

**PROBLEM I** Let  $X = \{X(n) : n \geq 0\}$  be Markov chain defined on probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ , with state space  $E$  and transition probability matrix  $P = (p(i, j) : i, j \in E)$ . Let  $a, b \in E$ ,  $\tau_0 = 0$ ,  $\sigma_k = \inf\{n \geq \tau_{k-1} : X(n) = b\}$ ,  $\tau_k = \inf\{n \geq \sigma_k : X(n) = a\}$ . Prove:  $\tau_n, \sigma_n, n \geq 1$  are all stopping time on  $(\mathcal{F}_n : n \geq 0)$ .

**SOLUTION**. Since  $\sigma_1 = \inf\{n \geq 0 : X(n) = b\}$ , then  $\{\sigma_1 = m\} = \{X(i) \neq b, 0 \leq i \leq m-1, X(m) = b\} \in \sigma\{X(i), 0 \leq i \leq m-1, X_m\} = \mathcal{F}_m$ . Then  $\sigma_1$  is stopping time. Next, we will prove  $\sigma_k, \tau_{k-1}, k \in \mathbb{N}$  are stopping time. Since  $\sigma_1, \tau_0$  are stopping time, which we have proved. Assume  $\sigma_k, \tau_{k-1}$  are stopping time, we will prove  $\sigma_{k+1}, \tau_k$  are stopping time. Let  $m \in \mathbb{N}^+$ ,  $\{\tau_k = m\} = \bigcup_{i \in [0, m-1] \cap \mathbb{N}} \{\sigma_k = i, X(i+l) \neq a, 1 \leq l \leq m-i-1, X(m) = a\} \in \mathcal{F}_m$ , since  $\{\sigma_k = i\} \in \mathcal{F}_i \subset \mathcal{F}_m, \forall 0 \leq i \leq m-1, \sigma(X_j) \subset \mathcal{F}_m, j \leq m$ . Let  $m \in \mathbb{N}^+$ ,  $\{\sigma_{k+1} = m\} = \bigcup_{i \in [0, m-1] \cap \mathbb{N}} \{\tau_k = i, X(i+l) \neq a, 1 \leq l \leq m-i-1, X(m) = b\} \in \mathcal{F}_m$ , since  $\{\tau_k = i\} \in \mathcal{F}_i \subset \mathcal{F}_m, \forall 0 \leq i \leq m-1, \sigma(X_j) \subset \mathcal{F}_m, j \leq m$ . Therefore,  $\sigma_{k+1}, \tau_k$  are stopping time.  $\square$

**PROBLEM II** Let  $(X_n : n \geq 0)$  is a one-dimension simple random walk starting at 1. Let  $e(n) = \{X_{n \wedge \tau_1} : n \geq 0\}$ , where  $\tau_1 = \inf\{n \geq 0 : X_n = 0\}$ . Find the distribution of  $\sup_{n \geq 0} e(n)$ .

**SOLUTION**. Assume  $\mathbb{P}(X_{n+1} - X_n = 1) = p, \mathbb{P}(X_{n+1} - X_n = -1) = q, p, q > 0, p + q = 1$ . Let  $E := \sup_{n \geq 0} e(n)$ . Let  $\gamma = \inf\{n \geq 0 : X_n = 0 \text{ or } X_n = m\} \geq 1$ , where  $m \in \mathbb{N}^+$ . First of all, if  $p = q = \frac{1}{2}$ : Easy to get that  $\gamma < \infty, a.s.$ , so  $X_{n \wedge \gamma} \xrightarrow{a.s.} X_\gamma$ . And  $0 \leq X_{n \wedge \gamma} \leq m$ , so  $\mathbb{E}(X_\gamma) = \mathbb{E}(X_{n \wedge \gamma}) = \mathbb{E}(X_0)$ . Noting that  $\{X_\gamma = 0\} \stackrel{a.s.}{=} \{E < m\}$  and  $\{X_\gamma = m\} \stackrel{a.s.}{=} \{E \geq m\}$ , we get two equations:

$$\begin{cases} \mathbb{P}(E < m) + \mathbb{P}(E \geq m) = 1 \\ 0\mathbb{P}(E < m) + m\mathbb{P}(E \geq m) = 1 \end{cases}.$$

Solve this equation, we get  $\mathbb{P}(E \geq m) = \frac{1}{m}$ . So  $\mathbb{P}(E = m) = \frac{1}{m(m+1)}$ , and easily  $\mathbb{P}(E = \infty) = 0$ .

