

PROBLEM I Assume $(\Omega, \mathcal{F}, \mathbb{P})$ is a probability space, and $C \in \mathcal{F}$ satisfy $\mathbb{P}(C) > 0$. Let $\mathbb{P}_C : \mathcal{F} \rightarrow \mathbb{R}$, $\mathbb{P}_C(X) = \frac{\mathbb{P}(C \cap X)}{\mathbb{P}(C)}$. Assume $A, B \in \mathcal{F}$, and $\mathbb{P}(B \cap C) > 0$, prove that $\mathbb{P}_C(A | B) = \mathbb{P}(A | B \cap C)$.

SOLUTION. Since $\mathbb{P}_C(B) = \frac{\mathbb{P}(B \cap C)}{\mathbb{P}(C)} > 0$, then $\mathbb{P}_C(A | B)$ is well-defined. So

$$\mathbb{P}_C(A | B) = \frac{\mathbb{P}_C(A \cap B)}{\mathbb{P}_C(B)} = \frac{\frac{\mathbb{P}(A \cap B \cap C)}{\mathbb{P}(C)}}{\frac{\mathbb{P}(B \cap C)}{\mathbb{P}(C)}} = \frac{\mathbb{P}(A \cap B \cap C)}{\mathbb{P}(B \cap C)} = \mathbb{P}(A | B \cap C)$$

□

PROBLEM II Assume that $(X_n : n \geq 0)$ is 1-dimensional simple symmetry random walk, prove that $(|X_n| : n \geq 0)$ is a Markov chain ranges in \mathbb{N} .

SOLUTION. By the definition of $(X_n : n \geq 0)$, we know that $(X_n : n \geq 0)$ is a Markov chain in \mathbb{Z} . Let $\mathcal{F}_n := \sigma(X_1, \dots, X_n)$, $\mathcal{G}_n := \sigma(|X_1|, \dots, |X_n|)$, then $\mathcal{G}_n \subset \mathcal{F}_n$. $\forall i \in \mathbb{N}$, $\mathbb{P}(|X_{n+1}| = i | \mathcal{F}_n) = \mathbb{P}(X_{n+1} = i | \mathcal{F}_n) + \mathbb{P}(X_{n+1} = -i | \mathcal{F}_n) = \mathbb{P}(X_{n+1} = i | X_n) + \mathbb{P}(X_{n+1} = -i | X_n) = \mathbb{P}(|X_{n+1}| = i | X_n) = \frac{1}{2} \mathbb{1}(|X_n - i| = 1)$. Noting $|X_n - i| = 1 \iff ||X_n| - i| = 1$, so $\mathbb{P}(|X_{n+1}| = i | \mathcal{F}_n) = \frac{1}{2} \mathbb{1}(|X_n| - i = 1)$ is measurable about $\sigma(|X_n|)$. Since $\sigma(|X_n|) \subset \mathcal{G}_n \subset \mathcal{F}_n$, so we finally get that $\mathbb{P}(|X_{n+1}| = i | \mathcal{G}_n) = \mathbb{P}(|X_{n+1}| = i | |X_n|)$. So $(|X_n| : n \geq 0)$ is a Markov chain on \mathbb{N} . □

PROBLEM III Assume $(X_n : n \geq 0)$ is a Markov chain ranges in E . Assume $\phi : E \rightarrow F$ is injection. Prove that $(\phi(X_n) : n \geq 0)$ is a Markov chain ranges in $\phi(E)$.

SOLUTION. Without loss of generality assume $F = \phi(E)$, then ϕ is bijection. Now let $\mathcal{F}_n := \sigma(X_1, \dots, X_n)$, (E, \mathcal{E}) , (Ω, \mathcal{C}) are measurable space, where $X_n : \Omega \rightarrow E$. Let (F, \mathcal{F}) , $\mathcal{F} := \phi(\mathcal{E})$. Since ϕ is bijection, then $\forall A \in \mathcal{E}, \exists B \in \mathcal{C}$, s.t. $X_n(B) = A$, then $\exists A' \in \mathcal{F}$, s.t. $A = \phi^{-1}(A')$, then $X_n(B) = \phi(A)$. Thus $\sigma(X_n) \subset \sigma(\phi(X_n))$, then in the same way, we easily get that $\sigma(X_n) = \sigma(\phi(X_n))$, so $\mathcal{F}_n = \sigma(\phi(X_1), \dots, \phi(X_n))$. Then $\mathbb{P}(\phi(X_{n+1}) = i | \mathcal{F}_n) = \mathbb{P}(X_{n+1} = \phi^{-1}(i) | \mathcal{F}_n) = \mathbb{P}(X_{n+1} = \phi^{-1}(i) | X_{n+1}) = \mathbb{P}(\phi(X_{n+1}) = i | \phi(X_n))$. So $(\phi(X_n) : n \geq 0)$ is Markov chain. □

PROBLEM IV Assume $(X_n : n \geq 0), (Y_n : n \geq 0)$ are two independent Markov chains on E, F respectively. Prove that $((X_n, Y_n) : n \geq 0)$ is Markov chain on $E \times F$.

