PROBABILITY & STATISTICS

Lecture 11 – Confidence Intervals

MLE FOR NORMAL DISTRIBUTIONS

estimating $\,\mu$ and σ

- X1, X2, ..., Xn samples from the normal distribution.
- Parameters μ and σ are unknown.

- X1, X2, ..., Xn samples from the normal distribution.
- Parameters μ and σ are unknown.

•
$$\hat{\mu} = ?$$
 $\hat{\sigma} = ?$

- X1, X2, ..., Xn samples from the normal distribution.
- Parameters μ and σ are unknown.

•
$$\hat{\mu} = ?$$
 $\hat{\sigma} = ?$

• Maximum likelihood.

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- $\hat{\mu} = ? \hat{\sigma} = ?$

$$L(\mu, \sigma) =$$

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- $\hat{\mu} = ? \hat{\sigma} = ?$

$$L(\mu, \sigma) = \prod_{i=1}^{n} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{X_i - \mu}{\sigma}\right)^2}$$

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- $\hat{\mu} = ? \hat{\sigma} = ?$

$$L(\mu, \sigma) = \prod_{i=1}^{n} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{X_i - \mu}{\sigma}\right)^2}$$

$$\log L(\mu, \sigma) =$$

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- $\hat{\mu} = ? \hat{\sigma} = ?$

$$L(\mu, \sigma) = \prod_{i=1}^{n} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{X_i - \mu}{\sigma}\right)^2}$$

$$\log L(\mu, \sigma) = -\sum_{i=1}^{n} \log(\sigma\sqrt{2\pi}) - \frac{1}{2}\sum_{i=1}^{n} \left(\frac{X_i - \mu}{\sigma}\right)^2 =$$

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- $\hat{\mu} = ? \hat{\sigma} = ?$

$$L(\mu, \sigma) = \prod_{i=1}^{n} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{X_i - \mu}{\sigma}\right)^2}$$

$$\log L(\mu, \sigma) = -\sum_{i=1}^{n} \log(\sigma\sqrt{2\pi}) - \frac{1}{2} \sum_{i=1}^{n} \left(\frac{X_i - \mu}{\sigma}\right)^2 =$$

$$= -n\log\sigma - n\log\sqrt{2\pi} - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (X_i - \mu)^2$$

$$\log L(\mu, \sigma) = -n \log \sigma - n \log \sqrt{2\pi} - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (X_i - \mu)^2$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) =$$

$$\frac{d}{d\sigma}\log L(\mu,\sigma) =$$

$$\log L(\mu, \sigma) = -n \log \sigma - n \log \sqrt{2\pi} - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (X_i - \mu)^2$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0$$

$$\frac{d}{d\sigma}\log L(\mu,\sigma) =$$

$$\log L(\mu, \sigma) = -n \log \sigma - n \log \sqrt{2\pi} - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (X_i - \mu)^2$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0$$

$$\frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow$$

$$\sum_{i=1}^{n} (X_i - \mu) = \sum_{i=1}^{n} X_i - n\mu = 0$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (X_i - \mu)^2 = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow$$

$$\sum_{i=1}^n (X_i - \mu) = \sum_{i=1}^n X_i - n\mu = 0$$

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i = \overline{X}$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow$$

$$\sum_{i=1}^n (X_i - \mu) = \sum_{i=1}^n X_i - n\mu = 0$$

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i = \overline{X}$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (X_i - \mu)^2 = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow \qquad -n\sigma^2 + \sum_{i=1}^n (X_i - \mu)^2 = 0$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow$$

$$\sum_{i=1}^{n} (X_i - \mu) = \sum_{i=1}^{n} X_i - n\mu = 0$$

$$\widehat{\mu} = \frac{1}{n} \sum_{i=1}^{n} X_i = \overline{X}$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$-n\sigma^2 + \sum_{i=1}^{n} (X_i + \mu)^2 = 0$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow$$

$$\sum_{i=1}^{n} (X_i - \mu) = \sum_{i=1}^{n} X_i - n\mu = 0$$

$$\widehat{\mu} = \frac{1}{n} \sum_{i=1}^{n} X_i = \overline{X}$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$-n\sigma^2 + \sum_{i=1}^{n} (X_i + \mu)^2 = 0$$

$$-n\sigma^2 + \sum_{i=1}^{n} (X_i - \bar{X})^2 = 0$$

$$\frac{d}{d\mu}\log L(\mu,\sigma) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (X_i - \mu) = 0 \qquad \frac{d}{d\sigma}\log L(\mu,\sigma) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$\frac{1}{\sigma^2} \neq 0 \Rightarrow \qquad -n\sigma^2 + \sum_{i=1}^{n} (X_i - \mu)^2 = 0$$

$$\sum_{i=1}^{n} (X_i - \mu) = \sum_{i=1}^{n} X_i - n\mu = 0 \qquad -n\sigma^2 + \sum_{i=1}^{n} (X_i - \bar{X})^2 = 0$$

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} X_i = \bar{X} \qquad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^{n} (X_i - \bar{X})^2 = S^2$$

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- Parameters μ and σ are unknown.

