

**A general model.** The Multivariate Normal linear model is

$$\begin{aligned} Y'_t &= X'_t B + u'_t \\ u'_t &\sim N(\mathbf{0}_{1 \times n}, \Sigma_u), \end{aligned} \tag{1}$$

where  $Y_t$  is a  $1 \times n$  row vector. This model nests within it many other models, including both VARs and VECMs, as well as seemingly-unrelated regressions (SURs). For a VAR of order  $p$ ,

$$X'_t = \left[ \mathbf{1}, \{Y'_{t-j}\}_{j=1}^p \right], \tag{2}$$

while, for a VECM,

$$X'_t = \left[ \mathbf{1}, Y_{t-1}, \{\Delta Y'_{t-j}\}_{j=1}^{p-1} \right], \tag{3}$$

where  $\Delta Y_t = Y_t - Y_{t-1}$ . The particular variables in  $X_t$  aren't particularly important for computing the forms of the posteriors, except for specifying an appropriate prior, so we'll be ignoring that until we look at specific priors. We'll just say that  $X'_t$  is an  $1 \times m$  row vector instead.

**Priors and likelihoods for the general model.** The conjugate priors we'll use for this model are Matrix Normal and Inverse Wishart:

$$\begin{aligned} B | \Sigma_u &\sim MN(\bar{B}, \Omega, \Sigma_u) \\ \Sigma_u &\sim IW(\Psi, d), \end{aligned} \tag{4}$$

where  $\Psi$  is an  $n \times n$  matrix and  $\Omega$  is an  $m \times m$  matrix.<sup>1</sup> Both have very standard forms. The Matrix Normal may be slightly less familiar, but it's closely related to the Multivariate Normal distribution, as, we'll see later.

The PDF of a  $B | \Sigma_u$ , given that it is distributed  $MN(\bar{B}, \Omega, \Sigma_u)$ , is

$$\pi(B | \Sigma_u) = (2\pi)^{-\frac{nm}{2}} |\Sigma_u|^{-\frac{m}{2}} |\Omega|^{-\frac{n}{2}} \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} (B - \bar{B})' \Omega^{-1} (B - \bar{B}) \right) \right), \tag{5}$$

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<sup>1</sup>For VARs and VECMs of order  $p$  with  $n$  variables,  $m = np + 1$ .

while the PDF of  $\Sigma_u$ , given that it is distributed  $IW(\Psi, d)$ , is

$$\pi(\Sigma_u) = 2^{-\frac{dn}{2}} \frac{|\Psi|^{\frac{d}{2}}}{\Gamma_n\left(\frac{d}{2}\right)} |\Sigma_u|^{-\frac{d+n+1}{2}} \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}\Psi\right)\right). \quad (6)$$

Let  $\mathbf{X}$  and  $\mathbf{Y}$  be the stacked values of  $X'_t$  and  $Y'_t$  for  $t = 0$  to  $T - 1$ , respectively. Then the distribution of  $\mathbf{Y}$  is

$$\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u \sim MN(\mathbf{XB}, \mathbf{I}_T, \Sigma_u), \quad (7)$$

which means it has the likelihood

$$P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) = (2\pi)^{-\frac{nT}{2}} |\Sigma_u|^{-\frac{T}{2}} \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}(\mathbf{Y} - \mathbf{XB})'(\mathbf{Y} - \mathbf{XB})\right)\right). \quad (8)$$

**Distribution of  $\mathbf{B} | \Sigma_u, \mathbf{X}, \mathbf{Y}$ .** By Bayes rule, we know

$$P(\mathbf{B} | \Sigma_u, \mathbf{X}, \mathbf{Y}) \propto P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) \quad (9)$$

To find the posterior, we first compute the product

$$\begin{aligned} P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) &= (2\pi)^{-\frac{n(m+T)}{2}} |\Sigma_u|^{-\frac{m+T}{2}} |\Omega|^{-\frac{n}{2}} \\ &\quad \times \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}(\mathbf{Y}'\mathbf{Y} - \mathbf{B}'\mathbf{X}'\mathbf{Y} - \mathbf{Y}'\mathbf{X}\mathbf{B} + \mathbf{B}'\mathbf{X}'\mathbf{X}\mathbf{B})\right)\right) \\ &\quad \times \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}(\bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \mathbf{B}'\Omega^{-1}\bar{\mathbf{B}} - \bar{\mathbf{B}}'\Omega^{-1}\mathbf{B} + \mathbf{B}'\Omega^{-1}\mathbf{B})\right)\right) \end{aligned} \quad (10)$$

$$\begin{aligned} &= (2\pi)^{-\frac{n(m+T)}{2}} |\Sigma_u|^{-\frac{m+T}{2}} |\Omega|^{-\frac{n}{2}} \\ &\quad \times \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}(\mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}})\right)\right) \\ &\quad \times \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}\left(-(\mathbf{Y}'\mathbf{X} + \bar{\mathbf{B}}'\Omega^{-1})\mathbf{B} - \mathbf{B}'\left(\Omega^{-1}\bar{\mathbf{B}} + \mathbf{X}'\mathbf{Y}\right)\right)\right)\right) \\ &\quad \times \exp\left(-\frac{1}{2}\text{tr}\left(\Sigma_u^{-1}\left(\mathbf{B}'\left(\Omega^{-1} + \mathbf{X}'\mathbf{X}\right)\mathbf{B}\right)\right)\right) \end{aligned} \quad (11)$$

