

Supplementary note - Unmeasured confounding

Gibran Hemani, Kate Tilling, George Davey Smith

Background

The MR Steiger method uses instrumental variables to infer the causal orientation between two variables. Once an orientation is inferred there is a question as to whether the inference is erroneous due to processes that can bias the MR Steiger method. Here we focus on unmeasured confounding. Suppose that X is inferred to be causal for Y, we ask whether some combination of confounding parameters could give rise to the inferred direction under a data generating model in which the reverse causal direction is true. Figure 1 illustrates the two competing causal directions.

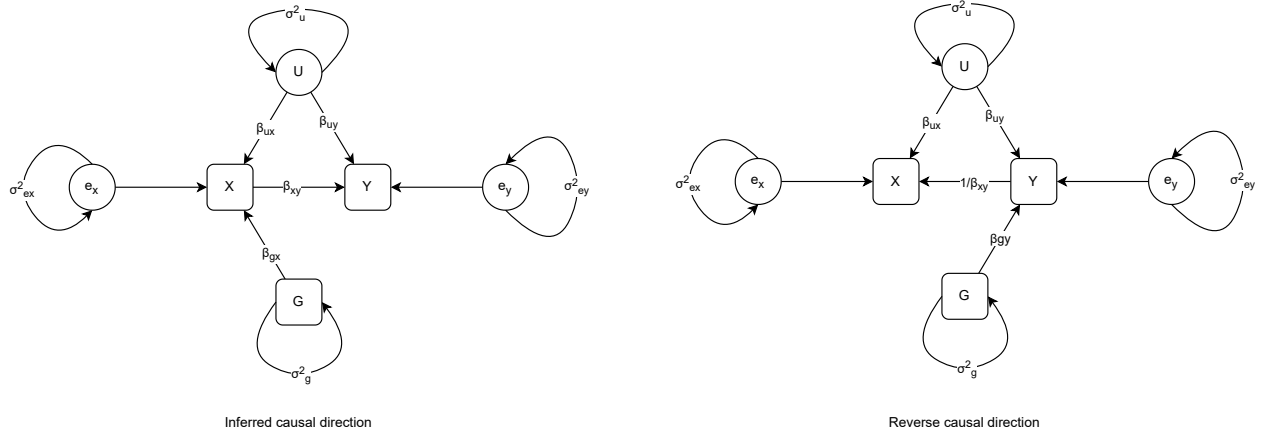


Figure 1: Two competing causal directions and their parameters represented as path diagrams.

According to the Inferred causal direction in figure 1,

$$X = \alpha_x + \beta_{gx}G + \beta_{ux}U + e_x$$

where SNP with allele frequency p has variance $\sigma_G^2 = 2p(1-p)$, $U \sim N(0, \sigma_u^2)$ is an unmeasured confounder, and $e_x \sim N(0, \sigma_{e_x}^2)$ is an error term. The variance of X will be

$$\sigma_x^2 = \beta_{gx}^2 \sigma_g^2 + \beta_{ux}^2 \sigma_u^2 + \sigma_{e_x}^2$$

We can also write

$$Y = \alpha_y + \beta_{xy}X + \beta_{uy}U + e_y$$

where $e_y \sim N(0, \sigma_{e_y}^2)$ is an error term. Going forwards intercept terms can be ignored. The variance of Y will be

$$\sigma_y^2 = \beta_{xy}^2 \beta_{gx}^2 \sigma_g^2 + \sigma_u^2 (\beta_{xy} \beta_{ux} + \beta_{uy})^2 + \beta_{xy}^2 \sigma_{e_x}^2 + \sigma_{e_y}^2$$

The variance explained in X by G will be

$$R_{gx}^2 = \frac{\beta_{gx}^2 \sigma_g^2}{\beta_{gx}^2 \sigma_g^2 + \beta_{ux}^2 \sigma_u^2 + \sigma_{e_x}^2}$$

