

For a random walk y , we let $D(n, y)$ be the number of downcrossings of y over n . **PROBLEM I** Let $S = (S_n : n \geq 0)$ be the one-dimensional symmetry simple random walk with $S_0 = c \geq 0$. Let $k \geq 1$ and τ be the time of the k -th downcrossing of 0. X_b is the times of $(S_{n \wedge \tau} : n \geq 0)$ downcrossing of b . Prove:

1. $(X_b : b \geq c - 1)$ is branch process. And offspring distribution is $Geo(\frac{1}{2})$
2. $(X_{-a} : a \geq 1)$ is branch process. And offspring distribution is $Geo(\frac{1}{2})$
3. $(X_b : 0 \leq b \leq c - 1)$ is migrating branch process. And offspring distribution is $Geo(\frac{1}{2})$

And the migrating distribution is concentrating on 1.

SOLUTION. 1. Fix $b \geq c - 1$. Let ϕ_0 be the journey from start point to $b + 1$. Let e_n be n -th journey from $b + 1$ to b . Let ε_n be n -th journey after ϕ_0 from b to $b + 1$. Then we know that e_n, ε_n are independent. Easy to get that $D(e_n, b) = 1$ and $D(\varepsilon_n, b) = 0, D(\varepsilon_n, b + 1) = 0$. Easy to get that $D((S_{n \wedge \tau} : n \in \mathbb{N}), b + 1) = \sum_{t=1}^{D((S_{n \wedge \tau} : n \in \mathbb{N}), b)} D(e_t, b + 1)$. Noting that $\forall d : c - 1 \leq d \leq b, D((S_{n \wedge \tau} : n \in \mathbb{N}), d) \in \sigma(\varepsilon_n : n \in \mathbb{N})$. We easily get that $D(e_t, b + 1) \perp \sigma(\varepsilon_n : n \in \mathbb{N})$. So X_b is Markov process. And to prove it's branch process, we only need to prove that $D(e_t, b + 1)$ are i.i.d. It has been proved that $D(e_t, b + 1)$ are i.i.d and $Geo(\frac{1}{2})$. So the offspring distribution is $Geo(\frac{1}{2})$.

2. Fix $a \geq 1$. Let ϕ_0 be the journey from start point to $-a$. Let e_n be n -th journey from $-a$ to $-a - 1$, and ε_n be n -th journey from $-a - 1$ to $-a$. Then easy to get that $X_{-a-1} = \sum_{t=1}^{X_{-a}} D(\varepsilon_t, -a - 1)$. For the same reason we easily get that $D(\varepsilon_t, -a - 1) \perp \sigma(e_n : n \in \mathbb{N})$. And by reflecting easy to get that $D(\varepsilon_t, -a - 1) \sim Geo(\frac{1}{2})$, too. So $(X_{-a} : a \geq 1)$ is branch process and offspring distribution is $Geo(\frac{1}{2})$

3. Fix $b < c - 1$. Let ϕ_0 be the journey from start point to $b + 1$. Let e_n be the n -th journey from $b + 1$ to b and ε_n be n -th journey from b to $b + 1$. Then easy to prove that $X_{b+1} = D(\phi_0, b + 1) + \sum_{t=1}^{X_b} D(e_n, b + 1)$. Noting that $D(\phi_0, b + 1) = 1$. So for the same reason, we get that $(X_b : 0 \leq b \leq c - 1)$ is migrating branch process, with offspring distribution $Geo(\frac{1}{2})$ and migrating distribution δ_1 .

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SOLUTION.

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PROBLEM II $c < b \in \mathbb{Z}_+$. Let $W = (W_n : n \geq 0)$ be the one-dimensional reflecting symmetry simple random walk with $W_0 = c \geq 0$ on $\mathbb{Z}^{0,b}$, whose transition matrix is $P^{0,b}$, where $a = 0, p, q > 0, p + q = 1$. Let $k \geq 1$ and τ is the time of the k -th downcrossing 0 on (W_n) . $0 \leq a \leq b$, X_a is the times of $(S_{n \wedge \tau} : n \geq 0)$ downcrossing a . Prove:

1. $(X_a : c - 1 \leq a \leq b - 1)$ is branch process. And offspring distribution is $Geo(p)$.
2. $(X_a : 0 \leq a \leq c - 1)$ is migrating branch process. And offspring distribution is $Geo(p)$. And the migrating distribution is concentrating on 1.

SOLUTION. 1. Fix a such that $c - 1 \leq a < b - 1$. Let ϕ_0 be the journey from start point to a . Let e_n be the n -th journey from a to $a + 1$, and ε_n be the n -th journey from $a + 1$ to a . For reflecting simple random walk, we can also prove that e_n, ε_n are independent. Noting that $X_{a+1} = \sum_{t=1}^{X_a} D(\varepsilon_t, a + 1)$, we easily get the conclusion.

