

PROBLEM I Assume $(N_t : t \geq 0)$ is a Poisson process with parameter α . Let $P(t) := \mathbb{P}(2 \nmid N_t)$, $Q(t) := \mathbb{P}(2 \mid N_t)$. Prove that $P(t) = e^{-\alpha t} \sinh(\alpha t)$, $Q(t) = e^{-\alpha t} \cosh(\alpha t)$.

SOLUTION. Since $(N_t : t \geq 0)$, then $\mathbb{P}(N_t = k) = \frac{(\alpha t)^k}{k!} e^{-\alpha t}$. Then $\mathbb{P}(2 \nmid N_t) = \mathbb{P}(\bigcup_{k \in \mathbb{N}} \{N_t = 2k + 1\}) = \sum_{k \in \mathbb{N}} \mathbb{P}(N_t = 2k + 1) = \sum_{k=0}^{\infty} \frac{(\alpha t)^{2k+1}}{(2k+1)!} e^{-\alpha t} = e^{-\alpha t} \sinh(\alpha t)$. Then $\mathbb{P}(2 \mid N_t) = \mathbb{P}(\bigcup_{k \in \mathbb{N}} \{N_t = 2k\}) = \sum_{k \in \mathbb{N}} \mathbb{P}(N_t = 2k) = \sum_{k=0}^{\infty} \frac{(\alpha t)^{2k}}{(2k)!} e^{-\alpha t} = e^{-\alpha t} \cosh(\alpha t)$. \square

PROBLEM II Assume $(N_t : t \geq 0)$ is a Poisson process with parameter α . Prove that $\lim_{t \rightarrow \infty} \frac{N_t}{t} = \alpha$, a.s..

SOLUTION. First of all, we prove $\mathbb{P}(\forall 0 \leq s \leq t, N_s \leq N_t) = 1$. For some $s, t \in \mathbb{Q}$, $0 \leq s \leq t$, we have $\mathbb{P}(N_s > N_t) = 0$ since $N_t - N_s \sim \text{Poisson}(\alpha(t-s))$. So we get $\mathbb{P}(\exists s, t \in \mathbb{Q}, 0 \leq s \leq t, N_s > N_t) = 0$. Now we will prove $\exists s, t \in \mathbb{R}, 0 \leq s \leq t, N_s > N_t \iff \exists a, b \in \mathbb{Q}, 0 \leq a \leq b, N_a > N_b$. Since $\exists a, b \in \mathbb{Q} \subset \mathbb{R}, 0 \leq a \leq b, N_a > N_b$, then we only need to prove $\exists s, t \in \mathbb{R}, 0 \leq s \leq t, N_s > N_t \implies \exists a, b \in \mathbb{Q}, 0 \leq a \leq b, N_a > N_b$. Let $a_n = \frac{[ns]}{n}, b_n = \frac{[nt]}{n}$. Then $\lim a_n = s, \lim b_n = t$. Easily $a_n \geq s, b_n \geq t$. So since N_t is continuous we get $\lim N_{a_n} = N_s, \lim N_{b_n} = N_t$. Since $N_s > N_t$, we get $\exists n, N_{a_n} > N_{b_n}$. Let $a = a_n, b = b_n$ will work. So $\mathbb{P}(\forall 0 \leq s \leq t, N_s \leq N_t) = 1 - \mathbb{P}(\exists 0 \leq s \leq t, N_s > N_t) = 1 - \mathbb{P}(\exists s, t \in \mathbb{Q}, 0 \leq s \leq t, N_s > N_t) = 1 - 0 = 1$.

Let $X_0 = N_0, X_k = N_k - N_{k-1}, k = 1, \dots, n, \dots$, then $N_0 = 0$, a.s., $X_k \sim P(\alpha), k \in \mathbb{N}$. Then $\frac{\sum_{i=0}^n X_i}{n} = \frac{N_n}{n} \rightarrow \alpha, n \rightarrow \infty$, as LLN. As we have proved, N_t is increasing almost sure, then $\frac{N_{[t]}}{[t]} \leq \frac{N_t}{t} \leq \frac{N_{[t]}}{[t]} \frac{[t]}{t}$. So $\frac{N_t}{t} \rightarrow \alpha$, a.s.. \square

PROBLEM III Assume $(N_t : t \geq 0)$ is a Poisson process with parameter $\alpha > 0$. Prove that $\frac{N_t - \alpha t}{\sqrt{\alpha t}} \xrightarrow{d} N(0, 1)$.

