

STA 360: Homework 9

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Problem 9.3

The Crime Data

```
set.seed(1)
data = read.table("crime.dat", header = T)
dataNames = names(data)[-1]
library(MASS);
Y.X <- data.matrix(data, rownames.force = NA)
X <- Y.X[, -1]
Y <- Y.X[, 1]
```

Functions

```
getSSR_g <- function(X,Y,g){
  result <- t(Y)%*(diag(length(X[,1]))-(g/(g+1))*X%*%solve(t(X)%*%X)%*%t(X))%*%Y;
}
getBetaOLS <- function(X,Y){
  return(solve(t(X)%*%X)%*%t(X)%*%Y);
}
getSSR <- function(B.ols,X,Y){
  return(t(Y)%*%Y - 2*t(B.ols)%*%t(X)%*%Y + t(B.ols)%*%t(X)%*%X%*%B.ols);
}
getSigma2OLS <- function(B.ols, X, Y){
  SSR = getSSR(B.ols,X,Y);
  s2ols = SSR/(length(X[,1]) - length(X[1,]));
  return(s2ols[1]);
}
getStandardErrors <- function(B, X, sigma2){
  xt.x.inverse <- solve(t(X)%*%X);
  return(sqrt(diag(xt.x.inverse)*sigma2));
}
getUnitInfoVariance<- function(B, X, sigma2, n){
  xt.x.inverse <- solve(t(X)%*%X);
  return(xt.x.inverse*sigma2*n);
}

g <- n <- 47; p <- 15;
nu0 = 2; s0.2 = 1;
beta.ols = getBetaOLS(X,Y);
sigma2.ols = getSigma2OLS(beta.ols, X, Y);
```

Monte Carlo Samples

```
Samples = 10000;
SSR.g = getSSR_g(X,Y,g);
SIGMA2 <- 1/rgamma(Samples, (nu0+n)/2, (nu0*sigma2.ols + SSR.g)/2);
xt.x.inverse <- solve(t(X)%*%X);
BETA <- matrix(0, nrow = length(SIGMA2), ncol = p);
for(i in 1:Samples){
  BETA[i,] <- mvrnorm(1, g*beta.ols/(g+1), g*SIGMA2[i]*xt.x.inverse/(g+1));
}
betaNames <- c()
for(k in 1:15){
  betaNames <- c(betaNames,paste("Beta", toString(k),": ", dataNames[k], sep = ""));
}

Summary = matrix(0, nrow = 15, ncol = 3)
colnames(Summary) = c("Mean", "2.5%", "97.5%")
rownames(Summary) = betaNames

for(k in 1:15){
  Summary[k, 1] <- mean(BETA[,k])
  Summary[k, 2:3] <- quantile(BETA[,k],c(.025, .975))
}
print("Posterior information")
```

```
## [1] "Posterior information"
```

```
print(Summary)
```

```
##              Mean      2.5%      97.5%
## Beta1: M      2.797176e-01 0.05050938 0.51296186
## Beta2: So     -1.517898e-05 -0.30774699 0.30622456
## Beta3: Ed      5.329830e-01 0.23007109 0.82722670
## Beta4: Po1     1.451424e+00 0.05377408 2.84009734
## Beta5: Po2     -7.784345e-01 -2.24513658 0.66757015
## Beta6: LF      -6.557984e-02 -0.32644299 0.19617413
## Beta7: M.F     1.297843e-01 -0.13902793 0.39025129
## Beta8: Pop     -6.735417e-02 -0.27704278 0.14394989
## Beta9: NW      1.099525e-01 -0.18816115 0.39707669
## Beta10: U1     -2.659403e-01 -0.59191958 0.06758902
## Beta11: U2      3.600585e-01 0.04797487 0.66411662
## Beta12: GDP     2.350488e-01 -0.20337761 0.67915579
## Beta13: Ineq    7.098331e-01 0.30746319 1.10507045
## Beta14: Prob   -2.807706e-01 -0.50927734 -0.05355299
## Beta15: Time   -6.198556e-02 -0.27816930 0.15991672
```

```
#Least Squares
```

```
standErrors <- getStandardErrors(beta.ols, X, sigma2.ols)
SummaryOLS = matrix(0, nrow = 15, ncol = 3)
colnames(SummaryOLS) = c("OLS_Mean", "2.5%", "97.5%")
rownames(SummaryOLS) = betaNames
for(k in 1:15){
  SummaryOLS[k,1] <- beta.ols[k];
  SummaryOLS[k,2] <- beta.ols[k]-2*standErrors[k];
  SummaryOLS[k,3] <- beta.ols[k]+2*standErrors[k];
}
```

```

}
print("OLS")

## [1] "OLS"

print(SummaryOLS)

##              OLS_Mean      2.5%      97.5%
## Beta1: M      0.2865177028  0.01919875  0.55383666
## Beta2: So     -0.0001179958 -0.36242447  0.36218848
## Beta3: Ed      0.5445161778  0.19095296  0.89807940
## Beta4: Po1     1.4716146465 -0.13594905  3.07917834
## Beta5: Po2    -0.7817757455 -2.45813939  0.89458790
## Beta6: LF     -0.0659672893 -0.36802793  0.23609335
## Beta7: M.F     0.1313002714 -0.17416954  0.43677009
## Beta8: Pop    -0.0702910179 -0.32013040  0.17954836
## Beta9: NW      0.1090590127 -0.22936245  0.44748048
## Beta10: U1    -0.2705407273 -0.65760644  0.11652499
## Beta11: U2     0.3687335028  0.01468306  0.72278395
## Beta12: GDP    0.2380580097 -0.27169225  0.74780827
## Beta13: Ineq   0.7262918200  0.26536648  1.18721716
## Beta14: Prob  -0.2852262729 -0.54859472 -0.02185783
## Beta15: Time  -0.0615771841 -0.31949841  0.19634404

```

Comparison of OLS and Bayesian Method

Generally speaking, the β vector obtained from the ordinary least squares method and the bayesian method look quite similar in terms of the expected values of $\beta_1, \dots, \beta_{15}$. Moreover, the confidence intervals obtained by both methods look quite similar, however the bayesian method produces narrower confidence intervals for most of the beta values. In addition to this, my Bayesian intervals show that 0 is not in the confidence interval of P_{01} , suggesting that it is a significant predictor of crime rates, however my ordinary least squares interval for P_{01} contains zero, suggesting otherwise.

Explanatory variables which seem to have the strongest relationship with crime rate are M(Percentage of males aged 14-24), Po1 (police expenditure in 1960), Ed (mean years of schooling), U2 (unemployment rate of urban males ages 35-39), Ineq (income inequality), and Prob (probability of imprisonment) with increased probability of imprisonment lowering the crime rate (on average) and increases in all of the other significant explanatory variables leading to an increase in crime rate. This is evidenced by the absense of zero in these explanatory variables' confidence intervals.

9.3b

i

```

#Dividing the data roughly in half
firstHalf <- sample(1:47, 23, replace = F, prob = rep(1/47, 47));
secondHalf = c(setdiff(1:47, firstHalf));
testX <- X[firstHalf,];
testY <- Y[firstHalf];
trainX <- X[secondHalf,];
trainY <- Y[secondHalf];

```

```
trainBeta.ols <- getBetaOLS(trainX, trainY);
print("OLS Betas from Training Data")
```

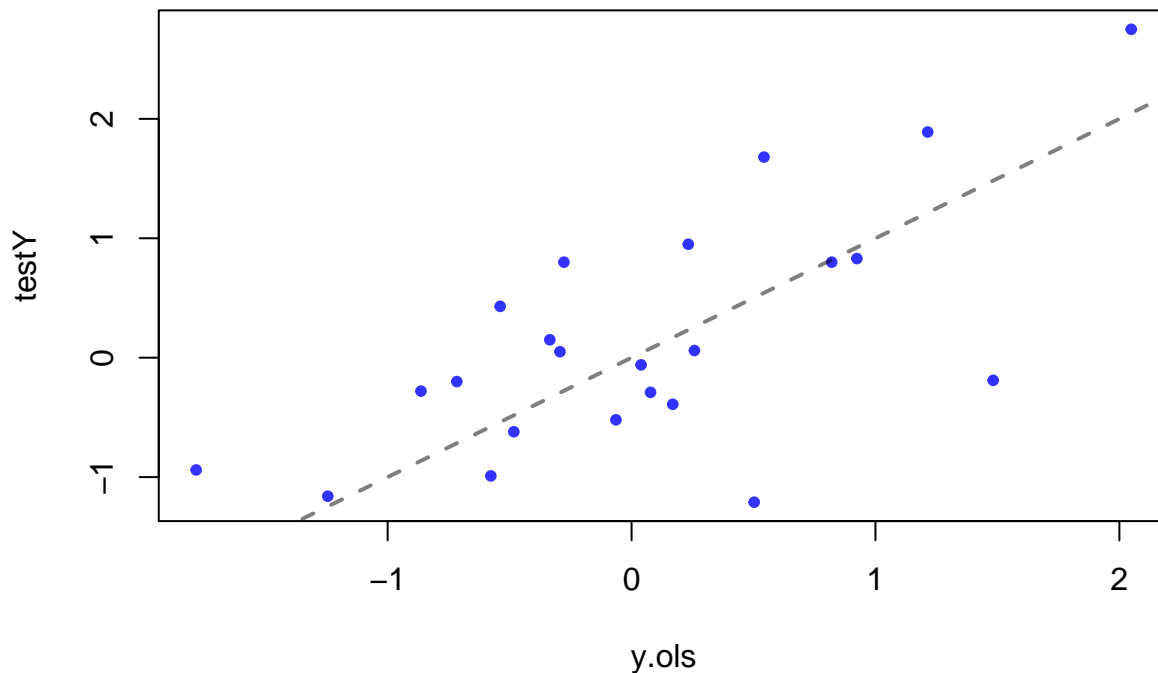
```
## [1] "OLS Betas from Training Data"
```

```
print(trainBeta.ols)
```

```
##           [,1]
## M      0.29513644
## So    -0.06115299
## Ed     0.38931920
## Po1    0.98165066
## Po2   -0.40682969
## LF     0.26909824
## M.F    0.08787996
## Pop   -0.04609962
## NW     0.49767913
## U1     0.06210397
## U2     0.16025710
## GDP    0.27653522
## Ineq   0.42708021
## Prob  -0.07273479
## Time   0.14060766
```

```
y.ols <- testX%%trainBeta.ols;
plot(y.ols, testY, col = rgb(0,0,1,.8), pch = 20, cex=1,
     main = "Y.ols vs Y.test")
abline(a = 0, b = 1, lty = 2, lwd = 2, col = rgb(0,0,0,.5))
```

Y.ols vs Y.test



```
print("OLS Prediction Error");
```

```
## [1] "OLS Prediction Error"
predError <- sum((y.ols - testY)^2)/length(testY);
print(predError)
```

```
## [1] 0.5726231
```

ii

```
g = length(trainY);
trainSigma2.ols <- getSigma2OLS(trainBeta.ols, trainX, trainY);
trainSSR.g = getSSR_g(trainX, trainY, g);
trainSIGMA2 <- 1/rgamma(Samples, (nu0+length(trainY))/2, (nu0*trainSigma2.ols + trainSSR.g)/2);
train.xt.x.inverse <- solve(t(trainX)%*%trainX);
trainBETA <- matrix(0, nrow = length(trainSIGMA2), ncol = p);
for(i in 1:Samples){
  trainBETA[i,] <- mvrnorm(1, g*trainBeta.ols/(g+1),
    g*trainSIGMA2[i]*train.xt.x.inverse/(g+1));
}
trainSummary = matrix(0, nrow = 15, ncol = 3)
colnames(trainSummary) = c("Mean", "2.5%", "97.5%")
rownames(trainSummary) = betaNames

for(k in 1:15){
  trainSummary[k, 1] <- mean(trainBETA[,k])
  trainSummary[k, 2:3] <- quantile(trainBETA[,k], c(.025, .975))
}
print("Trained Beta Values Summary")
```

```
## [1] "Trained Beta Values Summary"
```

