

## 5 Discrete Memoryless Channels

### 5.1 Homework 7

- Chapter 7, Problems 3,7,8,9,10,11

- Supplementary Problems:

1. Let  $X$  and  $Y$  be the input and output of a BSC. Show that if  $\epsilon = 0.5$ , then  $X$  and  $Y$  are independent.
2. Show in Example 7.8 that  $H(Y | E) = (1 - \gamma)h_b(a)$

**3. Memory increases capacity.** Consider a BSC with crossover probability  $0 < \epsilon < 1$  represented by  $X_i = Y_i + Z_i \bmod 2$ , where  $X_i, Y_i$ , and  $Z_i$  are respectively, the input, the output, and the noise variable at time  $i$ . Then

$$\Pr\{Z_i = 0\} = 1 - \epsilon \quad \text{and} \quad \Pr\{Z_i = 1\} = \epsilon$$

for all  $i$ . We assume that  $\{X_i\}$  and  $\{Z_i\}$  are independent, but we make no assumption that  $Z_i$  are i.i.d. so that the channel may have memory.

a) Prove that

$$I(\mathbf{X}; \mathbf{Y}) \leq n - h_b(\epsilon)$$

b) Show that the upper bound in (a) can be achieved by letting  $X_i$  be i.i.d. bits taking the values 0 and 1 with equal probability and  $Z_1 = Z_2 = \dots = Z_n$

c) Show that with the assumptions in (b),  $I(\mathbf{X}; \mathbf{Y}) > nC$ , where  $C = 1 - h_b(\epsilon)$  is the capacity of the BSC if it is memoryless.

Typo:  $Y_i = X_i + Z_i$ ,  $X_i + Y_i = (X_i + X_i) + Z_i = 0 + Z_i = Z_i$

**Answer** (a)

$$\begin{aligned} I(\mathbf{X}; \mathbf{Y}) &= H(\mathbf{Y}) - H(\mathbf{Y} | \mathbf{X}) \\ &\leq \sum_{i=1}^n H(Y_i) - \sum_{i=1}^n H(Y_i | \mathbf{Y}^{i-1}, \mathbf{X}) \\ &\leq n \cdot 1 - H(Y_1 | \mathbf{X}) \\ &= n - h_b(\epsilon) \end{aligned}$$

(b) In order to achieve this upper bound, we have to 1) make  $H(\mathbf{Y}) = \sum_{i=1}^n H(Y_i)$  and  $H(Y_i) = 1$ , i.e., the output distribution of the BSC is uniform. This can be done by letting  $p(X_i)$  be the uniform distribution on  $\{0, 1\}$ . 2)  $H(Y_i | \mathbf{Y}^{i-1}, \mathbf{X}) = 0, i \geq 2$ , that is the random variable  $Z_i, i \geq 2$  are fixed (same as  $Z_1$ ).

(c) With assumption in b holds

$$I(\mathbf{X}; \mathbf{Y}) = n - h_b(\epsilon) \geq n - nh_b(\epsilon) = n(1 - h_b(\epsilon)) = nC$$

7. Let

$$P(\epsilon) = \begin{bmatrix} 1 - \epsilon & \epsilon \\ \epsilon & 1 - \epsilon \end{bmatrix}$$

be the transition matrix for a BSC with crossover probability  $\epsilon$ . Define  $a * b = (1 - a)b + a(1 - b)$  for  $0 \leq a, b \leq 1$

a) Prove that a DMC with transition matrix  $P(\epsilon_1)P(\epsilon_2)$  is equivalent to a BSC with crossover probability  $\epsilon_1 * \epsilon_2$ . Such a channel is the cascade of two BSCs with crossover probabilities  $\epsilon_1$  and  $\epsilon_2$  respectively.

b) Repeat (a) for a DMC with transition matrix  $P(\epsilon_2)P(\epsilon_1)$ .

c) Prove that

$$1 - h_b(\epsilon_1 * \epsilon_2) \leq \min(1 - h_b(\epsilon_1), 1 - h_b(\epsilon_2))$$

This means that the capacity of the cascade of two BSCs is upper bounded by the capacity of either of the two BSCs.

d) Prove that a DMC with transition matrix  $P(\epsilon)^n$  is equivalent to a BSC with crossover probabilities  $\frac{1}{2}(1 - (1 - 2\epsilon)^n)$