The variance explained in Y by G will be

$$R_{gy}^2 = \frac{\beta_{gx}^2 \beta_{xy}^2 \sigma_g^2}{\beta_{xy}^2 \beta_{gx}^2 \sigma_g^2 + \sigma_u^2 (\beta_{xy} \beta_{ux} + \beta_{uy})^2 + \beta_{xy}^2 \sigma_{e_x}^2 + \sigma_{e_y}^2}$$

The variance explained in X by U will be

$$R_{ux}^2 = \frac{\beta_{ux}^2 \sigma_u^2}{\beta_{gx}^2 \sigma_g^2 + \beta_{ux}^2 \sigma_u^2 + \sigma_{e_x}^2}$$

The variance explained in Y by U will be

$$R_{uy}^2 = \frac{\sigma_u^2 (\beta_{uy} + \beta_{ux} \beta_{xy})^2}{\beta_{xy}^2 \beta_{gx}^2 \sigma_g^2 + \sigma_u^2 (\beta_{xy} \beta_{ux} + \beta_{uy})^2 + \beta_{xy}^2 \sigma_{e_x}^2 + \sigma_{e_y}^2}$$

Note that at this stage the asymptotic result of the computationally intensive simulations employed by Lutz et al could be derived analytically by observing that

$$\begin{aligned} R_{gx}^2 > R_{gy}^2 &\Rightarrow \frac{\beta_{gx}^2 \sigma_g^2}{\beta_{gx}^2 \sigma_g^2 + \beta_{ux}^2 \sigma_u^2 + \sigma_{e_x}^2} > \frac{\beta_{gx}^2 \beta_{xy}^2 \sigma_g^2}{\beta_{xy}^2 \beta_{gx}^2 \sigma_g^2 + \sigma_u^2 (\beta_{xy} \beta_{ux} + \beta_{uy})^2 + \beta_{xy}^2 \sigma_{e_x}^2 + \sigma_{e_y}^2} \\ &\Rightarrow \beta_{uy}^2 \sigma_u^2 + 2\beta_{ux} \beta_{uy} \beta_{xy} \sigma_u^2 + \sigma_{e_y}^2 > 0 \end{aligned}$$

is a second-degree polynomial inequality for β_{uy} which can be solved as

$$\begin{aligned} \Delta &= 4\beta_{ux}^2 \beta_{xy}^2 \sigma_u^4 - 4\sigma_u^2 \sigma_{e_y}^2 \\ &= 4\sigma_u^2 (\beta_{ux}^2 \beta_{xy}^2 - \sigma_{e_y}^2) \end{aligned}$$

This means that the comparison of the two R^2 values reduces to the sign of the quantity $\beta_{ux}^2 \beta_{xy}^2 - \sigma_{e_y}^2$. If the latter is negative, $R_{gx}^2 > R_{gy}^2$ holds regardless of the value of β_{uy} (and ignoring finite-sample variation). If it's positive, we will have $R_{gx}^2 < R_{gy}^2$ in the interval $[\beta_1, \beta_2]$ where

$$\begin{aligned} \beta_{1,2} &= \frac{-2\beta_{xu} \beta_{xy} \sigma_u^2 \pm \sqrt{\Delta}}{2\sigma_u^2} \\ &= \beta_{xu} \beta_{xy} \pm \sqrt{\beta_{xu}^2 \beta_{xy}^2 - \frac{\sigma_{e_y}^2}{\sigma_u^2}} \end{aligned}$$

As an application, for the parameter values used by Lutz et al. we obtain the reverse causal direction for β_{uy} values in the interval $[5 - 2\sqrt{6}, 5 + 2\sqrt{6}] \approx [0.1, 9.9]$.

However as described in the main text this does not serve as a sufficient sensitivity analysis because it allows observed quantities to vary considerably. We continue now to derive a more appropriate sensitivity analysis that does fix observed quantities.