Now, define

$$\widehat{\mathbf{B}} = \left( \Omega^{-1} + \mathbf{X}'\mathbf{X} \right)^{-1} \left( \Omega^{-1}\bar{\mathbf{B}} + \mathbf{X}'\mathbf{Y} \right) \quad (12)$$

$$\widehat{\mathbf{V}} = \left( \Omega^{-1} + \mathbf{X}'\mathbf{X} \right)^{-1}, \quad (13)$$

which allows us to rewrite (11) as

$$\begin{aligned} P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) &= (2\pi)^{-\frac{n(m+T)}{2}} |\Sigma_u|^{-\frac{m+T}{2}} |\Omega|^{-\frac{n}{2}} \\ &\times \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} \left( \mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} \right) \right) \right) \\ &\times \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} \left( -\widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\mathbf{B} - \mathbf{B}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}} + \mathbf{B}'\widehat{\mathbf{V}}^{-1}\mathbf{B} \right) \right) \right) \end{aligned} . \quad (14)$$

Completing the square, we obtain

$$\begin{aligned} P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) &= (2\pi)^{-\frac{n(m+T)}{2}} |\Sigma_u|^{-\frac{m+T}{2}} |\Omega|^{-\frac{n}{2}} \\ &\times \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} \left( \mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}} \right) \right) \right) \\ &\times \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} \left( \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}} - \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\mathbf{B} - \mathbf{B}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}} + \mathbf{B}'\widehat{\mathbf{V}}^{-1}\mathbf{B} \right) \right) \right) \end{aligned} \quad (15)$$

$$\begin{aligned} &= (2\pi)^{-\frac{n(m+T)}{2}} |\Sigma_u|^{-\frac{m+T}{2}} |\Omega|^{-\frac{n}{2}} \\ &\times \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} \left( \mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}} \right) \right) \right) \\ &\times \exp \left( -\frac{1}{2} \text{tr} \left( \Sigma_u^{-1} \left( (\mathbf{B} - \widehat{\mathbf{B}})' \widehat{\mathbf{V}}^{-1} (\mathbf{B} - \widehat{\mathbf{B}}) \right) \right) \right). \end{aligned} \quad (16)$$

We notice now that the third term is proportional to a Matrix Normal PDF for the distribution  $\text{MN}(\widehat{\mathbf{B}}, \widehat{\mathbf{V}}, \Sigma_u)$ . We, therefore, conclude that the the posterior for the coefficients is

$$\mathbf{B} | \Sigma_u, \mathbf{X}, \mathbf{Y} \sim \text{MN}(\widehat{\mathbf{B}}, \widehat{\mathbf{V}}, \Sigma_u).$$

We also note that

$$\begin{aligned} P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) &= (2\pi)^{-\frac{nT}{2}} |\Sigma_u|^{-\frac{T}{2}} |\widehat{\mathbf{V}}|^{\frac{n}{2}} |\Omega|^{-\frac{n}{2}} \\ &\times \exp\left(-\frac{1}{2} \text{tr}\left(\Sigma_u^{-1} (\mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}})\right)\right), \quad (17) \\ &\times P(\mathbf{B} | \Sigma_u, \mathbf{X}, \mathbf{Y}) \end{aligned}$$

which will be useful in computing the next posterior.

**Distribution of  $\Sigma_u | \mathbf{X}, \mathbf{Y}$ .** Turn to the posterior for  $\Sigma_u$ . By Bayes rule, again, we know that

$$P(\mathbf{B}, \Sigma_u | \mathbf{X}, \mathbf{Y}) \propto P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) \pi(\Sigma_u) \quad (18)$$

and, by the Law of Total Probability,

$$P(\Sigma_u | \mathbf{X}, \mathbf{Y}) = \int P(\mathbf{B}, \Sigma_u | \mathbf{X}, \mathbf{Y}) d\mathbf{B} \propto \left( \int P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) d\mathbf{B} \right) \pi(\Sigma_u). \quad (19)$$