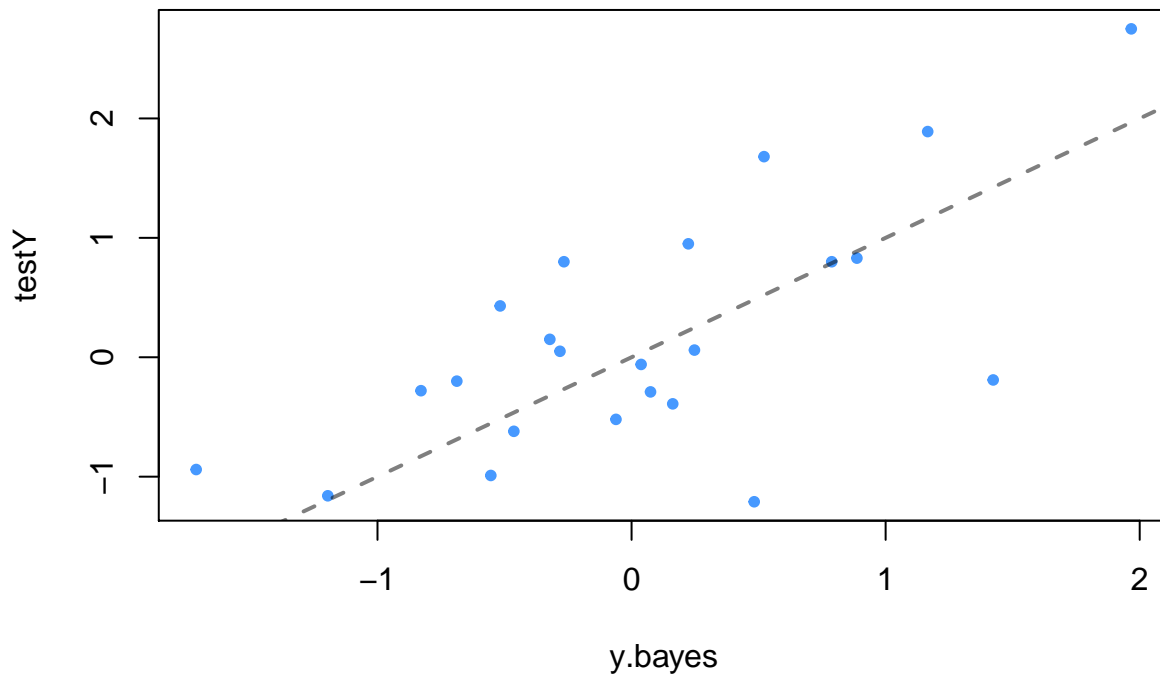
```
print(trainSummary)
```

```
##              Mean      2.5%      97.5%
## Beta1: M      0.27939040 -0.2754566 0.8434932
## Beta2: So     -0.05660868 -0.9116444 0.8375489
## Beta3: Ed      0.37720489 -0.3258930 1.0650089
## Beta4: Po1     0.93895602 -1.0905417 2.9628649
## Beta5: Po2    -0.38893861 -2.6395720 1.8620593
## Beta6: LF      0.25480780 -0.4385561 0.9651008
## Beta7: M.F     0.08590832 -0.3173931 0.4883991
## Beta8: Pop    -0.04346307 -0.4710220 0.3802389
## Beta9: NW      0.47750004 -0.4324989 1.4154862
## Beta10: U1     0.05993442 -0.5354867 0.6504630
## Beta11: U2     0.15300117 -0.3720993 0.6841563
## Beta12: GDP    0.27393239 -0.7878848 1.3056244
## Beta13: Ineq   0.41625927 -0.6109912 1.4145682
## Beta14: Prob  -0.06608849 -0.7237684 0.6005636
## Beta15: Time   0.13478909 -0.2507081 0.5295180
```

```
#Given the g prior
bayes.beta <- trainBeta.ols*g/(g+1)
y.bayes <- testX%*%bayes.beta;
plot(y.bayes, testY, col = rgb(.1,.5,1,.8), pch = 20, cex=1,
  main = "Y.ols vs Y.test")
```

```
abline(a = 0, b = 1, lty = 2, lwd = 2, col = rgb(0,0,0,.5))
```

Y.ols vs Y.test



```
print("Bayes Prediction Error");
```

```
## [1] "Bayes Prediction Error"
```

```
bayes.predError <- sum((y.bayes - testY)^2)/length(testY);
```

```
print(bayes.predError)
```

```
## [1] 0.5605371
```

My prediction error using the bayesian method is slightly lower (about 0.01) lower than my ordinary least squares error. This makes sense given that $\nu_0 = 2$ is quite small and $\frac{47}{48} \approx 1$.

9.3c

For Ordinary Least Squares

```
Runs = 1000;
GROUP_1 <- matrix(0, nrow = Runs, ncol = 23);
GROUP_2 <- matrix(0, nrow = Runs, ncol = 24);
OLS_BETAS <- matrix(0, nrow = Runs, ncol = 15);
PRED_ERROR <- matrix(0, nrow = Runs, ncol = 1);
for(run in 1:Runs){
  GROUP_1[run,] <- sample(1:47, 23, replace = F);
  GROUP_2[run,] <- c(setdiff(1:47, GROUP_1[run,]))
}
for(k in 1:Runs){
  OLS_BETAS[k,] <- getBetaOLS(X[GROUP_1[k,],], Y[GROUP_1[k,]])
```

```

    prediction      <- X[GROUP_2[k,],]%*%OLS_BETAS[k,]
    PRED_ERROR[k,] <- sum((prediction - Y[GROUP_2[k,]])^2/length(GROUP_2[1,]))
  }
ols.average = mean(PRED_ERROR[,1]);
print("Expected Value given OLS");

## [1] "Expected Value given OLS"

print(ols.average);

## [1] 1.00169

print("Confidence interval");

## [1] "Confidence interval"

CI95 = matrix(0, nrow = 1, ncol = 2);
CI95[1,] <- quantile(PRED_ERROR, c(0.025, 0.975));
colnames(CI95) <- c("2.5%", "97.5%");
print(CI95)

##           2.5%    97.5%
## [1,] 0.4292764 2.514804

```

For The Bayesian Approach

```

g = 23;
SmallerSamples <- 1000;
BAYES_ERROR    <- matrix(0, nrow = SmallerSamples, ncol = 1);
for(run in 1:Runs){
  localTrainX <- X[GROUP_1[run,],];
  localTrainY <- Y[GROUP_1[run,]];
  localTestX  <- X[GROUP_2[run,],];
  localTestY  <- Y[GROUP_2[run,]];
  localOlsBeta <- OLS_BETAS[run,];
  localBeta.bayes <- localOlsBeta*g/(g+1);
  pred.bayes <- localTestX*%*%localBeta.bayes
  BAYES_ERROR[run,1] <- sum((pred.bayes - localTestY)^2/length(GROUP_2[1,]))
}
print("Expected Error using Bayesian Method")

## [1] "Expected Error using Bayesian Method"

print(mean(BAYES_ERROR))

## [1] 0.947454

CIBayes95 <- matrix(0, nrow = 1, ncol = 2);
CIBayes95[1,] <- quantile(BAYES_ERROR, c(0.025, 0.975))
print("With CI")

## [1] "With CI"

colnames(CIBayes95) <- c("2.5%", "97.5%");
print(CIBayes95)

##           2.5%    97.5%

```

[1,] 0.415769 2.33486

Hierarchical Modeling

Part 1: The Derivation of Gibbs

Using priors of

$$\begin{aligned}\beta_0 &\sim MVN(\mu, \Lambda) \\ \Sigma_0^{-1} &\sim Wishart(\eta_0, S_0^{-1})\end{aligned}$$

and for each school j , sampling

$$\beta_j \sim MVN(\beta_0, \Sigma_0)$$

and approximating $Y_{i,j}$ (the score of the i^{th} student in the j^{th} school) using the model

$$Y_{i,j} \sim N(\beta_j^T X_{i,j}, \sigma^2)$$

which can be represented as

$$Y_{1:n_j,j} \sim MVN(X_{1:n_j,j} \beta_j, \sigma^2 I)$$

where $Y_{1:n_j,j} \in \mathbb{R}^{n_j \times 1}$ is vector of the scores of each student in school j , and $X_{1:n_j,j} \in \mathbb{R}^{1:n_j \times p}$ is the matrix where the k^{th} row represent the k^{th} student (one of n_j students in school j), of school j , and each column represents a different covariate (one of p) for that student. The $\sigma^2 I$ represents the identity matrix in $I \in \mathbb{R}^{n_j \times n_j}$. From now on, I shall represent $X_{1:n_j,j}$ simply as X_j .

the full conditional distributions of our unknown values can be calculated through the joint distribution

$$p(Y, X, \beta_{1:m}, \beta_0, \sigma^2, \Sigma_0) = \left[\prod_{j=1}^m p(Y_j | \beta_j, X_j, \sigma^2 I) p(\beta_j | \beta_0, \Sigma_0) \right] p(\sigma^2) p(\beta_0) p(\Sigma_0^{-1})$$

Full Conditional of β_j

$$\begin{aligned}p(\beta_j | \dots) &\propto p(Y, X, \beta_{1:m}, \beta_0, \sigma^2, \Sigma_0) \propto p(Y_j | \beta_j, X_j, \sigma^2 I) p(\beta_j | \beta_0, \Sigma_0) \\ &\propto \left(e^{-\frac{1}{2}(\beta_j - \beta_0)^T \Sigma_0^{-1} (\beta_j - \beta_0)} \right) \left(e^{-\frac{1}{2\sigma^2} (Y_j - X_j \beta_j)^T (Y_j - X_j \beta_j)} \right) \propto e^{-\frac{1}{2} \beta_j^T \Sigma_0^{-1} \beta_j + \beta_j^T \Sigma_0^{-1} \beta_0} e^{-\frac{1}{2} \beta_j^T \frac{X_j^T X_j}{\sigma^2} \beta_j + \beta_j^T \frac{X_j^T Y_j}{\sigma^2}} \\ &\propto e^{-\frac{1}{2} \beta_j^T \left(\Sigma_0^{-1} + X_j^T X_j / \sigma^2 \right) \beta_j + \beta_j^T \left(\Sigma_0^{-1} \beta_0 + X_j^T Y_j / \sigma^2 \right)} \\ &\implies (\beta_j | \dots) \sim MVN \left(\left[\Sigma_0^{-1} + X_j^T X_j / \sigma^2 \right]^{-1} \left[\Sigma_0^{-1} \beta_0 + X_j^T Y_j / \sigma^2 \right], \left[\Sigma_0^{-1} + X_j^T X_j / \sigma^2 \right]^{-1} \right)\end{aligned}$$

Full Conditional of β_0

$$\begin{aligned}
p(\beta_0 \mid \dots) &\propto \left[\prod_{j=1}^m p(\beta_j \mid \beta_0, \Sigma_0) \right] p(\beta_0) \\
&\propto \left[e^{-\frac{1}{2} \sum_{j=1}^m (\beta_j - \beta_0)^T \Sigma_0^{-1} (\beta_j - \beta_0)} \right] e^{-\frac{1}{2} (\beta_0 - \mu)^T \Lambda^{-1} (\beta_0 - \mu)} \propto \left(e^{-\frac{1}{2} m \beta_0^T \Sigma_0^{-1} \beta_0 + \beta_0^T m \Sigma_0^{-1} \bar{\beta}} \right) \left(e^{-\frac{1}{2} \beta_0^T \Lambda^{-1} \beta_0 + \beta_0 \Lambda^{-1} \mu} \right) \\
&\propto e^{-\frac{1}{2} \beta_0^T [\Lambda^{-1} + m \Sigma_0^{-1}] \beta_0 + \beta_0^T [\Lambda^{-1} \mu + m \Sigma_0^{-1} \bar{\beta}]} , \quad \bar{\beta} = \frac{1}{m} \sum_{j=1}^m \beta_j \\
&\implies (\beta_0 \mid \dots) \sim MVN([\Lambda^{-1} + m \Sigma_0^{-1}]^{-1} [\Lambda^{-1} \mu + m \Sigma_0^{-1} \bar{\beta}], [\Lambda^{-1} + m \Sigma_0^{-1}]^{-1})
\end{aligned}$$

