Ch1 - Estimation (*Point and Interval Estimation*)

ECO 204

Statistics For Business and Economics - II

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Outline

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- 1. Descriptive vs Inferential Statistics
- 2. Inferential Statistics Part I, Estimation
 - Estimation, Estimate and Estimator
 - Estimator and Sampling Distribution
 - Properties of Point Estimator
 - How to get Point Estimators
- 3. Interval Estimation of Population Mean μ
 - Basic idea of Interval Estimation
 - lacktriangle Interval Estimation First Example σ known case
 - \blacksquare Deriving Interval Estimator σ known case
 - Interval Estimator σ unknown case
 - Interval Estimator σ unknown case with large samples
- 4. Proportion Estimation

1. Descriptive vs Inferential Statistics

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Descriptive vs Inferential Statistics

Motivating Picture

...a picture may be worth a thousand words.. - [Djikstra]

Consider following picture,

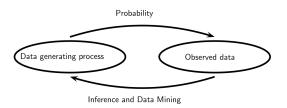


Figure 1: The figure is directly taken from Wasserman (2013), clearly explains what you did in ECO 104 (indicated by the arrow at the top going to right direction) and what you will do in ECO 204 (indicated by the arrow at the bottom going to left).

The work of *Probability Theory* is to describe - how the data / sample has been generated from a population, and the work of *Statistics* (or in particular Statistical Inference) is to make conclusions about the population using a sample.

- Welcome to ECO 204!
- ECO 204 is about Inferential Statistics (as opposed to Descriptive Statistics which you did in ECO 104, as a side note - ECO 104 was about two things - i) Descriptive Statistics and ii) Probability Theory)
- ► Any idea about *Inferential Statistics*...?
- Roughly, It's is a branch of Statistics that helps us to make conclusions about the population using a sample
- ▶ Now couple of questions from ECO 104
 - What is the population for any particular study.
 - ► What is *a sample*, and
 - ► How do you make *Inference* using example?
- Let's see one example and answer these questions systematically, suppose we have the following data from 5 students studying currently at EWU (this is a hypothetical data, perhaps randomly collected!).

	Gender	Monthly Family Income (tk)
1.	Male	70,150
2.	Female	20,755
3.	Male	44,758
4.	Female	38,790
5.	Male	20,579

You should already know that the columns are called *variables* and the rows are called *observations*. Let me stat with some questions.

Important Questions

Suppose our goal is to understand the Income and Gender of the current students at EWU, then consider following questions,

- Q1: What is the population of the study?
- Q2: Currently what proportion of students are female at EWU? Is it possible to calculate this?
- Q3: Currently what is the average income of all students at EWU? Is it possible to calculate this?
- ▶ Q4: Can we **estimate** proportion of students who are female at EWU? If yes then how?
- Q5: Can we estimate average monthly family income all students at EWU? If yes then how?
- Q6: What's the difference between sample and population quantities?
- Q7: If I have a population data, then do I need a sample? Why would I go for a sample?
- Q8: What are the possible issues in a sample?
- Q9: Are the sample quantities fixed or random?
- Q10: Do our predictions improve if we collect more samples?

Here are some answers.

- Ans 1: The population is the set of all units in the study, so in this case the population is all currently enrolled EWU students.
- Ans 2: We don't know this proportion exactly unless we have the data from all the students who are currently studying at EWU. In other words, we need data from the full population. Usually getting full population is very hard or time consuming. Although this is the quantity we are after, in other words this is one of our targets. but still we cannot calculate this exactly!
- Ans 3: Again the reasoning is same as Ans 2. Since we don't have the full population it's almost impossible to get the average income of all the current students.
- Ans 4: Yes, by using sample we can calculate **sample proportion** and then this sample proportion can work as an **estimate** for the **population proportion** of female students. In this sample proportion of female is 40%. So we can say, roughly, given that our sample represents the population, it is possible that the population proportion is close to 40%. Again, always remember here *population proportion* is for the entire population, and usually this is impossible to calculate. This is the first example of Statistical Inference, in particular we call this Statistical Estimation (more on this later)
- Ans 5: The answer is similar to the last answer. Yes we can estimate the **population mean** using sample mean of the students in the sample, in this case the sample mean is 39,006.4 taka and we can take this **sample mean** as an **estimate** of the **population mean**. Again to make it clear here **population mean** is the average of the income of all the current students. This is the second example of Statistical Inference and in particular Statistical Estimation.

- ▶ Ans 6: We already answered this, but let's make it concrete, sample quantities are estimates of the population quantities. Population quantities are always our targets. Definitely chances are very low that they will be exactly same, however with a "good" sample we might be close to our target.
- Ans 7: NO, I have all information I need, so no need for sampling. I go for sampling because usually I don't have access to the population.
- Ans 8: Obvious issue biased sample (sample doesn't represent population properly). Another issue small sample size. From now on, we will avoid issues related to "biased sample", and assume our sample is a fairly good representative of the population...but ...?
- Ans 9: Definitely random since if we have a different sample then sample proportion or sample mean both will change.
- Ans 10: Of course

- ▶ In the last example, we used a sample to calculate some sample quantities, in fact we calculated, *sample proportion* of female students in the sample, and *sample mean of family income*, and then we argued that we can use these objects to "predict", or or "conclude" or "infer" about the unknown population quantities.
- ► This was an example of *Inferential Statistics* or *Statistical Inference*, in particular this is what we call *Statistical Estimation!*. The formal definition of *Statistical Inference* would require us to carefully define many things, but perhaps informally we can say -

Definition: Statistical Inference

Statistical Inference is a procedure where we have a **target parameter** θ defined for the population, and then we use a sample to make conclusions regarding the target parameter. There are two key branches of Statistical Inference,

- ► Statistical Estimation (in short Estimation)
- Statistical Hypothesis Testing (in short Testing)

We can also categorize estimation as point and interval estimation. Typical examples of Statistical Inference include making conclusion about the population mean, population proportion, population variance, etc., using sample mean, sample proportion, sample variance.

