<4월 15일 meeting comment 정리>

```
library(boot)
library(tableone)
library(survey)
library(Hmisc)
library(tidyverse)
library(Matching)
delta<-seq(-2,-1,by=0.01) # delta0 sequence
prevalence<-matrix(0,nrow=length(delta),ncol=1)</pre>
## prevalence calculate by changing delta0
for(i in 1:length(delta)){
  delta0<-delta[i]
  delta_b<-0.01
  delta_c<-0.01
  #B,C,U,E generating
  set.seed(123*i)
  exposure_hat<-replicate(100000,expr={
    #set.seed(delta0)
    B < -rbinom(1,1,prob=0.5)
    C < -rnorm(1,0,1)
    U<-runif(1)
    p_z<-inv.logit(delta0+delta_b*B+delta_c*C)
    E < -rbinom(1,1,prob=p_z)
    c(B,C,U,E)
    })
  prevalence[i,]<-mean(exposure_hat[4,])</pre>
}
# exposure ratio
prevalence
min(prevalence)
[1] 0.11981
## Data generating
delta_rare<-delta[which(prevalence==min(prevalence))] # -2</pre>
delta_rare
[1] -2
# sample store
B_sample<-matrix(0,nrow=100,ncol=1000)
```

```
C_sample<-matrix(0,nrow=100,ncol=1000)
U_sample<-matrix(0,nrow=100,ncol=1000)
E_sample<-matrix(0,nrow=100,ncol=1000)
Y_sample<-matrix(0,nrow=100,ncol=1000)
effect < -c(log(1.2), log(1.5), log(2), log(2))
# 1000th replication data generating
for(repl in 1:1000){
  delta0<-delta_rare
  delta_b<-0.01
  delta_c<-0.01
  set.seed(123*repl)
  sample <- replicate (100, expr={
    B < -rbinom(1, 1, prob = 0.5)
    C < -rnorm(1,0,1)
    U<-runif(1)
    p_z<-inv.logit(delta0+delta_b*B+delta_c*C)
    E < -rbinom(1,1,prob=p_z)
    c(B,C,U,E)
    })
  B_sample[,repl]<-sample[1,]
  C_sample[,repl]<-sample[2,]
  U_sample[,repl]<-sample[3,]
  E_sample[,repl]<-sample[4,]
  X\_sample < -cbind (B\_sample[,repl], C\_sample[,repl], U\_sample[,repl], E\_sample[,repl])
  # Y random sampling - continuous
  Y_{\text{sample}}[\text{repl}] \leftarrow \text{as.matrix}(X_{\text{sample}}) * *\% \text{effect+rnorm}(\text{nrow}(X_{\text{sample}}), 0, 1)
## Data store ##
write.csv(B_sample,file="B_sample.csv",row.names=FALSE)
write.csv(C_sample,file="C_sample.csv",row.names=FALSE)
                                                                    ⇒ 생성학 Data
#write.csv(U_sample,file="U_sample.csv",row.names=FALSE)
                                                                      CSV File & 272
write.csv(E_sample,file="E_sample.csv",row.names=FALSE)
write.csv(Y_sample,file="Y_sample.csv",row.names=FALSE)
## Data import ## 3 2 25t Pata Import
B_sample<-read.csv("B_sample.csv",header=TRUE)
```

```
#U_sample<-read.csv("U_sample.csv",header=TRUE)
          E_sample<-read.csv("E_sample.csv",header=TRUE)
          Y_sample<-read.csv("Y_sample.csv",header=TRUE)
          ## Data export function - treat + covariance ##
          #data_export<-function(var_treat,var_cov,data){
             name<-c(var_treat,var_cov)</pre>
             idx_name<-rep(0,length(name))
          #
             for(n in 1:length(name)){
               idx_name[n]<-which(colnames(data)==name[n])
          #
             mydata<-data[idx_name]
             return(mydata)
          #}
          #############
          ## weight generation function ##
          weight_make<-function(var_treat,var_cov,estimate,data){
            result<-list()
            mydata < -data[c(var_treat, var_cov)
                                           만 가지? 오도록,,
            trt<-mydata[,var_treat]
            ps<-glm(trt~,data=mydata,family='binomial')$fitted.values : 4524 get
            #ps<-predigt(psmodel,type='response')</pre>
                        my formula = as formula (Sprin+f("%5~.
