

CH 4. Continuous Random Variables and Their Probability Distributions

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(Cumulative) Distribution Function (CDF)

Exists for any r.v.!!

$$\underline{F(y)} = F_Y(y) = \underline{P(Y \leq y)}, \quad \text{for } -\infty < y < \infty. \quad \star$$

- ▶ $F(-\infty) := \lim_{y \rightarrow -\infty} F(y) = 0$
- ▶ $F(\infty) := \lim_{y \rightarrow \infty} F(y) = 1$
- ▶ $F(y)$ is a non-decreasing function of y .

Y is continuous if $F(y)$ is continuous for $-\infty < y < \infty$.

(e.g.)

Y	1	2	3	4
$P(Y)$.4	.3	.2	.1

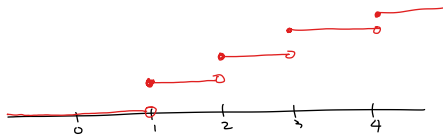
$$F(-1) = 0$$

\vdots

$$F(4) = 1$$

$$F(1) = .4 \quad F(2^-) = .4$$

$$F(2) = .7$$



- Step ft
- right-continuous.

different
unless
Y is conti

$$P(a \leq Y < b) = P(Y < b) - P(Y < a) = F(b-) - F(a-)$$
$$P(a < Y \leq b) = P(Y \leq b) - P(Y \leq a) = F(b) - F(a)$$

Some continuous RV, if their CDF is

absolutely continuous, then it has a density.
(pdf)

We focus on those only !!

Probability Density Function

$$\cancel{f(y) = P(Y=y)}$$

$$\left[\begin{aligned} f(y) &= \frac{dF(y)}{dy} = F'(y), \text{ or} \\ F(y) &= \int_{-\infty}^y f(t) dt = P(Y \leq y) \end{aligned} \right]$$

F : non-decreasing

- ▶ $f(y) \geq 0$ for any value of y . note: $f(y)$ can be > 1
it is NOT a probability !!
- ▶ $\int_{-\infty}^{\infty} f(y) dy = \underline{\underline{1}}$.

Probability can be found by:

• not bounded
• not continuous possible

$$P(a \leq Y \leq b) = \int_a^b f(y) dy$$

Expected Values for Continuous r.v.

$$E(Y) = \mu = \int_{-\infty}^{\infty} yf(y)dy,$$

provided that the integral exists.

$g(y)$: a real-valued function of Y . Then


$$E(g(Y)) = \int_{-\infty}^{\infty} g(y)f(y)dy$$

$$V(Y) = E[(Y - \mu)^2] = \int_{-\infty}^{\infty} (y - \mu)^2 f(y)dy = \sigma^2: \text{ Variance}$$

$$\sigma = \sqrt{V(Y)}: \text{ Standard deviation}$$

Uniform Probability Distribution

Flat
 Y has a continuous uniform probability distribution on the interval (θ_1, θ_2) iff the density of Y is

$$f(y) = \frac{1}{\theta_2 - \theta_1} \quad \theta_1 \leq y \leq \theta_2,$$


denoted as $Y \sim U(\theta_1, \theta_2)$.

$$\mu = E(Y) = \frac{\theta_1 + \theta_2}{2} \quad \text{and} \quad \sigma^2 = V(Y) \stackrel{\text{Check!!}}{=} \frac{(\theta_2 - \theta_1)^2}{12}.$$

Proof of Tchebyshev's inequality \longrightarrow (Markov Ineq.)

$$\forall k > 0, \quad P(|Y - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

$$\textcircled{::} \text{ LHS} = P(Y \leq \mu - k\sigma \text{ or } Y \geq \mu + k\sigma)$$

$$= \underbrace{\int_{-\infty}^{\mu - k\sigma} 1 \cdot f(y) dy}_{\substack{y \leq \mu - k\sigma \\ \Rightarrow y - \mu \leq -k\sigma \\ \Leftrightarrow \frac{y - \mu}{k\sigma} \leq -1 \\ \Rightarrow \frac{(y - \mu)^2}{k^2 \sigma^2} \geq 1}} + \int_{\mu + k\sigma}^{\infty} f(y) dy$$

$$\leq \int_{-\infty}^{\mu - k\sigma} \frac{(y - \mu)^2}{k^2 \sigma^2} f(y) dy + \int_{\mu + k\sigma}^{\infty} \frac{(y - \mu)^2}{k^2 \sigma^2} f(y) dy$$

$$\leq \frac{1}{k^2} \frac{1}{\sigma^2} \int_{-\infty}^{\infty} (y - \mu)^2 f(y) dy = \frac{1}{k^2}$$