From equation (17), we know

$$\begin{aligned} \int P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) d\mathbf{B} &= (2\pi)^{-\frac{nT}{2}} |\Sigma_u|^{-\frac{T}{2}} |\widehat{\mathbf{V}}|^{\frac{n}{2}} |\Omega|^{-\frac{n}{2}} \\ &\times \exp\left(-\frac{1}{2} \text{tr}\left(\Sigma_u^{-1} (\mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}})\right)\right) \quad (20) \\ &\times \underbrace{\int P(\mathbf{B} | \Sigma_u, \mathbf{X}, \mathbf{Y}) d\mathbf{B}}_{=1}, \end{aligned}$$

so

$$\begin{aligned} \left( \int P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) d\mathbf{B} \right) \pi(\Sigma_u) &= 2^{-\frac{dn}{2}} \frac{|\Psi|^{\frac{d}{2}}}{\Gamma_n(\frac{d}{2})} (2\pi)^{-\frac{nT}{2}} |\Omega|^{-\frac{n}{2}} |\widehat{\mathbf{V}}|^{\frac{n}{2}} |\Sigma_u|^{-\frac{T+d+n+1}{2}} \\ &\times \exp\left(-\frac{1}{2} \text{tr}\left(\Sigma_u^{-1} (\mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}'\widehat{\mathbf{V}}^{-1}\widehat{\mathbf{B}} + \Psi)\right)\right). \quad (21) \end{aligned}$$

If we define

$$\widehat{\mathbf{S}} = \mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}'\left(\Omega^{-1} + \mathbf{X}'\mathbf{X}\right)\widehat{\mathbf{B}} + \Psi, \quad (22)$$

we notice that this is proportional to the PDF for the distribution  $\text{IW}(\widehat{\mathbf{S}}, T+d)$ . We conclude that the posterior for the covariance matrix is

$$\Sigma_u | \mathbf{X}, \mathbf{Y} \sim \text{IW}(\widehat{\mathbf{S}}, T+d). \quad (23)$$

Moreover, we have (after pulling out appropriate constant terms)

$$\left( \int P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) d\mathbf{B} \right) \pi(\Sigma_u) = \pi^{-\frac{nT}{2}} \frac{\Gamma_n\left(\frac{T+d}{2}\right)}{\Gamma_n\left(\frac{d}{2}\right)} |\Psi|^{\frac{d}{2}} |\Omega|^{-\frac{n}{2}} |\widehat{\mathbf{V}}|^{\frac{n}{2}} |\widehat{\mathbf{S}}|^{-\frac{T+d}{2}} P(\Sigma_u | \mathbf{X}, \mathbf{Y}). \quad (24)$$

While the representation of  $\widehat{\mathbf{S}}$  we have above is particularly efficient, the following is another form we can use:

$$\widehat{\mathbf{S}} = (\mathbf{Y} - \mathbf{X}\widehat{\mathbf{B}})'(\mathbf{Y} - \mathbf{X}\widehat{\mathbf{B}}) + (\widehat{\mathbf{B}} - \bar{\mathbf{B}})' \Omega^{-1} (\widehat{\mathbf{B}} - \bar{\mathbf{B}}) + \Psi. \quad (25)$$

This is a particularly intuitive form, as it allows us to write the posterior mode of  $\Sigma_u$ , which is

$$\begin{aligned} \frac{\widehat{\mathbf{S}}}{T+d+n+1} &= \frac{T-m}{T+d+n+1} \frac{(\mathbf{Y} - \mathbf{X}\widehat{\mathbf{B}})'(\mathbf{Y} - \mathbf{X}\widehat{\mathbf{B}})}{T-m} \\ &\quad + \frac{m}{T+d+n+1} \frac{(\widehat{\mathbf{B}} - \bar{\mathbf{B}})' \Omega^{-1} (\widehat{\mathbf{B}} - \bar{\mathbf{B}})}{m} \\ &\quad + \frac{d+n+1}{T+d+n+1} \frac{\Psi}{d+n+1} \end{aligned} \quad (26)$$

as a weighted average of three constituent estimates of the variance: the sample error covariance, the covariance of deviations in coefficients from the prior mean, and the prior mean from the Inverse Wishart prior.

**The marginal likelihood.** The marginal distribution is equal to

$$P(\mathbf{Y} | \mathbf{X}) = \int \left( \int P(\mathbf{Y} | \mathbf{X}, \mathbf{B}, \Sigma_u) \pi(\mathbf{B} | \Sigma_u) d\mathbf{B} \right) \pi(\Sigma_u) d\Sigma_u \quad (27)$$

$$= \pi^{-\frac{nT}{2}} \frac{\Gamma_n\left(\frac{T+d}{2}\right)}{\Gamma_n\left(\frac{d}{2}\right)} |\Psi|^{\frac{d}{2}} |\Omega|^{-\frac{n}{2}} |\widehat{\mathbf{V}}|^{\frac{n}{2}} |\widehat{\mathbf{S}}|^{-\frac{T+d}{2}} \underbrace{\int P(\Sigma_u | \mathbf{X}, \mathbf{Y}) d\Sigma_u}_{=1} \quad (28)$$

$$= \pi^{-\frac{nT}{2}} \frac{\Gamma_n\left(\frac{T+d}{2}\right)}{\Gamma_n\left(\frac{d}{2}\right)} |\Psi|^{\frac{d}{2}} |\Omega|^{-\frac{n}{2}} |\widehat{\mathbf{V}}|^{\frac{n}{2}} |\widehat{\mathbf{S}}|^{-\frac{T+d}{2}}. \quad (29)$$