•
$$\hat{\mu} = ?$$
 $\hat{\sigma} = ?$

Maximum likelihood:

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- Parameters μ and σ are unknown.

•
$$\hat{\mu} = ?$$
 $\hat{\sigma} = ?$

- Maximum likelihood:
 - $\hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} X_i = \overline{X}$ MLE for μ is sample mean

- $X_1, X_2, ..., X_n$ samples from the normal distribution.
- Parameters μ and σ are unknown.

•
$$\hat{\mu} = ?$$
 $\hat{\sigma} = ?$

- Maximum likelihood:
 - $\hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} X_i = \overline{X}$ MLE for μ is sample mean
 - $\widehat{\sigma^2} = \frac{1}{n} \sum_{i=1}^n (X_i \overline{X})^2 = S^2$ MLE for σ^2 is sample variance

Bias, variance and consistency

- Parameter estimation:
 - Given samples $X_1, X_2, ... X_n$
 - Obtain estimate $\hat{\theta}$ of an unknown parameter θ .

• Parameter estimation:

Given samples $X_1, X_2, ... X_n$ Obtain estimate $\hat{\theta}$ of an unknown parameter θ .

•
$$\hat{\theta} = T(X_1, ..., X_n)$$
, $T(X)$ — estimator.

- Parameter estimation:
 - Given samples $X_1, X_2, ... X_n$ Obtain estimate $\hat{\theta}$ of an unknown parameter θ .
- $\hat{\theta} = T(X_1, ..., X_n)$, T(X) estimator.
- So far, we only constructed ML estimators.

- Parameter estimation:
 - Given samples $X_1, X_2, \dots X_n$ Obtain estimate $\hat{\theta}$ of an unknown parameter θ .
- $\hat{\theta} = T(X_1, ..., X_n)$, T(X) estimator.
- So far, we only constructed ML estimators. But *anything* can be an estimator:

- Parameter estimation:
 - Given samples $X_1, X_2, ... X_n$
 - Obtain estimate $\hat{\theta}$ of an unknown parameter θ .
- $\hat{\theta} = T(X_1, ..., X_n)$, T(X) estimator.
- So far, we only constructed ML estimators. But *anything* can be an estimator:

$$\hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n X_i, \qquad \hat{\theta}_2 = \frac{X_1 + X_n}{2}, \qquad \hat{\theta}_3 = \max(X_1, \dots, X_n), \qquad \dots$$

- Parameter estimation:
 - Given samples $X_1, X_2, ... X_n$
 - Obtain estimate $\hat{\theta}$ of an unknown parameter θ .
- $\hat{\theta} = T(X_1, ..., X_n)$, T(X) estimator.
- So far, we only constructed ML estimators. But anything can be an estimator:

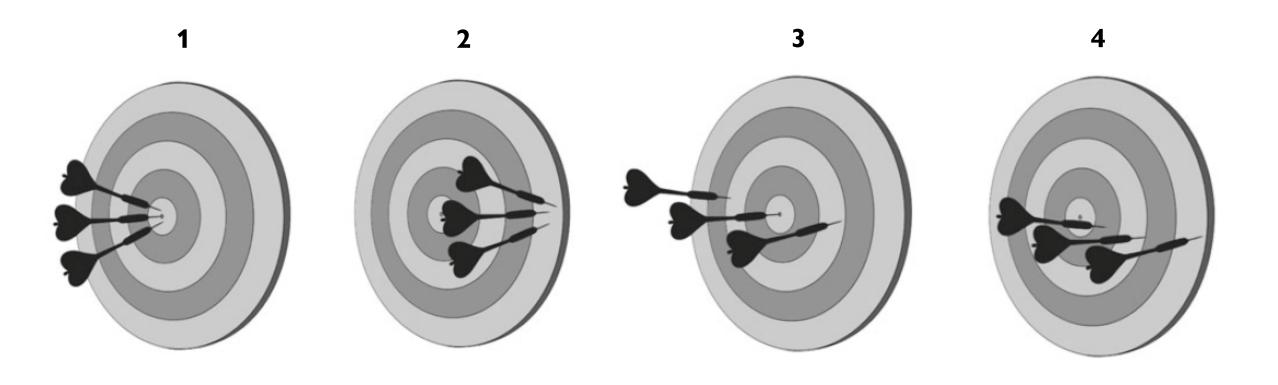
$$\hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n X_i$$
, $\hat{\theta}_2 = \frac{X_1 + X_n}{2}$, $\hat{\theta}_3 = \max(X_1, \dots, X_n)$, ...

• We need to compare different estimators.

