

**PROBLEM I** Assume  $(N_t : t \geq 0)$  is a Poisson process with parameter  $\alpha \geq 0$  and initial value 0.  $\{\xi_n : n \in \mathbb{N}\}$  are i.i.d. r.v. with distribution  $\mu$  and independent with  $(N_t : t \geq 0)$ . Let  $X_t = \sum_{n=0}^{N_t} \xi_n, t \geq 0, \forall s \geq 0$ ,

1.  $(N_{s+t} - N_s : t \geq 0)$  is a Poisson process with parameter  $\alpha$ .
2.  $\{\xi_{N_s+n} : n \in \mathbb{N}^+\}$  are i.i.d. with distribution  $\mu$  and are independent with  $(N_{s+t} - N_s : t \geq 0)$ .
3.  $(X_t : t \geq 0)$  satisfies  $\forall 0 = t_0 < t_1 < \dots < t_n, X_{t_1}, X_{t_k} - X_{t_{k-1}}, k = 2, \dots, n$  are independent.

**SOLUTION.** 1. Suppose  $(N_t : t \geq 0)$  satisfies  $N_t - N_s \sim \text{Poisson}(\alpha(t-s)), \forall t \geq s \geq 0$ , then  $(N_{t+r} - N_r) - (N_{s+r} - N_r) = N_{t+r} - N_{s+r} \sim \text{Poisson}(\alpha(t-s))$ . Besides,  $\forall 0 = t_0 < t_1 < t_2 < \dots < t_n, D_t := N_{t+r} - N_r$ , then  $D_{t_k} - D_{t_{k-1}} = N_{t_k+r} - N_{t_{k-1}+r}, \forall k = 1, \dots, n$ . So  $D_{t_k} - D_{t_{k-1}} = N_{s_k} - N_{s_{k-1}}, \forall k = 1, \dots, n$ , where  $s_k = t_k + r$ , are independent to each other. Besides,  $N_{0+r} - N_r = 0$ , obviously  $D_{t_k} - D_{t_{k-1}}$  is independent to  $N_{0+r} - N_0$ . Last, since the orbit of  $(N_t : t \geq 0)$  is continuous, then  $N_{t+r} - N_r$  is continuous for any  $t \geq 0$ . Thus, by the definition of Poisson process, we get  $(N_{t+r} - N_r : t \geq 0)$  is Poisson process.

2. Assume  $(\Omega, \mathcal{F}), (\mathbb{N}, \mathcal{P}(\mathbb{N})), (E, \mathcal{E})$  are sigma algebra.  $\xi_n : \Omega \rightarrow E, N_t : \Omega \rightarrow \mathbb{N}$ . First of all, we prove that  $\forall n \in \mathbb{N}^+, \xi_{N_s+n}$  has contribution  $\mu$ :  $\forall A \in \mathcal{E}, \mathbb{P}(\{\xi_{N_s+n} \in A\}) = \mathbb{P}(\bigcup_{k=0}^{\infty} \{\xi_{k+n} \in A, N_s = k\}) = \sum_{k=0}^{\infty} \mathbb{P}(\xi_{k+n} \in A, N_s = k) = \sum_{k=0}^{\infty} \mathbb{P}(\xi_{k+n} \in A) \mathbb{P}(N_s = k) = \sum_{k=0}^{\infty} \mathbb{P}(\xi_1 \in A) \mathbb{P}(N_s = k) = \mathbb{P}(\xi_1 \in A)$ .

Secondly, we prove that  $\{\xi_{N_s+n} : n \in \mathbb{N}^+\}$  are independent:  $\forall J \subset \mathbb{N}^+, \text{card}(J) < \infty, \{A_i \in \mathcal{E} : i \in J\}$ , then  $\mathbb{P}(\bigcap_{i \in J} \{\xi_{N_s+i} \in A_i\}) = \mathbb{P}(\bigcup_{k=0}^{\infty} (\bigcap_{i \in J} \{\xi_{k+i} \in A_i, N_s = k\})) = \sum_{k=0}^{\infty} \mathbb{P}(\bigcap_{i \in J} \{\xi_{k+i} \in A_i\} \cap \{N_s = k\}) = \sum_{k=0}^{\infty} \mathbb{P}(\bigcap_{i \in J} \{\xi_{k+i} : i \in A_i\}) \mathbb{P}(N_s = k) = \sum_{k=0}^{\infty} \mathbb{P}(\bigcap_{i \in J} \{\xi_{1+i} \in A_i\}) \mathbb{P}(N_s = k) = \mathbb{P}(\bigcap_{i \in J} \{\xi_{1+i} \in A_i\}) = \prod_{i \in J} \mathbb{P}(\xi_{1+i} \in A_i) = \prod_{i \in J} \mathbb{P}(\xi_{N_s+i} \in A_i)$ .