In practice we tend to observe the following quantities: β_{gx} , σ_g^2 , σ_x^2 , σ_y^2 . The analysis is used to estimate β_{xy} . We can often obtain estimates of β_{OLS} . We do not know σ_u^2 , β_{ux} or β_{uy} , but given estimates of β_{OLS}

and β_{xy} we can obtain possible values for these confounder parameters. The observational association in this system will be

$$\beta_{OLS} = \frac{\beta_{gx}^2 \beta_{xy} \sigma_g^2 + \beta_{ux}^2 \beta_{xy} \sigma_u^2 + \beta_{ux} \beta_{uy} \sigma_u^2 + \beta_{xy} \sigma_{e_x}^2}{\sigma_g^2 \beta_{gx}^2 + \sigma_u^2 \beta_{ux}^2 + \sigma_{e_x}^2}$$

Hence the association between X and Y due to confounding will be

$$\begin{aligned} \beta_C &= \beta_{OLS} - \beta_{xy} \\ &= \frac{\beta_{ux} \beta_{uy} \sigma_u^2}{\sigma_g^2 \beta_{gx}^2 + \sigma_u^2 \beta_{ux}^2 + \sigma_{e_x}^2} \end{aligned}$$

The key question is this: **If β_{gx} , σ_g^2 , σ_x^2 , σ_y^2 , $\hat{\beta}_{OLS}$ and $\hat{\beta}_{xy}$ are fixed, are there values of R_{ux} and R_{uy} that can satisfy either X being causal for Y or Y being causal for X?** We approach this question by analytically exploring this possible confounding parameter space. The possible range of U-X confounding is

$$\beta_{ux} \in \left\{ -\sqrt{\frac{\sigma_x^2 - \beta_{gx}^2 \sigma_g^2}{\sigma_u^2}}, \sqrt{\frac{\sigma_x^2 - \beta_{gx}^2 \sigma_g^2}{\sigma_u^2}} \right\}$$

which means that for any particular value of β_{ux} within this range the values of

$$\sigma_{e_x}^2 = \sigma_x^2 - \beta_{gx}^2 \sigma_g^2 - \beta_{ux}^2 \sigma_u^2$$

and

$$\beta_{uy} = \beta_C \frac{\beta_{gx}^2 \sigma_g^2 + \beta_{ux}^2 \sigma_u^2 + \sigma_{e_x}^2}{\beta_{ux} \sigma_u^2}$$

and

$$\sigma_{e_y}^2 = \sigma_y^2 - \beta_{xy}^2 \beta_{gx}^2 \sigma_g^2 - \sigma_u^2 (\beta_{xy} \beta_{ux} + \beta_{uy})^2 - \beta_{xy}^2 \sigma_{e_x}^2$$

can be inferred directly. Overall, through this set of equations, we can obtain confounding values that could give rise to the observed quantities under either the inferred causal direction or the reverse causal direction. In the case of the reverse causal direction the value of $\beta_{xy,rev} = 1/\beta_{xy}$ and $\beta_{OLS,rev} = \beta_{OLS} \sigma_x^2 / \sigma_y^2$, $\beta_{gx,rev} = \beta_{gx} \beta_{xy}$. The sensitivity analysis proceeds by finding the total confounding parameter space across models for the inferred causal direction and the reverse causal direction, and then calculating the fraction of that parameter space that agrees with the inferred causal direction. A proportion close to 1 will suggest that there is relatively little chance of the inferred direction being incorrect due to unmeasured confounding. If the OLS estimate is unknown then a range of plausible values can be evaluated.

One further component to this approach is the option to weight the possible parameter space. We might consider it less plausible that large fractions of the variance in X and Y are explained by confounding variables, and so the contribution of scenarios that have confounding values that explain more of the variance can be downweighted. The weighting is obtained by

$$w = \phi_{0,s}(R_{ux}^2) \phi_{0,s}(R_{uy}^2)$$

where $\phi_{0,s}$ is the normal density function with mean 0 and standard deviation s , the scaling parameter. Smaller s will lead to more downweighting of larger confounding variances.

Finally, as an alternative to reporting an overall probability, analysts can evaluate the confounding parameters that would be required to give rise to the opposing causal direction to evaluate their plausibility.

