

**PROBLEM I** Prove that if  $(X_n : n \geq 0)$  is a simple random walk, then so is  $(-X_n : n \geq 0)$ .

**SOLUTION.** Since  $(X_n : n \geq 0)$  is a simple random walk, then  $\exists(\xi_i : i \geq 0)$  are i.i.d. r.v.  $X_0$  is a r.v. which is independent with  $\xi_1$  such that  $X_n = X_0 + \sum_{i=0}^n \xi_i$ ,  $\mathbb{P}(|\xi_1| = 1) = 1$ . So let  $Y_n = -X_n$ ,  $Y_0 = -X_0$  is r.v.,  $\varepsilon_i = -\xi_i, i \geq 0$ , then  $(\varepsilon_i, i \geq 0)$  are i.i.d. and is independent with  $Y_0$ , and  $\mathbb{P}(|\varepsilon_1| = 1) = 1$ . So  $(Y_n : n \geq 0)$  is a simple random walk.  $\square$

**PROBLEM II** Let  $(X_n : n \geq 0)$  be a  $d$ -dimensional random walk, and  $\mathbb{P}(|\xi_1| \geq 1) > 0$ , prove that  $\mathbb{P}(\sup_n |X_n| = \infty) = 1$ .

**SOLUTION.** Since  $\mathbb{P}(|\xi_1| \geq 1) > 0$ , then  $\exists t \in \mathbb{R}^d$ , such that  $\mathbb{P}(\xi_1 = t) > 0$ . Besides,

$$\begin{aligned} & \mathbb{P}(\sup_n |X_n| = \infty) \\ &= \mathbb{P}\left(\bigcap_{k \in \mathbb{N}} \left\{ \sup_n |X_n| \geq k \right\}\right) \\ &= \lim_{k \rightarrow \infty} \mathbb{P}(\sup_n |X_n| \geq k) \\ &= \inf_{k \in \mathbb{N}} \mathbb{P}(\sup_n |X_n| \geq k) \end{aligned} \tag{1}$$

Then, to prove  $\mathbb{P}(\sup_n |X_n| = \infty) = 1$  is equal to prove  $\forall k \in \mathbb{N}, \mathbb{P}(\sup_n |X_n| \geq k) = 1$ . Let  $v > 4 \frac{k}{|t|}$  and let  $A_u = \{\omega \in \Omega : \xi_{uv+1} = t, \dots, \xi_{uv+v} = t\}$ , so  $\forall \omega \in A_u, |X_{uv+v} - X_{uv}| = |\sum_{m=uv+1}^{uv+v} \xi_m| = |vt| = v|t| \geq 4k$ . Then  $2 \max\{|X_{uv+v}|, |X_{uv}|\} \geq |X_{uv+v}| + |X_{uv}| \geq |\sum_{m=uv+1}^{uv+v} \xi_m| \geq 4k$ , so  $\max\{|X_{uv+v}|, |X_{uv}|\} \geq 2k > k$ . Thus  $\sup_n |X_n| \geq k$ . Besides, since  $\xi_i, i \in \mathbb{N}^+$  is i.i.d., then  $\mathbb{P}(A_u) = \mathbb{P}(\xi_1 = 1)^v$ . And it is obvious that  $A_u, u \in \mathbb{N}^+$  is independent,  $\sum_{i=0}^{\infty} \mathbb{P}(A_i) = \infty$ , by BC theorem, we can get that  $\mathbb{P}(\bigcap_{i=1}^{\infty} \bigcup_{j=i}^{\infty} A_j) = 1$ . Since  $\bigcap_{i=1}^{\infty} \bigcup_{j=i}^{\infty} A_j \subset \bigcup_{i=1}^{\infty} A_i \subset \{\sup_n |X_n| = \infty\}$ , then  $\mathbb{P}(\{\sup_n |X_n| = \infty\}) = 1$ .  $\square$

**PROBLEM III** Let  $(X_n : n \geq 0)$  be a symmtry simple random walk with  $X_0 = 0$ , for  $d = 2$ , prove that

$$\mathbb{P}(S_{2n} = 0) = \frac{1}{4^{2n}} \left( \frac{(2n)!}{(n!)^2} \right)^2$$

For  $d = 3$ , prove that

$$\mathbb{P}(S_{2n} = 0) = \frac{1}{2^{2n}} \frac{(2n)!}{(n!)^2} \sum_{i+j+k=n} \left( \frac{1}{3^n} \frac{n!}{i!j!k!} \right)^2$$

*SOLUTION*. 1.  $d = 2$ ,

$$\begin{aligned}
 & \mathbb{P}(S_{2n} = 0) \\
 &= \left(\frac{1}{4^{2n}}\right) \left(\sum_{k=0}^n \binom{2n}{k} \binom{2n-k}{k} \binom{2n-2k}{n-k}\right) \\
 &= \frac{1}{4^{2n}} \sum_{k=0}^n \frac{(2n)!}{(k!)^2((n-k)!)^2} \\
 &= \frac{1}{4^{2n}} \frac{(2n)!}{(n!)^2} \sum_{k=0}^n \frac{(n!)^2}{(k!)^2((n-k)!)^2} \\
 &= \frac{1}{4^{2n}} \frac{(2n)!}{(n!)^2} \sum_{k=0}^n \binom{n}{k} \binom{n}{n-k} \\
 &= \frac{1}{4^{2n}} \frac{(2n)!}{(n!)^2} \binom{2n}{n} \\
 &= \frac{1}{4^{2n}} \left(\frac{(2n)!}{(n!)^2}\right)^2
 \end{aligned} \tag{2}$$

2.  $d = 3$ ,

$$\begin{aligned}
 & \mathbb{P}(S_{2n} = 0) \\
 &= \frac{1}{6^{2n}} \left(\sum_{k+j=0}^n \binom{2n}{k} \binom{2n-k}{k} \binom{2n-2k}{j} \binom{2n-2k-j}{j} \binom{2n-2k-2j}{n-k-j}\right) \\
 &= \frac{1}{6^{2n}} \left(\sum_{j+k=0}^n \frac{(2n)!}{(k!)^2(j!)^2((n-k-j)!)^2}\right) \\
 &= \frac{1}{6^{2n}} \frac{(2n)!}{(n!)^2} \sum_{j+k=0}^n \frac{(n!)^2}{(k!)^2(j!)^2((n-k-j)!)^2} \\
 &= \frac{1}{2^{2n}} \frac{(2n)!}{(n!)^2} \sum_{i+j+k=n} \left(\frac{1}{3^n} \frac{n!}{i!j!k!}\right)^2
 \end{aligned} \tag{3}$$

□

**PROBLEM IV** Assume  $(S_n : n \geq 0)$  is a symmetry simple random walk with  $S_0 = i \in \mathbb{Z}$ . Prove that  $\forall a \in \mathbb{Z}$ , let  $\tau_a := \min\{n \in \mathbb{N} : S_n = a\}$ , then  $\mathbb{P}(\tau_a < \infty) = 1$ .

*SOLUTION*. By the theorem 1.2.2 of textbook, it is obvious that  $P(\tau_a < \infty) = \lim_{b \rightarrow \infty} P_i(\tau_a < \tau_b) = \frac{b-i}{b-a} = 1$ . □