Last, we will prove that  $\{\xi_{N_s+n} : n \in \mathbb{N}^+\}$  are independent with  $(N_{t+s} - N_s : t \geq 0)$ .  $\forall \{A_n \in \mathcal{E} : n \in \mathbb{N}^+\}, k \in \mathbb{N}$ , then  $\mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{N_s+n} \in A_n\} \cap \{N_{t+s} - N_s = k\}) = \mathbb{P}(\bigcup_{i \in \mathbb{N}} (\bigcap_{n \in \mathbb{N}^+} \{\xi_{i+n} \in A_n\} \cap \{N_{t+s} = k+i, N_s = i\})) = \sum_{i \in \mathbb{N}} \mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{i+n} \in A_n\} \cap \{N_{t+s} = k+i, N_s = i\}) = \sum_{i \in \mathbb{N}} \mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{i+n} \in A_n\}) \mathbb{P}(N_{t+s} = k+i, N_s = i) = \prod_{n \in \mathbb{N}^+} \mathbb{P}(\xi_{1+n} \in A_n) \mathbb{P}(N_{t+s} = k+i, N_s = i) = \prod_{n \in \mathbb{N}^+} \mathbb{P}(\xi_{1+n} \in A_n) \mathbb{P}(N_{t+s} - N_s = k) = \prod_{n \in \mathbb{N}^+} \mathbb{P}(\xi_{N_s+n} \in A_n) \mathbb{P}(N_{t+s} - N_s = k) = \mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{N_s+n} \in A_n\}) \mathbb{P}(N_{t+s} - N_s = k)$ .

3.  $\forall 0 = t_0 < t_1 < \dots < t_n$ , then  $X_{t_1} = \sum_{i=1}^{N_{t_1}} \xi_i, X_{t_k} - X_{t_{k-1}} = \sum_{i=1}^{N_{t_k} - N_{t_{k-1}}} \xi_{N_{t_{k-1}}+i} \xi_i, k =$

$2, \dots, n$ , then  $\forall \{A_k \in \mathcal{E} : k = 1, \dots, n\}$ ,

$$\begin{aligned}
& \mathbb{P}\left(\bigcap_{k=1}^n \sum_{i=1}^{N_{t_k}-N_{t_{k-1}}} \xi_{i+N_{t_{k-1}}} \in A_k\right) \\
&= \mathbb{P}\left(\bigcup_{0 \leq u_1 \leq \dots \leq u_n} \left\{ \sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k, N_{t_k} = u_k, k = 1, \dots, n \right\}\right) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \mathbb{P}\left(\sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k, k = 1, \dots, n \mid N_{t_k} = u_k, k = 1, \dots, n\right) \mathbb{P}(N_{t_k} = u_k, k = 1, \dots, n) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \mathbb{P}\left(\sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k, k = 1, \dots, n\right) \mathbb{P}(N_{t_k} = u_k, k = 1, \dots, n) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k\right) \prod_{j=1}^n \mathbb{P}(N_{t_j} = u_j) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k\right) \prod_{j=1}^n \mathbb{P}(N_{t_j} - N_{t_{j-1}} = u_j - u_{j-1}) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k\right) \mathbb{P}(N_{t_k} - N_{t_{k-1}} = u_k - u_{k-1}) \\
&= \sum_{u_1-u_0 \in \mathbb{N}} \dots \sum_{u_n-u_{n-1} \in \mathbb{N}} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k, N_{t_k} - N_{t_{k-1}} = u_k - u_{k-1}\right) \\
&= \prod_{k=1}^n \sum_{u_k-u_{k-1} \in \mathbb{N}} \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k, N_{t_k} - N_{t_{k-1}} = u_k - u_{k-1}\right) \\
&= \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{N_{t_k}-N_{t_{k-1}}} \xi_{i+N_{t_{k-1}}} \in A_k\right)
\end{aligned} \tag{1}$$

□

**PROBLEM II**  $X$  is a poisson random measure on  $(E, \mathcal{E})$  with intensity  $\mu$ , where  $\mu$  is  $\sigma$  finite measure. Prove  $\forall f \in (E, \mathcal{E}), f \geq 0$ ,

$$\mathbb{E}e^{-X(f)} = \exp \left\{ - \int_E (1 - e^{-f(x)}) \mu(dx) \right\}.$$