We will **generally use** the Greek letter θ to represent the target parameter. In the examples above,

- for the first one θ is the population proportion. If we write population proportion with p, then in this case $\theta = p$
- for the second one θ is the population mean. If we write population mean with μ , then in this case $\theta=\mu$

From the next section, first we will focus on **Estimation** and then in the next chapter we will move to **Testing**. In both cases, we will always start with the some *numerical techniques* that you have already learned before, for example, sample mean \overline{x} , sample proportion \overline{p} , etc. But our goal is one step more - that is making *inference* about the population.

- ► The practice of Statistics falls broadly into two categories *Descriptive and Inferential*
- Descriptive Statistics is about describing the data using both numerical and graphical techniques / methods,
 - ▶ Numerical Methods: Calculating Sample Mean, Median, Mode, Variance, Standard Deviation, etc.
 - ► Graphical Methods: Looking at Bar Charts, Pie Charts, Histograms, etc
- ► The goal in this case is to describe the data, not to make any inference (or predict) about the population. This is what you did in ECO 104. You will do some recap exercises in the first problem set (which I will have to prepare, sorry not done yet!)

- However <u>Inferential Statistics</u> goes one step more, we try to make "good" conclusions about the population using a sample.
- ► There are two major themes of Inferential Statistics,
 - ► Statistical Estimation, in short we say Estimation
 - ► Hypothesis Testing, in short we say *Testing*
- ▶ You have already seen two examples of Estimation, we will see more. First we will focus on Estimation and then we will move to Hypothesis Testing.
- In both cases we will almost always start with the same numerical techniques that you have already learned in ECO104, that is we will use
 - sample mean,
 - sample median,
 - sample mode,
 - sample variance or sample standard deviation
 - sample quantiles or percentiles,
 - sample covariance, or sample correlation etc,

but our goal is in this course will be one step more - that is making *inference* about the population.

In the next section we will talk about *Estimation*.

2. Inferential Statistics - Part I. Estimation

- Estimation, Estimate and Estimator
- Estimator and Sampling Distribution
- Properties of Point Estimator
- How to get Point Estimators
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4. Proportion Estimation

Inferential Statistics - Part I, Estimation

Estimation, Estimate and Estimator

We have already seen two examples of Statistical Estimation, in particular we saw

- ightharpoonup we can use sample proportion \bar{p} to estimate the unknown population proportion p.
- ightharpoonup also, we can use sample mean \bar{x} to estimate the unknown population mean μ .

Below we clearly define Statistical Estimation, in particular we will define what is *the target* parameter, an estimate and an estimator, point estimation and interval estimation,

Definition (Statistical Estimation)

Statistical estimation is a statistical inference procedure which assigns numerical values to the unknown population parameters θ (e.g. mean μ , proportion p, variance σ^2) using data from a random sample X_1, X_2, \ldots, X_n . There are two types of Statistical Estimation,

- **Point estimation:** Provide a single best-guess number $\hat{\theta}$ for θ .
- ▶ Interval estimation: Provide an interval of possible values that, in repeated sampling, contains θ with a specified coverage probability (e.g. a 95 % confidence interval).

Any function of the random sample (X_1, \ldots, X_n) is called a *statistic*. When a statistic is used to infer a parameter it is called an *estimator*. Both Statistic and Estimator are random variables, and their values change from sample to sample. For a fixed sample the value of the estimator is called an *estimate*. The probability distribution of a Statistic (or an estimator) is called *sampling distribution*.

There are several things to understand in the definition,

Key Terms in the Definition

- ▶ Q1. What is a **Target Parameter** θ ?
- ▶ Q2. What is a **Random Sample** $(X_1, ..., X_n)$?
- ▶ Q3. What is a **Statistic**?
- Q4. What is an Estimator?
- ▶ Q5. What is an **Estimate**?
- Q6. What is a Sampling distribution of a Statistic / Estimator?
- ▶ Q7. What is **Point Estimation** and **Interval Estimation**.

Again we will answer these questions and also understand the definition using an example. Consider the following sample, this is the same sample as above but just with one variable that is - Monthly family Income

sl.	Monthly Family Income (tk)	R.V.
1.	70,150	$X_1 = ?$
2.	20,755	$X_2 = ?$
3.	44,758	$X_3 = ?$
4.	38,790	$X_4 = ?$
5.	20,579	$X_5 = ?$

Remarks on Notation: Usually for fixed numbers we will use lower case letters $x_1, x_2, x_3, \ldots, x_n$, rather than numbers, just to make it more general...and for random variables we use upper case letters $X_1, X_2, X_3, \ldots, X_n$. Also generally when we think about n random variables, we write $X_1, X_2, X_3, \ldots, X_n$, and similarly for n fixed numbers we write $x_1, x_2, x_3, \ldots, x_n$.

▶ Ans 1 : In this case the target parameter θ can be population mean of all current students at EWU. Since it's a mean in this case we often use μ for the target parameter.

Ans 2: The idea of the random sample is when we think each row as a random data point. How this is random? The idea is before sampling it is possible to have different values in each row. So we can think we have a random variable $X_1 = ?$ at first row, X_2 at the second row and so on... so in principle each row can take different values from the population, and each value before sampling is a random variable. Also note after the sampling, when we have **observed the sample**, a random data becomes a fixed number and in this case the random variable X_1 has already taken a value, so we get $X_1 = 70, 150$, and similarly for $X_2 = 20, 755$. Here we don't have any randomness, we call it **observed data or realized data**. Important, there is no randomness after we have observed the sample! So in the sample we have 5 random variables, they are X_1, X_2, X_3, X_4, X_5 , together we call it a random sample. And also we have 5 fixed numbers, and they are 70, 150; 20, 755; 44, 758; 38, 790 and 20, 579, together we call it an observed sample or a realized sample.