Full Conditional for σ^2

$$\begin{aligned}
p(1/\sigma^2 \mid \dots) &\propto \left[\prod_{j=1}^m p(Y_j \mid \beta_j, X_j, \sigma^2 I) \right] p(1/\sigma^2) \\
&\propto \left(|I|^{-\frac{\sum_{j=1}^m n_j}{2}} (1/\sigma^2)^{\frac{\sum_{j=1}^m n_j}{2}} e^{-\frac{1}{\sigma^2} \sum_{j=1}^m (Y_j - X_j \beta_j)^T (Y_j - X_j \beta_j)} \right) \left((1/\sigma^2)^{\frac{\nu_0}{2} - 1} e^{-\frac{\nu_0 \sigma_0^2}{2 \sigma^2}} \right) \\
&\propto (1/\sigma^2)^{\frac{\nu_0 + \sum_{j=1}^m n_j}{2} - 1} e^{-\left(\frac{\nu_0 \sigma_0^2 + \sum_{j=1}^m (Y_j - X_j \beta_j)^T (Y_j - X_j \beta_j)}{2} \right) \frac{1}{\sigma^2}} \\
&\implies (\sigma^2 \mid \dots) \sim \text{Inv-Gamma} \left(\frac{\nu_0 + \sum_{j=1}^m n_j}{2}, \frac{\nu_0 \sigma_0^2 + \sum_{j=1}^m (Y_j - X_j \beta_j)^T (Y_j - X_j \beta_j)}{2} \right)
\end{aligned}$$

Full Conditional for Σ_0^{-1}

$$\begin{aligned}
p(\Sigma_0^{-1} \mid \dots) &\propto \left[\prod_{j=1}^m p(\beta_j \mid \beta_0, \Sigma_0) \right] p(\Sigma_0^{-1}) \\
&\propto \left(|\Sigma_0^{-1}|^{\frac{m}{2}} e^{-\frac{1}{2} \sum_{j=1}^m (\beta_j - \beta_0)^T \Sigma_0^{-1} (\beta_j - \beta_0)} \right) \left(|\Sigma_0^{-1}|^{\frac{\eta_0 - p - 1}{2}} e^{-\frac{1}{2} \text{tr}(S_0 \Sigma_0^{-1})} \right) \propto |\Sigma_0^{-1}|^{\frac{(\eta_0 + m) - p - 1}{2}} e^{-\frac{1}{2} \text{tr}([S_0 + S_\beta] \Sigma_0^{-1})} \\
&\quad S_\beta = \sum_{j=1}^m (\beta_j - \beta_0)(\beta_j - \beta_0)^T \\
&\implies (\Sigma_0^{-1} \mid \dots) \sim \text{Wishart}(\eta_0 + m, [S_0 + S_\beta]^{-1})
\end{aligned}$$

Gibbs

Starting with a gift from an oracle of initial values of $(\sigma^2)^{(0)}$, $(\Sigma_0^{-1})^{(0)}$, and $\beta_0^{(0)}$

(1) For j in $1 : m$

$$\rightarrow \beta_j^{(s+1)} \sim \text{MVN} \left([(\Sigma_0^{-1})^{(s)} + X_j^T X_j / (\sigma^2)^{(s)}]^{-1} [(\Sigma_0^{-1})^{(s)} \beta_0^{(s)} + X_j^T Y_j / (\sigma^2)^{(s)}], [(\Sigma_0^{-1})^{(s)} + X_j^T X_j / (\sigma^2)^{(s)}]^{-1} \right)$$

(2)

$$\rightarrow (\sigma^2)^{(s+1)} \sim \text{Inv-Gamma}\left(\frac{\nu_0 + \sum_{j=1}^m n_j}{2}, \frac{\nu_0 \sigma_0^2 + \sum_{j=1}^m (Y_j - X_j \beta_j^{(s+1)})^T (Y_j - X_j \beta_j^{(s+1)})}{2}\right)$$

(3)

$$\rightarrow \beta_0^{(s+1)} \sim \text{MVN}([\Lambda^{-1} + m(\Sigma_0^{-1})^{(s)}]^{-1}[\Lambda^{-1}\mu + m(\Sigma_0^{-1})^{(s)}(\bar{\beta})^{(s+1)}], [\Lambda^{-1} + m(\Sigma_0^{-1})^{(s)}]^{-1})$$

(4)

$$\rightarrow (\Sigma_0^{-1})^{(s+1)} \sim \text{Wishart}(\eta_0 + m, [S_0 + S_\beta^{(s+1)}]^{-1})$$

Part II: The Implementation

Functions

```
library(MASS)
generateX <- function(Nk, p){
  Covariance <- matrix(.2, nrow = p, ncol = p) + diag(p)*.8;
  X <- matrix(0, nrow = 1, ncol = p);
  for(i in 1:Nk){
    Row_i <- mvrnorm(1, rep(0, p), Covariance);
    X <- rbind(X, Row_i);
  }
  X <- X[-1,]
  X[,1] = 1;
  return(X);
}
generateY <- function(X, B, Nk, sigma.2){
  XB <- X%*%B;
  p = length(X[,1]);
  Y <- mvrnorm(1, XB, sigma.2*diag(p));
  return(Y);
}
calcualeS_b <- function(beta0, MCMC.Betas, m, s){
  S_b <- matrix(0, nrow = length(beta0), ncol = length(beta0));
  for(j in 1:m){
    Bj <- MCMC.Betas[[j]][s,];
    S_b <- S_b + (Bj - beta0)%*%t(Bj - beta0);
  }
  return(S_b)
}
updateGamma <- function(dataY, dataX, MCMC.Betas, m, s){
  Build <- 0;
  for(j in 1:m){
    Xj <- dataX[[j]];
    Yj <- dataY[[j]];
    Bj <- MCMC.Betas[[j]][s,];
```

```

    Build <- Build + t(Yj - Xj**Bj)**(Yj - Xj**Bj);
  }
  return(Build);
}

```

Generate Data

```

N <- c(1:10)*10; p <- 10; m <- 10;
beta0 <- rep(0, p); sigma.2 <- 1; Sigma0 = diag(p);
dataX <- list();
dataBetas <- matrix(0, nrow = m, ncol = p);
dataY <- list();
for(j in 1:m){
  Xj <- generateX(N[j],p);
  dataX[[j]] <- Xj;
  Bj <- mvrnorm(1, beta0, Sigma0);
  dataBetas[j,] <- Bj;
  Yj <- generateY(Xj, Bj, N[j], sigma.2);
  dataY[[j]] <- Yj;
}

```

Running the Gibbs Sampler

```

Lambda.inv <- diag(p); eta0 <- 2; S0 <- diag(p); nu0 <- 1; sigma0.2 <- 1;
mu <- rep(0, p);
Samples <- 10000;
MCMC.Beta0 <- matrix(0, nrow = Samples, ncol = p);
MCMC.sigma2 <- rep(0, Samples);
MCMC.Sigma0 <- matrix(0, nrow = Samples, ncol = p*p);
MCMC.Betas <- list();
for(j in 1:m){
  MCMC.Betas[[j]] <- matrix(0, nrow = Samples, ncol = p);
}
#Initial Values
sigma.2 <- 1; beta0 <- rep(0, p); Sigma0 <- diag(p);
#Run the Gibbs
for(s in 1:Samples){
  #Betas 1, ..., m
  Sigma0.inv <- solve(Sigma0);
  betaSums <- rep(0, p);
  for(j in 1:m){
    Xj <- dataX[[j]]; Yj <- dataY[[j]];
    Var.j <- solve(Sigma0.inv + t(Xj)**Xj/sigma.2);
    Mean.j <- Var.j**(Sigma0.inv**beta0 + t(Xj)**Yj/sigma.2);
    Beta.j <- mvrnorm(1, Mean.j, Var.j)
    betaSums <- betaSums + Beta.j;
    MCMC.Betas[[j]][s,] <- Beta.j;
  }
  betaAverage <- betaSums/m;
}

```

```

#sigma.2
A <- (nu0 + sum(N))/2;
B <- (nu0*sigma0.2 + updateGamma(dataY, dataX, MCMC.Betas, m, s))/2;
sigma.2 <- 1/rgamma(1, A, B);
MCMC.sigma2[s] <- sigma.2;
#Beta0
beta0.var <- solve(Lambda.inv + m*Sigma0.inv);
beta0.mean <- beta0.var%*%(Lambda.inv%*%mu+m*Sigma0.inv%*%betaAverage);

beta0 <- mvrnorm(1, beta0.mean, beta0.var);
MCMC.Beta0[s,] <- beta0;
#Sigma0
Sigma0 <- solve(rWishart(1,eta0+m, solve(S0 + calculteS_b(beta0, MCMC.Betas,m, s)))[,1])
prepSigma0 <- matrix(Sigma0, nrow = 1, ncol = p*p);
MCMC.Sigma0[s,] <- prepSigma0;
}

```

Part III: Assessing the Convergence

Beta Values of each school

Below, I plot the MCMC average and an MCMC 95% confidence interval for each of the beta values 1,...,10 for each school 1,...,10. I also plot what beta value was actually used to obtain the data (titled “actual”) and the difference between my MCMC expected value and the actual value, and the standard error of each variable as the square root of the variance of the sample divided by the effective sample size. As shown in the boxplots, the beta values seem to have reached a stationary distribution given that the distributions are not experiences large movements or changes over time.

```

library(coda)
betaNames <- c()
for(k in 1:m){
  betaNames <- c(betaNames,paste("Beta", toString(k), sep = ""));
}
for(j in 1:m){
  Summary = matrix(0, nrow = p, ncol = 6)
  colnames(Summary) = c("Actual", "Mean", "2.5%", "97.5%", "Difference %", "Standard Error")
  rownames(Summary) = betaNames
  for(k in 1:p){
    Summary[k, 1] <- dataBetas[j,k];
    Summary[k, 2] <- mean(MCMC.Betas[[j]][,k]);
    Summary[k, 3:4] <- quantile(MCMC.Betas[[j]][,k],c(.025, .975));
    Summary[k, 5] <- (Summary[k, 1] - Summary[k,2])/Summary[k,1]*100;
    Summary[k, 6] <- sqrt(var(MCMC.Betas[[j]][,k])/effectiveSize(MCMC.Betas[[j]][,k]));
  }
  par(mfrow = c(2,5));
  print(paste("For School ", toString(j)));
  print(Summary)
  print(paste("Boxplots of the convergence for B_", toString(m), ", For each specific Beta"))
  for(k in 1:p){
    boxplot(MCMC.Betas[[j]][,p][1:2000], MCMC.Betas[[j]][,p][2001:4000], MCMC.Betas[[j]][,p][4001:6000])
  }
}