SOLUTION. Let $X_0 = N_0, X_k = N_k - N_{k-1}, k = 1, \dots, n, \dots$, then $N_0 = 0$, a.s., $X_k \sim P(\alpha), k \in \mathbb{N}$. So $\mathbb{E}(X_k) = \alpha, \mathbb{V}(X_k) = \alpha$, then by CLT, $\frac{N_n - \alpha n}{\sqrt{\alpha n}} = \frac{\sum_{k=0}^n X_k - \alpha n}{\sqrt{n\alpha}} \xrightarrow{d} N(0, 1)$. Noting $\frac{N_t - \alpha t}{\sqrt{\alpha t}} = \frac{N_{[t]} - \alpha[t]}{\sqrt{\alpha[t]}} + \frac{N_t - N_{[t]} - \alpha(t - [t])}{\sqrt{\alpha t}}$. So $t \rightarrow \infty, [t] \rightarrow \infty$, and $[t] \sim t$. Since $N_t - N_{[t]} \stackrel{d}{=} N_{t - [t]}$, and $t - [t] \leq 1$, we easily get $\mathbb{P}(N_t - N_{[t]} = n) = \frac{((t - [t])\alpha)^n}{n!} e^{-(t - [t])\alpha} \rightarrow 0, t \rightarrow \infty$. Then $\frac{N_t - N_{[t]}}{\sqrt{\alpha t}} \xrightarrow{d} 0$. Easily $\frac{\alpha(t - [t])}{\alpha t} \rightarrow 0$, so finally we get that $\frac{N_t - \alpha t}{\sqrt{\alpha t}} \xrightarrow{d} N(0, 1)$ \square

PROBLEM IV Assume $(X_t : t \geq 0), (Y_t : t \geq 0)$ are two independent Poisson processes with parameter α, β respectively. Prove that $(X_t + Y_t : t \geq 0)$ is Poisson process with parameter $\alpha + \beta$.

SOLUTION. Let $Z_t := X_t + Y_t, t \geq 0$. First we prove $Z_{t+s} - Z_s \sim \text{Poisson}((\alpha + \beta)t)$. Since $X_{t+s} - X_s \sim \text{Poisson}(\alpha t), Y_{t+s} - Y_s \sim \text{Poisson}(\beta t)$, and $X_{t+s} - X_s \perp Y_{t+s} - Y_s$, by the additional property of Poisson, we can get $Z_{t+s} - Z_s = X_{t+s} - X_s + Y_{t+s} - Y_s \sim \text{Poisson}((\alpha + \beta)t)$.

Second we prove $\forall 0 = t_0 < t_1 < \dots < t_n, Z_{t_{k+1}} - Z_{t_k}, k = 1, \dots, n-1, Z_0$ are independent. Easily $Z_{t_{k+1}} - Z_{t_k} = X_{t_{k+1}} - X_{t_k} + Y_{t_{k+1}} - Y_{t_k}$ and $X_{t_{k+1}} - X_{t_k}, X_0, Y_{t_{k+1}} - Y_{t_k}, Y_1$ are independent. Then $X_{t_{k+1}} - X_{t_k} + Y_{t_{k+1}} - Y_{t_k} \in \sigma(\{X_{t_{k+1}} - X_{t_k} : k = 1, \dots, n\} \cup \{Y_{t_{k+1}} - Y_{t_k} : k = 1, \dots, n\}), Z_0$ are independent.

Finally, we prove that $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} Z_s = Z_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} Z_s \in \mathbb{R}) = 1$. Since $Z_t = X_t + Y_t$, and $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} Y_s = Y_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} Y_s \in \mathbb{R}) = 1$,

$\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} X_s = X_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} X_s \in \mathbb{R}) = 1$, then we can easily get $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} Z_s = Z_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} Z_s \in \mathbb{R}) = 1$.

All in all, $(X_t + Y_t : t \geq 0)$ is a Poisson process with parameter $\alpha + \beta$. \square

PROBLEM V Assume $(\xi_n : n \in \mathbb{N}^+)$ is a sequence of i.i.d. random variable ranging in \mathbb{Z}^d . Let $X_n = X_0 + \sum_{k=1}^n \xi_k$, and $X_0 \perp (\xi_n : n \in \mathbb{N}^+)$ ranging in \mathbb{Z}^d , too. Assume $(N_t : t \geq 0)$ is a Poisson process with parameter $\alpha > 0$. Discuss $\frac{X_{N_t}}{t}$ when $t \rightarrow \infty$.

SOLUTION. First we prove that $\lim_{t \rightarrow \infty} N_t = \infty, a.s.$. Since $\mathbb{P}(\sup_t N_t \geq n) \geq \mathbb{P}(N_t \geq n), \forall t, \forall n \in \mathbb{N}$. and $\lim_{t \rightarrow \infty} \mathbb{P}(N_t \geq n) = 1 - \lim_{t \rightarrow \infty} \sum_{i=1}^{n-1} \frac{(\alpha t)^i}{i!} e^{-\alpha t} = 1$, so $\mathbb{P}(\sup_t N_t \geq n) = 1, \forall n \in \mathbb{N}$, then $\mathbb{P}(\sup_t N_t = \infty) = 1$. Since Problem II we know N_t is increasing almost sure, then we can get $\mathbb{P}(\lim_{t \rightarrow \infty} N_t = \infty) = 1$.

Since $\frac{X_{N_t}}{t} = \frac{X_{N_t}}{N_t} \frac{N_t}{t}$ and we have proved that $\frac{N_t}{t} \rightarrow \alpha, a.s.$ in Problem II, so we only need to find $\frac{X_{N_t}}{N_t}$. Since $N_t \rightarrow \infty, a.s.$, we only need to find $\frac{X_n}{n}$ when $n \rightarrow \infty$.

If $\mathbb{E}(\xi_1)$ exists, then by LLN $\frac{X_n}{n} \rightarrow \mathbb{E}(\xi_1), a.s.$. Then we easily get $\frac{X_{N_t}}{t} \rightarrow \alpha \mathbb{E}(\xi_1), a.s.$ \square