- Ans 3: A statistic is simply a function of the random sample, for example following are examples of statistic,
 - $ightharpoonup \sum_{i=1}^n X_i$
 - $\overline{X} = \frac{1}{n}(X_1 + X_2 + \dots, X_n) = \frac{1}{n}\sum_{i=1}^n X_i$
 - $ightharpoons rac{\overline{X}-\mu}{\sigma/\sqrt{n}}$, where μ and σ are simply constants, μ is population mean and σ is population standard deviation
 - ▶ $\frac{\overline{X}-\mu}{S/\sqrt{n}}$ where S is the sample standard deviation and $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i \overline{X})^2$ and $S = \sqrt{S^2}$ In a nutshell a statistic is any quantity made from the random sample X_1, X_2, \ldots, X_n
- ▶ Ans 4: When a Statistic is used for Estimation, we call it an **Estimator**. If Population mean μ is the target quantity, we can use \overline{X} to calculate random sample mean. In this case \overline{X} is a statistic and also it's an estimator.
- ▶ Ans 5: If we have a fixed sample then the value of the estimator becomes fixed, and on that case the numerical value of the estimator is called an **estimate**. For example if μ (or the population mean) is the target parameter then \overline{X} is the estimator (this is the random sample mean) and for fixed sample we will get a value of \overline{X} which will be an **estimate**. For example for this particular sample the estimate is $\overline{X} = 39,006$ taka. Again an estimate is a fixed number for a specific sample and an estimator is a random variable.

- Ans 6: Now since a statistic or an estimator are random variables and their values will be different from sample to sample, it will have many possible values. The probability distribution over these values is called sampling distribution. For a concrete example, think about \overline{X} , this is an estimator when the target parameter is μ , and this will have many possible values for different samples. So we can think about a probability distribution over the values of \overline{X} , and this is the sampling distribution of \overline{X} . We will talk about sampling distribution detail in the next section, but there is a theorem which says if population is normally distributed with mean μ and variane σ^2 , then the distribution of \overline{X} is also normal with mean μ and variance σ^2/n . Using notation we write this as $\overline{X} \sim \mathcal{N}(\mu, \sigma^2/n)$. Again sampling distribution of sample mean \overline{X} means distribution of different sample means, you should keep the shown picture (on the left) in your mind.
- ▶ Ans 6: When the parameter is μ , point estimation means we propose one best value, for example in this case sample mean \bar{x} for a fixed sample. On the other hand interval estimation means we propose an interval of possible values with certain probability, such that in repeated sampling the parameter μ will be inside the interval. We will talk about interval estimation in the coming sections. But before that we need to discuss about sampleing distribution of sample means.

Inferential Statistics - Part I, Estimation

Estimator and Sampling Distribution

- Before continuing from here you are supposed to have a look at the recap slides of ECO 104, if not, then please do so before you proceed,
- \blacktriangleright Now when it comes to random variable \overline{X} , we are interested in 3 important questions,
 - 1. What is the Expectation of the random variable \overline{X} , written as $\mathbb{E}(\overline{X})$?
 - 2. What is the variance of the random variable \overline{X} , written $\mathbb{V}(\overline{X})$?
 - 3. What is the probability distribution of \overline{X} (this is what we call *sampling distribution!*), for example is it the case that

$$\overline{X} \sim \mathcal{N}(?,?)$$

- ▶ The answer to the third question is what we call *Sampling Distribution of Sample Means*. Note that, this is the distribution of sample means \overline{x} , that we get from repeated sampling!. This is possibly the most important object for now,Definitely if we know the answer of 3, we know the answers of 1 and 2 (why?)
- Related to this question, will now see three important results,

Theorem 1.1: (Mean and Variance of \overline{X} with only i.i.d assumption)

If we have i.i.d random variables X_1, X_2, \ldots, X_n with the same mean μ and same variance σ^2 , then

$$i) \quad \mathbb{E}\left(\overline{X}\right) = \mu \tag{1}$$

$$ii) \quad \mathbb{V}\left(\overline{X}\right) = \frac{\sigma^2}{n} \tag{2}$$

In Words: This says if all the random variables X_1, X_2, \ldots, X_n are identically and independently distributed (in short we say i.i.d. random variables) with same mean and variance, this means

$$\mathbb{E}(X_1) = \mathbb{E}(X_2) = \ldots = \mathbb{E}(X_n) = \mu$$
 and $\mathbb{V}(X_1) = \mathbb{V}(X_2) = \ldots = \mathbb{V}(X_n) = \sigma^2$

then no matter what is the distribution of the random variables, the expectation of the sample mean \overline{X} will be equal to the population mean μ , i.e., $\mathbb{E}(\overline{X}) = \mu$, and the variance of the sample mean is $\mathbb{V}(\overline{X}) = \sigma^2/n$, (where n is the sample size. The key thing here is i.i.d. which means identically and independently distributed!).

Note the mean and variance could come from different distributions, not just normal. For example if the population follows Normal distribution $\mathcal{N}(\mu,\sigma^2)$ then the mean is μ and variance is σ^2 , in this case theorem says

$$\mathbb{E}\left(\overline{X}\right) = \mu$$

$$\mathbb{V}\left(\overline{X}\right) = \frac{\sigma^2}{n}$$

If the data follows $\operatorname{Bern}(p)$, then the mean is p and variance is p(1-p), so the theorem says

$$\mathbb{E}(\overline{X}) = p$$

$$\mathbb{V}(\overline{X}) = \frac{p(1-p)}{n}$$

So in this case the μ becomes p and σ^2 is p(1-p). Important is the theorem makes no assumption whether the data is Normal or Bernoulli.