```

```

MCMC.Betas[[j]][,p][6001:8000], MCMC.Betas[[j]][,p][8000:10000])
}
}

```

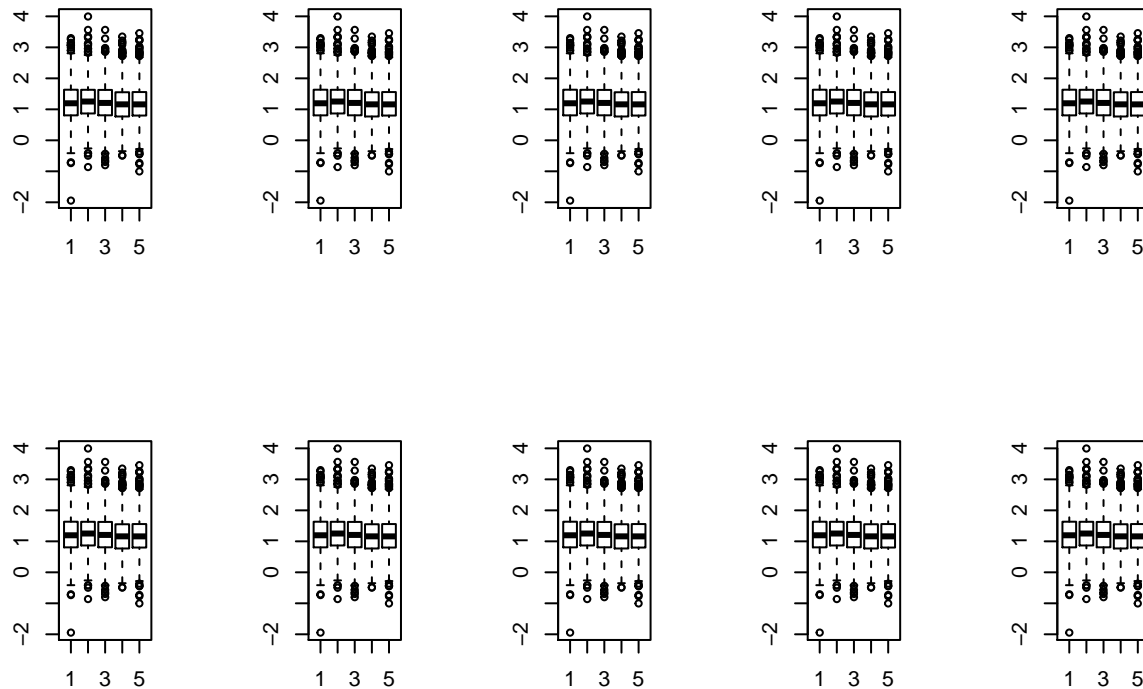
```
## [1] "For School 1"
```

	Actual	Mean	2.5%	97.5%	Difference %
Beta1	0.519232900	0.92731027	0.1748990	1.66800764	-78.59236
Beta2	0.088852486	-1.59686956	-3.1512436	-0.15910118	1897.21428
Beta3	-1.126787356	-0.91714508	-1.6881570	-0.17283148	18.60531
Beta4	-1.427552520	0.39997336	-0.7813500	1.52544890	128.01812
Beta5	-0.005831412	0.05697641	-1.0643559	1.07009332	1077.06028
Beta6	0.480801410	0.35432865	-0.5899421	1.33623195	26.30457
Beta7	1.749163934	1.03853446	0.3358297	1.75599344	40.62681
Beta8	0.189239838	-1.58087278	-3.3433568	0.09416411	935.38054
Beta9	0.856832333	0.75826848	-0.1423133	1.76738004	11.50328
Beta10	0.937914161	1.21588506	0.1116477	2.43496515	-29.63714

```
## Standard Error
```

Beta1	0.004993955
Beta2	0.011992020
Beta3	0.005557258
Beta4	0.009786467
Beta5	0.009388476
Beta6	0.007605691
Beta7	0.004824568
Beta8	0.012957533
Beta9	0.008148100
Beta10	0.008666417

```
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```

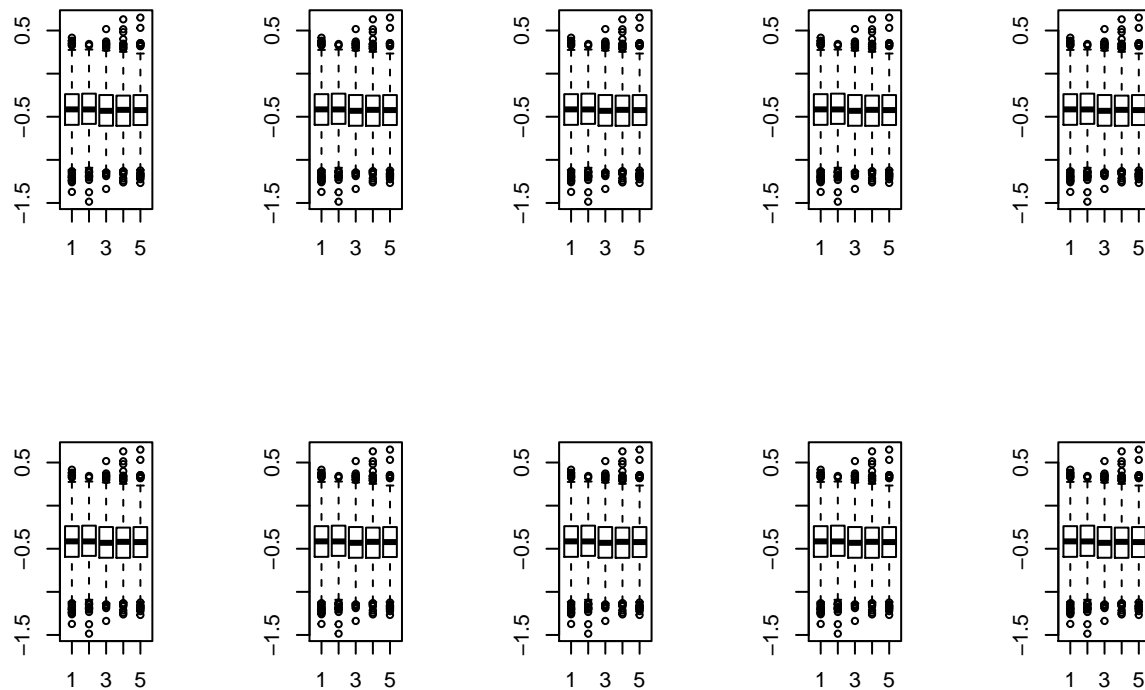


```
## [1] "For School 2"
```

	Actual	Mean	2.5%	97.5%	Difference %
Beta1	-0.9348537	-1.13053472	-1.67586067	-0.58505650	-20.93173

```
## Beta2 -0.6292481 -0.79869642 -1.53880138 -0.06565357 -26.92870
## Beta3 0.7973415 0.56751447 0.01707696 1.13409088 28.82416
## Beta4 -1.0961474 -0.43818598 -1.15732083 0.27611031 60.02490
## Beta5 -0.1383049 0.07829095 -0.72727338 0.89287320 156.60752
## Beta6 0.5669420 0.71183842 0.05339262 1.37853325 -25.55755
## Beta7 -1.6522887 -1.82479703 -2.79582529 -0.87196355 -10.44057
## Beta8 0.6589718 0.74307257 0.28821794 1.18264636 -12.76242
## Beta9 -0.0861814 -0.10817320 -0.53733910 0.32292986 -25.51803
## Beta10 -0.3548020 -0.42328662 -0.94841650 0.08553807 -19.30221
## Standard Error
## Beta1 0.003190836
## Beta2 0.005570460
## Beta3 0.003990478
## Beta4 0.005207830
## Beta5 0.006603572
## Beta6 0.004467097
## Beta7 0.007162751
## Beta8 0.002794076
## Beta9 0.002843435
## Beta10 0.003177251
```

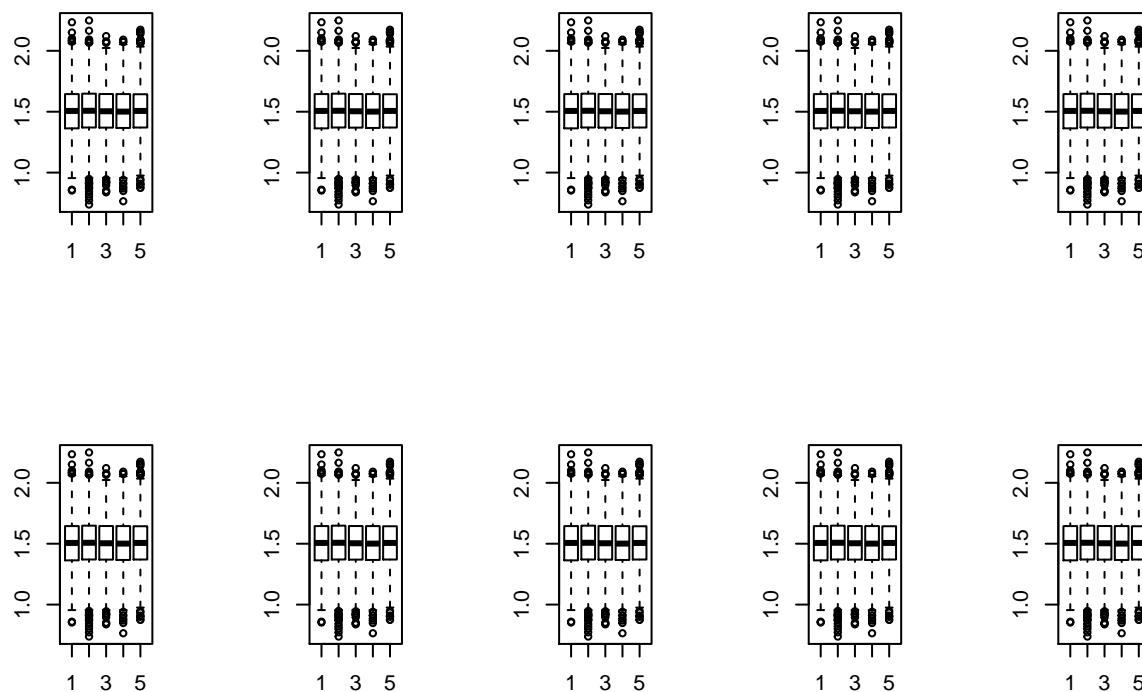
```
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```



```
## [1] "For School 3"
```

