	1.212516
age:children	10.680614	1	3.268121
age:smoker	8.978807	1	2.996466
children:smoker	2.157729	1	1.468921

```
> sqrt(vif(m5))>10
```

	GVIF	Df	GVIF^(1/(2*Df))
age	FALSE	FALSE	FALSE
sex	FALSE	FALSE	FALSE
children	FALSE	FALSE	FALSE
smoker	FALSE	FALSE	FALSE
smoker:bmiOver30	FALSE	FALSE	FALSE

age:children	FALSE	FALSE	FALSE
age:smoker	FALSE	FALSE	FALSE
children:smoker	FALSE	FALSE	FALSE

=> pass!