**Answer** (a) The composite transition probability is

$$p(Y|X) = \sum_{Z=0,1} p(Y|Z)p(Z|X)$$

Given  $p(Y|Z) = P(\epsilon_2)$ ,  $p(Z|X) = P(\epsilon_1)$  the above could be write down in matrix format, that is

$$\begin{aligned} p(\epsilon) &= P(\epsilon_2)P(\epsilon_1) = \begin{bmatrix} 1-\epsilon_2 & \epsilon_2 \\ \epsilon_2 & 1-\epsilon_2 \end{bmatrix} \begin{bmatrix} 1-\epsilon_1 & \epsilon_1 \\ \epsilon_1 & 1-\epsilon_1 \end{bmatrix} \\ &= \begin{bmatrix} (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2 & (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2 \\ (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2 & (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2 \end{bmatrix} \\ &= \begin{bmatrix} (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2 & (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2 \\ (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2 & (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2 \end{bmatrix} \end{aligned}$$

Note that, denoting  $\epsilon^* = (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2$ , and

$$1 - \epsilon^* = 1 - ((1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2) = (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2$$

We obtain

$$p(\epsilon^*) = \begin{bmatrix} (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2 & (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2 \\ (1-\epsilon_2)\epsilon_1 + (1-\epsilon_1)\epsilon_2 & (1-\epsilon_2)(1-\epsilon_1) + \epsilon_1 * \epsilon_2 \end{bmatrix} = \begin{bmatrix} 1-\epsilon^* & \epsilon^* \\ \epsilon^* & 1-\epsilon^* \end{bmatrix}$$

therefore equivalent to a BSC with crossover probability  $\epsilon^* = \epsilon_1 * \epsilon_2$

(b) As seen from the above matrix multiplication, the formula is symmetric about  $\epsilon_1$  and  $\epsilon_2$ . Therefore the conclusion holds for  $p(\epsilon_1)p(\epsilon_2)$

(c)

$$C_1 = 1 - H_b(\epsilon_1) \quad C_2 = 1 - H_b(\epsilon_2) \quad C_3 = 1 - H_b(\epsilon_1 * \epsilon_2)$$

From Markov Chain  $X - Z - Y$  we can see that

$$C_3 = I(X; Y) \leq I(X; Z, Y) = I(X; Z) = C_1$$

$$C_3 = I(X; Y) \leq I(X, Z; Y) = I(Z; Y) = C_2$$

Thus

$$C_3 = \min \leq \{C_1, C_2\}$$

which is exactly same as

$$1 - h_b(\epsilon_1 * \epsilon_2) \leq \min(1 - h_b(\epsilon_1), 1 - h_b(\epsilon_2))$$

(d) For  $n = 1$ , the  $\frac{1}{2}(1 - (1 - 2\epsilon)^n) = \epsilon$  trivially holds. Suppose it holds for  $n = k - 1$ ,  $k \geq 2$ , then for  $n = k$ ,

$$\begin{aligned} \Pr(Y_k = 1 | X = 0) &= \Pr(Y_k = 1 | Y_{k-1} = 0, X = 0) \Pr(Y_{k-1} = 0 | X = 0) + \\ &\quad \Pr(Y_k = 1 | Y_{k-1} = 1, X = 0) \Pr(Y_{k-1} = 1 | X = 0) \\ &= \Pr(Y_k = 1 | Y_{k-1} = 0) \Pr(Y_{k-1} = 0 | X = 0) + \Pr(Y_k = 1 | Y_{k-1} = 1) \Pr(Y_{k-1} = 1 | X = 0) \\ &= \epsilon \left( 1 - \frac{1}{2}(1 - (1 - 2\epsilon)^{k-1}) \right) + (1 - \epsilon) \frac{1}{2}(1 - (1 - 2\epsilon)^{k-1}) \\ &= \frac{1}{2}(1 - (1 - 2\epsilon)^k) \end{aligned}$$

We proved by mathematical induction.

**8. Symmetric channel.** A DMC is symmetric if the rows of the transition matrix  $p(y | x)$  are permutations of each other and so are the columns. Determine the capacity of such a channel. See Section 4.5 in Gallager [129] for a more general discussion.