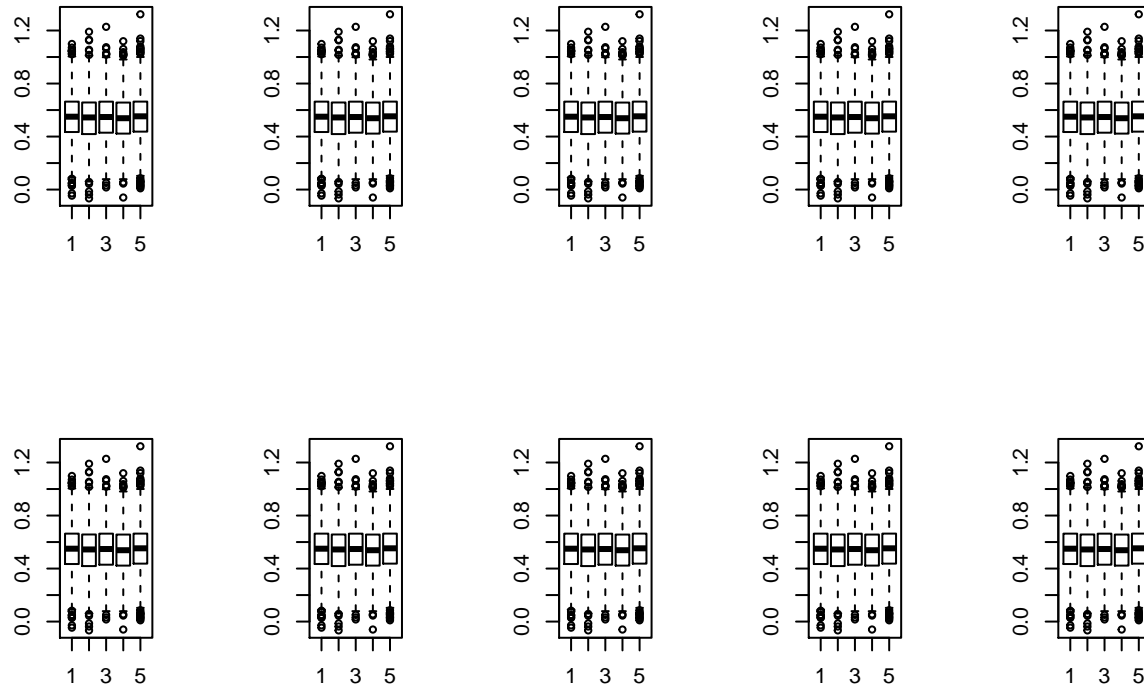
	Actual	Mean	2.5%	97.5%	Difference %
Beta1	0.1457298	0.06956212	-0.3558259	0.5079099	52.266378
Beta2	0.9601534	0.98680820	0.5692211	1.4053055	-2.776099
Beta3	0.4485389	0.32813607	-0.1843375	0.8315370	26.843339
Beta4	-0.2379138	-0.48788697	-0.8759205	-0.0980955	-105.068833
Beta5	-0.5147511	-0.71264750	-1.2958631	-0.1450362	-38.445058
Beta6	0.4493605	0.71214935	0.3363172	1.0931337	-58.480635
Beta7	1.1414098	1.31820271	0.8307345	1.8021782	-15.489000
Beta8	-1.3798826	-1.20795995	-1.5556789	-0.8511520	12.459225

```
## Beta9 -1.4105962 -1.92170248 -2.4087193 -1.4388710 -36.233351
## Beta10 1.5377685 1.50608235 1.1015927 1.9126763 2.060528
##      Standard Error
## Beta1 0.002516622
## Beta2 0.002367522
## Beta3 0.002985884
## Beta4 0.002164961
## Beta5 0.003376872
## Beta6 0.002195447
## Beta7 0.002769760
## Beta8 0.001983600
## Beta9 0.002824575
## Beta10 0.002159828
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```

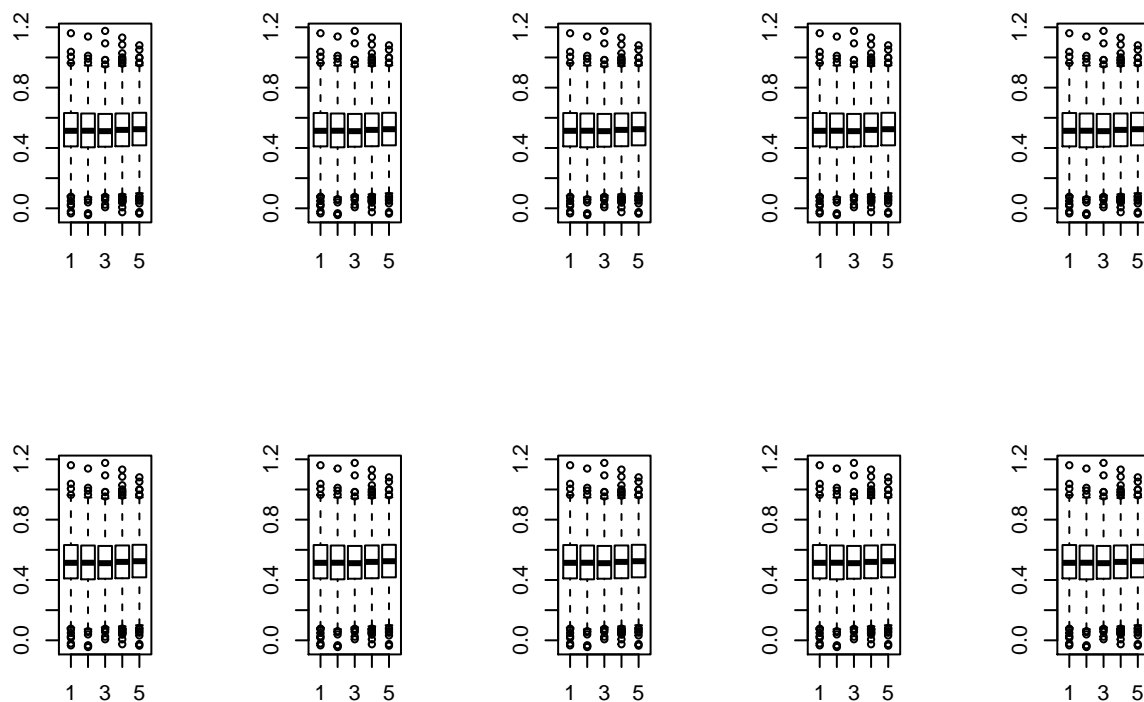