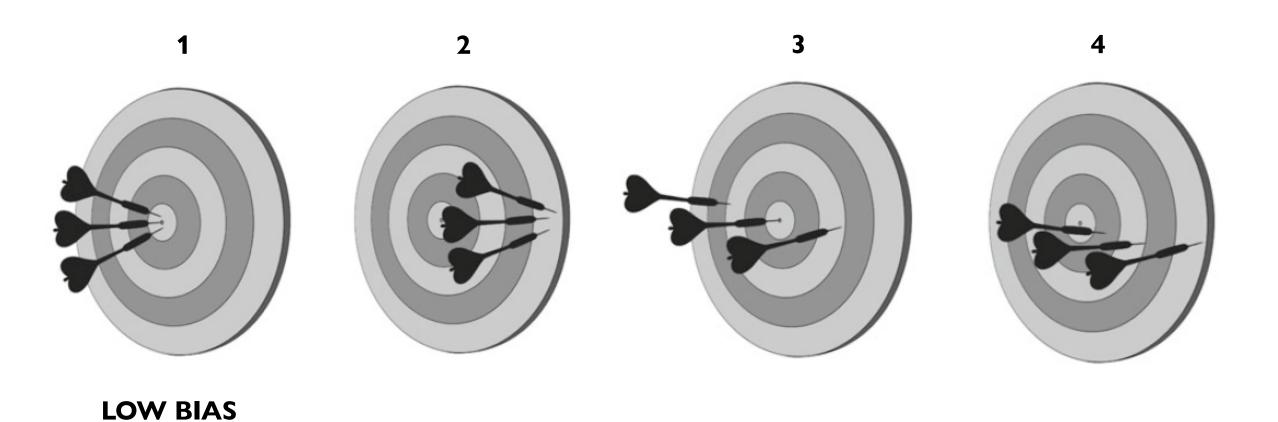
Bias

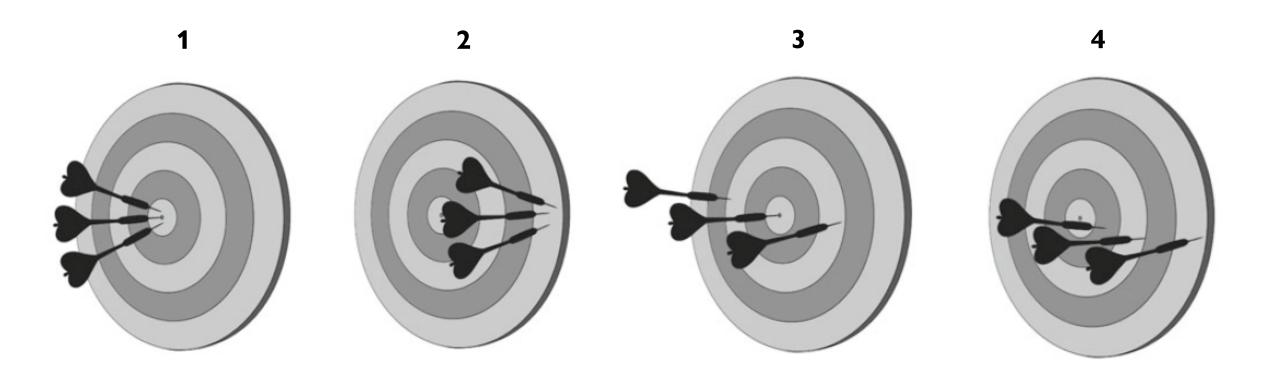
Variance

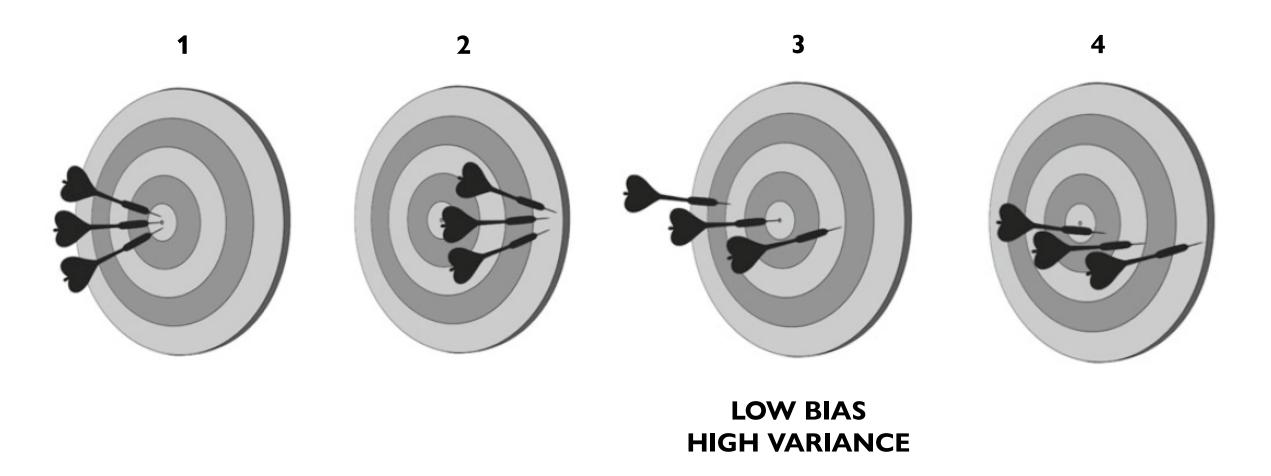
Consistency

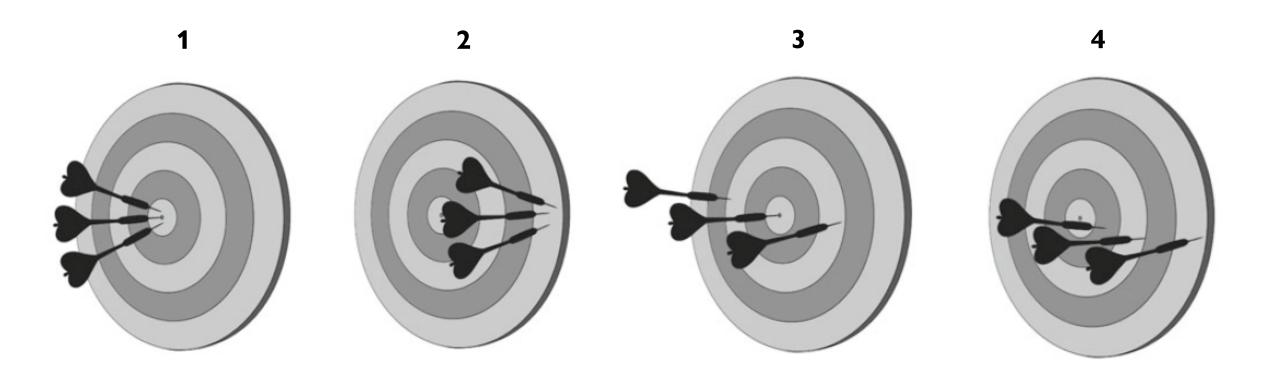


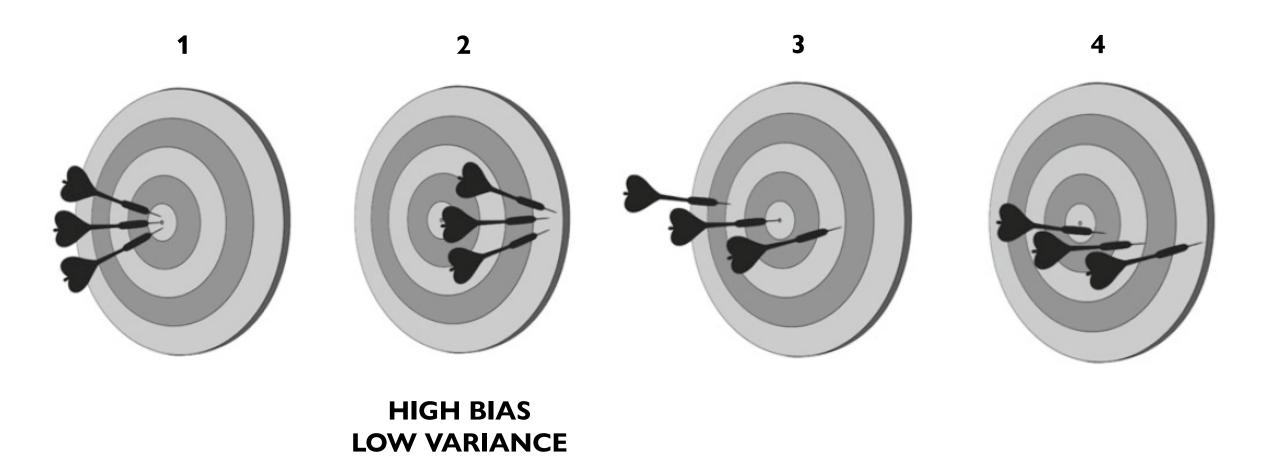
LOW VARIANCE

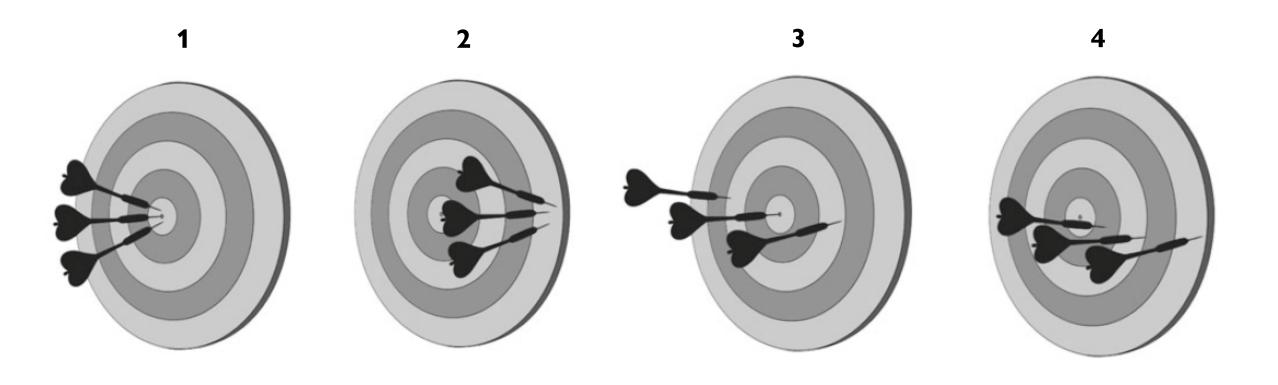


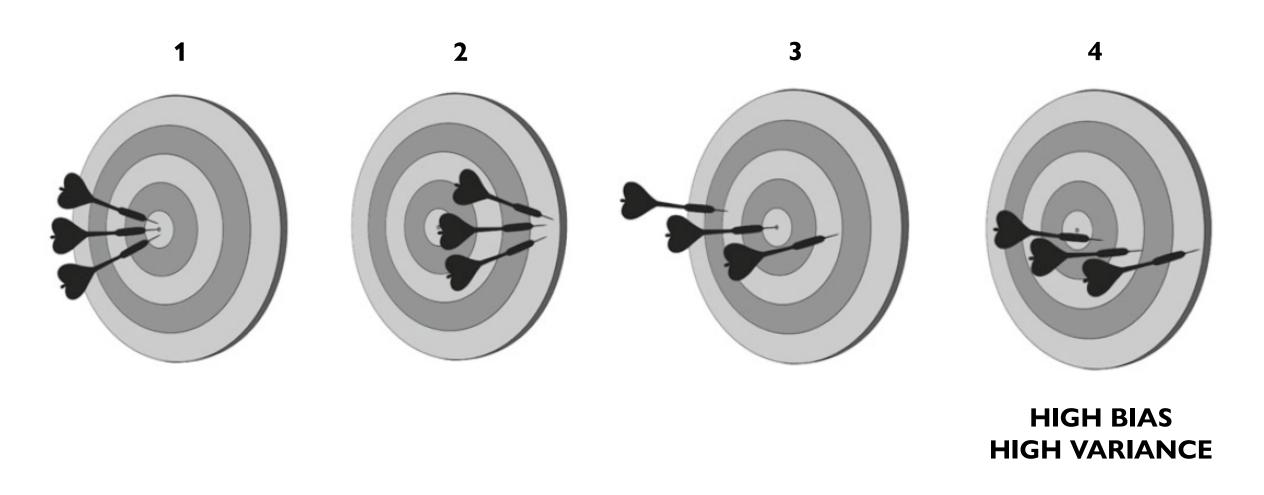


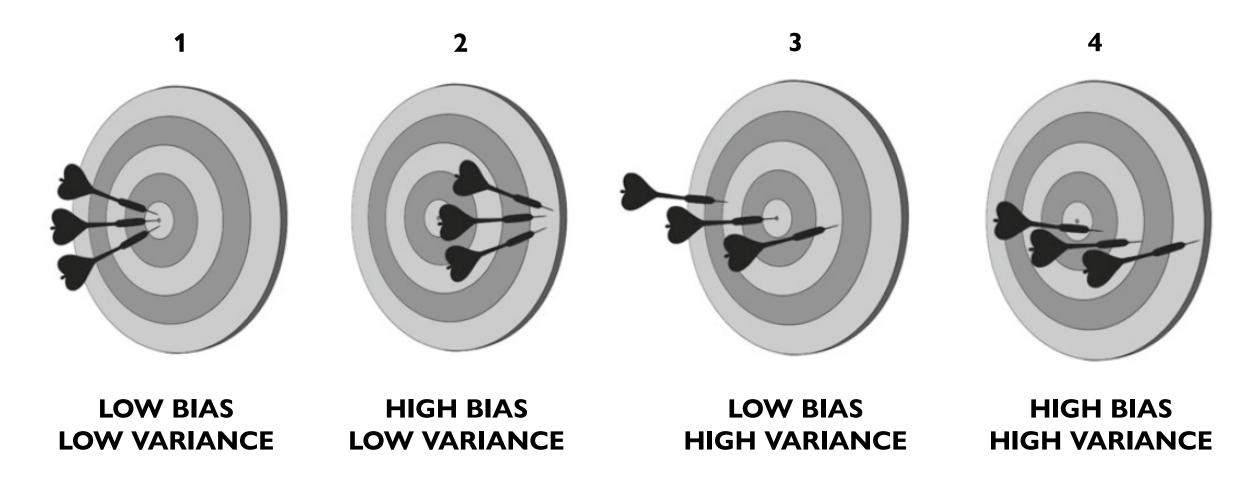








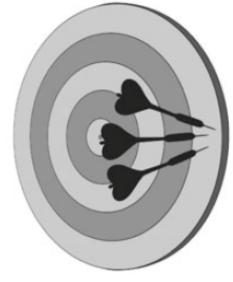




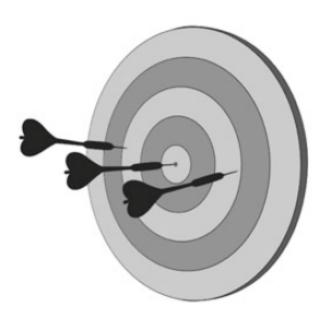




2



HIGH BIAS LOW VARIANCE



LOW BIAS HIGH VARIANCE

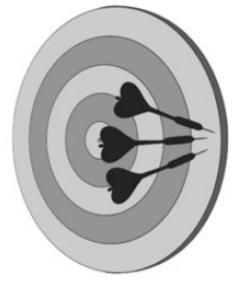


HIGH BIAS HIGH VARIANCE

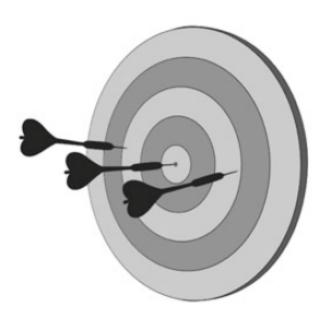




2



HIGH BIAS LOW VARIANCE



LOW BIAS HIGH VARIANCE



HIGH BIAS HIGH VARIANCE



BIAS

An estimator T(X) is unbiased of θ if

$$E(T(X)) = \theta$$

BIAS

An estimator T(X) is unbiased of θ if

$$E(T(X)) = \theta$$

Bias is defined as

$$bias(T(X)) = \theta - E(T(X))$$

$$X_1, ..., X_n \sim Bernoulli(p)$$