**Answer** Note that

$$\begin{aligned} H(Y | X) &= - \sum_x \sum_y P(y | x) P(x) \log P(y | x) \\ &= - \sum_x P(x) \left( \sum_y P(y | x) \log P(y | x) \right) \end{aligned}$$

Note that by symmetric assumption,  $H(\Pi) = - \sum_y P(y | x) \log P(y | x)$  is independent of  $x$ . Thus,

$$\begin{aligned} H(Y | X) &= - \sum_x P(x) \left( \sum_y P(y | x) \log P(y | x) \right) \\ &= \sum_x P(x) H(\Pi) = H(\Pi) \end{aligned}$$

Since the capacity is

$$\begin{aligned} C &= \max_x I(X; Y) = \max_x H(Y) - H(Y|X) \\ &= \max_x H(Y) - H(\Pi) \end{aligned}$$

Recall that In Theorem 2.43, we have proved that for any random variable

$$H(X) \leq \log |\mathcal{X}|$$

Therefore

$$C \leq \log |\mathcal{Y}| - H(\Pi)$$

where  $H(\Pi) = - \sum_y P(y | x) \log P(y | x)$  and the equality is attained when  $Y$  is uniform.

9. Let  $C_1$  and  $C_2$  be the capacities of two DMCs with transition matrices  $P_1$  and  $P_2$ , respectively, and let  $C$  be the capacity of the DMC with transition matrix  $P_1 P_2$ . Prove that  $C \leq \min(C_1, C_2)$

**Answer** Think of the constructed Markov Chain  $X - Z - Y$ , where  $Z$  denoted the DMC in the middle phase

$$C_3 = I(X; Y) \leq I(X; Z, Y) = I(X; Z) = C_1$$

$$C_3 = I(X; Y) \leq I(X, Z; Y) = I(Z; Y) = C_2$$

Similar to proof in Problem 7.

$$C_3 = \min \leq \{C_1, C_2\}$$

**10. Two parallel channels.** Let  $C_1$  and  $C_2$  be the capacities of two DMCs  $p_1(y_1 | x_1)$  and  $p_2(y_2 | x_2)$ , respectively. Determine the capacity of the DMC

$$p(y_1, y_2 | x_1, x_2) = p_1(y_1 | x_1) p_2(y_2 | x_2)$$

Hint: Prove that

$$I(X_1, X_2; Y_1, Y_2) \leq I(X_1; Y_1) + I(X_2; Y_2)$$

if  $p(y_1, y_2 | x_1, x_2) = p_1(y_1 | x_1) p_2(y_2 | x_2)$

**Answer**

$$\begin{aligned}
I(X_1, X_2; Y_1, Y_2) &= \sum_{x_1, x_2, y_1, y_2} p(x_1, x_2, y_1, y_2) \log \frac{p(x_1, x_2, y_1, y_2)}{p(x_1, x_2)p(y_1, y_2)} \\
&= \sum_{x_1, x_2, y_1, y_2} p(x_1, x_2, y_1, y_2) \log \frac{(y_1, y_2 | x_1, x_2) p(x_1, x_2)}{p(x_1, x_2)p(y_1, y_2)} \\
&= \sum_{x_1, x_2, y_1, y_2} p(x_1, x_2, y_1, y_2) \log \frac{p_1(y_1 | x_1) p_2(y_2 | x_2)}{p(y_1, y_2)} \\
&= \sum_{x_1, x_2, y_1, y_2} p(x_1, x_2, y_1, y_2) \log p_1(y_1 | x_1) + \sum_{x_1, x_2, y_1, y_2} p(x_1, x_2, y_1, y_2) \log p_2(y_2 | x_2) - \\
&\quad \sum_{x_1, x_2, y_1, y_2} p(x_1, x_2, y_1, y_2) \log p(y_1, y_2) \\
&= \sum_{x_1, y_1} p(x_1, y_1) \log p_1(y_1 | x_1) + \sum_{x_2, y_2} p(x_2, y_2) \log p_2(y_2 | x_2) - \sum_{y_1, y_2} p(y_1, y_2) \log p_1(y_1, y_2) \\
&= I(X_1; Y_1) + I(X_2; Y_2) + H(Y_1, Y_2) - H(Y_1) - H(Y_2) \\
&\leq I(X_1; Y_1) + I(X_2; Y_2)
\end{aligned}$$