```
## [1] "For School 4"
##      Actual      Mean      2.5%      97.5% Difference %
## Beta1 2.20078106 2.0363121 1.6845794 2.3830286 7.473207
## Beta2 -1.34872828 -1.3925066 -1.7938357 -0.9964340 -3.245894
## Beta3 -0.65423519 -0.7758996 -1.1529016 -0.4050029 -18.596437
## Beta4 0.68051268 0.7519082 0.3801733 1.1202888 -10.491432
## Beta5 1.41392145 1.4283886 0.8439261 2.0135778 -1.023196
## Beta6 -0.05019354 -0.2029185 -0.6015503 0.2055100 -304.272063
## Beta7 1.57945954 1.6424671 1.1337033 2.1536104 -3.989183
## Beta8 -0.71364665 -0.7485914 -1.0841493 -0.4125378 -4.896644
## Beta9 0.65923422 0.9903267 0.6697265 1.3040419 -50.223802
## Beta10 0.67426770 0.5448073 0.1975371 0.8822639 19.200145
##      Standard Error
## Beta1 0.001994208
## Beta2 0.002593176
## Beta3 0.002312574
## Beta4 0.002371436
```

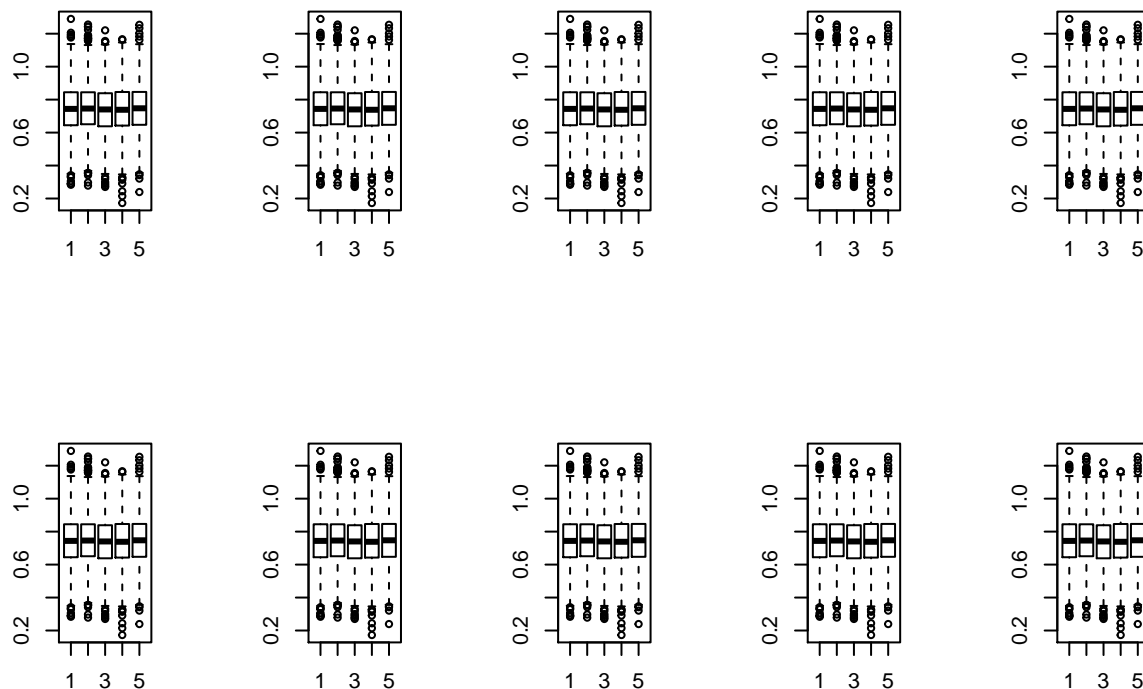
```
## Beta5      0.004228346
## Beta6      0.002434239
## Beta7      0.003623331
## Beta8      0.001837888
## Beta9      0.001922467
## Beta10     0.002085276
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```



```
## [1] "For School 5"
##      Actual      Mean      2.5%      97.5% Difference %
## Beta1 -2.1075154 -2.17599153 -2.4668178 -1.88325574 -3.249140
## Beta2  1.2629342  1.10798307  0.7880575  1.43359089  12.269138
## Beta3  0.6688175  0.73074705  0.4070425  1.05888971 -9.259555
## Beta4  1.2886718  1.32657195  0.9601763  1.68313648 -2.941025
## Beta5  1.1261540  1.06828296  0.7821985  1.34646401  5.138822
## Beta6  1.5187252  1.53879151  1.2504664  1.82697042 -1.321258
## Beta7  0.3945978  0.53320743  0.2197399  0.84327982 -35.126799
## Beta8 -0.6389493 -0.39370942 -0.7214326 -0.06216811  38.381739
## Beta9 -0.3126930 -0.02973844 -0.3513043  0.29894481  90.489576
## Beta10 0.3663076  0.51921734  0.1999914  0.84071954 -41.743537
##      Standard Error
## Beta1      0.001618933
## Beta2      0.001834312
## Beta3      0.001995988
## Beta4      0.002035877
## Beta5      0.001581895
## Beta6      0.001594858
## Beta7      0.001794524
## Beta8      0.001791070
## Beta9      0.001833565
## Beta10     0.001986575
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```

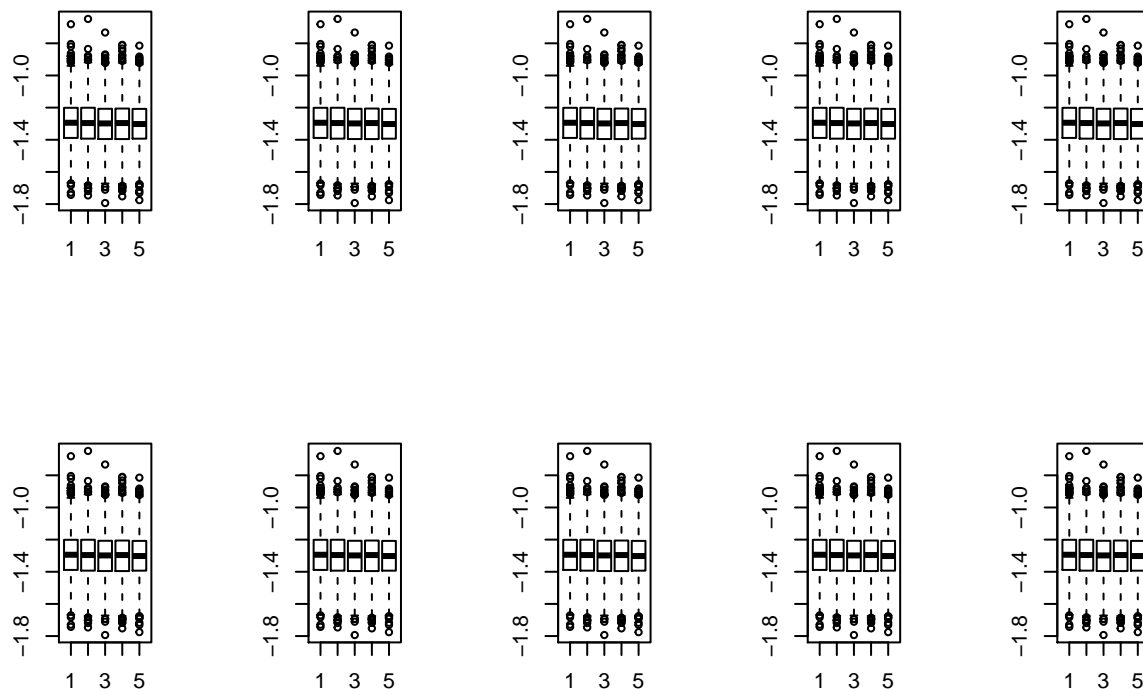
```
## [1] "For School 6"
##           Actual      Mean      2.5%      97.5% Difference %
## Beta1  -0.4321404 -0.2817322 -0.5595491 -0.001358495  34.8054048
## Beta2   0.2953645  0.4749042  0.1957570  0.756605380 -60.7857825
## Beta3   0.5838039  0.5161877  0.2824079  0.752560097  11.5820111
## Beta4  -0.8630841 -0.5961750 -0.8614140 -0.330057289  30.9250359
## Beta5   1.2180149  0.9385340  0.6433931  1.235207874  22.9456014
## Beta6  -0.3273892 -0.3193704 -0.5995298 -0.043225349   2.4493303
## Beta7   1.1433698  1.1525943  0.8604335  1.444620395  -0.8067846
## Beta8   0.3815491  0.4330258  0.1574677  0.715813637 -13.4914854
## Beta9  -1.1144790 -0.9207046 -1.1850086 -0.649081337  17.3869989
## Beta10  0.8292392  0.7448627  0.4586898  1.035083683  10.1751773
##           Standard Error
## Beta1      0.001545908
## Beta2      0.001517713
## Beta3      0.001271350
## Beta4      0.001461964
## Beta5      0.001705992
## Beta6      0.001467553
## Beta7      0.001541088
## Beta8      0.001433680
## Beta9      0.001433258
## Beta10     0.001585299
## [1] "Boxplots of the convergence for B_ 10 , For each specific Beta"
```



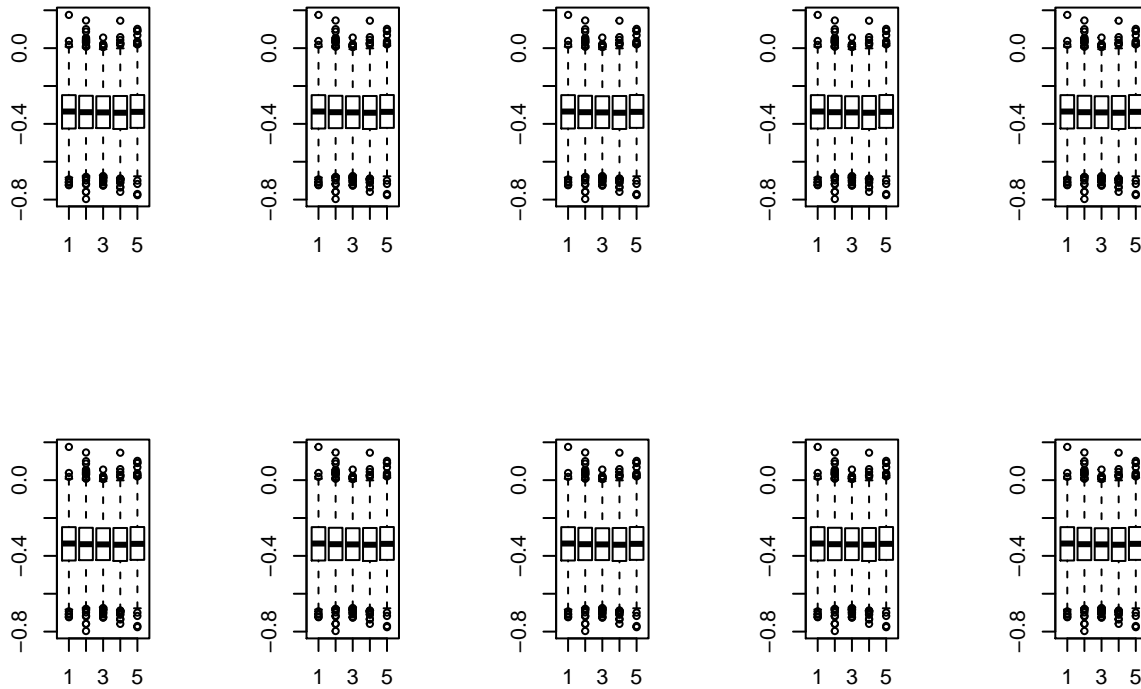
```
## [1] "For School 7"
```

	Actual	Mean	2.5%	97.5%	Difference %
Beta1	-0.7423270	-0.8135814	-1.0597069	-0.5620083	-9.598793551
Beta2	1.7150315	1.5582345	1.2939556	1.8195122	9.142516136
Beta3	-1.5917220	-1.4412941	-1.6965854	-1.1900649	9.450643639
Beta4	-0.7642565	-0.8083004	-1.0816178	-0.5378244	-5.762966193
Beta5	-2.2065053	-2.3026394	-2.5502435	-2.0539116	-4.356850431
Beta6	-0.7349218	-0.5724435	-0.8579993	-0.2872428	22.108246322
Beta7	1.0006059	0.8955869	0.5819896	1.2133049	10.495534709
Beta8	1.0082431	0.9931948	0.7068564	1.2761697	1.492532903
Beta9	-1.8777272	-1.8776119	-2.1716942	-1.5840906	0.006141473
Beta10	-1.2321601	-1.2980507	-1.5686687	-1.0222377	-5.347565963

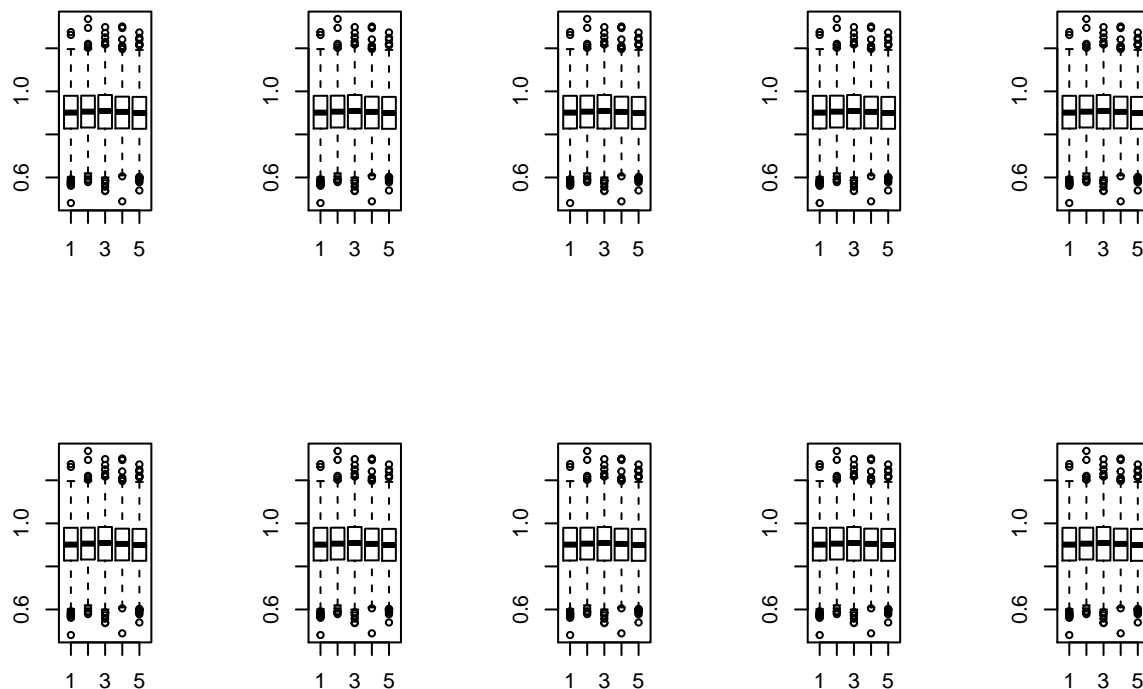
```
##      Standard Error
## Beta1      0.001287807
## Beta2      0.001477026
## Beta3      0.001363746
## Beta4      0.001490657
## Beta5      0.001417927
## Beta6      0.001515486
## Beta7      0.001797570
## Beta8      0.001510914
## Beta9      0.001627103
## Beta10     0.001483780
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```



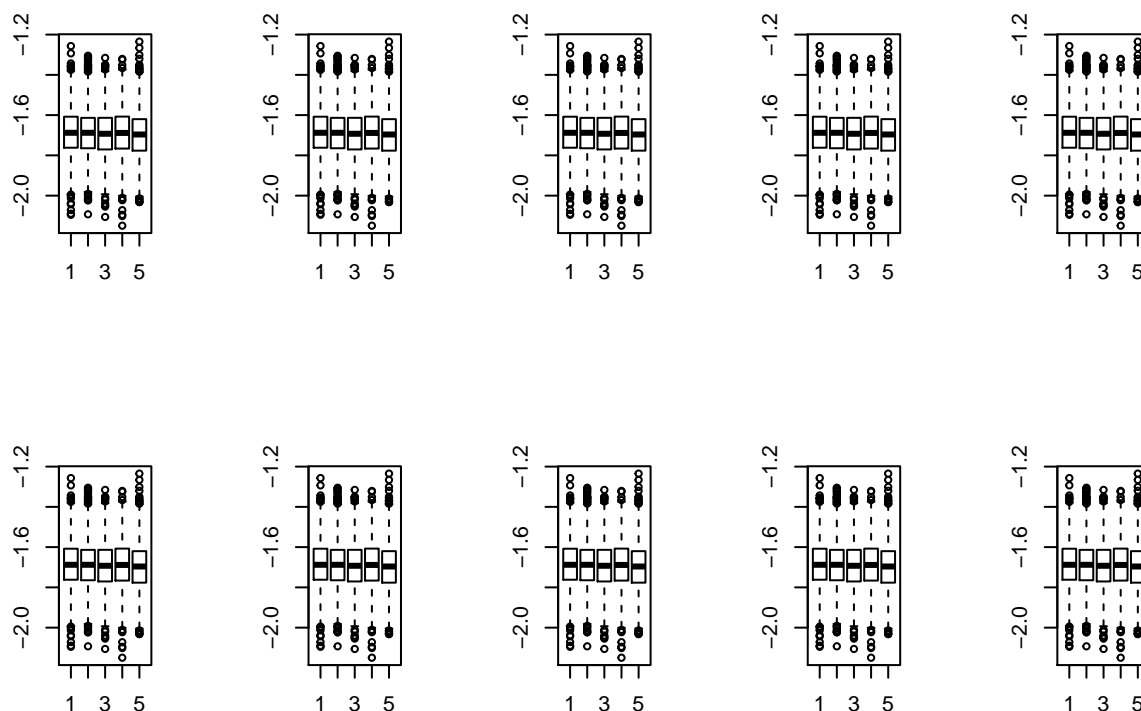
```
## [1] "For School 8"
##           Actual      Mean      2.5%      97.5% Difference %
## Beta1  -0.95737569 -1.11537370 -1.3380809 -0.895153447  -16.503240
## Beta2  -0.23390465 -0.23964805 -0.4792352 -0.004060667   -2.455441
## Beta3   0.17016795 -0.09019615 -0.3479227  0.166140879  153.004196
## Beta4   0.52633096  0.62045360  0.4087780  0.829427531  -17.882787
## Beta5   0.86562423  0.94004577  0.7182160  1.163194304   -8.597441
## Beta6  -0.74917472 -0.77306884 -1.0193983 -0.524140180   -3.189393
## Beta7  -0.04272028 -0.14711651 -0.3917135  0.095228251 -244.371643
## Beta8  -1.20318984 -1.21536039 -1.4729020 -0.955065182   -1.011524
## Beta9   0.82550669  0.64985396  0.4135003  0.891310643   21.278171
## Beta10 -0.24719869 -0.33724668 -0.5904039 -0.085407559  -36.427374
##           Standard Error
## Beta1      0.001161961
## Beta2      0.001259121
## Beta3      0.001442220
## Beta4      0.001094801
## Beta5      0.001172922
## Beta6      0.001377053
## Beta7      0.001292998
## Beta8      0.001347581
## Beta9      0.001323609
## Beta10     0.001359888
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```