Theorem 1.2: (Distribution of \overline{X} with normality and i.i.d assumption)

If we have i.i.d random variables X_1, X_2, \dots, X_n where they all are distributed with $\mathcal{N}(\mu, \sigma^2)$, then

i)
$$\mathbb{E}(\overline{X}) = \mu$$
 , ii) $\mathbb{V}(\overline{X}) = \frac{\sigma^2}{n}$
iii) $\overline{X} \sim \mathcal{N}(\mu, \frac{\sigma^2}{n})$, iv) $Z \sim \mathcal{N}(0, 1)$ where $Z = \frac{\overline{X} - \mu}{\frac{\sigma}{\sqrt{n}}}$ (3)

v)
$$T \sim t_{n-1}$$
 where $T = \frac{\overline{X} - \mu}{\frac{S}{\sqrt{n}}}$ and $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \overline{X})^2$ (4)

In Words: First note, we added one more condition here that is now we are assuming i.i.d. and also assuming Normality in the data, this gives the distribution of \overline{X} (which is more than what we got before) and the distribution is normal with mean μ and variance σ^2/n . This is a very important result, and we will do some problems related to this. The number (iv) is a direct consequence of the relationship between standard normal and normal distributions in general. The last one is a result for T statistic and t distribution, here n-1 is the parameter of the distribution, we call it degrees of freedom

Let's talk about the T statistic,

• first of all, note that T is a random variable, since \overline{X} is a random variable, (and also S is a random variable!)

$$T = \frac{\overline{X} - \mu}{\frac{S}{\sqrt{n}}}$$

number v) says the distribution of this random variable is t_{n-1} , where n-1 is the degrees of freedom (which is the parameter of the distribution)

Theorem 1.3: (Central Limit Theorem (CLT) and related results)

Let X_1, X_2, \ldots, X_n be i.i.d random variables with population mean μ and variance σ^2 . Then for large n (technically we need $n \to \infty$), we get following results:

i)
$$Z \stackrel{approx}{\sim} \mathcal{N}(0,1)$$
 where $Z = \frac{\overline{X} - \mu}{\frac{\sigma}{\sqrt{n}}}$ [CLT] (5)

$$ii)$$
 $\overline{X} \stackrel{approx}{\sim} \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$

iii)
$$T \stackrel{approx}{\sim} \mathcal{N}\left(0,1\right)$$
 where $T = \frac{\overline{X} - \mu}{\frac{S}{\sqrt{n}}}$

This theorem is similar to the first one, it says if we have i.i.d. random variables, with mean μ and variance σ^2 , then when *n* is *large*,

$$Z \stackrel{\textit{approx}}{\sim} \mathcal{N}\left(0,1\right) \text{ where } Z = \dfrac{\overline{X} - \mu}{\dfrac{\sigma}{\sqrt{n}}}$$

here again this is distribution free result, for example when we have $\mathcal{N}(\mu, \sigma^2)$, Z is same as above, but when the data is $\mathsf{Bern}(p)$, then

$$Z \stackrel{approx}{\sim} \mathcal{N}\left(0,1
ight) \text{ where } Z = rac{\overline{X} - p}{\sqrt{rac{p(1-p)}{n}}}$$

In general, note that we can always write

$$Z = \frac{\overline{X} - \mathbb{E}(\overline{X})}{\sqrt{\mathbb{V}(\overline{X})}} \text{ where } Z \stackrel{\textit{approx}}{\sim} \mathcal{N}\left(0,1\right)$$

Similarly for T Statistic we can write,

$$T = \frac{\overline{X} - \mathbb{E}(\overline{X})}{\sqrt{\widehat{\mathbb{V}}(\overline{X})}}$$

And in general we can write,

$$\widehat{\mathbb{V}}(\overline{X}) = rac{S^2}{n}$$
, where $S^2 = rac{1}{n-1} \sum_{i=1}^n (X_i - \overline{X})^2$

But for Bernoulli, we can also write,

$$\widehat{\mathbb{V}}(\overline{X}) = \frac{\bar{p}(1-\bar{p})}{n}$$

Note \bar{p} is same as \overline{X} which is sample proportion.

▶ Interestingly what happens here is this T statistic (which we wrote below) becomes approximately Normal when the sample size becomes very large, this is what we wrote in Theorem 1.3 (iii) i.e.,

$$T = rac{ar{p} - p}{\sqrt{rac{ar{p}(1 - ar{p})}{n}}} \overset{\mathsf{approx.}}{\sim} \mathcal{N}(0, 1)$$

▶ This is an application of Central Limit Theorem (CLT), and also another theorem known as Continuous Mapping Theorem.

Note that in repeated sampling, T is a random variable (this is because both \overline{X} and S is random), but you will see sometimes people use lower case t to denote the t-statistic. It doesn't matter as long as you understand when we are talking about random variable and when we are talking about a fixed number (i.e., realized value of the random variable). We will use uppercase T to denote this random variable and lowercase t to denote the realized value of the random variable.

Statistic, Point Estimator and Sampling Distribution - Important Remarks

- ► So we understood that *the idea of the sampling distribution is a repeated sampling idea*. In real life you can only have one sample, so you can never calculate this using a sample data.
- ▶ And the last three results tell us that, we can only know the sampling distribution of means under certain assumptions (in particular we need either normality or large sample size)
- ▶ If we assume normality (this means our data is normally distributed), then the distribution of the sample means is also normal and this result is for any sample size! This is called the exact distribution!
- ightharpoonup If we don't assume normality for the population, then usually we have no hope, except for large n.

