

PROBLEM I Assume $(N_t : t \geq 0)$ is a Poisson process with parameter $\alpha \geq 0$ and initial value 0. $\{\xi_n : n \in \mathbb{N}\}$ are i.i.d. r.v. with distribution μ and independent with $(N_t : t \geq 0)$. Let $X_t = \sum_{n=0}^{N_t} \xi_n, t \geq 0, \forall s \geq 0$,

1. $(N_{s+t} - N_s : t \geq 0)$ is a Poisson process with parameter α .
2. $\{\xi_{N_s+n} : n \in \mathbb{N}^+\}$ are i.i.d. with distribution μ and are independent with $(N_{s+t} - N_s : t \geq 0)$.
3. $(X_t : t \geq 0)$ satisfies $\forall 0 = t_0 < t_1 < \dots < t_n, X_{t_1}, X_{t_k} - X_{t_{k-1}}, k = 2, \dots, n$ are independent.

SOLUTION. 1. Suppose $(N_t : t \geq 0)$ satisfies $N_t - N_s \sim \text{Poisson}(\alpha(t-s)), \forall t \geq s \geq 0$, then $(N_{t+r} - N_r) - (N_{s+r} - N_r) = N_{t+r} - N_{s+r} \sim \text{Poisson}(\alpha(t-s))$. Besides, $\forall 0 = t_0 < t_1 < t_2 < \dots < t_n, D_t := N_{t+r} - N_r$, then $D_{t_k} - D_{t_{k-1}} = N_{t_k+r} - N_{t_{k-1}+r}, \forall k = 1, \dots, n$. So $D_{t_k} - D_{t_{k-1}} = N_{s_k} - N_{s_{k-1}}, \forall k = 1, \dots, n$, where $s_k = t_k + r$, are independent to each other. Besides, $N_{0+r} - N_r = 0$, obviously $D_{t_k} - D_{t_{k-1}}$ is independent to $N_{0+r} - N_0$. Last, since the orbit of $(N_t : t \geq 0)$ is continuous, then $N_{t+r} - N_r$ is continuous for any $t \geq 0$. Thus, by the definition of Poisson process, we get $(N_{t+r} - N_r : t \geq 0)$ is Poisson process.

2. Assume $(\Omega, \mathcal{F}), (\mathbb{N}, \mathcal{P}(\mathbb{N})), (E, \mathcal{E})$ are sigma algebra. $\xi_n : \Omega \rightarrow E, N_t : \Omega \rightarrow \mathbb{N}$. First of all, we prove that $\forall n \in \mathbb{N}^+, \xi_{N_s+n}$ has contribution μ : $\forall A \in \mathcal{E}, \mathbb{P}(\{\xi_{N_s+n} \in A\}) = \mathbb{P}(\bigcup_{k=0}^{\infty} \{\xi_{k+n} \in A, N_s = k\}) = \sum_{k=0}^{\infty} \mathbb{P}(\xi_{k+n} \in A, N_s = k) = \sum_{k=0}^{\infty} \mathbb{P}(\xi_{k+n} \in A) \mathbb{P}(N_s = k) = \sum_{k=0}^{\infty} \mathbb{P}(\xi_1 \in A) \mathbb{P}(N_s = k) = \mathbb{P}(\xi_1 \in A)$.

Secondly, we prove that $\{\xi_{N_s+n} : n \in \mathbb{N}^+\}$ are independent: $\forall J \subset \mathbb{N}^+, \text{card}(J) < \infty, \{A_i \in \mathcal{E} : i \in J\}$, then $\mathbb{P}(\bigcap_{i \in J} \{\xi_{N_s+i} \in A_i\}) = \mathbb{P}(\bigcup_{k=0}^{\infty} (\bigcap_{i \in J} \{\xi_{k+i} \in A_i, N_s = k\})) = \sum_{k=0}^{\infty} \mathbb{P}(\bigcap_{i \in J} \{\xi_{k+i} \in A_i \in A_i\} \cap \{N_s = k\}) = \sum_{k=0}^{\infty} \mathbb{P}(\bigcap_{i \in J} \{\xi_{k+i} : i \in A_i\}) \mathbb{P}(N_s = k) = \sum_{k=0}^{\infty} \mathbb{P}(\bigcap_{i \in J} \{\xi_{1+i} \in A_i\}) \mathbb{P}(N_s = k) = \mathbb{P}(\bigcap_{i \in J} \{\xi_{1+i} \in A_i\}) = \prod_{i \in J} \mathbb{P}(\xi_{1+i} \in A_i) = \prod_{i \in J} \mathbb{P}(\xi_{N_s+i} \in A_i)$.