Analysis

```
library(dplyr)

##
## Attaching package: 'dplyr'

## The following objects are masked from 'package:stats':
##
##   filter, lag

## The following objects are masked from 'package:base':
##
##   intersect, setdiff, setequal, union

library(ggplot2)
library(latex2exp)
library(ieugwasr)

## API: public: http://gwas-api.mrcieu.ac.uk/

library(TwoSampleMR)

## TwoSampleMR version 0.5.6
## [>] New: Option to use non-European LD reference panels for clumping etc
## [>] Some studies temporarily quarantined to verify effect allele
## [>] See news(package='TwoSampleMR') and https://gwas.mrcieu.ac.uk for further details

##
## Attaching package: 'TwoSampleMR'

## The following object is masked from 'package:ieugwasr':
##
##   ld_matrix

This function obtains the rsq values given fixed parameters

get_calcs <- function(bxy, bgx, bux, buy, vg, vu, vex, vey) {
  args <- as.list(environment())
  bxyo <- ((bgx^2*bxy*vg + bux^2*bxy*vu + bux*buy*vu + bxy*vex) / (vg*bgx^2 + vu*bux^2 + vex))
  vx <- bgx^2 * vg + bux^2 * vu + vex
  vy <- bxy^2*bgx^2*vg + (bxy*bux+buy)^2*vu + bxy^2*(vex) + vey
  conf <- bux * vu * buy / (vg * bgx^2 + vu * bux^2 + vex)
  rsqxyo <- bxyo^2 * vx / vy
  rsqxyos <- rsqxyo * sign(bxyo)
  rsqxy <- bxy^2 * vx / vy
  rsqxys <- rsqxy * sign(bxy)
  rsqgx <- bgx^2*vg / (bgx^2 * vg + bux^2 * vu + vex)
  rsqgy <- bgx^2*bxy^2*vg / (bxy^2*bgx^2*vg + (bxy*bux+buy)^2*vu + bxy^2*(vex) + vey)
  rsqux <- bux^2*vu / (bgx^2 * vg + bux^2 * vu + vex)
  rsquy <- (buy + bux * bxy)^2 * vu / (bxy^2*bgx^2*vg + (bxy*bux+buy)^2*vu + bxy^2*(vex) + vey)
  rsquxs <- rsqux * sign(bux)
  rsquys <- rsquy * sign(buy)
  return(c(args, list(
    vx=vx,
```

```

    vy=vy,
    bxyo=bxyo,
    conf=conf,
    rsqgx=rsqgx,
    rsqgy=rsqgy,
    rsqux=rsqux,
    rsquy=rsquy,
    rsqxy=rsqxy,
    rsqxyo=rsqxyo,
    rsqxyos=rsqxyos,
    rsquxs=rsquxs,
    rsquys=rsquys,
    rsqxys=rsqxys
  )))
}

```

Check by comparing to simulated individual level data

```

fn <- function(x,y, z)
{
  a <- 10
}
fn(1,c(2,3,4,5),3) %>% as_tibble()

```

```

## # A tibble: 1 x 1
##   value
##   <dbl>
## 1     10

```

```

get_calcs_id <- function(bxy, bgx, bux, buy, vg, vu, vex, vey, n=500000){
  u <- rnorm(n, sd = sqrt(vu))
  g <- rnorm(n, sd = sqrt(vg))
  ex <- rnorm(n, sd = sqrt(vex))
  ey <- rnorm(n, sd = sqrt(vey))
  x <- u * bux + g * bgx + ex
  y <- u * buy + x * bxy + ey
  res <- tibble(
    bxyo=cov(x,y)/var(x),
    rsqux=cor(u,x)^2,
    rsquy=cor(u,y)^2,
    rsqxyo=cor(x,y)^2,
    rsqxyos=rsqxyo*sign(bxyo),
    rsquxs=rsqux*sign(bux),
    rsquys=rsquy*sign(buy),
    rsqgx = cor(g,x)^2,
    rsqgy = cor(g,y)^2
  )
  return(res)
}
bind_rows(get_calcs(0.1, 1, 1, 1, 0.5, 0.1, 0.4, 0.9-0.1^2) %>% as_tibble(),
get_calcs_id(0.1, 1, 1, 1, 0.5, 0.1, 0.4, 0.9-0.1^2) %>% as_tibble())