$$\hat{p}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i - \text{(un)biased?}$$

$$X_1, ..., X_n \sim Bernoulli(p)$$

$$\hat{p}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i - \text{(un)biased?}$$

$$E(\hat{p}_{ML}) =$$

$$X_1, ..., X_n \sim Bernoulli(p)$$

$$\hat{p}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i - \text{(un)biased?}$$

$$E(\hat{p}_{ML}) = E\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) =$$

$$X_1, ..., X_n \sim Bernoulli(p)$$

$$\hat{p}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i - \text{(un)biased?}$$

$$E(\hat{p}_{ML}) = E\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) = \frac{1}{n}\sum_{i=1}^{n}E(X_{i}) = 0$$

$$X_1, ..., X_n \sim Bernoulli(p)$$

$$\hat{p}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i - \text{(un)biased?}$$

$$E(\hat{p}_{ML}) = E\left(\frac{1}{n}\sum_{i=1}^{n} X_i\right) = \frac{1}{n}\sum_{i=1}^{n} E(X_i) = \frac{1}{n} \cdot np = p$$

$$X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i$$
 – (un)biased?

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i - \text{(un)biased?}$$

$$E(\hat{\mu}_{ML}) = E\left(\frac{1}{n}\sum_{i=1}^{n} X_i\right) = \frac{1}{n}\sum_{i=1}^{n} E(X_i) = \frac{1}{n} \cdot n\mu = \mu$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\sigma}_{ML}^2 = s^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 - \text{(un)biased?}$$

$$X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\sigma}_{ML}^2 = s^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 - \text{(un)biased?}$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 =$$

$$X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\sigma}_{ML}^2 = s^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 - \text{(un)biased?}$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \left(\frac{1}{n} \sum_{i=1}^n X_i^2\right)^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 + \frac{1}{n} \sum_{i=$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\sigma}_{ML}^2 = s^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 - \text{(un)biased?}$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \left(\frac{1}{n} \sum_{i=1}^n X_i^2\right)^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) =$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \frac{1}{n} \sum_{i=1}^n EX_i^2 - E\bar{X}^2$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \frac{1}{n} \sum_{i=1}^n EX_i^2 - E\bar{X}^2$$

$$EX_i^2 = ?$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \frac{1}{n} \sum_{i=1}^n EX_i^2 - E\bar{X}^2$$

$$EX_i^2 = ?$$

$$Var(X_i) = \sigma^2 =$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \frac{1}{n} \sum_{i=1}^n EX_i^2 - E\bar{X}^2$$

$$EX_i^2 = ?$$

$$Var(X_i) = \sigma^2 = EX_i^2 - (EX_i)^2 =$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \frac{1}{n} \sum_{i=1}^n EX_i^2 - E\bar{X}^2$$

$$EX_i^2 = ?$$

$$Var(X_i) = \sigma^2 = EX_i^2 - (EX_i)^2 = EX_i^2 - \mu^2$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \frac{1}{n} \sum_{i=1}^n EX_i^2 - E\bar{X}^2$$