Therefore we know that the new  $C \leq C_1 + C_2$

11. In the system below, there are two channels with transition matrices  $p_1(y_1 | x)$  and  $p_2(y_2 | x)$ . These two channels have a common input alphabet  $\mathcal{X}$  and output alphabets  $\mathcal{Y}_1$  and  $\mathcal{Y}_2$ , respectively, where  $\mathcal{Y}_1$  and  $\mathcal{Y}_2$  are disjoint. The position of the switch is determined by a random variable  $Z$  which is independent of  $X$ , where  $\Pr\{Z = 1\} = \lambda$

a) Show that

$$I(X; Y) = \lambda I(X; Y_1) + (1 - \lambda) I(X; Y_2)$$

b) The capacity of the system is given by  $C = \max_{p(x)} I(X; Y)$ . Show that  $C \leq \lambda C_1 + (1 - \lambda) C_2$ , where  $C_i = \max_{p(x)} I(X; Y_i)$  is the capacity of the channel with transition matrix  $p_i(y_i | x)$ ,  $i = 1, 2$

c) If both  $C_1$  and  $C_2$  can be achieved by a common input distribution, show that  $C = \lambda C_1 + (1 - \lambda) C_2$

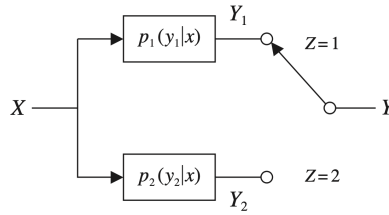


Figure 16: Problem 11.

**Answer** (a)

$$\begin{aligned}
E(Y) &= \lambda E(Y_1) + (1 - \lambda) E(Y_2) \\
I(X; Y) &= E \log \frac{p(X, Y)}{p(X)p(Y)} \\
&= E \log \frac{p(X, (Y_1, Y_2, Z))}{p(X)p(Y_1, Y_2, Z)} \\
&= E \left( \mathbb{1}_{Z=1} \log \frac{p(X, Y_1)}{p(X)p(Y_1)} \right) + E \left( \mathbb{1}_{Z=2} \log \frac{p(X, Y_2)}{p(X)p(Y_2)} \right) \\
&= \lambda I(X; Y_1) + (1 - \lambda) I(X; Y_2)
\end{aligned}$$

(b)

$$\begin{aligned}
\max_{p(x)} I(X; Y) &= \max_{p(x)} \lambda I(X; Y_1) + (1 - \lambda) I(X; Y_2) \\
&\leq \max_{p(x)} \lambda I(X; Y_1) + \max_{p(x)} \lambda I(X; Y_2) \\
&= \lambda C_1 + (1 - \lambda) C_2
\end{aligned}$$

(c) The equality is achieved when  $\arg\max_{p(x)} I(X; Y_1) = \arg\max_{p(x)} I(X; Y_2)$ , otherwise the two maximum values cannot be achieved simultaneously.

Let  $X$  and  $Y$  be the input and output of a BSC. Show that if  $\epsilon = 0.5$ , then  $X$  and  $Y$  are independent.

**Answer**

$$I(X; Y) = H(Y) - H(Y|X) = H(Y) - h_b(0.5) = 1 - 1 = 0$$

Therefore,  $X, Y$  independent

Show in Example 7.8 that  $H(Y | E) = (1 - \gamma)h_b(a)$ , where

$$E = \begin{cases} 0 & \text{if } Y \neq e \\ 1 & \text{if } Y = e \end{cases}$$

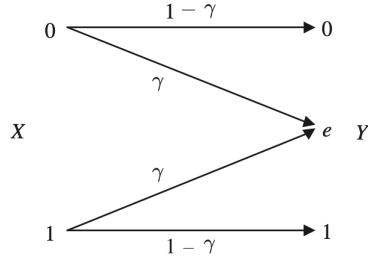


Figure 17: Example 7.8.

$$\begin{aligned}
H(Y | E) &= - \sum_{e,y} p(e, y) \log \frac{p(e, y)}{p(e)} \\
&= -p(e = 1, y = e) \log \frac{p(e = 1, y = e)}{p(e = 1)} - p(e = 0, y = 1) \log \frac{p(e = 0, y = 1)}{p(e = 0)} - p(e = 0, y = 0) \log \frac{p(e = 0, y = 0)}{p(e = 0)} \\
&= -0 - (1 - a)(1 - \gamma) \log(1 - a) - a(1 - \gamma) \log(a) \\
&= (1 - \gamma)h_b(a)
\end{aligned}$$