```

```

## # A tibble: 2 x 22
##   bxy  bgx  bux  buy  vg  vu  vex  vey  vx  vy  bxyo  conf  rsqgx

```

```
## <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1  0.1    1    1    1  0.5  0.1  0.4  0.89    1  1.02 0.2    0.1 0.5
## 2  NA     NA     NA     NA NA     NA     NA     NA     NA NA     0.201 NA    0.501
## # ... with 9 more variables: rsqgy <dbl>, rsqux <dbl>, rsquy <dbl>,
## #   rsqxy <dbl>, rsqxyo <dbl>, rsqxyos <dbl>, rsquxs <dbl>, rsquys <dbl>,
## #   rsqxys <dbl>
```

Parameter ranges used by Lutz et al

```
get_calcs(1, 1, -5, seq(0,11,by=1), 1, 1, 1, 1) %>% as_tibble() %>% dplyr::select(bux, buy, vx, vy, rsq
```

```
## # A tibble: 12 x 6
##   bux   buy   vx   vy rsqgx rsqgy
##   <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1   -5    0   27   28 0.0370 0.0357
## 2   -5    1   27   19 0.0370 0.0526
## 3   -5    2   27   12 0.0370 0.0833
## 4   -5    3   27    7 0.0370 0.143
## 5   -5    4   27    4 0.0370 0.25
## 6   -5    5   27    3 0.0370 0.333
## 7   -5    6   27    4 0.0370 0.25
## 8   -5    7   27    7 0.0370 0.143
## 9   -5    8   27   12 0.0370 0.0833
## 10  -5    9   27   19 0.0370 0.0526
## 11  -5   10   27   28 0.0370 0.0357
## 12  -5   11   27   39 0.0370 0.0256
```

Calculate the values of rux and ruy that would satisfy the fixed parameters

```
sens <- function(bxy=0.1, bxyo=0.2, bgx=0.5, vx=1, vy=1, vu=1, vg=0.5, simsize=100) {
  # vx <- bgx^2 * vg + p$bx_vec^2 * vu + vex
  bux_lim <- sqrt((vx - bgx^2 * vg)/vu)
  bux_vec <- seq(-bux_lim, bux_lim, length.out=simsize)
  # Allow causal effect to vary by +/- 200%
  vex <- vx - bgx^2 * vg - bux_vec^2 * vu
  conf <- bxyo - bxy
  buy_vec <- conf * (bgx^2*vg + bux_vec^2*vu + vex) / (bux_vec * vu)
  vey <- vy - (bxy^2*bgx^2*vg + (bxy*bux_vec+buy_vec)^2*vu + bxy^2*vex)
  # vy <- bxy^2*bgx^2*vg + (bxy*bux_vec+buy_vec)^2*vu + bxy^2*vex + vey
  bux_vec * vu * buy_vec / (vg * bgx^2 + vu * bux_vec^2 + vex)
  res <- get_calcs(bxy, bgx, bux_vec, buy_vec, vg, vu, vex, vey) %>%
    as_tibble() %>%
    mutate(bxy=bxy, bgx=bgx, bux=bux_vec, buy=buy_vec, vg=vg, vu=vu, vex=vex, vey=vey)
  return(res)
}
```