SOLUTION. Let $\mathcal{F}_n = \sigma(X_0, \dots, X_n)$ and $\mathcal{G}_n = \sigma(Y_0, \dots, Y_n)$, Let $\mathcal{H}_n = \sigma((X_0, Y_0), \dots, (X_n, Y_n))$. Obviously, $\mathcal{H}_n = \sigma(\mathcal{F}_n, \mathcal{G}_n)$.

$$\begin{aligned} \mathbb{P}(X_{n+1} = i, Y_{n+1} = j | \mathcal{H}_n) &= \mathbb{E}(\mathbb{P}(X_{n+1} = i, Y_{n+1} = j | \mathcal{H}_n, X_{n+1}) | \mathcal{H}_n) \\ &= \mathbb{E}(\mathbb{P}(X_{n+1} = i) \mathbb{P}(Y_{n+1} = j | \mathcal{H}_n, X_{n+1}) | \mathcal{H}_n) \\ &= \mathbb{E}(\mathbb{1}_i(X_{n+1}) \mathbb{P}(Y_{n+1} | Y_n) | \mathcal{H}_n) \\ &= \mathbb{P}(Y_{n+1} | Y_n) \mathbb{P}(X_{n+1} = i | \mathcal{H}_n) \\ &= \mathbb{P}(Y_{n+1} | Y_n) \mathbb{P}(X_{n+1} = i | X_n) \end{aligned}$$

So $\mathbb{P}(X_{n+1} = i, Y_{n+1} = j \mid \mathcal{H}_n) \in \sigma(X_n, Y_n) \subset \mathcal{H}_n$. And $\mathbb{P}(X_{n+1} = i, Y_{n+1} = j \mid X_n, Y_n) = \mathbb{E}(\mathbb{P}(X_{n+1} = i, Y_{n+1} = j \mid X_n, Y_n, \mathcal{H}_n) \mid \mathcal{H}_n) = \mathbb{P}(X_{n+1} = i, Y_{n+1} = j \mid \mathcal{H}_n)$. Therefore, $\mathbb{P}(X_{n+1} = i, Y_{n+1} = j \mid X_n, Y_n) = \mathbb{P}(X_{n+1} \mid X_n)\mathbb{P}(Y_{n+1} \mid Y_n)$. So $((X_n, Y_n) : n \geq 0)$ is Markov chain. \square

PROBLEM V Assume $(X_n : n \geq 0), (Y_n : n \geq 0)$ are two independent Markov chains on E, F respectively. Let $\mathcal{H}_n := \sigma((X_0, Y_0), \dots, (X_n, Y_n))$. Prove that $(X_n : n \geq 0)$ is Markov chain over $(\mathcal{H}_n : n \geq 0)$.

SOLUTION. Take $\mathcal{F}_n, \mathcal{G}_n$ as above in Problem IV. Obviously $X_n \in \mathcal{F}_n \subset \mathcal{H}_n$. Easily $\mathbb{P}(X_{n+1} = i \mid \mathcal{H}_n) = \mathbb{P}(X_{n+1} = i \mid \mathcal{F}_n, \mathcal{G}_n) = \mathbb{P}(X_{n+1} = i \mid \mathcal{F}_n) = \mathbb{P}(X_{n+1} \mid X_n)$. So $(X_n : n \geq 0)$ is Markov chain over $(\mathcal{H}_n : n \geq 0)$. \square

PROBLEM VI Let μ_0 be a probability distribution on \mathbb{N} . For $n \geq 1$, let

$$\mu_n(0) = \mu_{n-1}^{*2}(0) + \mu_{n-1}^{*2}(1), \mu_n(j) = \mu_{n-1}^{*2}(j+1), \forall j \geq 1$$

Where $\mu^{*2} = \mu * \mu$. Let F_n be distribution function of μ_n . Let $F_{n-1}^{-1}(y) := \inf\{x \geq 0 : y \leq F_{n-1}(x)\}$ for $y \in [0, 1]$. Assume $X_0 \sim \mu_0$, and $(U_n : n \geq 0)$ are i.i.d r.v. with distribution $U(0, 1)$. Let $X_{n+1} := \max\{0, X_n + F_n^{-1}(U_n) - 1\}$. Then $(X_n : n \geq 0)$ is Markov chain.

SOLUTION. Let $\mathcal{F}_n := \sigma(X_0, \dots, X_n)$. For $i > 0$, we have

$$\begin{aligned} \mathbb{P}(X_{n+1} = i \mid \mathcal{F}_n) &= \mathbb{P}(X_n + F_n^{-1}(U_n) - 1 = i \mid \mathcal{F}_n) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{P}(X_n = k, F_n^{-1}(U_n) = i + 1 - k \mid \mathcal{F}_n) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{1}_k(X_n) \mathbb{P}(F_n^{-1}(U_n) = i + 1 - k) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{1}_k(X_n) \mathbb{P}(F_n^{-1}(U_n) = i + 1 - k \mid X_n) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{P}(X_n = k, F_n^{-1}(U_n) = i + 1 - k \mid X_n) \\ &= \mathbb{P}(X_n + F_n^{-1}(U_n) - 1 = i \mid X_n) = \mathbb{P}(X_{n+1} = i \mid X_n) \end{aligned}$$

For $i = 0$, we have

$$\begin{aligned} \mathbb{P}(X_{n+1} = 0 \mid \mathcal{F}_n) &= \mathbb{P}(X_n + F_n^{-1}(U_n) - 1 \leq 0 \mid \mathcal{F}_n) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{P}(X_n = k, F_n^{-1}(U_n) \leq 1 - k \mid \mathcal{F}_n) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{1}_k(X_n) \mathbb{P}(F_n^{-1}(U_n) \leq 1 - k) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{1}_k(X_n) \mathbb{P}(F_n^{-1}(U_n) \leq 1 - k \mid X_n) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{P}(X_n = k, F_n^{-1}(U_n) \leq 1 - k \mid X_n) \\ &= \mathbb{P}(X_n + F_n^{-1}(U_n) - 1 \leq 0 \mid X_n) = \mathbb{P}(X_{n+1} = 0 \mid X_n) \end{aligned}$$

So $(X_n : n \geq 0)$ is Markov chain on \mathcal{F} \square