Secondly,  $p \neq q$ : Let  $Y_n := (\frac{q}{p})^{X_n}$ , then  $\mathbb{E}(Y_{n+1}) = \mathbb{E}(\mathbb{E}(Y_{n+1} | Y_n)) = \mathbb{E}((\frac{q}{p})^{X_n+1}p + (\frac{q}{p})^{X_n-1}q) = \mathbb{E}((\frac{q}{p})^{X_n}) = \mathbb{E}(Y_n)$ . Obviously,  $\gamma$  is stopping time. Then  $\mathbb{E}(Y_{n+1 \wedge \gamma}) = \mathbb{E}(\mathbb{E}(Y_{n+1 \wedge \gamma} | Y_{n \wedge \gamma})) = \mathbb{E}(\mathbb{1}_{n < \gamma} \mathbb{E}(Y_{n+1} | Y_n) + \mathbb{1}_{n \geq \gamma} \mathbb{E}(Y_\gamma | Y_\gamma)) = \mathbb{E}(\mathbb{1}_{n < \gamma} Y_n + \mathbb{1}_{n \geq \gamma} \mathbb{E}(Y_\gamma)) = \mathbb{E}(Y_{\gamma \wedge n})$ . Then  $\mathbb{E}(Y_{n \wedge \gamma}) = \mathbb{E}(Y_0) = \frac{q}{p}$ . Besides,  $\gamma < \infty, a.s.$ ,  $Y_{n \wedge \gamma} \xrightarrow{a.s.} Y_\gamma$ ,  $Y_n < \infty, a.s.$ . Then  $\mathbb{E}(Y_{n \wedge \gamma}) \rightarrow \mathbb{E}(Y_\gamma) = \mathbb{P}(T_0 < T_m) + (\frac{q}{p})^m \mathbb{P}(T_0 > T_m)$ , where  $T_i = \inf\{k \geq 1 : X_k = i\}, i \in \mathbb{N}$ . Additionally,  $\{T_0 < T_m\} \stackrel{a.s.}{=} \{E < m\}, \{T_0 > T_m\} \stackrel{a.s.}{=} \{E \geq m\}$ . Then  $\mathbb{P}(E < m) + (\frac{q}{p})^m \mathbb{P}(E \geq m) = \frac{q}{p}, \mathbb{P}(E < m) + \mathbb{P}(E \geq m) = 1$ . Thus,  $\mathbb{P}(E \geq m) = \frac{p^m - qp^{m-1}}{p^m - q^m}, \mathbb{P}(E < m) = \frac{qp^{m-1} - q^m}{p^m - q^m}$ . Then  $\mathbb{P}(E = m) =$

$$\frac{(\frac{p}{q})^m (\frac{p}{q} - 1)}{((\frac{p}{q})^m - 1)((\frac{p}{q})^{m+1} - 1)}. \text{ Furthermore, easily } \mathbb{P}(E = \infty) = \lim_{m \rightarrow \infty} \mathbb{P}(E \geq m) = \begin{cases} 0 & \frac{q}{p} > 1 \\ 1 - \frac{q}{p} & \frac{q}{p} < 1 \end{cases}. \quad \square$$

**PROBLEM III** Prove:

1. When  $0 < p \leq q$ , the reflecting random walk with transition matrix  $Q_+^a$  is recurrent.
2. When  $0 < q \leq p$ , the reflecting random walk with transition matrix  $Q_-^a$  is recurrent.