$$EX_i^2 = ?$$

$$Var(X_i) = \sigma^2 = EX_i^2 - (EX_i)^2 = EX_i^2 - \mu^2$$

$$E(X_i^2) = \sigma^2 + \mu^2$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

$$Var(\bar{X}) =$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

 $Var(\bar{X}) = E\bar{X}^2 - (E\bar{X})^2 =$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

$$Var(\bar{X}) = E\bar{X}^2 - (E\bar{X})^2 = E\bar{X}^2 - \mu^2$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

$$Var(\bar{X}) = E\bar{X}^2 - (E\bar{X})^2 = E\bar{X}^2 - \mu^2$$

$$Var(\bar{X}) =$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

$$Var(\bar{X}) = E\bar{X}^2 - (E\bar{X})^2 = E\bar{X}^2 - \mu^2$$

$$Var(\bar{X}) = Var\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) =$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$
$$E\bar{X}^2 = ?$$

$$Var(\bar{X}) = E\bar{X}^2 - (E\bar{X})^2 = E\bar{X}^2 - \mu^2$$

$$Var(\bar{X}) = Var\left(\frac{1}{n}\sum_{i=1}^{n} X_i\right) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - E\bar{X}^2$$

$$E\bar{X}^2 = ?$$

$$Var(\bar{X}) = E\bar{X}^2 - (E\bar{X})^2 = E\bar{X}^2 - \mu^2$$

$$Var(\bar{X}) = Var\left(\frac{1}{n}\sum_{i=1}^{n} X_i\right) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

$$E\bar{X}^2 = \frac{\sigma^2}{n} + \mu^2$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 =$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 = \frac{(n-1)\sigma^2}{n}$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 = \frac{(n-1)\sigma^2}{n}$$

Sample variance $S^2 = \frac{1}{n} \sum_{i=1}^{n} (X_i - \bar{X})^2$ is a **biased** estimator for the parameter σ^2 of the Normal distribution.

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 = \frac{(n-1)\sigma^2}{n}$$

Sample variance $S^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$ is a **biased** estimator for the parameter σ^2 of the Normal distribution.

How to make it unbiased?

$$S_{unbiased}^2 =$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 = \frac{(n-1)\sigma^2}{n}$$

Sample variance $S^2 = \frac{1}{n} \sum_{i=1}^{n} (X_i - \bar{X})^2$ is a **biased** estimator for the parameter σ^2 of the Normal distribution.

How to make it unbiased?

$$S_{unbiased}^2 = \frac{n}{(n-1)} \cdot S^2 =$$

$$E(\hat{\sigma}_{ML}^2) = \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 = \frac{(n-1)\sigma^2}{n}$$

Sample variance $S^2 = \frac{1}{n} \sum_{i=1}^{n} (X_i - \bar{X})^2$ is a **biased** estimator for the parameter σ^2 of the Normal distribution.

How to make it unbiased?

$$S_{unbiased}^2 = \frac{n}{(n-1)} \cdot S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \overline{X})^2$$

VARIANCE

Variance of an estimator T(X) is defined as

$$Var(T(X)) = E\{T(X) - E(T(X))\}^{2}$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i, = \bar{X}, \qquad Var(\hat{\mu}_{ML}) = ?$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i, = \bar{X}, \qquad Var(\hat{\mu}_{ML}) = ?$$

$$Var(\bar{X}) =$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i, = \overline{X}, \qquad Var(\hat{\mu}_{ML}) = ?$$

$$Var(\bar{X}) = Var\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) =$$

$$X_1, \ldots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} X_i, = \overline{X}, \qquad Var(\hat{\mu}_{ML}) = ?$$

$$Var(\bar{X}) = Var\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) = \frac{n\sigma^{2}}{n^{2}} = \frac{\sigma^{2}}{n}$$

• Can a biased and an unbiased estimator have the same variance?

- Can a biased and an unbiased estimator have the same variance?
- Example:

$$\hat{\mu}_{ML} = \bar{X}$$
 — unbiased

$$\hat{\mu}_2 = \hat{\mu}_{ML} + 1$$
 -biased

- Can a biased and an unbiased estimator have the same variance?
- Example:

$$\widehat{\mu}_{ML} = \overline{X} - \text{unbiased}$$

$$Var(\widehat{\mu}_{ML}) = \frac{\sigma^2}{n}$$

$$\hat{\mu}_2 = \hat{\mu}_{ML} + 1$$
 -biased

- Can a biased and an unbiased estimator have the same variance?
- Example:

$$\widehat{\mu}_{ML} = \overline{X} - \text{unbiased}$$

$$Var(\widehat{\mu}_{ML}) = \frac{\sigma^2}{n}$$

$$\hat{\mu}_2 = \hat{\mu}_{ML} + 1$$
 —biased but

$$Var(\hat{\mu}_2) =$$

- Can a biased and an unbiased estimator have the same variance?
- Example:

