

Instructions:

- You have until Tuesday May 5th 2:59pm PST to finish and upload your answers to CCLE. For the complete regulations, see the announcement on CCLE.
- On the first page of your submission, write down your full name, University ID, and the following academic integrity statement, then sign and date.

I agree to the UCLA Student Code of Conduct for academic integrity. I agree to use NO resource other than lecture notes and the textbook for this exam. I agree to communicate with NO ONE regarding this exam. Violations of this code will result in immediate FAILURE of this course.

- Use blank sheets of paper to write down answers with requested supporting work.

Problem 1

Given a data set

$$\mathcal{D} = \{(\mathbf{x}^{(i)}, t^{(i)})\}_{i=1}^N, \quad \mathbf{x}^{(i)} \in \mathbb{R}^D, t^{(i)} \in \mathbb{R}.$$

Fixing an integer $M \in \mathbb{N}$ and a basis function $\phi : \mathbb{R}^D \rightarrow \mathbb{R}^M$. The regression problem is about finding the best parameter $\mathbf{w} \in \mathbb{R}^M$ so

$$t^{(i)} = \mathbf{w}^T \phi(\mathbf{x}^{(i)}) + \epsilon^{(i)}$$

where $\epsilon^{(i)} \sim \mathcal{N}(0, \beta^{-1})$ are independent identical (unbiased) Gaussian noise.

- (a) Write down a formula for the likelihood function $p(\mathcal{D}|\mathbf{w}, \beta)$. (10 points)

$$p(\mathcal{D}|\mathbf{w}, \beta) = \prod_{i=1}^N \mathcal{N}(t^{(i)}|\mathbf{w}^T \phi(\mathbf{x}^{(i)}), \beta^{-1}).$$

- (b) Show that the maximum likelihood solution

$$\mathbf{w}_\beta^* = \arg \max_{\mathbf{w}} p(\mathcal{D}|\mathbf{w}, \beta)$$

for any value of $\beta > 0$ is the same as the least square solution

$$\bar{\mathbf{w}} = \arg \min_{\mathbf{w}} \sum_{i=1}^N \frac{1}{2} |t^{(i)} - y(x^{(i)}, \mathbf{w})|^2. \quad (10 \text{ points})$$

The least square error function $\frac{1}{2} |t_i - y(x^{(i)}, \mathbf{w})|^2$ is the same as the negative log-likelihood divided by $\beta > 0$ (up to a constant).

- (c) Fixing the model complexity $M \in \mathbb{N}$, give three examples of the basis function $\phi(\mathbf{x})$. (10 points)

Linear $(1, \mathbf{x})$, spline functions, Gaussian, sigmoidal, etc.

Problem 2

Suppose a data set $\mathcal{D} = \{(\mathbf{x}^{(i)}, t^{(i)})\}_{i=1}^N$ is given. $\mathbf{x}^{(i)} \in \mathbb{R}^D, t^{(i)} \in \mathbb{R}$ for $i = 1, \dots, N$.

- (a) Show that the optimal solution $\mathbf{w}^* = \arg \min J(\mathbf{w})$ for a regularized sum-of-squares error function

$$J(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^N (t^{(i)} - \mathbf{w}^T \phi(\mathbf{x}^{(i)}))^2 + \frac{\lambda}{2} \|\mathbf{w}\|^2,$$

where $\lambda > 0$, is a linear combination of the vectors $\{\phi(\mathbf{x}^{(i)})\}_{i=1}^N$. In other words, show that

$$\mathbf{w}^* = \sum_{i=1}^N a^{(i)} \phi(\mathbf{x}^{(i)})$$

for some scalars $a^{(i)} \in \mathbb{R}, i = 1, \dots, N$. (10 points)

From the optimal condition

$$\begin{aligned} 0 &= \nabla J(\mathbf{w}^*) = \sum_{i=1}^N (t^{(i)} - (\mathbf{w}^*)^T \phi(\mathbf{x}^{(i)}))(-\phi(\mathbf{x}^{(i)})) + \lambda \mathbf{w}^* \\ \mathbf{w}^* &= \frac{1}{\lambda} \sum_{i=1}^N (t^{(i)} - (\mathbf{w}^*)^T \phi(\mathbf{x}^{(i)})) \phi(\mathbf{x}^{(i)}) = \sum_{i=1}^N a^{(i)} \phi(\mathbf{x}^{(i)}) \\ a^{(i)} &= \frac{1}{\lambda} (t^{(i)} - (\mathbf{w}^*)^T \phi(\mathbf{x}^{(i)})). \end{aligned}$$

(I let go of missing negative sign for this problem.) Alternatively, one can show that

$$\begin{aligned} \mathbf{w}^* &= (\mathbf{K} + \lambda \mathbf{I})^{-1} \Phi^T \mathbf{t} \\ &= \Phi^T (\mathbf{K} + \lambda \mathbf{I})^{-1} \mathbf{t} \\ &= \sum_{i=1}^N \phi(\mathbf{x}^{(i)})^T (i\text{-th entry of } (\mathbf{K} + \lambda \mathbf{I})^{-1} \mathbf{t}) \\ &= \Phi^T \mathbf{a}. \end{aligned}$$

This requires showing that Φ^T commutes with the matrix $(\mathbf{K} + \lambda \mathbf{I})^{-1}$ but will solve problem 2c together.