**SOLUTION.** 1. First, we consider  $\mu$  is finite and  $f = \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x)$ : Since  $X$  is a poisson random measure, then  $\exists \eta \xi_1, \dots, \xi_n, \dots$ , where  $\eta \sim \text{Poisson}(\mu(E)), \{\xi_n : n \in \mathbb{N}^+\}$  are i.i.d.,  $\xi_1 \sim \bar{\mu} :=$

$\mu(E)^{-1}\mu$ .

$$\begin{aligned}
\mathbb{E}e^{-X(f)} &= \sum_{m=0}^{\infty} \mathbb{E}(\exp \left\{ -\sum_{j=1}^m \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(\xi_j) \right\}) \mathbb{P}(\eta = m) \\
&= \sum_{m=0}^{\infty} e^{-\mu(E)} \frac{\mu(E)^m}{m!} (\mathbb{E}(\exp \left\{ -\sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(\xi_1) \right\}))^m \\
&= \sum_{m=0}^{\infty} e^{-\mu(E)} \frac{1}{m!} \left( \int_E \exp \left\{ -\sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\} \mu(dx) \right)^m \\
&= \exp \left\{ -\mu(E) + \int_E \exp \left\{ -\sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\} \mu(dx) \right\} \\
&= \exp \left\{ \int_E \exp \left\{ -\sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\} - 1 \mu(dx) \right\} \\
&= \exp \left\{ \int_E -(1 - \exp \left\{ -\sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\}) \mu(dx) \right\} \\
&= \exp \left\{ \int_E -(1 - \exp \{-f(x)\}) \mu(dx) \right\}
\end{aligned} \tag{2}$$

2. Secondly, we consider  $\mu$  is finite and  $0 \leq f \in (E, \mathbb{E})$ : Then  $\exists f_j \geq 0, j \in \mathbb{N}$  is simple measurable function, such that  $f_j \rightarrow f, j \rightarrow \infty, \forall \omega \in \Omega$ . So by LCDT, we get  $\mathbb{E}e^{-X(f)} = \exp \left\{ \int_E -(1 - \exp \{-f(x)\}) \mu(dx) \right\}$ .
3. Lastly, we consider  $\mu$  is  $\sigma$  finite and  $0 \leq f \in (E, \mathbb{E})$ : Then  $E = \bigcup_{i=1}^{\infty} E_i, \forall i, \mu(E_i) < \infty$ . Let  $X_i = X \mathbb{1}_{E_i}(x)$ , so  $\mathbb{E}e^{-X(f)} = \mathbb{E}e^{-\sum_{i=1}^{\infty} X_i(f)} = \mathbb{E} \exp \left\{ \sum_{i=2}^{\infty} X_i(f) \right\} \exp \left( \int_{E_1} -(1 - \exp \{-f(x)\}) \mu(dx) \right) = \exp \left\{ \sum_{i=1}^{\infty} \int_{E_i} -(1 - \exp \{-f(x)\}) \mu(dx) \right\}$ .

□

**PROBLEM III**  $\mu$  is a finite measure,  $X$  is a poisson random measure with intensity  $\mu$  on  $(E, \mathcal{E})$ .  $\phi : (E, \mathcal{E}) \rightarrow (F, \mathcal{F})$  is measurable. Prove :  $X \circ \phi^{-1}$  is a poisson random measurable with intensity  $\mu \circ \phi^{-1}$  on  $(F, \mathcal{F})$ .

**SOLUTION**. Obviously,  $\mu \circ \phi^{-1}$  is measurable on  $(\mathcal{F})$  and finite.

1. First, we prove  $X \circ \phi^{-1}(B) \sim \text{Poisson}(\mu \circ \phi^{-1}(B)), \forall B \in \mathcal{F}$ : Since  $\forall B \in \mathcal{F}, \phi^{-1}(B) \in \mathcal{E}$ , then  $X \circ \phi^{-1}(B) = X(\phi^{-1}(B)) \sim \text{Poisson}(\mu(\phi^{-1}(B))) = \text{Poisson}(\mu \circ \phi^{-1}(B))$ .
2. Secondly,  $\forall B_i \in \mathcal{F}, i \in \mathbb{N}, B_i \cap B_j = \emptyset, i \neq j$ , Then  $\phi^{-1}(B_i) \cap \phi^{-1}(B_j) = \emptyset$ , then  $X(\phi^{-1}(B_i)), i \in \mathbb{N}$  are independent. Besides,  $X \circ \phi^{-1}(\bigcup_{i \in \mathbb{N}} B_i) = X(\phi^{-1}(\bigcup_{i \in \mathbb{N}} B_i)) = X(\bigcup_{i \in \mathbb{N}} \phi^{-1}(B_i)) = \sum_{i \in \mathbb{N}} X(\phi^{-1}(B_i)) = \sum_{i \in \mathbb{N}} X \circ \phi^{-1}(B_i)$ .