2. Fix $a : 0 \leq a < c - 1$. Let ϕ_0 be the journey from start point to $a + 1$. Let e_n be the n -th journey from $a + 1$ to a and ε_n be n -th journey from a to $a + 1$. Then easy to prove that $X_{a+1} = D(\phi_0, a + 1) + \sum_{t=1}^{X_a} D(e_n, a + 1)$. Noting that $D(\phi_0, a + 1) = 1$. So for the same reason, we get that $(X_a : 0 \leq a \leq c - 1)$ is migrating branch process, with offspring distribution $Geo(\frac{1}{2})$ and migrating distribution δ_1 . □

PROBLEM III Let $W = (W_n : n \geq 0)$ be the one-dimensional simple random walk with $W_0 = 0$, whose transition matrix P given by equation (4.4.3) on textbook, $0 < p < q < 1$. X_a is the times of $(W_{n \wedge \tau} : n \geq 0)$ downcrossing a . $r = \frac{p}{q}$. Prove:

1. $\mathbb{P}(X_0 = i) = r^i(1 - r), i \geq 0$;
2. $a \geq 0, \mathbb{P}(X_a = 0) = 1 - r^{a+1}, \mathbb{P}(X_a = i) = r^{a+i}(1 - r), i \geq 1$.

SOLUTION. 1. Since $p < q$, then $W_n \rightarrow -\infty, n \rightarrow \infty$. Let $\tau_0 = 0, \forall k \geq 1, \sigma_k = \inf\{n \geq \tau_{k-1} : W_n = 1\}, \tau_k = \inf\{n \geq \sigma_k : W_n = 0\}$.

- (a) If $i = 0$, then $\{X_0 = i\} \stackrel{\text{a.s.}}{=} \{\sigma_1 = \infty\}$. Then $\mathbb{P}(X_0 = i) = \mathbb{P}(\sigma_1 = \infty) = r$.
- (b) If $i \geq 1$, then $\{X_0 = i\} \stackrel{\text{a.s.}}{=} \{\sigma_i < \infty, \sigma_{i+1} = \infty\}$. Since $\{\tau_i < \infty\} \subset \{\sigma_i < \infty\}, \mathbb{P}(\sigma_i < \infty, \tau_i = \infty) = 0$, then by strong markov property,

$$\begin{aligned} \mathbb{P}(\sigma_{i+1} < \infty \mid \sigma_i < \infty) &= \mathbb{P}(\sigma_{i+1} < \infty \mid \sigma_i < \infty, \tau_i < \infty) \\ &= \mathbb{P}(\sigma_{i+1} < \infty \mid \tau_i < \infty) \\ &= \mathbb{P}(\sigma_{i+1} < \infty \mid \tau_i < \infty, W_{\tau_i} = 0) \\ &= \mathbb{P}(\sigma_1 < \infty) = r \end{aligned}$$

Therefore,

$$\mathbb{P}(\sigma_{i+1} < \infty) = \mathbb{P}(\sigma_{i+1} < \infty \mid \sigma_i < \infty) \mathbb{P}(\sigma_i < \infty)$$

Then $\mathbb{P}(\sigma_i < \infty) = r^i$. Therefore, $\mathbb{P}(X_0 = i) = \mathbb{P}(\sigma_i < \infty, \sigma_{i+1} = \infty) = \mathbb{P}(\sigma_i < \infty) \mathbb{P}(\sigma_{i+1} = \infty \mid \sigma_i < \infty) = r^i(1 - r)$.

2. Let $D_a = \inf(n \geq 0 : W_n = a)$, then $\mathbb{P}(D_a < \infty) = r^a$. By strong markov property, $(W_{D_a+n-a} : n \geq 0)$ is a random walk starting from 0 under $\mathbb{P}(\cdot \mid D_a < \infty) = \mathbb{P}(\cdot \mid D_a < \infty, W_{D_a} = a)$. By the conclusion in 1, $\mathbb{P}(X_a = i \mid D_a < \infty) = r^i(1 - r), i \geq 0$. Then

$$\begin{aligned} \mathbb{P}(X_a = 0) &= \mathbb{P}(D_a = \infty) + \mathbb{P}(D_a < \infty, X_a = 0) \\ &= 1 - r^a + \mathbb{P}(D_a < \infty) \mathbb{P}(X_a = 0 \mid D_a < \infty) \\ &= 1 - r^a + r^a(1 - r) = 1 - r^{a+1} \end{aligned}$$

$\forall i \geq 1,$

$$\begin{aligned}\mathbb{P}(X_a = i) &= \mathbb{P}(D_a < \infty, X_a = i) \\ &= \mathbb{P}(D_a < \infty) \mathbb{P}(X_a = i \mid D_a < \infty) \\ &= r^a r^i (1 - r) = r^{a+i} (1 - r)\end{aligned}$$

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PROBLEM IV Let $W = (W_n : n \geq 0)$ be the one-dimensional simple random walk with $W_0 = 0$, whose transition matrix P given by equation (4.4.3) on textbook, $0 < p < q < 1$. X_a is the times of $(W_{n \wedge \tau} : n \geq 0)$ downcrossing a . $r = \frac{p}{q}$. Prove: if $a \leq -1$, then $X_a - 1 \sim \text{Geo}(1 - r)$, i.e. $\mathbb{P}(X_a - 1 = i) = r^{i-1} (1 - r), i \geq 1$.

SOLUTION. Let $D_a = \inf\{n \geq 0 : X_n = a\}$. Then $\lim_{n \rightarrow \infty} W_n = -\infty$, so $\mathbb{P}(D_a < \infty) = 1$. By strong markov property, $(W_{D_a+n} - a : n \geq 0)$ is a random walk starting at 0 with the downcrossing $X_a - 1$ on 0. By the conclusion of III, we get $X_a - 1 \sim G(1 - r)$. □