Statistic, Point Estimator and Sampling Distribution - Important Remarks

- ► The standard deviation of sampling distribution is called *standard error*! This is standard deviation, but this name is special for sampling distribution.
- ▶ In general any function of the random sample is called a "Statistic", so an estimator is also a Statistic. The difference is Estimator is a type of Statistic where we are estimating some target! A statistic might not have any goal, it's just a function of random variables $X_1, X_2, X_3, \ldots, X_n$! The distribution of statistic is called sampling distribution.
- ▶ For example, \overline{X} , Z are both examples statistics but \overline{X} is an estimator for μ , Z is just a statistic.
- ▶ Another example is S^2 , where $S^2 = \frac{1}{n-1} \sum_{i=1}^n \left(X_i \overline{X}\right)^2$. This is a statistic since it's a function of the random sample. And this is also an estimator for σ^2 , because it is targeting population variance σ^2 . Note that S^2 is just a sample variance.

Inferential Statistics - Part I, Estimation

Properties of Point Estimator

Properties of Point Estimators

- ▶ Why did we take the average to estimate μ ? Why not median, or maximum? These all are examples of estimators for μ , so why sample mean \overline{X} ?
- lacktriangle The answer is, the sample average is a "good" estimator for the population mean μ
- ► What do we mean by "good"?
- ▶ One answer is it is an "unbiased" and a "consistent estimator"?
- ▶ What does "unbiasedness" mean? In notation this means

$$\mathbb{E}(\overline{X}) = \mu$$

- ▶ The interpretation is, if we calculate, sample means many times, *on average* we are not doing a bad job, even if our sample size *n* is not that big.
- Draw the dart picture.... on board
- Notice this result does not depend on the sample sizes, so we say unbiasedness is a finite sample property of an estimator.

Properties of Point Estimators

Now let's focus on *consistency*, we say an estimator is a *consistent estimator* then, if we have $n \to \infty$ then there is a very high probability that \overline{X} will approach to μ . So we can say

if $n \to \infty$ then $\overline{X} \to \mu$ happens with very high probability

- So this says, even if for small sample our sample mean is doing a bad job, as we increase the sample size we will eventually go very close to μ .
- ▶ If you contrast consistency to unbiasedness, you will notice that this is a limiting property or we say *asymptotic property* (because we are saying $n \to \infty$, recall limit...), as opposed to finite sample property!
- ▶ The estimator sample mean X_n is both an unbiased and a consistent estimator for the population mean μ .

Properties of Point Estimators

- So we talked about unbiasedness and consistency of an estimator, there is another thing called *variance* of an estimator, for sample mean this is $\mathbb{V}(\overline{X}) = \frac{\sigma^2}{n}$ (In Anderson you will see the notation $\sigma_{\overline{x}}^2$ to represent the same object, but I will not use this notation).
- You will not see too much discussion of the variance of an estimator here, but in higher courses this theme will come a lot!
- ▶ Looking at the variance of estimators is useful if we want to compare two or more estimators.
- lt is possible that two estimators are unbiased, one has very high variance.
- ▶ This means on average both are doing fine, but one has very high uncertainty!
- Again the dart picture!

Inferential Statistics - Part I, Estimation

How to get Point Estimators

How to get Estimators

- There are many techniques to get estimators. For example, here are some common techniques,
 - Method of Least Squares
 - ► Method of Maximum Likelihood
 - Method of Moments
- Sadly in this course we will not cover systematically any of these technique, but in higher courses you will see all of these methods.
- ▶ But we will talk about *method of least squares* when we talk about regression, and already we have seen some examples of *method of moment* techniques, ... roughly you can think the word "moment" is same as "expectation"...
- ▶ In the method of moment idea, we can get estimators by *replacing population expectation with averages*.

 $\mathbb{E}(X) = \mu$, if this is our target object

$$\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$$
 , replace the expectation, then we have an estimator

▶ Can you think about an estimator of σ^2 , recall σ^2 is actually the population variance of X, so $Var(X) = \sigma^2$, and for variance we have the following formula

$$\sigma^2 = \operatorname{Var}(X) = \mathbb{E}\left[\left(X - \mathbb{E}\left(X\right)\right)^2\right]$$

How to get Estimators

▶ This should be sample variance S^2 , where

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n \left(X_i - \overline{X} \right)^2$$

- ▶ Again note that, we replaced Expectation with averages.
- ▶ S^2 is an unbiased estimator of σ^2 .
- ▶ What if we divide by n rather than n-1? this is also an estimator of σ^2 , but unfortunately this is a biased estimator!

1. Descriptive vs Inferential Statistic

- 2. Inferential Statistics Part I, Estimation
 - Estimation, Estimate and Estimator
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 - Properties of Point EstimatorHow to get Point Estimators

3. Interval Estimation of Population Mean μ

- Basic idea of Interval Estimation
- \blacksquare Interval Estimation First Example σ known case
- Deriving Interval Estimator σ known case
- Interval Estimator σ unknown case
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4. Proportion Estimation

Interval Estimation of Population Mean μ

Interval Estimation of Population Mean μ

Basic idea of Interval Estimation

- Before we proceed, here is the recap of the terminologies,
 - $ightharpoonup \overline{X}$ is the *point estimator* of the population mean μ (note this is a random object!)
 - ightharpoonup When we calculate this for one fixed sample, then this is what we call *an estimate* of the unknown parameter μ .
 - ► The whole process is called estimation.
 - The probability distribution of the sample means (in repeated sampling) is what we call the sampling distribution of the sample means.