Alternatively, plugging back in our expressions for  $\widehat{\mathbf{V}}$  and  $\widehat{\mathbf{S}}$ , the marginal distribution may also be written as

$$\begin{aligned} P(\mathbf{Y} | \mathbf{X}) &= \pi^{-\frac{nT}{2}} \frac{\Gamma_n\left(\frac{T+d}{2}\right)}{\Gamma_n\left(\frac{d}{2}\right)} |\Psi|^{\frac{d}{2}} |\Omega|^{-\frac{n}{2}} \left| \Omega^{-1} + \mathbf{X}'\mathbf{X} \right|^{-\frac{n}{2}} \\ &\quad \times \left| \mathbf{Y}'\mathbf{Y} + \bar{\mathbf{B}}'\Omega^{-1}\bar{\mathbf{B}} - \widehat{\mathbf{B}}' \left( \Omega^{-1} + \mathbf{X}'\mathbf{X} \right) \widehat{\mathbf{B}} + \Psi \right|^{-\frac{T+d}{2}}. \end{aligned} \quad (30)$$

These are generic results that we use to conduct inference in VAR models of all sizes.

**Making draws from the posteriors.** Posterior inference on the parameters of the linear model involves draws from the joint distribution  $\mathbf{B}, \Sigma | \mathbf{X}, \mathbf{Y}$ , which can be achieved by drawing from  $\Sigma_u | \mathbf{X}, \mathbf{Y}$  and  $\mathbf{B} | \Sigma_u, \mathbf{X}, \mathbf{Y}$  in sequence. Then, to obtain N draws from the posterior, for each j from 1 to N, we cycle through two steps:

1. Draw  $\Sigma_u^{(j)}$  from  $IW(\widehat{\mathbf{S}}, T + d)$ , where  $\widehat{\mathbf{S}}$  is given by (22) and  $\widehat{\mathbf{B}}$  is given by (12). This can be achieved by the following steps:
  - (a) Draw an  $n \times (T + d)$  matrix Q, where each entry of Q is i.i.d.  $N(0, 1)$ ;
  - (b) Let  $Chol(\widehat{\mathbf{S}})$  be the lower triangular matrix in the Cholesky decomposition of  $\widehat{\mathbf{S}}$ .<sup>2</sup> Then  $\Sigma_u^{(j)} = Chol(\widehat{\mathbf{S}})(Q'Q)^{-1}Chol(\widehat{\mathbf{S}})'$ ;
2. Draw  $\mathbf{B}^{(j)}$  from  $MN(\widehat{\mathbf{B}}, \widehat{\mathbf{V}}, \Sigma_u^{(j)})$ , where  $\widehat{\mathbf{B}}$  is given by (12) and  $\widehat{\mathbf{V}}$  is given by (13). This can be achieved by the following steps:
  - (a) Draw an  $m \times n$  matrix W, where each entry of W is i.i.d.  $N(0, 1)$ ;

<sup>2</sup>That is the convention we'll keep going forward: that  $Chol(\cdot)$  refers to the lower-triangular matrix in the decomposition.

(b) Compute  $B^{(j)} = \widehat{B} + \text{Chol}(\widehat{V})W \text{Chol}(\Sigma_u^{(j)})'$ .

Compared to the implementation where we work with the vectorized coefficients, this offers multiple ways to speed up the calculations. For example, only  $W \text{Chol}(\Sigma_u)'$  needs to be recomputed with each draw. That's cheap as  $\Sigma_u$  is never very large.

**An alternative form for the posterior  $B | \Sigma_u, \mathbf{X}, \mathbf{Y}$ .** In the literature, it's more common to specify the priors as Multivariate Normal than Matrix Normal, though this is less efficient from a computational standpoint. In that case, the prior is

$$\text{vec}(B) | \Sigma_u \sim N(\text{vec}(\bar{B}), \Sigma_u \otimes \Omega), \quad (31)$$

by standard results on the relationship between the Matrix Normal and Multivariate Normal distributions. By the same results, the posterior for  $\text{vec}(B)$  is

$$\text{vec}(B) | \Sigma_u, \widehat{\mathbf{X}}, \widehat{\mathbf{Y}} \sim N(\text{vec}(\widehat{B}), \Sigma_u \otimes \widehat{V}). \quad (32)$$