Test sensitivity analysis

```
u_sensitivity <- function(bxy, bxyo, bgx, vx, vy, vg, vu = 1, simsize=10000, scaling=1, plot=TRUE)
{
  o <- params <- bind_rows(
    sens(bxy=bxy, bxyo=bxyo, bgx=bgx, vx=vx, vy=vy, vu=vu, vg=vg, simsize=simsize) %>%
      mutate(direction="inferred"),
    sens(bxy=1/bxy, bxyo=bxyo * vx / vy, bgx=bgx * bxy, vx=vy, vy=vx, vu=vu, vg=vg, simsize=simsize) %>%
      mutate(direction="reverse")
  ) %>%
  filter(
```

```

    vex >= 0 &
    vey >= 0 &
    rsquy >= 0 & rsquy <= 1 &
    rsqux >= 0 & rsqux <= 1 &
    rsqgx >= 0 & rsqgx <= 1 &
    rsqgy >= 0 & rsqgy <= 1
  ) %>%
  group_by(direction) %>%
  do({
    x <- .
    x1 <- x$rsqux[-1]
    x2 <- x$rsqux[-length(x$rsqux)]
    y1 <- x$rsquy[-1]
    y2 <- x$rsquy[-length(x$rsquy)]
    d <- sqrt((x1-x2)^2 + (y1-y2)^2)
    d[d > quantile(d, na.rm=T, probs=0.99)*4] <- NA
    x$d <- c(NA, d)
    x$weight <- dnorm(x$rsqux, sd=scaling) * dnorm(x$rsquy, sd=scaling)
    x
  })

  w <- o$d * o$weight
  w1 <- w[o$direction=="inferred"]
  prop <- sum(w1, na.rm=T) / sum(w, na.rm=T)
  ret <- list(result=o, prop=prop)
  if(plot) {
    ret$pl <- ggplot(o, aes(x=rsquxs, y=rsquys)) +
      geom_point(aes(colour=direction, size=weight))
  }
  return(ret)
}

```