- ▶ Point estimator is nice, but it is rather crude! we are just giving one number as a guess. Now we will discuss another type of estimation, known as <u>Interval estimation</u>. Here also we will have <u>Interval estimators</u> (which is a random interval) and an <u>Interval estimate</u> for a fixed sample.
- Interval estimators is a little bit flexible, because it gives a range of possible values of the parameter (not just one value)! In particular, given α , where $0 < \alpha < 1$, we say a $100(1-\alpha)\%$ interval estimator for μ is a random interval [L,U] such that

$$\mathbb{P}(L \le \mu \le U) = 1 - \alpha \tag{6}$$

For example, if we want to construct a 95% interval estimator, then $1-\alpha=.95$, and we want to find L, and U such that, there is a 95% possibility that the true parameter will fall in this interval. So this means,

$$\mathbb{P}(L \le \mu \le U) = .95 \tag{7}$$

What is the interpretation

- ► Recall the *the frequency interpretation of probability*.
- ▶ Using the frequency interpretation means, if we construct the interval [L.U], then around 100 times, roughly 95 times the true μ fall in this interval.
- ▶ We can only think about probability when we have a random object, what is random here?
- First of all note that, μ here is not random (in classical statistics the parameter is never a random object, it is always fixed!), so what is random inside the probability?
- \triangleright Actually, we will see that the L and U are random in repeated sampling.
- ▶ In fact, the random L and U depends on the random sample $X_1, X_2, X_3, \ldots, X_n$, so we should write, $L(X_1, X_2, \ldots, X_n)$ and $U(X_1, X_2, \ldots, X_n)$. But just to make our life easier, we will use L and U. You should understand that these are functions of the random sample.
- ▶ We will see that we will construct interval estimator of the type,

$$\mathbb{P}(L \le \mu \le U) = 1 - \alpha$$

The interpretation is,

"In a repeated sampling, 95 out 100 times the interval constructed using [L,U] will contain the true parameter μ "

- ightharpoonup "Ideally" the interval [L, U] should have two properties:
 - ▶ $\mathbb{P}(L \leq \mu \leq U)$ should be high, that is, the true parameter $\mu \in [L, U]$ will happen with high probability.
 - ightharpoonup The length of the interval [L, U] should be relatively narrow on average.
- ▶ How do we find a interval estimator? There are different methods, but definitely we need to use a statistic and the distribution of the statistic (i.e., the sampling distribution)

Interval Estimation of Population Mean μ

Interval Estimation - First Example σ known case

 σ known

Let's do an example first where we will calculate interval estimate for a fixed sample.

Example 1.4: (Interval Estimator and Interval Estimate/Confidence Interval)

Suppose we have $\overline{x}=82$, population standard deviation $\sigma=20$, sample size n=100, and we are asked to compute the 95% confidence interval or interval estimate of the population mean μ , then since $z_{1-\alpha/2}=z_{0.975}=1.96$ (this is $1-\alpha/2$ quantile of the standard normal distribution and you can calculate this using R function qnorm(.975)), the interval estimator is

$$[\overline{X} - 1.96 \frac{20}{\sqrt{100}}, \quad \overline{X} + 1.96 \frac{20}{\sqrt{100}}]$$
 (8)

The interval estimate or confidence interval is

$$[82 - 1.96 \frac{20}{\sqrt{100}}, \quad 82 + 1.96 \frac{20}{\sqrt{100}}]$$

$$= [82 - 3.92, \quad 82 + 3.92]$$

$$= [78, \quad 85.92] \tag{9}$$

Now note, the first one at (8) is a *random interval* since \overline{X} is random but the second one (9) is a deterministic interval (there is no randomness here!), this is the interval estimate, Anderson et al. (2020) called this *confidence interval*.

So in the second one either our population mean μ is there or it is not there. If you say that there is a 95% probability that true parameter μ will fall inside [78, 85.92], this is a *wrong interpretation*. We can say "for this particular sample, *the interval estimate* is [78, 85.92]".

- So what is the correct interpretation? You should say if we construct these kinds of intervals 100 times, then roughly 95 times our true parameter will fall inside.
- So now we have a probabilistic interpretation.
- ▶ A Side Note: Note that when we constructed the interval estimate, we added and subtracted the following same number with \overline{x}

$$\frac{\sigma}{\sqrt{n}} \times z_{1-\alpha/2}$$

Here σ/\sqrt{n} is the standard error and the whole term is called the *margin of error* of the point estimate.

- ► The idea is \overline{x} is our point estimate, but of course there might be some error, so we say that with $1-\alpha$ confidence roughly the error is $\frac{\sigma}{\sqrt{n}} \times z_{1-\alpha/2}$
- Let's see how we can do the whole calculation of Example 1.5 in **Q**
- First note, we have following information
 - n = 100
 - $\overline{x} = 82$
 - $ightharpoonup \alpha = 0.05$ (this is because we are asked to construct 90% confidence interval)
 - $\sigma = 20$

Interval Estimate / Confidence Interval

 σ known

Rcode - sigma known (confidence interval)

```
# First create some objects with the information given
n <- 100
xbar <- 82
alpha <- 0.05
sigma <- 20
# calculate sderror and moe and save them as objects
sderror <- sigma/sqrt(n)</pre>
moe <- gnorm(1 - alpha/2) * sderror
# upper limit
xbar + moe
# [1] 85.91993
# lower limit
xbar - moe
# [1] 78.08007
```

► So the interval estimate or the confidence interval in this case is (78.08, 85.91)

Interval Estimation of Population Mean μ

Deriving Interval Estimator - σ known case

- ▶ Now how did we get that interval?
- ► Suppose we already know that our population data is normally distributed,
- ▶ This means all the random variables X_1, X_2, \ldots, X_n are also normally distributed with the distribution $\mathcal{N}(\mu, \sigma^2)$. Additionally assume they are independent.
- ► This means we have [applying Theorem 1.3 (iii)]

$$\overline{X} \sim \mathcal{N}(\mu, \sigma^2/n)$$

▶ But this also means [applying Theorem 1.3 (iv)] (this is just doing standardization)

$$Z = rac{\overline{X} - \mu}{rac{\sigma}{\sqrt{n}}} \sim \mathcal{N}(0, 1)$$

- ▶ Recall \overline{X} is a statistic and an estimator of μ . In this case Z is also a statistic, the benefit of transforming \overline{X} to Z is we can now use standard normal. Why did we do this? We will see that here Z also plays an important role to find the interval estimator for μ .
- Now, let's derive the interval estimator for μ. You can skip the derivation for exam but I recommend you to do it at least once in your lifetime, actually this is not difficult at all.