            if(estimate=="ATE"){
이렇게 작성하면
              weight<-ifelse(trt==1,1/ps,1/(1-ps))
errora भिषेत
28年28日
            else if(estimate=='ATT'){
              weight<-ifelse(trt==1,1,ps/(1-ps))
            }
                               data[, var-treat] 0/9
              weight<-ifelse(trt==1,(1-ps)/ps,1)
            result[[1]]<-weight
                    ) result $ untrimmed + weight
          result[[2]]<-ifelse(weight>quantile(weight,prob=0.99),quantile(weight,prob=0.99),weight)
            return(result) result $ trimmed &
        trunation 격용한 WeTaht 와 하기 않은 WeTaht 2기기 Version listoil 자칭
```

C_sample<-read.csv("C_sample.csv",header=TRUE)

```
## weighted data export ##
weighteddata < - function (data, weight) {
  require(survey) अने भेगिक्स function
  data<-svydesign(ids=~1,data=data,weights=~weight)
                         resultコト lTst 胡fin
  mydata<-data[[7]]
                           ⇒ 7th on weighted Data >1 918
  return(mydata)
              > 하수사회보다는 계산한 Weight 값을
                       Data의 학명로 추가 ⇒ Hardcoding으로 가중명한, 분산계산 ⇒ SMD 공식 시방방방
                                                                Balance check!
## balance check function - no function ##
#balance_check<-function(var_treat,var_cov,weight,data){
   before_table<-matrix(0,nrow=length(var_cov),ncol=2)
   rownames(before_table)<-var_cov
#
colnames(before_table)<-c("IPTW_before_mean_difference","IPTW_before_variance_rati
0")
#
   after_table<-matrix(0,nrow=length(var_cov),ncol=2)
   rownames(after_table)<-var_cov
#
colnames(after_table)<-c("IPTW_after_mean_difference","IPTW_after_variance_ratio")
# ind_treat<-which(colnames(data)==var_treat)</pre>
# treat<-data[,ind_treat]</pre>
  cov<-data[-ind_treat]
  for(i in 1:length(cov)){
#
     var_cov<-cov[,i]
#
     std_var_cov = (var_cov - mean(var_cov))/sd(var_cov)
     simple_M1 = mean(std_var_cov[treat==1])
#
#
     simple_M0 = mean(std_var_cov[treat==0])
     simple_V1 = var(std_var_cov[treat==1])
#
#
     simple_V0 = var(std_var_cov[treat==0])
#
     wgted_M1 = Hmisc::wtd.mean(x=std_var_cov[treat==1],weights=weight[treat==1])
#
     wgted_M0 = Hmisc::wtd.mean(x=std_var_cov[treat==0],weights=weight[treat==0])
#
     wgted_V1 = Hmisc::wtd.var(x=std_var_cov[treat==1],weights=weight[treat==1])
#
     wgted_V0 = Hmisc::wtd.var(x=std_var_cov[treat==0],weights=weight[treat==0])
#
     before_table[i,1] = simple_M1 - simple_M0
#
     before_table[i,2] = simple_V1/simple_V0
#
     after_table[i,1] = wgted_M1 - wgted_M0
#
     after_table[i,2] = wgted_V1/wgted_V0
# }
```

```
# result<-cbind(before_table,after_table)</pre>
# return(result)
#}
## balance check function - function yes ##
balance_check<-function(var_treat,var_cov,data){
                         les String offen les raw Data
 require(tableone)
 table < - Create Table One (vars = var_cov, strata = var_treat, data = data, test = FALSE)
 print(table,smd=TRUE)
}
## check ## set gut by the Example
E<-E_sample[,1]
B<-B_sample[,1]
#U<-U_sample[,1]
C<-C_sample[,1]
Y<-Y_sample[,1]
data<-data.frame("E"=E,"B"=B,"C"=C,"Y"=Y)
cov<-c("B","C")
#mydata<-data_export("E",cov,data)</pre>
#head(mydata)
tableone_before<-balance_check(var_treat="E",var_cov=cov,data=data)
Weight 409 7! confounder
           Stratified by E
            0
                     1
                              SMD
                       12
              88
 B (mean (SD)) 0.56 (0.50) 0.33 (0.49) 0.451
 C (mean (SD)) -0.04 (0.99) 0.31 (1.10) 0.335
weight_result<-weight_make("E",cov,estimate='ATE',data=data)</pre>
weight_result
[[1]] * Question) Weight가 모두 1이라는 것은 PS-Score 가 매우 각다는 경을 의미 → how to Solve?