$$\widehat{\mu}_{ML} = \overline{X} - \text{unbiased}$$

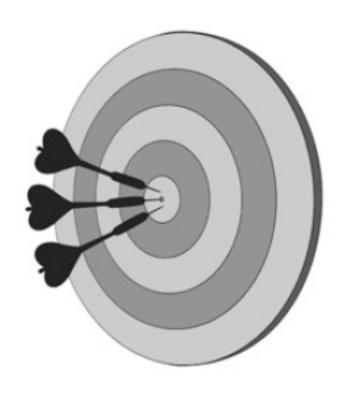
$$Var(\widehat{\mu}_{ML}) = \frac{\sigma^2}{n}$$

$$\hat{\mu}_2 = \hat{\mu}_{ML} + 1$$
 —biased but $Var(\hat{\mu}_2) = \frac{\sigma^2}{n}$

BIAS-VARIANCE TRADE-OFF

• Impossible to simultaneously optimize bias and variance.

• Related to *under-* and *overfitting* in Machine Learning.



LOW BIAS
LOW VARIANCE

CONSISTENCY

For a good estimator, as the sample size increases, the values of the estimator should get closer to the parameter being estimated.

CONSISTENCY

For a good estimator, as the sample size increases, the values of the estimator should get closer to the parameter being estimated.

Let $T_1, T_2, ... T_n$ be a sequence of estimators for $\theta, T_k = T(X_1, ..., X_k)$.

CONSISTENCY

For a good estimator, as the sample size increases, the values of the estimator should get closer to the parameter being estimated.

Let $T_1, T_2, ... T_n$ be a sequence of estimators for $\theta, T_k = T(X_1, ..., X_k)$.

Then $\{T_n\}$ is consistent if $\forall \epsilon > 0$

$$\lim_{n\to\infty} P(|T_n - \theta| < \epsilon) = 1$$

$$\hat{\mu} = \bar{X}$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

Let's chose
$$\epsilon = \frac{c\sigma}{\sqrt{n}}$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

Let's chose
$$\epsilon = \frac{c\sigma}{\sqrt{n}}$$

$$P\left(|T_n - \mu| \ge \frac{c\sigma}{\sqrt{n}}\right) \le$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

Let's chose
$$\epsilon = \frac{c\sigma}{\sqrt{n}}$$

$$P\left(|T_n - \mu| \ge \frac{c\sigma}{\sqrt{n}}\right) \le$$

Chebyshev's inequality:

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

Let's chose
$$\epsilon = \frac{c\sigma}{\sqrt{n}}$$

$$P\left(|T_n - \mu| \ge \frac{c\sigma}{\sqrt{n}}\right) \le \frac{1}{c^2} =$$

Chebyshev's inequality:

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

Let's chose
$$\epsilon = \frac{c\sigma}{\sqrt{n}}$$

$$P\left(|T_n - \mu| \ge \frac{c\sigma}{\sqrt{n}}\right) \le \frac{1}{c^2} = \frac{\sigma^2}{n\epsilon^2}$$

Chebyshev's inequality:

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

$$\hat{\mu} = \bar{X}, \qquad E\hat{\mu} = \mu, \qquad Var(\hat{\mu}) = \sigma^2/n$$

Let's show that
$$\forall \epsilon > 0$$
 $\lim_{n \to \infty} P(|T_n - \mu| \ge \epsilon) = 0$

Let's chose
$$\epsilon = \frac{c\sigma}{\sqrt{n}}$$

Chebyshev's inequality:
$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

$$P\left(|T_n - \mu| \ge \frac{c\sigma}{\sqrt{n}}\right) \le \frac{1}{c^2} = \frac{\sigma^2}{n\epsilon^2} \to 0 \quad n \to \infty$$

1. The MLE is consistent:

1. The MLE is consistent:

$$\hat{\theta} \stackrel{P}{\rightarrow} \theta$$

1. The MLE is **consistent**:

$$\hat{\theta} \stackrel{P}{\rightarrow} \theta$$

2. The MLE is **equivariant**:

1. The MLE is consistent:

$$\hat{\theta} \stackrel{P}{\rightarrow} \theta$$

2. The MLE is **equivariant**:

If $\hat{\theta}$ is MLE for θ than $g(\hat{\theta})$ is MLE for $g(\theta)$

1. The MLE is consistent:

$$\hat{\theta} \stackrel{P}{\rightarrow} \theta$$

2. The MLE is **equivariant**:

If
$$\hat{\theta}$$
 is MLE for θ than $g(\hat{\theta})$ is MLE for $g(\theta)$

3. MLE is asymptotically efficient:

1. The MLE is consistent:

$$\hat{\theta} \stackrel{P}{\to} \theta$$

2. The MLE is **equivariant**:

If
$$\hat{\theta}$$
 is MLE for θ than $g(\hat{\theta})$ is MLE for $g(\theta)$

3. MLE is asymptotically efficient: roughly speaking, among well-behaved estimators, it has the smallest variance, at least for large samples.