Deriving Interval Estimators

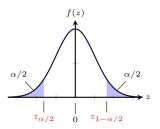
 σ known

We will construct two-sided interval (it is also possible to construct one sided interval, I will write something in the Appendix, but you can skip it for exam!).

To get the two sided interval, first fix α , let's say $\alpha=5\%$ then $1-\alpha$ is what we call *confidence coefficient / confidence level / nominal coverage*. Now since $Z\sim\mathcal{N}(0,1)$, we can write

$$\mathbb{P}\left(z_{\alpha/2} \le Z \le z_{1-\alpha/2}\right) = 1 - \alpha \tag{10}$$

Visually this means,



Here $z_{\alpha/2}$ is a value such that $\mathbb{P}(Z < z_{\alpha/2}) = \alpha/2$ and $z_{1-\alpha/2}$ is a value such that $\mathbb{P}(Z < z_{1-\alpha/2}) = 1 - \alpha/2$. It is important to mention that because of the *symmetry* of the Normal distribution always we will have $z_{\alpha/2} = -z_{1-\alpha/2}$ (note the two tail probabilities are equal, and it is $\alpha/2$).

Now we will do some algebra with the term inside the probability in (10), recall we had $z_{\alpha/2} \leq Z \leq z_{1-\alpha/2}$,

$$\begin{split} z_{\alpha/2} &\leq Z \leq z_{1-\alpha/2} = -z_{1-\alpha/2} \leq Z \leq z_{1-\alpha/2} \text{ [using symmetry of the normal]} \\ &= -z_{1-\alpha/2} \leq \frac{\overline{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \leq z_{1-\alpha/2} \\ &= -\frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \leq \overline{X} - \mu \leq \frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \text{ [multiplying all sides by } \sigma/\sqrt{n} \text{]} \\ &= \frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \geq -\overline{X} + \mu \geq -\frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \text{ [multiplying all sides by } -1 \text{]} \\ &= -\frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \leq -\overline{X} + \mu \leq \frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \text{ [rewriting the inequalities]} \\ &= \overline{X} - \frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \leq \mu \leq \overline{X} + \frac{\sigma}{\sqrt{n}} \ z_{1-\alpha/2} \text{ [adding } \overline{X} \text{ to all sides]} \end{split}$$

 σ known

So this means, writing

$$\mathbb{P}\left(z_{\alpha/2} \le Z \le z_{1-\alpha/2}\right) = 1 - \alpha$$

is same as

$$\mathbb{P}\left(\overline{X} - \frac{\sigma}{\sqrt{n}}z_{1-\alpha/2} \leq \mu \leq \overline{X} + \frac{\sigma}{\sqrt{n}}\ z_{1-\alpha/2}\right) = 1 - \alpha$$

So we have found our *upper and lower confidence limits*, these are

$$L = \overline{X} - rac{\sigma}{\sqrt{n}} z_{1-lpha/2}$$
 and $U = \overline{X} + rac{\sigma}{\sqrt{n}} \ z_{1-lpha/2}$

So the interval estimator is

$$[\overline{X} - \frac{\sigma}{\sqrt{n}} z_{1-\alpha/2}$$
 , $\overline{X} + \frac{\sigma}{\sqrt{n}} z_{1-\alpha/2}]$

Now if we calculate this for a fixed sample we will call it an interval estimate which will be

$$\left[\overline{x} - \frac{\sigma}{\sqrt{n}} z_{1-\alpha/2} \quad , \quad \overline{x} + \frac{\sigma}{\sqrt{n}} z_{1-\alpha/2}\right]$$

For an interval estimate, there is no probabilistic interpretation.

But for the interval estimator, can think about the frequency interpretation of probability, that is, if we do repeated sampling 100 times, then 95 out 100 times the intervals that we constructed will contain the true parameter μ

Interval Estimation of Population Mean μ

Interval Estimator - σ unknown case

 σ unknown

- ightharpoonup Can we construct intervals when we do not know the population standard deviation σ . The answer is YES!
- ▶ We need to use the statistic *T* from the Theorem 1.3 (v). This is called *t-statistic*, on the other hand when we used *Z*, that is called *z-statistic*.
- Note that in this case the statistic T follows a new sampling distribution, it is called t-distribution, with parameter n-1 (where n is the sample size!), there is a special name of this parameter, it is called degrees of freedom.
- ▶ The idea is if we use the sample standard deviation S, which is possible to calculate using the sample. Then we get a new Statistic T, which is distributed with t-distribution with n-1 degrees of freedom (Again to emphasize, degrees of freedom (df) is a parameter for the t-distribution).

T or the t-statistic is written as,

$$T = \frac{\overline{X} - \mu}{\frac{S}{\sqrt{n}}} \tag{11}$$

And according to the Theorem 1.3 (v), we have

$$T \sim t_{(n-1)}$$

- ▶ This means T is distributed with t distribution with parameter (n-1), or degrees of freedom (n-1).
- ▶ Note that in (11) S is the sample standard deviation, Recall S^2 is

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \overline{X})^2$$
 and $S = \sqrt{S}$

- ▶ Now how do we get an interval estimator in this case? The steps are actually same as page 28, except now you need to use quantile from t distribution with parameter n-1.
- If you do, then you should get the following *interval estimator* using t_{n-1} distribution,

$$\left[\overline{X} - \frac{S}{\sqrt{n}} t_{(n-1),1-\alpha/2} \quad \text{,} \quad \overline{X} + \frac{S}{\sqrt{n}} t_{(n-1),1-\alpha/2}\right]$$

► The *interval estimate* in this case is,

$$\left[\overline{x} - \frac{s}{\sqrt{n}} t_{(n-1), 1-\alpha/2} \quad , \quad \overline{x} + \frac{s}{\sqrt{n}} t_{(n-1), 1-\alpha/2}\right]$$
 (12)

▶ Here $t_{(n-1),1-\alpha/2}$ is the $1-\alpha/2$ quantile of the t distribution with parameter (n-1), and \overline{x} is the sample mean, and s is the sample standard deviation.