Last, we will prove that $\{\xi_{N_s+n} : n \in \mathbb{N}^+\}$ are independent with $(N_{t+s} - N_s : t \geq 0)$. $\forall \{A_n \in \mathcal{E} : n \in \mathbb{N}^+\}, k \in \mathbb{N}$, then $\mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{N_s+n} \in A_n\} \cap \{N_{t+s} - N_s = k\}) = \mathbb{P}(\bigcup_{i \in \mathbb{N}} (\bigcap_{n \in \mathbb{N}^+} \{\xi_{i+n} \in A_n\} \cap \{N_{t+s} = k+i, N_s = i\})) = \sum_{i \in \mathbb{N}} \mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{i+n} \in A_n\} \cap \{N_{t+s} = k+i, N_s = i\}) = \sum_{i \in \mathbb{N}} \mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{i+n} \in A_n\}) \mathbb{P}(N_{t+s} = k+i, N_s = i) = \prod_{n \in \mathbb{N}^+} \mathbb{P}(\xi_{1+n} \in A_n) \mathbb{P}(N_{t+s} = k+i, N_s = i) = \prod_{n \in \mathbb{N}^+} \mathbb{P}(\xi_{1+n} \in A_n) \mathbb{P}(N_{t+s} - N_s = k) = \prod_{n \in \mathbb{N}^+} \mathbb{P}(\xi_{N_s+n} \in A_n) \mathbb{P}(N_{t+s} - N_s = k) = \mathbb{P}(\bigcap_{n \in \mathbb{N}^+} \{\xi_{N_s+n} \in A_n\}) \mathbb{P}(N_{t+s} - N_s = k)$.

3. $\forall 0 = t_0 < t_1 < \dots < t_n$, then $X_{t_1} = \sum_{i=1}^{N_{t_1}} \xi_i, X_{t_k} - X_{t_{k-1}} = \sum_{i=1}^{N_{t_k} - N_{t_{k-1}}} \xi_{N_{t_{k-1}}+i} \xi_i, k =$

$2, \dots, n$, then $\forall \{A_k \in \mathcal{E} : k = 1, \dots, n\}$,

$$\begin{aligned}
& \mathbb{P}\left(\bigcap_{k=1}^n \sum_{i=1}^{N_{t_k}-N_{t_{k-1}}} \xi_{i+N_{t_{k-1}}} \in A_k\right) \\
&= \mathbb{P}\left(\bigcup_{0 \leq u_1 \leq \dots \leq u_n} \left\{ \sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k, N_{t_k} = u_k, k = 1, \dots, n \right\}\right) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \mathbb{P}\left(\sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k, k = 1, \dots, n \mid N_{t_k} = u_k, k = 1, \dots, n\right) \mathbb{P}(N_{t_k} = u_k, k = 1, \dots, n) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \mathbb{P}\left(\sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k, k = 1, \dots, n\right) \mathbb{P}(N_{t_k} = u_k, k = 1, \dots, n) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=u_{k-1}+1}^{u_k} \xi_i \in A_k\right) \prod_{j=1}^n \mathbb{P}(N_{t_j} = u_j) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k\right) \prod_{j=1}^n \mathbb{P}(N_{t_j} - N_{t_{j-1}} = u_j - u_{j-1}) \\
&= \sum_{0 \leq u_1 \leq \dots \leq u_n} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k\right) \mathbb{P}(N_{t_k} - N_{t_{k-1}} = u_k - u_{k-1}) \\
&= \sum_{u_1-u_0 \in \mathbb{N}} \dots \sum_{u_n-u_{n-1} \in \mathbb{N}} \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k, N_{t_k} - N_{t_{k-1}} = u_k - u_{k-1}\right) \\
&= \prod_{k=1}^n \sum_{u_k-u_{k-1} \in \mathbb{N}} \mathbb{P}\left(\sum_{i=1}^{u_k-u_{k-1}} \xi_{u_{k-1}+i} \in A_k, N_{t_k} - N_{t_{k-1}} = u_k - u_{k-1}\right) \\
&= \prod_{k=1}^n \mathbb{P}\left(\sum_{i=1}^{N_{t_k}-N_{t_{k-1}}} \xi_{i+N_{t_{k-1}}} \in A_k\right)
\end{aligned} \tag{1}$$