□

**PROBLEM IV**  $\alpha \geq 0$  is constant,  $\mu$  is probability on  $\mathbb{R}$  and  $\mu(\{0\}) = 0$ . Let  $N(ds, dz, du)$  is a poisson random measure on  $(0, \infty) \times \mathbb{R} \times (0, \infty)$  with intensity  $ds\mu(dz)du$ .  $Y_0$  is independent with

$N(ds, dz, du)$ . Let

$$Y_t = Y_0 + \int_0^t \int_{\mathbb{R}} \int_0^\alpha z N(ds, dz, du), t > 0.$$

Prove:  $(Y_t : t \geq 0)$  is a compound poisson random process with rate  $\alpha$  and jumping distribution  $\mu$

**SOLUTION.** 1. Obviously, since  $N(ds, dz, du)$  is a poisson random measure, then  $\forall 0 = t_0 < t_1 < \dots < t_n$ ,  $Y_{t_0}, Y_{t_k} - Y_{t_{k-1}}, k = 2, \dots, n$  are independent.

2.  $\forall s, t \geq 0, \theta \in \mathbb{R}$ , then

$$\begin{aligned} \mathbb{E}e^{i\theta(Y_{s+t}-Y_s)} &= \mathbb{E} \exp \left\{ \int_s^{s+t} \int_{\mathbb{R}} \int_0^\alpha i\theta z N(ds, du, dz) \right\} \\ &= \exp \left\{ t\alpha \int_{\mathbb{R}} (e^{i\theta z} - 1) \mu(dz) \right\} \\ &= \exp(-t\alpha) \sum_{k=0}^{\infty} \frac{1}{k!} (t\alpha \int_{\mathbb{R}} e^{i\theta z} \mu(dz))^k \\ &= e^{-t\alpha} \sum_{k=0}^{\infty} \frac{(\alpha k)^k}{k!} \int_{\mathbb{R}} e^{i\theta z} \mu^{*k}(dz) \end{aligned} \tag{3}$$

□

**PROBLEM**  $\mu$  is a finite measure,  $X$  is a poisson random measure with intensity  $\mu$  on  $(E, \mathcal{E})$ . Prove:

1.  $\mathbb{E}[X(f)e^{-X(g)}] = \mu(fe^{-g})\mathbb{E}[e^{-X(g)}]$
2.  $\mathbb{E}[X(f)^2e^{-X(g)}] = [\mu(f^2e^{-g}) + \mu(fe^{-g})^2]\mathbb{E}[e^{-X(g)}]$

1.  $\forall \theta \geq 0$ , then  $\mathbb{E}e^{-X(\theta f+g)} = \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\}$ . Since  $e^{-X(\theta f+g)} \geq 0$ , then  $\mathbb{E}[X(f)e^{-X(\theta f+g)}] = \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \int_E f(x) e^{-\theta f(x)-g(x)} \mu(dx)$ . Thus, when  $\theta = 0$ , we can get  $\mathbb{E}[X(f)e^{-X(g)}] = \mu(fe^{-g})\mathbb{E}[e^{-X(g)}]$ .

2.  $\forall \theta \geq 0$ , then  $\mathbb{E}[X(f)^2e^{-X(\theta f+g)}] = \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \int_E f(x)^2 e^{-\theta f(x)-g(x)} \mu(dx)$ . Since  $f \geq 0, e^{-X(\theta f+g)} \geq 0$ , then

$$\begin{aligned} &\mathbb{E}[X(f)^2e^{-X(\theta f+g)}] \\ &= \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \left( \int_E f(x) e^{-\theta f(x)-g(x)} \mu(dx) \right)^2 \\ &\quad + \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \int_E f(x)^2 e^{-\theta f(x)-g(x)} \mu(dx). \end{aligned}$$

Thus, when  $\theta = 0$ , we can get  $\mathbb{E}[X(f)^2e^{-X(g)}] = [\mu(f^2e^{-g}) + \mu(fe^{-g})^2]\mathbb{E}[e^{-X(g)}]$ .