Example 1 - bxy and bxyo are similar, and confounders that explain more of the variance are strongly downweighted

```

r <- u_sensitivity(bxy=0.1, bxyo=0.2, bgx=0.5, vx=1, vy=1, vg=0.5, scaling=0.1)
r

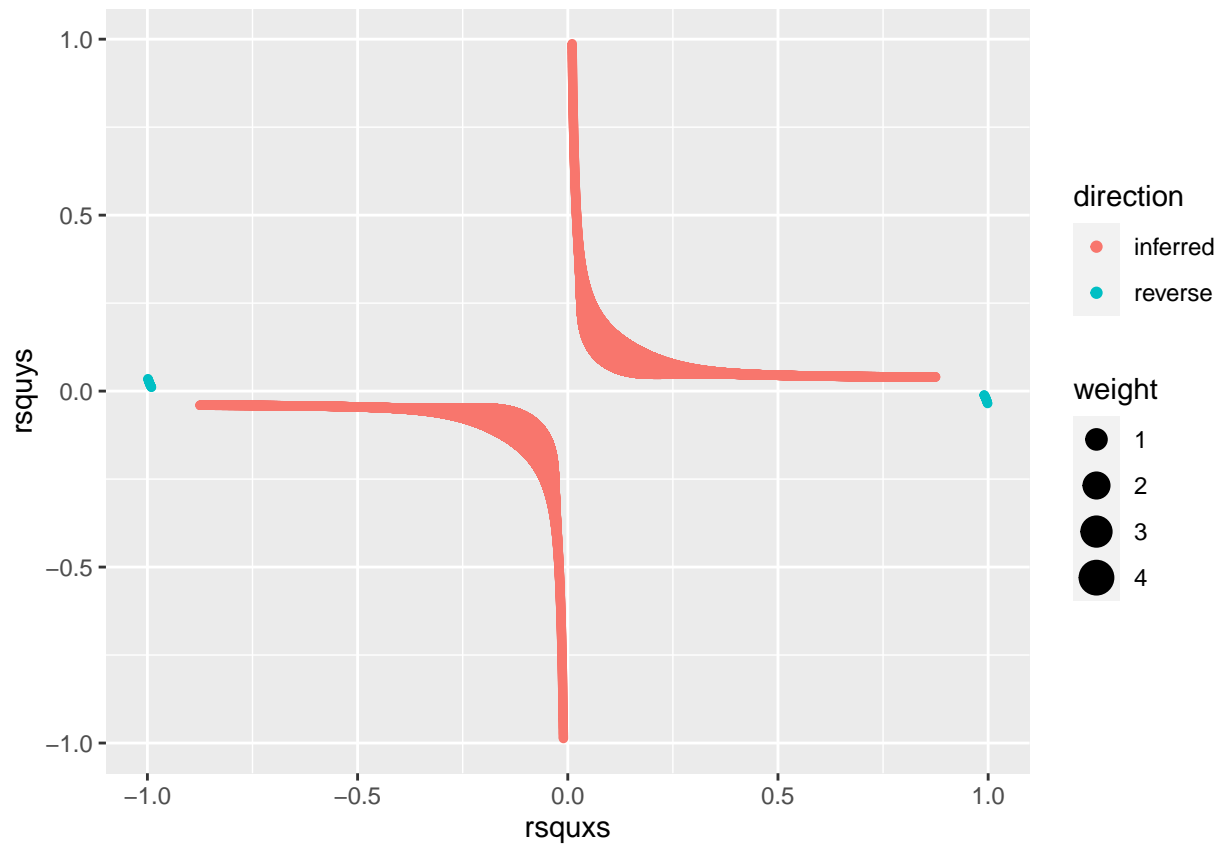
```

```

## $result
## # A tibble: 8,958 x 25
## # Groups:   direction [2]
##   bxy  bgx  bux  buy  vg  vu  vex  vey  vx  vy  bxyo  conf
##   <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1  0.1  0.5 -0.935 -0.107 0.5  1 0      0.959  1  1  0.2  0.1
## 2  0.1  0.5 -0.935 -0.107 0.5  1 0.000350 0.959  1  1  0.2  0.1
## 3  0.1  0.5 -0.935 -0.107 0.5  1 0.000700 0.959  1  1  0.2  0.1
## 4  0.1  0.5 -0.935 -0.107 0.5  1 0.00105 0.959  1  1  0.2  0.1
## 5  0.1  0.5 -0.935 -0.107 0.5  1 0.00140 0.959  1  1  0.2  0.1
## 6  0.1  0.5 -0.934 -0.107 0.5  1 0.00175 0.959  1  1  0.2  0.1
## 7  0.1  0.5 -0.934 -0.107 0.5  1 0.00210 0.959  1  1  0.2  0.1
## 8  0.1  0.5 -0.934 -0.107 0.5  1 0.00245 0.959  1  1  0.2  0.1
## 9  0.1  0.5 -0.934 -0.107 0.5  1 0.00280 0.959  1  1  0.2  0.1
## 10 0.1  0.5 -0.934 -0.107 0.5  1 0.00315 0.959  1  1  0.2  0.1
## # ... with 8,948 more rows, and 13 more variables: rsqgx <dbl>, rsqgy <dbl>,
## #   rsqux <dbl>, rsquy <dbl>, rsqxy <dbl>, rsqxyo <dbl>, rsqxyos <dbl>,