□

PROBLEM II X is a poisson random measure on (E, \mathcal{E}) with intensity μ , where μ is σ finite measure. Prove $\forall f \in (E, \mathcal{E}), f \geq 0$,

$$\mathbb{E}e^{-X(f)} = \exp \left\{ - \int_E (1 - e^{-f(x)}) \mu(dx) \right\}.$$

SOLUTION. 1. First, we consider μ is finite and $f = \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x)$: Since X is a poisson random measure, then $\exists \eta \xi_1, \dots, \xi_n, \dots$, where $\eta \sim \text{Poisson}(\mu(E)), \{\xi_n : n \in \mathbb{N}^+\}$ are i.i.d., $\xi_1 \sim \bar{\mu} :=$

$\mu(E)^{-1}\mu.$

$$\begin{aligned}
\mathbb{E}e^{-X(f)} &= \sum_{m=0}^{\infty} \mathbb{E}(\exp \left\{ - \sum_{j=1}^m \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(\xi_j) \right\}) \mathbb{P}(\eta = m) \\
&= \sum_{m=0}^{\infty} e^{-\mu(E)} \frac{\mu(E)^m}{m!} (\mathbb{E}(\exp \left\{ - \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(\xi_1) \right\}))^m \\
&= \sum_{m=0}^{\infty} e^{-\mu(E)} \frac{1}{m!} \left(\int_E \exp \left\{ - \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\} \mu(dx) \right)^m \\
&= \exp \left\{ -\mu(E) + \int_E \exp \left\{ - \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\} \mu(dx) \right\} \\
&= \exp \left\{ \int_E \exp \left\{ - \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\} - 1 \mu(dx) \right\} \\
&= \exp \left\{ \int_E -(1 - \exp \left\{ - \sum_{i=1}^n \theta_i \mathbb{1}_{B_i}(x) \right\}) \mu(dx) \right\} \\
&= \exp \left\{ \int_E -(1 - \exp \{-f(x)\}) \mu(dx) \right\}
\end{aligned} \tag{2}$$

2. Secondly, we consider μ is finite and $0 \leq f \in (E, \mathbb{E})$: Then $\exists f_j \geq 0, j \in \mathbb{N}$ is simple measurable function, such that $f_j \rightarrow f, j \rightarrow \infty, \forall \omega \in \Omega$. So by LCDT, we get $\mathbb{E}e^{-X(f)} = \exp \left\{ \int_E -(1 - \exp \{-f(x)\}) \mu(dx) \right\}$.
3. Lastly, we consider μ is σ finite and $0 \leq f \in (E, \mathbb{E})$: Then $E = \bigcup_{i=1}^{\infty} E_i, \forall i, \mu(E_i) < \infty$. Let $X_i = X \mathbb{1}_{E_i}(x)$, so $\mathbb{E}e^{-X(f)} = \mathbb{E}e^{-\sum_{i=1}^{\infty} X_i(f)} = \mathbb{E} \exp \left\{ \sum_{i=2}^{\infty} X_i(f) \right\} \exp \left(\int_{E_1} -(1 - \exp \{-f(x)\}) \mu(dx) \right) = \exp \left\{ \sum_{i=1}^{\infty} \int_{E_i} -(1 - \exp \{-f(x)\}) \mu(dx) \right\}$.

□

PROBLEM III μ is a finite measure, X is a poisson random measure with intensity μ on (E, \mathcal{E}) . $\phi : (E, \mathcal{E}) \rightarrow (F, \mathcal{F})$ is measurable. Prove : $X \circ \phi^{-1}$ is a poisson random measurable with intensity $\mu \circ \phi^{-1}$ on (F, \mathcal{F}) .

SOLUTION. Obviously, $\mu \circ \phi^{-1}$ is measurable on (\mathcal{F}) and finite.

1. First, we prove $X \circ \phi^{-1}(B) \sim \text{Poisson}(\mu \circ \phi^{-1}(B)), \forall B \in \mathcal{F}$: Since $\forall B \in \mathcal{F}, \phi^{-1}(B) \in \mathcal{E}$, then $X \circ \phi^{-1}(B) = X(\phi^{-1}(B)) \sim \text{Poisson}(\mu(\phi^{-1}(B))) = \text{Poisson}(\mu \circ \phi^{-1}(B))$.
2. Secondly, $\forall B_i \in \mathcal{F}, i \in \mathbb{N}, B_i \cap B_j = \emptyset, i \neq j$, Then $\phi^{-1}(B_i) \cap \phi^{-1}(B_j) = \emptyset$, then $X(\phi^{-1}(B_i)), i \in \mathbb{N}$ are independent. Besides, $X \circ \phi^{-1}(\bigcup_{i \in \mathbb{N}} B_i) = X(\phi^{-1}(\bigcup_{i \in \mathbb{N}} B_i)) = X(\bigcup_{i \in \mathbb{N}} \phi^{-1}(B_i)) = \sum_{i \in \mathbb{N}} X(\phi^{-1}(B_i)) = \sum_{i \in \mathbb{N}} X \circ \phi^{-1}(B_i)$.