- (b) We define the Gram matrix

$$\mathbf{K} = [K_{ij}] = [\phi(\mathbf{x}^{(i)})^T \phi(\mathbf{x}^{(j)})] \in \mathbb{R}^{N \times N}.$$

Show that \mathbf{K} is symmetric semi-positive definite. (A matrix $\mathbf{A} \in \mathbb{R}^{N \times N}$ is symmetric semi-positive definite if and only if $\mathbf{x}^T \mathbf{A} \mathbf{x} \geq 0$ for all vector $\mathbf{x} \in \mathbb{R}^N$.) (10 points)

Take any vector $\mathbf{y} \in \mathbb{R}^N$,

$$\begin{aligned} \mathbf{y}^T \mathbf{K} \mathbf{y} &= \sum_{i=1}^N \sum_{j=1}^N y_i \phi(\mathbf{x}^{(i)})^T \phi(\mathbf{x}^{(j)}) y_j \\ &= \left(\sum_{i=1}^N y_i \phi(\mathbf{x}^{(i)}) \right)^T \left(\sum_{j=1}^N y_j \phi(\mathbf{x}^{(j)}) \right) \geq 0. \end{aligned}$$

(c) Show that the coefficients from part (a) satisfy

$$(\mathbf{K} + \lambda \mathbf{I}_N) \begin{bmatrix} a^{(1)} \\ a^{(2)} \\ \vdots \\ a^{(N)} \end{bmatrix} = \begin{bmatrix} t^{(1)} \\ t^{(2)} \\ \vdots \\ t^{(N)} \end{bmatrix}. \quad (10 \text{ points})$$

Since $\mathbf{w}^* = \Phi \mathbf{a}$, where

$$\Phi = \begin{bmatrix} \phi(\mathbf{x}^{(1)})^T \\ \phi(\mathbf{x}^{(2)})^T \\ \vdots \\ \phi(\mathbf{x}^{(N)})^T \end{bmatrix} \in \mathbb{R}^{N \times M},$$

we have

$$\begin{aligned} \lambda a^{(i)} &= -(t^{(i)} - (\mathbf{w}^*)^T \phi(\mathbf{x}^{(i)})) \\ &= -(t^{(i)} - \mathbf{a}^T \Phi^T \phi(\mathbf{x}^{(i)})) \\ \lambda \mathbf{I}_N \mathbf{a} &= \mathbf{t} - \Phi \Phi^T \mathbf{a} = \mathbf{t} - \mathbf{K} \mathbf{a}. \end{aligned}$$

Problem 3

Consider the two-class classification problem. Denote the data set $\mathcal{D} = \{(\mathbf{x}^{(i)}, t^{(i)})\}_{i=1}^N$ where

$$t^{(i)} = \begin{cases} 1, & i \in C_{[1]} \\ 0, & i \in C_{[2]} \end{cases}$$

is the target variable encoding the class membership.

- (a) Suppose $p(\mathbf{x}|C_{[1]}) \sim \mathcal{N}(\boldsymbol{\mu}_{[1]}, \boldsymbol{\Sigma})$ and $p(\mathbf{x}|C_{[2]}) \sim \mathcal{N}(\boldsymbol{\mu}_{[2]}, \boldsymbol{\Sigma})$, that is, data from two classes scatter around different class-specific mean but share the same covariance matrix. Denote $p(C_{[1]}) = \pi$, hence $p(C_{[2]}) = 1 - \pi$. The likelihood function is given by

$$p(\mathcal{D}|\pi, \boldsymbol{\mu}_{[1]}, \boldsymbol{\mu}_{[2]}, \boldsymbol{\Sigma}) = \left(\prod_{i \in C_{[1]}} \pi \mathcal{N}(\mathbf{x}^{(i)}|\boldsymbol{\mu}_{[1]}, \boldsymbol{\Sigma}) \right) \cdot \left(\prod_{i \in C_{[2]}} (1 - \pi) \mathcal{N}(\mathbf{x}^{(i)}|\boldsymbol{\mu}_{[2]}, \boldsymbol{\Sigma}) \right)$$

Show that the maximum likelihood estimate of the class probability π is given by the fraction of data points in $C_{[1]}$, i.e.

$$\arg \max_{\pi} p(\mathcal{D}|\pi, \boldsymbol{\mu}_{[1]}, \boldsymbol{\mu}_{[2]}, \boldsymbol{\Sigma}) = \frac{\#\{i : i \in C_{[1]}\}}{N}. \quad (10 \text{ points})$$