*SOLUTION*. By symmetry, only need to prove the first question. Without loss of generality we can assume  $a = 0$ . We consider the equation

$$y_0 = y_1, \forall i \geq 1, y_i = qy_{i-1} + py_{i+1}$$

Only need to prove its all bounded solution are all constant. Easy to get  $y_{i+2} = \frac{1}{p}y_{i+1} - \frac{q}{p}y_i$ . Consider the characteristic equation of this sequence,  $x^2 - \frac{x}{p} + \frac{q}{p} = 0$ . We get  $x_1 = 1, x_2 = \frac{q}{p} \geq 1$ . If  $x_2 > 1$ , then  $y_n = c_1x_1^n + c_2x_2^n$  is bounded  $\iff c_2 = 0$ , so  $y_n = c_1x_1^n = c_1$  is constant. Else,  $x_2 = x_1 = 1$ , then  $y_n = (an + b)x_1^n = an + b$  is bounded  $\iff a = 0$ , so  $y_n = b$  is constant. So the Markov chain is recurrent.  $\square$

**PROBLEM IV** Let  $\phi_0(n : n \in \mathbb{N}^+)$  be simple random walk begin at  $\phi_0(0) \geq a + 1$ , let  $\zeta_0 := \inf\{m : \phi_0(m) = a + 1\}$ , let  $(W_n : n \in \mathbb{N})$  be reflecting simple random walk on  $\mathbb{Z}_+^a$ , starting at  $a + 1$ , independent with  $\phi_0$ . Let  $X_n := \begin{cases} \phi_0(n) & n \leq \zeta_0 \\ W_n - \zeta_0 & n \geq \zeta_0 \end{cases}$ . Prove that  $(X_n : n \in \mathbb{N})$  is reflecting random walk on  $\mathbb{Z}_+^a$  begin at  $\phi_0(0)$ .

*SOLUTION*. Now we consider  $n \in \mathbb{N}^+$  and  $i_0, i_1, i_2, \dots, i_{n+1} \in \mathbb{Z}_+^a$ .

1. If  $\forall k : 1 \leq k \leq n, i_k \neq a + 1$ , then we have

$$\begin{aligned} \mathbb{P}(X_0 = i_0, \dots, X_{n+1} = i_{n+1}) &= \mathbb{P}(\phi_0(0) = i_0, \dots, \phi_0(n+1) = i_{n+1}) \\ &= \mathbb{P}(\phi_0(0) = i_0, \dots, \phi_0(n) = i_n) \mathbb{P}(\phi_0(n+1) = i_{n+1} \mid \phi_0(n) = i_n) \\ &= \mathbb{P}(X_0 = i_0, \dots, X_n = i_n) q_+^a(i_n, i_{n+1}) \end{aligned}$$

2. Else, we let  $k := \inf\{m : 1 \leq m \leq n, i_m = a + 1\}$ . Then we have

$$\begin{aligned} &\mathbb{P}(X_0 = i_0, \dots, X_{n+1} = i_{n+1}) \\ &= \mathbb{P}(\phi_0(0) = i_0, \dots, \phi_0(k) = i_k, W_0 = a + 1, W_{k+2+i} = i_{k+i}, i = 1, \dots, n - k + 1) \\ &= \mathbb{P}(\phi_0(0) = i_0, \dots, \phi_0(k) = i_k) \mathbb{P}(W_0 = a + 1, W_{k+2+i} = i_{k+i}, i = 1, \dots, n - k + 1) \\ &= \mathbb{P}(\phi_0(0) = i_0, \dots, \phi_0(k) = i_k) \mathbb{P}(W_0 = a + 1, W_{k+2+i} = i_{k+i}, i = 1, \dots, n - k) q_+^a(i_n, i_{n+1}) \\ &= \mathbb{P}(X_0 = i_0, \dots, X_n = i_n) q_+^a(i_n, i_{n+1}) \end{aligned}$$

So we get  $(X_n : n \geq 0)$  is reflecting simple random walk on  $\mathbb{Z}_+^a$ .  $\square$