□

PROBLEM IV $\alpha \geq 0$ is constant, μ is probability on \mathbb{R} and $\mu(\{0\}) = 0$. Let $N(ds, dz, du)$ is a poisson random measure on $(0, \infty) \times \mathbb{R} \times (0, \infty)$ with intensity $ds\mu(dz)du$. Y_0 is independent with

$N(ds, dz, du)$. Let

$$Y_t = Y_0 + \int_0^t \int_{\mathbb{R}} \int_0^\alpha z N(ds, dz, du), t > 0.$$

Prove: $(Y_t : t \geq 0)$ is a compound poisson random process with rate α and jumping distribution μ

SOLUTION. 1. Obviously, since $N(ds, dz, du)$ is a poisson random measure, then $\forall 0 = t_0 < t_1 < \dots < t_n$, $Y_{t_0}, Y_{t_k} - Y_{t_{k-1}}, k = 2, \dots, n$ are independent.

2. $\forall s, t \geq 0, \theta \in \mathbb{R}$, then

$$\begin{aligned} \mathbb{E}e^{i\theta(Y_{s+t}-Y_s)} &= \mathbb{E} \exp \left\{ \int_s^{s+t} \int_{\mathbb{R}} \int_0^\alpha i\theta z N(ds, du, dz) \right\} \\ &= \exp \left\{ t\alpha \int_{\mathbb{R}} (e^{i\theta z} - 1) \mu(dz) \right\} \\ &= \exp(-t\alpha) \sum_{k=0}^{\infty} \frac{1}{k!} (t\alpha \int_{\mathbb{R}} e^{i\theta z} \mu(dz))^k \\ &= e^{-t\alpha} \sum_{k=0}^{\infty} \frac{(\alpha k)^k}{k!} \int_{\mathbb{R}} e^{i\theta z} \mu^{*k}(dz) \end{aligned} \tag{3}$$

□

PROBLEM μ is a finite measure, X is a poisson random measure with intensity μ on (E, \mathcal{E}) . Prove:

1. $\mathbb{E}[X(f)e^{-X(g)}] = \mu(fe^{-g})\mathbb{E}[e^{-X(g)}]$
2. $\mathbb{E}[X(f)^2e^{-X(g)}] = [\mu(f^2e^{-g}) + \mu(fe^{-g})^2]\mathbb{E}[e^{-X(g)}]$

1. $\forall \theta \geq 0$, then $\mathbb{E}e^{-X(\theta f+g)} = \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\}$. Since $e^{-X(\theta f+g)} \geq 0$, then $\mathbb{E}[X(f)e^{-X(\theta f+g)}] = \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \int_E f(x) e^{-\theta f(x)-g(x)} \mu(dx)$. Thus, when $\theta = 0$, we can get $\mathbb{E}[X(f)e^{-X(g)}] = \mu(fe^{-g})\mathbb{E}[e^{-X(g)}]$.

2. $\forall \theta \geq 0$, then $\mathbb{E}[X(f)^2e^{-X(\theta f+g)}] = \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \int_E f(x) e^{-\theta f(x)-g(x)} \mu(dx)$. Since $f \geq 0, e^{-X(\theta f+g)} \geq 0$, then

$$\begin{aligned} &\mathbb{E}[X(f)^2e^{-X(\theta f+g)}] \\ &= \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \left(\int_E f(x) e^{-\theta f(x)-g(x)} \mu(dx) \right)^2 \\ &\quad + \exp \left\{ - \int_E (1 - e^{-\theta f(x)-g(x)}) \mu(dx) \right\} \int_E f(x)^2 e^{\theta f(x)-g(x)} \mu(dx). \end{aligned}$$

Thus, when $\theta = 0$, we can get $\mathbb{E}[X(f)^2e^{-X(g)}] = [\mu(f^2e^{-g}) + \mu(fe^{-g})^2]\mathbb{E}[e^{-X(g)}]$.