Log-likelihood function boils down to

$$E(\pi) = \sum_{i \in C_{[1]}} \log \pi + \sum_{i \in C_{[2]}} \log(1 - \pi) + (\text{const w.r.t. } \pi).$$

The first optimality condition gives

$$\begin{aligned} E'(\pi) &= \sum_{i \in C_{[1]}} \frac{1}{\pi} - \sum_{i \in C_{[2]}} \frac{1}{1 - \pi} = 0 \\ \pi(1 - \pi)E'(\pi) &= N_1(1 - \pi) - N_2\pi = N_1 - (N_1 + N_2)\pi = 0 \\ \pi &= \frac{N_1}{N_1 + N_2}. \end{aligned}$$

Here $N_1 = \#\{i : i \in C_{[1]}\}$, $N_2 = \#\{i : i \in C_{[2]}\}$.

- (b) The *logistic sigmoid* function is defined by

$$\sigma(b) = \frac{1}{1 + e^{-b}}.$$

Show that (i) $\sigma(-b) = 1 - \sigma(b)$, (ii) σ is a monotonically increasing function, and (iii) σ maps all of \mathbb{R} onto the interval $(0, 1)$. (15 points)

$$(i) \quad 1 - \sigma(b) = 1 - \frac{1}{1 + e^{-b}} = \frac{e^{-b}}{1 + e^{-b}} = \frac{1}{e^b + 1} = \sigma(-b).$$

$$(ii) \quad \sigma'(b) = \frac{-(-e^{-b})}{(1 + e^{-b})^2} > 0 \text{ for any } b \in \mathbb{R}.$$

(iii) $\lim_{b \rightarrow -\infty} \sigma(b) = 0$, $\lim_{b \rightarrow +\infty} \sigma(b) = 1$. Since σ is monotonically increasing, any input between $\pm\infty$ must fall in the interval $(0, 1)$.

- (c) In an approach different from part (a), we suppose that $p(C_{[1]}|\mathbf{x}) = \sigma(\mathbf{w}^T \phi(\mathbf{x}))$ where $\mathbf{w} \in \mathbb{R}^M$ is a coefficient vector to be trained. According to part (b), $p(C_{[2]}|\mathbf{x}) = \sigma(-\mathbf{w}^T \phi(\mathbf{x}))$. The likelihood function is then given by this different formula,

$$p(\mathcal{D}|\mathbf{w}) = \left(\prod_{i \in C_{[1]}} \sigma(\mathbf{w}^T \phi(\mathbf{x}^{(i)})) \right) \cdot \left(\prod_{i \in C_{[2]}} \sigma(-\mathbf{w}^T \phi(\mathbf{x}^{(i)})) \right).$$

Define the error function $E(\mathbf{w}) = -\log p(\mathcal{D}|\mathbf{w})$ to be the negative logarithm of the likelihood. Compute $\nabla E(\mathbf{w})$ and explain why the maximum likelihood estimate $\nabla E(\mathbf{w}^*) = 0$ doesn't have an analytical solution. (15 points)

$$\frac{d}{dt} \log \sigma(t) = \frac{\sigma'(t)}{\sigma(t)} = \frac{(1 + e^{-t})e^{-t}}{(1 + e^{-t})^2} = 1 - \sigma(t) = \sigma(-t)$$

$$\begin{aligned} E(\mathbf{w}) &= - \left(\sum_{i \in C_{[1]}} \log \sigma(\mathbf{w}^T \phi(\mathbf{x}^{(i)})) \right) - \left(\sum_{i \in C_{[2]}} \log \sigma(-\mathbf{w}^T \phi(\mathbf{x}^{(i)})) \right) \\ \nabla E(\mathbf{w}) &= - \left(\sum_{i \in C_{[1]}} \sigma(-\mathbf{w}^T \phi(\mathbf{x}^{(i)})) \phi(\mathbf{x}^{(i)}) \right) - \left(\sum_{i \in C_{[2]}} \sigma(\mathbf{w}^T \phi(\mathbf{x}^{(i)})) (-\phi(\mathbf{x}^{(i)})) \right) \\ &= \left(\sum_{i \in C_{[1]}} (\sigma(\mathbf{w}^T \phi(\mathbf{x}^{(i)})) - 1) \phi(\mathbf{x}^{(i)}) \right) + \left(\sum_{i \in C_{[2]}} \sigma(\mathbf{w}^T \phi(\mathbf{x}^{(i)})) \phi(\mathbf{x}^{(i)}) \right) \\ &= \sum_{i=1}^N (\sigma(\mathbf{w}^T \phi(\mathbf{x}^{(i)})) - t^{(i)}) \phi(\mathbf{x}^{(i)}) \end{aligned}$$

Dependence on σ makes the expression nonlinear and therefore have no analytical solution.

- (d) **(Bonus)** Provide a strategy to compute the optimal solution \mathbf{w}^* numerically. (up to 10 points)

Any clearly stated suggestion for Newton's method or Gradient Descend will receive up to 10 points.