 [[2]]
```

```
weighteddata1<-weighteddata(data=data,weight=weight_result[[1]]) #weighteddata[[7]]
         weighteddata2<-weighteddata(data=data,weight=weight_result[[2]])</pre>
         tableone_after1<-balance_check(var_treat="E",var_cov=cov,data=weighteddata1)
                         Stratified by E
                          0
                                       1
                                                    SMD
                             88
                                          12
           n
           B (mean (SD)) 0.56 (0.50) 0.33 (0.49) 0.451
           C (mean (SD)) -0.04 (0.99) 0.31 (1.10) 0.335
         tableone_after2<-balance_check(var_treat="E",var_cov=cov,data=weighteddata2)
                         Stratified by E
                          0
                                                    SMD
                             88
                                          12
           n
           B (mean (SD)) 0.56 (0.50) 0.33 (0.49) 0.451
           C (mean (SD)) -0.04 (0.99) 0.31 (1.10) 0.335
         #### IPTW ATE & ATT ####
         result_ATE<-matrix(0,nrow=1000,ncol=2)</pre>
         colnames(result_ATE)<-c("before_truncation_ATE", "after_truncation_ATE")
         result_ATE<-as.data.frame(result_ATE)
         result_ATT<-matrix(0,nrow=1000,ncol=2)
科龄 化
         colnames(result_ATT)<-c("before_truncation_ATT", "after_truncation_ATT")
거강위해
         result_ATT<-as.data.frame(result_ATT)
matrix 对对
         ## 1000th simulated ##
         for(j in 1:1000){
           E<-E_sample[,j]
           B<-B_sample[,j]
           C<-C_sample[,j]
           #U<-U_sample[,j]
           Y<-Y_sample[,j]
           data<-data.frame("E"=E,"B"=B,"C"=C,"Y"=Y)
           cov<-c("B","C")
              S Confounder 1949 273
           #mydata<-data_export("E",cov,data)</pre>
           weight_ATE<-weight_make("E",cov,estimate='ATE',data=data)</pre>
            weight_ATT<-weight_make("E",cov,estimate='ATT',data=data)</pre>
     Wetgh
           weighteddata1_ATE<-weighteddata(data=data,weight=weight_ATE[[1]])
                              truncation old weight older
                                               pseudo-population Ayy
```

weighted data

```
weighteddata2_ATE<-weighteddata(data=data,weight=weight_ATE[[2]])
  weighteddata1_ATT<-weighteddata(data=data,weight=weight_ATT[[1]])
                                                                       Pseudo-population ANY
  weighteddata2_ATT<-weighteddata(data=data,weight=weight_ATT[[2]])
  result_ATE[j,1]
lm(Y~E,data=weighteddata1_ATE)$coef['E']
  result_ATE[j,2]<-lm(Y~E,data=weighteddata2_ATE)$coef['E']
  result_ATT[j,1]<-lm(Y~E,data=weighteddata1_ATT)$coef['E'
  result_ATT[j,2]<\m/Y~E,data=weighteddata1_ATT)$coef['E']
                        Weighted Data NETHA
}
mean(result_ATE[,1])
[1] 0.6957256
mean(result_ATE[,2])
[1] 0.6957256
mean(result_ATT[,1])
[1] 0.6957256
mean(result_ATT[,2])
[1] 0.6957256
log(2)
[1] 0.6931472
# plotting - ATE #
par(mfrow=c(2,1))
plot(result_ATE[,1],main="ATE estimators using IPTW")
abline(h=log(2),col='red')
plot(result_ATE[,2],main="ATE estimators using IPTW_truncated version")
abline(h=log(2),col='red')
# plotting - ATT #
par(mfrow=c(2,1))
plot(result_ATT[,1],main="ATT estimators using IPTW")
abline(h=log(2),col='red')
plot(result_ATT[,2],main="ATT estimators using IPTW_truncated version")
abline(h=log(2),col='red')
############
```

> truncation to

```
## matched data export ## , Matching & matching & Data return 5124 function
matcheddata<-function(var_treat,var_cov,data,PSM,M){
                           String tolen
                                                 S PS Matching
  require(Matching)
                                                     시행여부
  if(PSM==TRUE){
    mydata<-data[c(var_treat,var_cov)]
    trt<-mydataf, var_treat] > IPTW # my formula 08!
