

MarkovProcess 2

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PROBLEM I Prove that if $(X_n : n \geq 0)$ is a simple random walk, then so is $(-X_n : n \geq 0)$.

SOLUTION. Since $(X_n : n \geq 0)$ is a simple random walk, then $\exists(\xi_i : i \geq 0)$ are i.i.d. r.v. X_0 is a r.v. which is independent with ξ_1 such that $X_n = X_0 + \sum_{i=0}^n \xi_i$, $\mathbb{P}(|\xi_1| = 1) = 1$. So let $Y_n = -X_n$, $Y_0 = -X_0$ is r.v., $\varepsilon_i = -\xi_i, i \geq 0$, then $(\varepsilon_i, i \geq 0)$ are i.i.d. and is independent with Y_0 , and $\mathbb{P}(|\varepsilon_1| = 1) = 1$. So $(Y_n : n \geq 0)$ is a simple random walk. \square

PROBLEM II Let $(X_n : n \geq 0)$ be a d -dimensional random walk, and $\mathbb{P}(|\xi_1| \geq 1) > 0$, prove that $\mathbb{P}(\sup_n |X_n| = \infty) = 1$.

SOLUTION. Since $\mathbb{P}(|\xi_1| \geq 1) > 0$, then $\exists t \in \mathbb{R}^d$, such that $\mathbb{P}(\xi_1 = t) > 0$. Besides,

$$\begin{aligned} & \mathbb{P}(\sup_n |X_n| = \infty) \\ &= \mathbb{P}\left(\bigcap_{k \in \mathbb{N}} \left\{ \sup_n |X_n| \geq k \right\}\right) \\ &= \lim_{k \rightarrow \infty} \mathbb{P}(\sup_n |X_n| \geq k) \\ &= \inf_{k \in \mathbb{N}} \mathbb{P}(\sup_n |X_n| \geq k) \end{aligned} \tag{1}$$

Then, to prove $\mathbb{P}(\sup_n |X_n| = \infty) = 1$ is equal to prove $\forall k \in \mathbb{N}, \mathbb{P}(\sup_n |X_n| \geq k) = 1$. Let $v > 4 \frac{k}{|t|}$ and let $A_u = \{\omega \in \Omega : \xi_{uv+1} = t, \dots, \xi_{uv+v} = t\}$, so $\forall \omega \in A_u, |X_{uv+v} - X_{uv}| = |\sum_{m=uv+1}^{uv+v} \xi_m| = |vt| = v|t| \geq 4k$. Then $2 \max\{|X_{uv+v}|, |X_{uv}|\} \geq |X_{uv+v}| + |X_{uv}| \geq |\sum_{m=uv+1}^{uv+v} \xi_m| \geq 4k$, so $\max\{|X_{uv+v}|, |X_{uv}|\} \geq 2k > k$. Thus $\sup_n |X_n| \geq k$. Besides, since $\xi_i, i \in \mathbb{N}^+$ is i.i.d., then $\mathbb{P}(A_u) = \mathbb{P}(\xi_1 = 1)^v$. And it is obvious that $A_u, u \in \mathbb{N}^+$ is independent, $\sum_{i=0}^{\infty} \mathbb{P}(A_i) = \infty$, by BC theorem, we can get that $\mathbb{P}(\bigcap_{i=1}^{\infty} \bigcup_{j=i}^{\infty} A_j) = 1$. Since $\bigcap_{i=1}^{\infty} \bigcup_{j=i}^{\infty} A_j \subset \bigcup_{i=1}^{\infty} A_i \subset \{\sup_n |X_n| = \infty\}$, then $\mathbb{P}(\{\sup_n |X_n| = \infty\}) = 1$. \square

PROBLEM III Let $(X_n : n \geq 0)$ be a symmetry simple random walk with $X_0 = 0$, for $d = 2$, prove

that

$$\mathbb{P}(S_{2n} = 0) = \frac{1}{4^{2n}} \left(\frac{(2n)!}{(n!)^2} \right)^2$$

For $d = 3$, prove that

$$\mathbb{P}(S_{2n} = 0) = \frac{1}{2^{2n}} \frac{(2n)!}{(n!)^2} \sum_{i+j+k=n} \left(\frac{1}{3^n} \frac{n!}{i!j!k!} \right)^2$$

SOLUTION. 1. $d = 2$,

$$\begin{aligned} & \mathbb{P}(S_{2n} = 0) \\ &= \left(\frac{1}{4^{2n}} \right) \left(\sum_{k=0}^n \binom{2n}{k} \binom{2n-k}{k} \binom{2n-2k}{n-k} \right) \\ &= \frac{1}{4^{2n}} \sum_{k=0}^n \frac{(2n)!}{(k!)^2 ((n-k)!)^2} \\ &= \frac{1}{4^{2n}} \frac{(2n)!}{(n!)^2} \sum_{k=0}^n \frac{(n!)^2}{(k!)^2 ((n-k)!)^2} \\ &= \frac{1}{4^{2n}} \frac{(2n)!}{(n!)^2} \sum_{k=0}^n \binom{n}{k} \binom{n}{n-k} \\ &= \frac{1}{4^{2n}} \frac{(2n)!}{(n!)^2} \binom{2n}{n} \\ &= \frac{1}{4^{2n}} \left(\frac{(2n)!}{(n!)^2} \right)^2 \end{aligned} \tag{2}$$

2. $d = 3$,

$$\begin{aligned} & \mathbb{P}(S_{2n} = 0) \\ &= \frac{1}{6^{2n}} \left(\sum_{k+j=0}^n \binom{2n}{k} \binom{2n-k}{k} \binom{2n-2k}{j} \binom{2n-2k-j}{j} \binom{2n-2k-2j}{n-k-j} \right) \\ &= \frac{1}{6^{2n}} \left(\sum_{j+k=0}^n \frac{(2n)!}{(k!)^2 (j!)^2 ((n-k-j)!)^2} \right) \\ &= \frac{1}{6^{2n}} \frac{(2n)!}{(n!)^2} \sum_{j+k=0}^n \frac{(n!)^2}{(k!)^2 (j!)^2 ((n-k-j)!)^2} \\ &= \frac{1}{2^{2n}} \frac{(2n)!}{(n!)^2} \sum_{i+j+k=n} \left(\frac{1}{3^n} \frac{n!}{i!j!k!} \right)^2 \end{aligned} \tag{3}$$

□

PROBLEM IV Assume $(S_n : n \geq 0)$ is a symmetry simple random walk with $S_0 = i \in \mathbb{Z}$. Prove that $\forall a \in \mathbb{Z}$, let $\tau_a := \min\{n \in \mathbb{N} : S_n = a\}$, then $\mathbb{P}(\tau_a < \infty) = 1$.

SOLUTION. By the theorem 1.2.2 of textbook, it is obvious that $P(\tau_a < \infty) = \lim_{b \rightarrow \infty} P_i(\tau_a < \tau_b) = \frac{b-i}{b-a} = 1$. □