```
## [1] "For School 9"
##           Actual           Mean           2.5%           97.5% Difference %
## Beta1    0.42909843  0.502469732  0.2877525  0.7158266  -17.098944
## Beta2   -1.31081463 -1.380512491 -1.6213831 -1.1374663   -5.317141
## Beta3   -0.47988434 -0.515246433 -0.7176376 -0.3060739   -7.368879
## Beta4    0.08737966 -0.001661287 -0.2542578  0.2459222  101.901229
## Beta5   -0.98818192 -1.238202095 -1.5045531 -0.9754524  -25.301027
## Beta6    0.46749524  0.487206381  0.2646199  0.7033931   -4.216331
## Beta7   -1.24364121 -1.039028439 -1.2689484 -0.8056929   16.452717
## Beta8   -1.67134690 -1.688462280 -1.9047320 -1.4722299   -1.024047
## Beta9   -0.96277190 -0.990078752 -1.2030279 -0.7764755   -2.836274
## Beta10   0.71726403  0.903161707  0.6870032  1.1170965  -25.917607
##           Standard Error
## Beta1      0.001136709
## Beta2      0.001321243
## Beta3      0.001125366
## Beta4      0.001338520
## Beta5      0.001486354
## Beta6      0.001141948
## Beta7      0.001245238
## Beta8      0.001154183
## Beta9      0.001183860
## Beta10     0.001130383
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```



```
## [1] "For School 10"
##           Actual      Mean      2.5%      97.5% Difference %
## Beta1    0.25985556  0.26687257  0.07325226  0.4643073   -2.700351
## Beta2   -0.21924510  0.04002418 -0.18567538  0.2661480  118.255448
## Beta3   -0.68082173 -0.61479119 -0.88377047 -0.3500978   9.698654
## Beta4    0.90294177  0.97652353  0.74537122  1.2090221  -8.149115
## Beta5    0.17856163  0.23334177  0.03766459  0.4261578 -30.678565
## Beta6   -0.88852376 -1.17141349 -1.36936387 -0.9706997 -31.838173
## Beta7    0.04947061 -0.08070644 -0.28220169  0.1264649 263.140167
## Beta8    1.45787754  1.33862768  1.11125031  1.5595188   8.179690
## Beta9   -0.05710358  0.09953941 -0.15476453  0.3565885 274.313780
## Beta10  -1.45505113 -1.68926522 -1.91315913 -1.4597998 -16.096623
##           Standard Error
## Beta1      0.001037128
## Beta2      0.001213928
## Beta3      0.001496707
## Beta4      0.001236008
## Beta5      0.000991652
## Beta6      0.001057761
## Beta7      0.001121266
## Beta8      0.001187538
## Beta9      0.001380943
## Beta10     0.001233325
## [1] "Boxplots of the convergence for B_10 , For each specific Beta"
```



β_0

```
Summary = matrix(0, nrow = p, ncol = 5)
colnames(Summary) = c("Actual", "Mean", "2.5%", "97.5%", "Standard Error")
rownames(Summary) = betaNames
for(k in 1:p){
  Summary[k, 1] <- 0;
  Summary[k, 2] <- mean(MCMC.Beta0[,k]);
  Summary[k, 3:4] <- quantile(MCMC.Beta0[,k],c(.025, .975));
  Summary[k, 5] <- sqrt(var(MCMC.Beta0[,k])/effectiveSize(MCMC.Beta0[,k]));
}
print("For Beta_0");
```

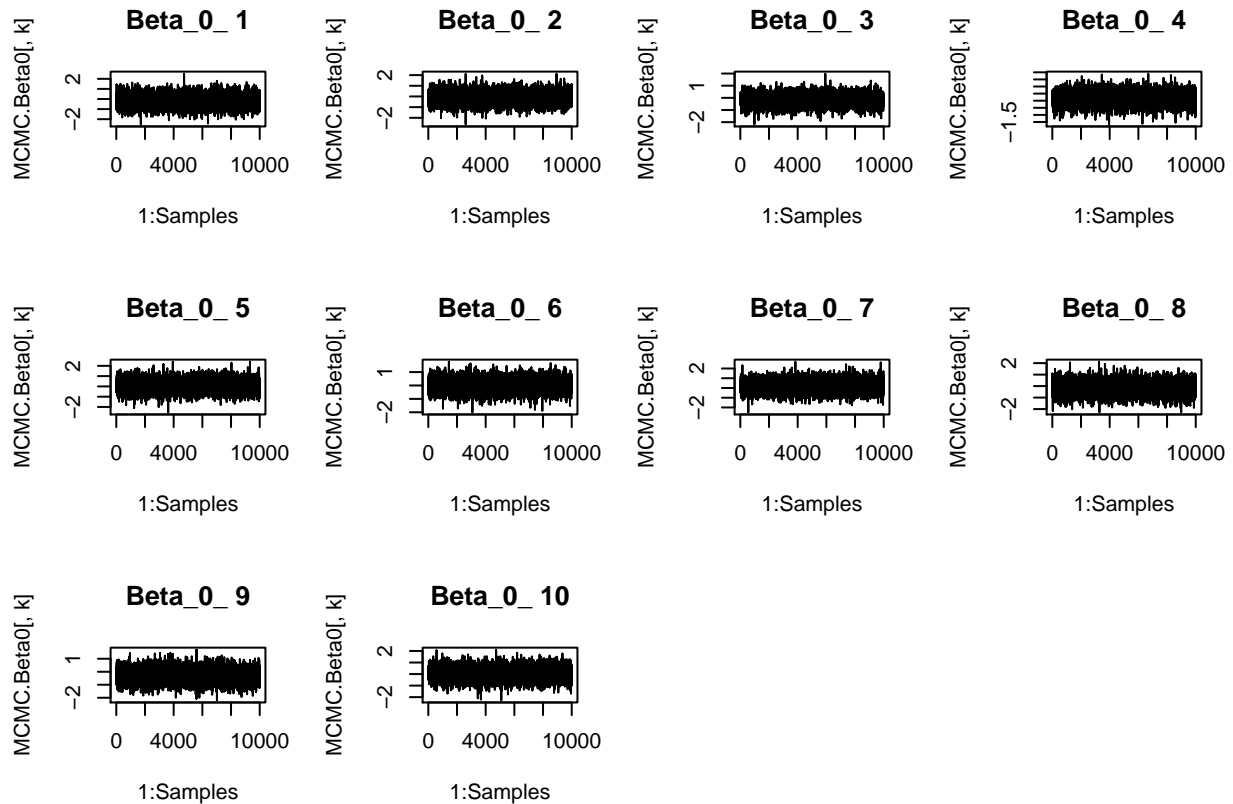
```
## [1] "For Beta_0"
```

```
print(Summary)
```

##	Actual	Mean	2.5%	97.5%	Standard Error
## Beta1	0	-0.19096254	-1.2193292	0.8452030	0.005287115
## Beta2	0	-0.12999749	-1.0939260	0.8501786	0.004923669
## Beta3	0	-0.20054704	-0.9596543	0.5592323	0.003972032
## Beta4	0	0.16503459	-0.6123830	0.9208262	0.004135720
## Beta5	0	0.05782741	-0.9199941	1.0879735	0.005506188
## Beta6	0	0.03851145	-0.7960740	0.8519223	0.004313926
## Beta7	0	0.24797811	-0.7855246	1.2548444	0.005378243
## Beta8	0	-0.21675474	-1.1960567	0.7899825	0.005109209
## Beta9	0	-0.29543627	-1.1695508	0.5947194	0.004659723
## Beta10	0	0.06878413	-0.8687981	1.0025311	0.004924480

```
par(mfrow = c(3,4));
for(k in 1:p){
```

```
plot(1:Samples, MCMC.Beta0[,k], type = 'l', main = paste("Beta_0_", toString(k)))
}
```



σ^2

```
Summary = matrix(0, nrow = 1, ncol = 5)
colnames(Summary) = c("Actual", "Mean", "2.5%", "97.5%", "Standard Error")
rownames(Summary) = c("Sigma^2")
for(k in 1:1){
  Summary[k, 1] <- 1;
  Summary[k, 2] <- mean(MCMC.sigma2);
  Summary[k, 3:4] <- quantile(MCMC.sigma2, c(.025, .975));
  Summary[k, 5] <- sqrt(var(MCMC.sigma2)/effectiveSize(MCMC.sigma2));
}
print("For Sigma^2");
```

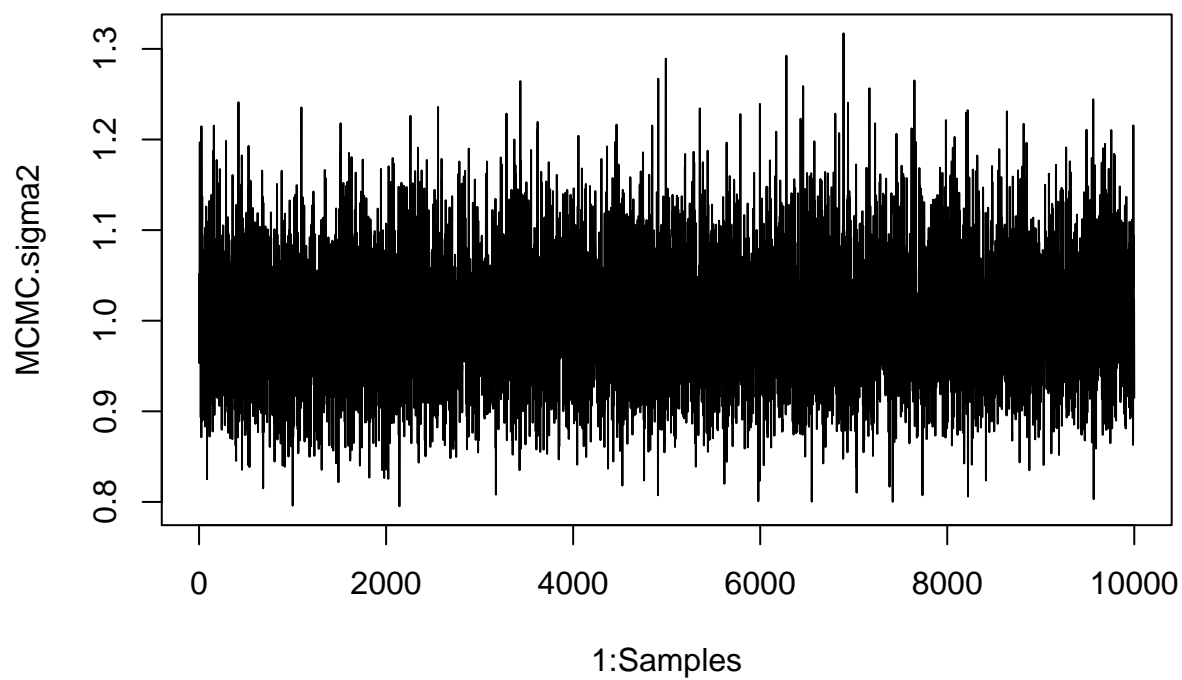
```
## [1] "For Sigma^2"
```

```
print(Summary)
```