    ps<-glm(tpt~.,data=mydata,family='binomial')$fitted.values # Ps ***
    m.out<-Match(Tr=data[,var_treat],M=M,X=logit(ps),replace=FALSE)
  }
 else{
    m.out<-Match(Tr=data[,var_treat],M=M,X=data[,var_cov],replace=FALSE)
  matched<-data[unlist(m.out[c("index.treated", "index.control")]),]
  return(matched)
                   > Matched 된 Data 말인, Index treated, Index control
                                          Dz return > Dataor Column으로 字r (matching old )
                                                                        OHUE 0으로 (oding)
> ## check ## ! 함 할인 위한 Example
E<-E_sample[,1]
B<-B_sample[,1]
#U<-U_sample[,1]
C<-C_sample[,1]
Y<-Y_sample[,1]
data<-data.frame("E"=E,"B"=B,"C"=C,"Y"=Y)
cov<-c("B","C")
# matched data #
matchdata_greedy<-matcheddata("E",cov,data=data,PSM=FALSE,M=1))
matchdata_PSM<-matcheddata("E",cov,data=data,PSM=TRUE,M=1)
# matched data balance check #
balance_PSM<-balance_check("E",cov,matchdata_PSM)
               Stratified by E
                ()
                                          SMD
                    12
                                12
  n
  B (mean (SD)) 0.50 (0.52) 0.33 (0.49) 0.328
  C (mean (SD)) -0.48 (0.80) 0.31 (1.10) 0.825
balance_greedy<-balance_check("E",cov,matchdata_greedy)
               Stratified by E
                0
                                         SMD
                               12
                   12
  n
```

```
OK!
  B (mean (SD)) 0.33 (0.49) 0.33 (0.49) <0.001 ( greedy matching Est
  C (mean (SD)) 0.21 (0.86) 0.31 (1.10) 0.105
                                                    group 간 balance 정도 맞추어권것 할만!
#### PSM matching & greedy matching ####
PSM_ATE<-c()
greedy ATE<-c()
## 1000th simulated ##
for(j in 1:1000){
  E<-E_sample[,j]
  B<-B_sample[,j]
  C<-C_sample[,j]
  #U<-U_sample[,j]
  Y<-Y_sample[,j]
  data<-data.frame("E"=E,"B"=B,"C"=C,"Y"=Y)
  cov<-c("B","C")
 # matched data generation #
  matchdata_greedy<-matcheddata("E",cov,data=data,PSM=FALSE,M=1)
  matchdata_PSM<-matcheddata("E",cov,data=data,PSM=TRUE,M=1)
PSM_ATE[j]<-matchdata_PSM$Y[matchdata_PSM$E==1]-matchdata_PSM$Y[matchdata
_{PSM}=0
            E[YIE=1] - F[YIE=0] 계산
greedy_ATE[j]<-matchdata_greedy$Y[matchdata_greedy$E==1]-matchdata_greedy$Y[</pre>
matchdata_greedy$E==0]
}
plot(PSM_ATE,main="ATE estimators using PSM")
abline(h=log(2),col='red')
plot(PSM_ATE,main="ATE estimators using greedy matching")
abline(h=log(2),col='red')
#### Regression ####
ATE_regression<-c()
for(j in 1:1000){
  E<-E_sample[,j]
```

```
B<-B_sample[,j]
C<-C_sample[,j]
#U<-U_sample[,j]
Y<-Y_sample[,j]
data<-data.frame("E"=E,"B"=B,"C"=C,"Y"=Y)
cov<-c("B","C")

#mydata<-data_export("E",cov,data)
ATE_regression[j]<-lm(Y~.,data=data)$coef['E']
}

par(mfrow=c(1,1))
plot(ATE_regression,main="ATE estimators using regression")
abline(h=log(2),col='red')
```

어 등은 국로 Bias 와 Variance로 평가 / 수정한 code 마다 저강, 맨위에 수정으로 무엇을 바꾸었는지 간단하게 작성,