- Let's see a concrete example.
- ▶ Suppose in Example 1.5, we don't know σ , rather we know sample standard deviation s=18.5.
- ightharpoonup This means everything is same except the information of σ is not known to us,
 - n = 100
 - $\overline{x} = 82$
 - $\sim \alpha = 0.05$ (this is because we are asked to construct 90% confidence interval)
 - ► *s* = 18.5
- Now we will do the calculation in **Q**
- \blacktriangleright You will see following important differences compared to σ known case.
 - As mentioend we need to use t_{n-1} distribution
 - We need to use s, which is the sample standard deviation
 - ▶ Because we don't know σ , we cannot calculate the standard error σ/\sqrt{n} , however we can calculate the estimate of the standard error, which is $\frac{s}{\sqrt{n}}$

 σ unknown

Rcode - sigma unknown (confidence interval)

```
# First create some objects with the information given
n <- 100
xbar <- 82
alpha <- 0.05
s <- 18.5
# calculate the estimate of the sderror and moe and save them as objects
sderror_est <- s/sqrt(n)</pre>
moe <- qt(1 - alpha/2, n-1) * sderror_est
# upper limit
xbar + moe
# [1] 85.6708
# lower limit
xbar - moe
# [1] 78.3292
```

▶ So the 95% interval estimate or the confidence interval in this case is (78.33, 85.67)

Interval Estimation of Population Mean μ

Interval Estimator - $\boldsymbol{\sigma}$ unknown case with large samples

Large Sample results for t statistic

- ▶ There is one last important result before we go to the next section.
- ▶ Recall, from the last section we learned that, when we use S, rather than σ , we get a new statistic, that is what we called t statistic,

$$T = \frac{\overline{X} - \mu}{\frac{S}{\sqrt{n}}}$$

- And we already learned that this statistic is distributed with t distribution.
- ▶ Now there is a very interesting result. Have a look at the result in Theorem 1.4 (iii), this says if we have a very large sample, then

$$T \stackrel{approx}{\sim} \mathcal{N}\left(0,1\right)$$

- ▶ This means for very large *n*, the *t* statistic is approximately normally distributed!
- ▶ This means, if the sample size n is large then, we can forget about anything called t distribution, and just use the normal distribution with t statistic.

Large Sample results for t statistic

- ▶ What's the implication of this result for our confidence interval construction?
- The answer is, for large sample size we can construct the confidence interval in the following way,

$$\left[\overline{x} - \frac{s}{\sqrt{n}} z_{1-\alpha/2} \quad , \quad \overline{x} + \frac{s}{\sqrt{n}} z_{1-\alpha/2}\right]$$
 (13)

- Now if you compare (12) and (13), you will understand the difference.
- ► What's the difference?
- Again why does this happen?

1. Descriptive vs Inferential Statistic

2. Inferential Statistics - Part I, Estimation

- Estimation, Estimate and Estimator
- Estimator and Sampling Distribution
- Properties of Point Estimator
- How to get Point Estimators

3. Interval Estimation of Population Mean µ

- Basic idea of Interval Estimation
- Interval Estimation First Example σ known case
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4. Proportion Estimation

- ▶ The idea of the proportion estimation is same as mean estimation.
- If we do point estimation, in this case our *target object* is population proportion p and the point estimate is the sample proportion \overline{x} , which we sometimes also write as \overline{p}
- For interval estimation, the general formula is,

$$\overline{p} \pm z_{1-\alpha/2} \times \sqrt{\frac{\overline{p}(1-\overline{p})}{n}}$$

▶ Where $z_{1-\alpha/2}$ is the $1-\alpha/2$ quantile of the standard normal distribution, and \overline{p} is the sample proportion and $\sqrt{\frac{\overline{p}(1-\overline{p})}{n}}$ is the estimate of the standard error, since in this case the standard error is

$$\sqrt{\mathbb{V}(\overline{\rho})} = \sqrt{\frac{p(1-\rho)}{n}}$$

If we use \overline{p} then we have an estimate of the standard error, which is

$$\sqrt{\widehat{\mathbb{V}}(\overline{\rho})} = \sqrt{\frac{\overline{p}(1-\overline{\rho})}{n}}$$

► The code is very simple, here is an example

Rcode - Confidence Interval for Population Proportion

```
# First create some objects with the information given
n <- 100
pbar <- 82/100 # you can also write this as xbar
alpha <- 0.05
# calculate the estimate of the sderror and moe and save them as objects
sderror_est <- sqrt((pbar*(1-pbar))/n) # this is for Bernoulli</pre>
moe <- qnorm(1 - alpha/2) * sderror est
# upper limit
pbar + moe
# lower limit
pbar - moe
```

Although a theoretical matter, but what happens here is, the statistic (which we wrote below) becomes approximately Normal when the sample size becomes very large, i.e.,

$$\frac{\bar{p}-p}{\sqrt{\frac{\bar{p}(1-\bar{p})}{n}}} \overset{\mathsf{approx.}}{\sim} \mathcal{N}(0,1)$$

▶ This is the reason we use the standard normal distribution in the above code, and this is what we explained in page 27 and 28.

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