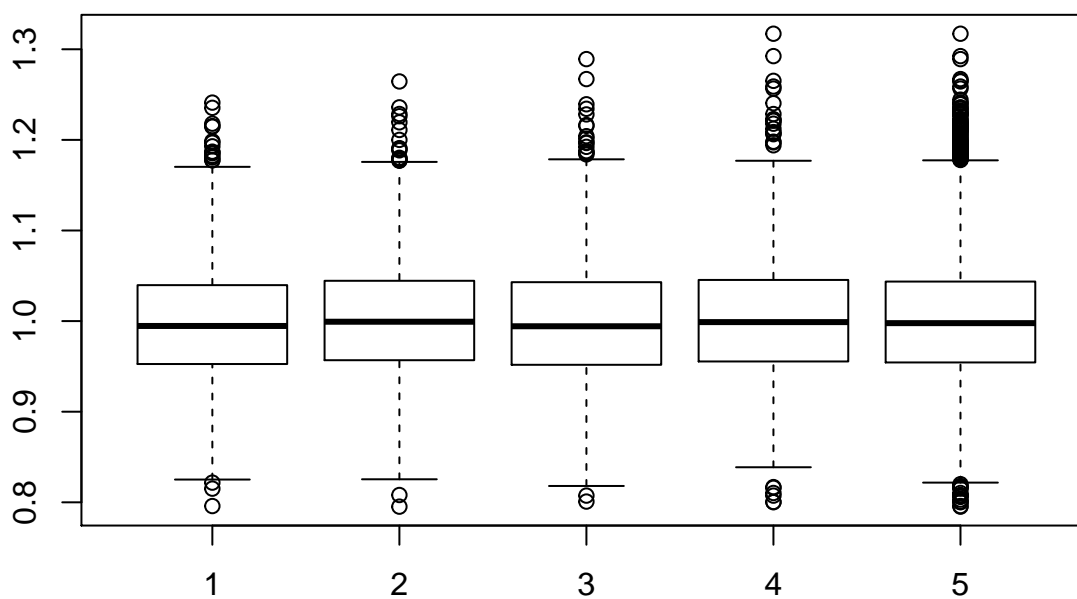
```
##      Actual    Mean    2.5%   97.5% Standard Error
## Sigma^2      1 1.00111 0.8788458 1.143155    0.0008024523
```

```
plot(1:Samples, MCMC.sigma2, type = 'l', main = paste("Beta_0_", toString(k)))
```

Beta_0_1



```
boxplot(MCMC.sigm2[1:2000], MCMC.sigm2[2001:4000], MCMC.sigm2[4001:6000],  
        MCMC.sigm2[6001:8000], MCMC.sigm2[8:10000])
```



The sampler seems to have done an excellent job reaching the stationary distribution of σ^2 as shown in the consistency of the distribution of σ^2 values of the boxplots above and the martingale like nature of the movement of the σ^2 value in the traceplot. Also, the standard error is strikingly small and the expected value is close to 1, the actual value which was used to produce the data.

Σ_0

To assess this, I have converted the p by p covariance matrix into a single vector of length p^2 . I will see how the specific values converge and what their distributions look like. The reader may find this large table of 100 row difficult to look at, but allow me to point out two interesting features of the table. First, recall that the actual Σ_0 that was used to generate the β_j values was an identity matrix. The first thing I would like to point out, is that all of the values which correspond to a location on the diagonal of the true Σ_0 are positive and their confidence intervals do not contain zero. Secondly, almost all of the off diagonal locations contain zero near the center of their interval. This being said, I feel comfortable with how the Gibbs sampler treated this covariance matrix and am satisfied with its convergence.

```
Summary = matrix(0, nrow = p*p, ncol = 5)
colnames(Summary) = c("Actual", "Mean", "2.5%", "97.5%", "Standard Error")
Covariance <- matrix(.2, nrow = p, ncol = p) + diag(p)*.8
actualSigma0 <- matrix(diag(p), nrow = p*p, ncol = 1, byrow = T);
Summary[,1] <- actualSigma0
for(k in 1:(p*p)){
  Summary[k, 2] <- mean(MCMC.Sigma0[,k]);
  Summary[k, 3:4] <- quantile(MCMC.Sigma0[,k],c(.025, .975));
  Summary[k, 5] <- sqrt(var(MCMC.Sigma0[,k])/effectiveSize(MCMC.Sigma0[,k]));
}
print("For Sigma_0");
```

```
## [1] "For Sigma_0"
```

```
print(Summary)
```

##	Actual	Mean	2.5%	97.5%	Standard Error
## [1,]	1	16.52219593	1.6654846	82.104413	0.7221185
## [2,]	0	-8.15899422	-48.4535861	6.269990	0.4251332
## [3,]	0	-5.11630817	-34.1206328	6.524359	0.4019561
## [4,]	0	0.26876686	-16.9967281	18.924125	0.2835359
## [5,]	0	0.79034838	-24.8677386	29.231185	0.3497616
## [6,]	0	-3.11561483	-28.7217948	11.340000	0.3769940
## [7,]	0	5.63298119	-13.0314887	39.402125	0.6446985
## [8,]	0	-3.15538589	-36.3177535	16.520796	0.3744213
## [9,]	0	3.26380879	-14.6364591	32.410624	0.3528708
## [10,]	0	3.15857214	-15.7205880	32.602031	0.4909101
## [11,]	0	-8.15899422	-48.4535861	6.269990	0.4251332
## [12,]	1	15.06418697	1.5369735	75.666044	0.4934390
## [13,]	0	1.35932654	-12.6605788	20.918444	0.2809712
## [14,]	0	-1.98290660	-23.2747109	11.938613	0.2637192
## [15,]	0	-3.39624646	-36.0519342	16.376525	0.3035081
## [16,]	0	-0.61267985	-18.6693742	18.788442	0.3095603
## [17,]	0	3.93267938	-15.6125073	35.899214	0.4129615
## [18,]	0	6.34133887	-10.7312657	41.439111	0.3872689
## [19,]	0	-7.66238047	-46.6521304	4.806391	0.3564555

##	[20,]	0	-3.98761163	-32.6490134	14.667043	0.3701713
##	[21,]	0	-5.11630817	-34.1206328	6.524359	0.4019561
##	[22,]	0	1.35932654	-12.6605788	20.918444	0.2809712
##	[23,]	1	7.83162547	0.7080814	37.098766	0.4748745
##	[24,]	0	-0.10870467	-12.2202374	12.611852	0.2996679
##	[25,]	0	3.74982135	-7.2426108	30.096417	0.3410979
##	[26,]	0	3.82916929	-5.5346342	22.466044	0.5041994
##	[27,]	0	-2.90950338	-24.2580249	11.516432	0.5453853
##	[28,]	0	-0.20935837	-17.2760096	17.672509	0.3244592
##	[29,]	0	-0.78670930	-16.2789231	14.136680	0.3435323
##	[30,]	0	3.07590741	-10.0051032	22.898525	0.5078646
##	[31,]	0	0.26876686	-16.9967281	18.924125	0.2835359
##	[32,]	0	-1.98290660	-23.2747109	11.938613	0.2637192
##	[33,]	0	-0.10870467	-12.2202374	12.611852	0.2996679
##	[34,]	1	7.58876756	0.7409899	34.151831	0.4523006
##	[35,]	0	5.20288244	-5.4641030	32.701211	0.3155328
##	[36,]	0	-0.28711704	-12.4482407	12.852193	0.4289371
##	[37,]	0	0.45121529	-16.7354974	17.468735	0.3202565
##	[38,]	0	-1.70617016	-21.7121255	13.402270	0.4641623
##	[39,]	0	5.69422602	-3.9408550	32.312112	0.4619285
##	[40,]	0	-0.93367893	-17.1789679	14.754457	0.4466434
##	[41,]	0	0.79034838	-24.8677386	29.231185	0.3497616
##	[42,]	0	-3.39624646	-36.0519342	16.376525	0.3035081
##	[43,]	0	3.74982135	-7.2426108	30.096417	0.3410979
##	[44,]	0	5.20288244	-5.4641030	32.701211	0.3155328
##	[45,]	1	15.10575515	1.5902337	74.908321	0.4590672
##	[46,]	0	-0.05212875	-19.4288696	19.958681	0.3184805
##	[47,]	0	2.10322764	-20.0037898	31.750048	0.4349129
##	[48,]	0	-1.65148561	-29.5762609	22.131518	0.4180054
##	[49,]	0	8.47733665	-3.5575341	50.259054	0.3526385
##	[50,]	0	2.21325903	-19.2199495	30.991363	0.2663872
##	[51,]	0	-3.11561483	-28.7217948	11.340000	0.3769940
##	[52,]	0	-0.61267985	-18.6693742	18.788442	0.3095603
##	[53,]	0	3.82916929	-5.5346342	22.466044	0.5041994
##	[54,]	0	-0.28711704	-12.4482407	12.852193	0.4289371
##	[55,]	0	-0.05212875	-19.4288696	19.958681	0.3184805
##	[56,]	1	9.18224588	0.8721524	40.153141	0.6840601
##	[57,]	0	-1.62642402	-22.4428156	15.762349	0.5485833
##	[58,]	0	-4.04726759	-27.6189389	10.427902	0.4717054
##	[59,]	0	-1.22793969	-19.0728271	15.322804	0.5041016
##	[60,]	0	6.13475651	-5.6265584	32.281229	0.7204304
##	[61,]	0	5.63298119	-13.0314887	39.402125	0.6446985
##	[62,]	0	3.93267938	-15.6125073	35.899214	0.4129615
##	[63,]	0	-2.90950338	-24.2580249	11.516432	0.5453853
##	[64,]	0	0.45121529	-16.7354974	17.468735	0.3202565
##	[65,]	0	2.10322764	-20.0037898	31.750048	0.4349129
##	[66,]	0	-1.62642402	-22.4428156	15.762349	0.5485833
##	[67,]	1	15.44391041	1.4395657	75.360364	0.8445346
##	[68,]	0	-1.23633304	-28.1729837	20.170012	0.3640382
##	[69,]	0	-0.45020242	-22.6026663	22.067768	0.3927891
##	[70,]	0	3.33610943	-13.8517047	32.340598	0.5659277
##	[71,]	0	-3.15538589	-36.3177535	16.520796	0.3744213
##	[72,]	0	6.34133887	-10.7312657	41.439111	0.3872689
##	[73,]	0	-0.20935837	-17.2760096	17.672509	0.3244592

##	[74,]	0	-1.70617016	-21.7121255	13.402270	0.4641623
##	[75,]	0	-1.65148561	-29.5762609	22.131518	0.4180054
##	[76,]	0	-4.04726759	-27.6189389	10.427902	0.4717054
##	[77,]	0	-1.23633304	-28.1729837	20.170012	0.3640382
##	[78,]	1	15.28767746	1.5269473	71.430542	0.7830595
##	[79,]	0	-2.22531395	-27.0944971	17.503326	0.5354935
##	[80,]	0	-9.89546276	-50.8592072	2.669979	0.5666460
##	[81,]	0	3.26380879	-14.6364591	32.410624	0.3528708
##	[82,]	0	-7.66238047	-46.6521304	4.806391	0.3564555
##	[83,]	0	-0.78670930	-16.2789231	14.136680	0.3435323
##	[84,]	0	5.69422602	-3.9408550	32.312112	0.4619285
##	[85,]	0	8.47733665	-3.5575341	50.259054	0.3526385
##	[86,]	0	-1.22793969	-19.0728271	15.322804	0.5041016
##	[87,]	0	-0.45020242	-22.6026663	22.067768	0.3927891
##	[88,]	0	-2.22531395	-27.0944971	17.503326	0.5354935
##	[89,]	1	12.66531632	1.2747366	61.852964	0.5779488
##	[90,]	0	-0.73224465	-21.7574331	19.390544	0.5542875
##	[91,]	0	3.15857214	-15.7205880	32.602031	0.4909101
##	[92,]	0	-3.98761163	-32.6490134	14.667043	0.3701713
##	[93,]	0	3.07590741	-10.0051032	22.898525	0.5078646
##	[94,]	0	-0.93367893	-17.1789679	14.754457	0.4466434
##	[95,]	0	2.21325903	-19.2199495	30.991363	0.2663872
##	[96,]	0	6.13475651	-5.6265584	32.281229	0.7204304
##	[97,]	0	3.33610943	-13.8517047	32.340598	0.5659277
##	[98,]	0	-9.89546276	-50.8592072	2.669979	0.5666460
##	[99,]	0	-0.73224465	-21.7574331	19.390544	0.5542875
##	[100,]	1	13.92777606	1.3309641	63.747743	0.8985910