

# 1 Distributions and Densities

## 1.1 Types of Continuous Probability Distributions

The following are some of the most important continuous probability distributions, some of which we've seen before.

- Uniform
- Exponential
- Gamma
- Normal
- Cauchy

### 1.1.1 Uniform Distribution

#### Definition 1.1: Uniform Distribution

$U$  has **Uniform** $[a, b]$  distribution, for  $a < b$ , if its PDF is

$$f(u) = \frac{1}{b-a}$$

for  $a \leq u \leq b$ .

**Remark:** Note that  $b - a$  is the *length* of  $[a, b]$ .

**Note:** There are also higher-dimensional uniform distributions, but then we replace length with area of volume.

### 1.1.2 Exponential Distribution

#### Definition 1.2: Exponential Distribution

$X$  is **Exponential** $(\lambda)$  with rate  $\lambda > 0$  if its PDF is

$$f(x) = \lambda e^{-\lambda x}$$

for  $x > 0$ .

Note that there is an important connection between the Exponential and the Poisson, which we will now describe.

(Example: Busy Server.) Suppose that a single server queue (e.g. call center, bank, etc.) is very busy, so that there is always someone in the queue. Suppose that service times are independent and **Exponential** $(\lambda)^a$ . As soon as someone has been served, the next person in the queue starts being served immediately. Let  $X_1, X_2, \dots$  be an IID sequence of **Exponential** $(\lambda)$  random variables. Then, the time  $T_n$  at which the point the  $n$ th person has been served is distributed as

$$\sum_{i=1}^n X_i.$$

This sum of  $n$  IID **Exponential** $(\lambda)$  random variables has a special distribution, called the **Gamma** $(n, \lambda)$  distribution. This has PDF<sup>b</sup>

$$g(x) = \lambda e^{-\lambda x} \frac{(\lambda x)^{n-1}}{(n-1)!}$$

for  $x > 0$ .

<sup>a</sup>As you wait longer, the time that you wait “decays,” or decreases.

<sup>b</sup>When  $n = 1$ , this reduces to  $\lambda e^{-\lambda x}$ , the PDF of a single Exponential( $\lambda$ ).

The Poisson distribution arises when we ask: What is the distribution for the number  $N_t$  of people served by time  $t > 0$ ? At any given time  $t > 0$ , we will (with probability 1, since the service times are continuous) be in the middle of serving someone. This person does not count towards  $N_t$ .

Note that  $(N_t : t > 0)$  is called the **Poisson process**<sup>1</sup>. This is a *collection*, indexed by time, of random variables<sup>2</sup>. In particular, the random variable  $N_t$  has the Poisson( $\lambda t$ ) distribution.

*Proof.* Note that  $N_t = k$  if and only if the  $k$ th person is served at some time  $T_k = s \leq t$ , and then the next service  $X_{k+1} > t - s$ . In other words, we need to have finished serving  $k$  people and be in the middle of serving the  $(k + 1)$ th person. Since

$$T_k = \sum_{i=1}^k X_i$$

and  $X_{k+1}$  are independent, it follows that

$$\mathbb{P}(N_t = k) = \int_0^t f_{T_k}(s)[1 - F_{X_{k+1}}(t - s)]ds.$$

Since  $T_k$  is Gamma( $k, \lambda$ ) and  $X_{k+1}$  is Exponential( $\lambda$ ), it follows that

$$\begin{aligned} \mathbb{P}(N_t = k) &= \int_0^t f_{T_k}(s)[1 - F_{X_{k+1}}(t - s)]ds \\ &= \int_0^t \lambda e^{-\lambda s} \frac{(\lambda s)^{k-1}}{(k-1)!} \cdot e^{-\lambda(t-s)} ds \\ &= e^{-\lambda t} \frac{\lambda^k}{(k-1)!} \int_0^t s^{k-1} ds \\ &= e^{-\lambda t} \frac{(\lambda t)^k}{k!}. \end{aligned}$$

Hence, this is the PMF of a Poisson, as claimed. □

### 1.1.3 Normal/Gaussian Distribution

Recall that a Binomial converges to a Poisson if

$$p = \lambda/n \mapsto 0$$

as

$$n \mapsto \infty.$$

On the other hand, if  $p$  is *fixed* (not converging to 0), the Binomial approaches a different distribution as  $n \mapsto \infty$  called the **Normal** or **Gaussian** distribution. Indeed, as  $n$  goes to infinity, we see a bell-shaped curve.

<sup>1</sup>This is a fascinating mathematical object with many properties and applications, which won't be covered here.

<sup>2</sup>Such an object is called a **stochastic process**.

**Definition 1.3: Normal Distribution**

$X$  is **Normal** $(\mu, \sigma^2)$  if its PDF is

$$\frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

for  $-\infty < x < \infty$ .

When  $\mu = 0$  and  $\sigma^2 = 1$ ,  $X$  is called a standard normal, and its PDF is given the special notation

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}.$$

If  $X$  is not standard, then you can “standardize” by taking  $Z = (X - \mu)/\sigma$ . Then,  $X$  is Normal $(\mu, \sigma^2)$  if and only if  $Z$  is Normal $(0, 1)$ .

**1.1.4 Cauchy Distribution**

Now, suppose that  $X$  and  $Y$  are two independent standard Normal random variables. A very interesting distribution arises if we consider the ratio

$$Z = X/Y.$$

Since  $X$  and  $Y$  are independent,

$$f_Z(z) = \int_{S_z} f_{X,Y}(x,y) dx dy,$$

where  $S_z = \{(x,y) \mid x/y = z\}$ . We make a change of variables  $x = uz$  and  $y = u$ . Then, as  $u$  varies over  $\mathbb{R}$ , we get the whole set  $S_z$ . The Jacobian of this transformation is  $|u|$ , so

$$f_Z(z) = \int_{S_z} f_{X,Y}(x,y) dx dy = \int_{-\infty}^{\infty} |u| f_{X,Y}(uz, u) du.$$

This is the same as

$$2 \left( \frac{1}{\sqrt{2\pi}} \right)^2 \int_0^{\infty} u e^{-(uz)^2/2 - u^2/2} du.$$

It can be shown that

$$\int_0^{\infty} x e^{-cx^2} dx = \frac{1}{2c}.$$

Hence,

$$f_Z(z) = \frac{1}{\pi} \int_0^{\infty} u e^{-\frac{u^2(1+z^2)}{2}} du = \frac{1}{\pi(1+z^2)}.$$

A random variable with this PDF is called a (standard) **Cauchy** random variable.

Note that the Cauchy distribution has some interesting properties. In particular, it has no expected/average value. So, if you take an IID sequence  $X_1, X_2, \dots$  of Cauchy random variables, the limit

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i$$

does not exist.