```

```
## # rsquxs <dbl>, rsquys <dbl>, rsqxys <dbl>, direction <chr>, d <dbl>,
## # weight <dbl>
##
## $prop
## [1] 1
##
## $pl
```



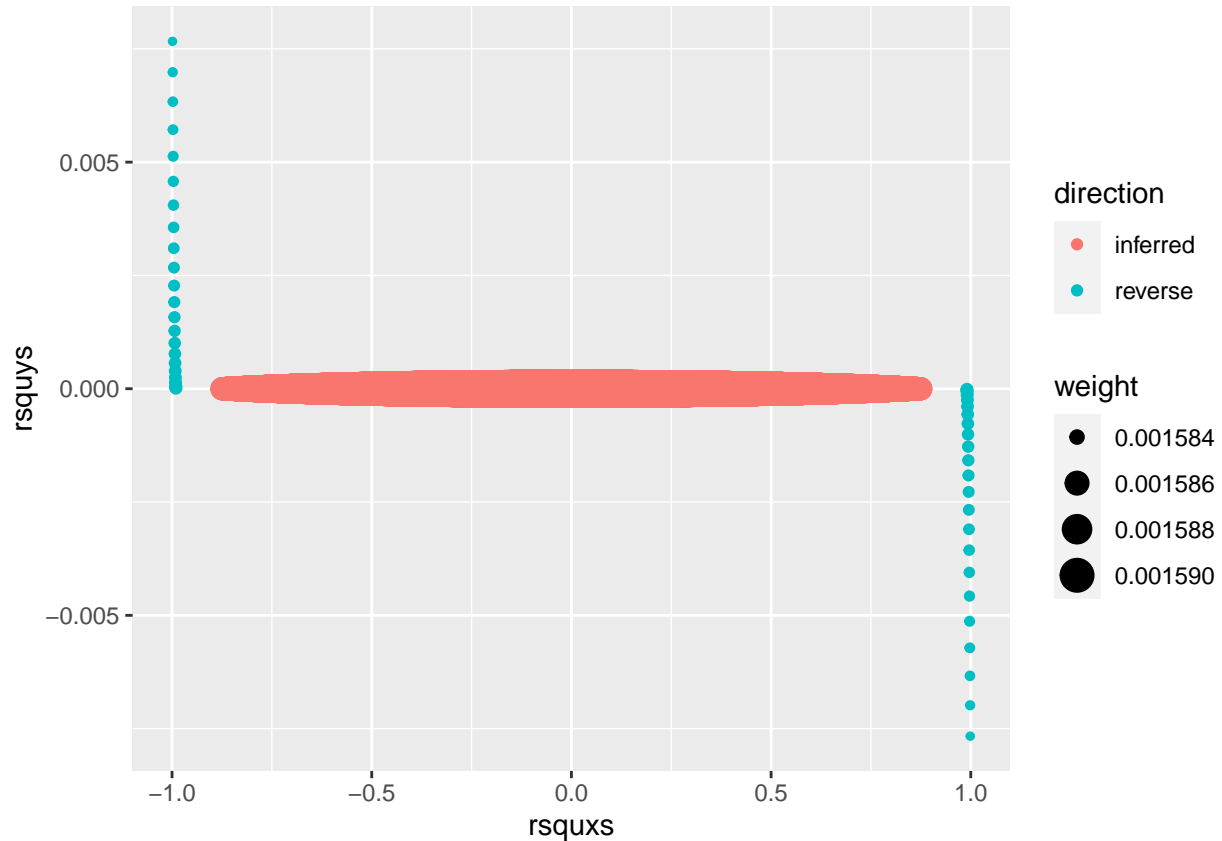
Same example but with more neutral weighting

```
r <- u_sensitivity(bxy=0.1, bxyo=0.1, bgx=0.5, vx=1, vy=1, vg=0.5, scaling=10)
r
```

```
## $result
## # A tibble: 10,044 x 25
## # Groups:   direction [2]
##   bxy  bgx  bux  buy  vg  vu  vex  vey  vx  vy  bxyo  conf
##   <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1  0.1  0.5 -0.935  0  0.5  1  0  0.99  1  1  0.1  0
## 2  0.1  0.5 -0.935  0  0.5  1  0.000350  0.99  1  1  0.1  0
## 3  0.1  0.5 -0.935  0  0.5  1  0.000700  0.99  1  1  0.1  0
## 4  0.1  0.5 -0.935  0  0.5  1  0.00105  0.99  1  1  0.1  0
## 5  0.1  0.5 -0.935  0  0.5  1  0.00140  0.99  1  1  0.1  0
## 6  0.1  0.5 -0.934  0  0.5  1  0.00175  0.99  1  1  0.1  0
## 7  0.1  0.5 -0.934  0  0.5  1  0.00210  0.99  1  1  0.1  0
## 8  0.1  0.5 -0.934  0  0.5  1  0.00245  0.99  1  1  0.1  0
## 9  0.1  0.5 -0.934  0  0.5  1  0.00280  0.99  1  1  0.1  0
```



```
## 10  0.1  0.5 -0.934    0  0.5    1 0.00315  0.99    1    1  0.1    0
## # ... with 10,034 more rows, and 13 more variables: rsqgx <dbl>, rsqgy <dbl>,
## #   rsqux <dbl>, rsquy <dbl>, rsqxy <dbl>, rsqxyo <dbl>, rsqxyos <dbl>,
## #   rsquxs <dbl>, rsquys <dbl>, rsqxys <dbl>, direction <chr>, d <dbl>,
## #   weight <dbl>
##
## $prop
## [1] 0.9867898
##
## $pl
```

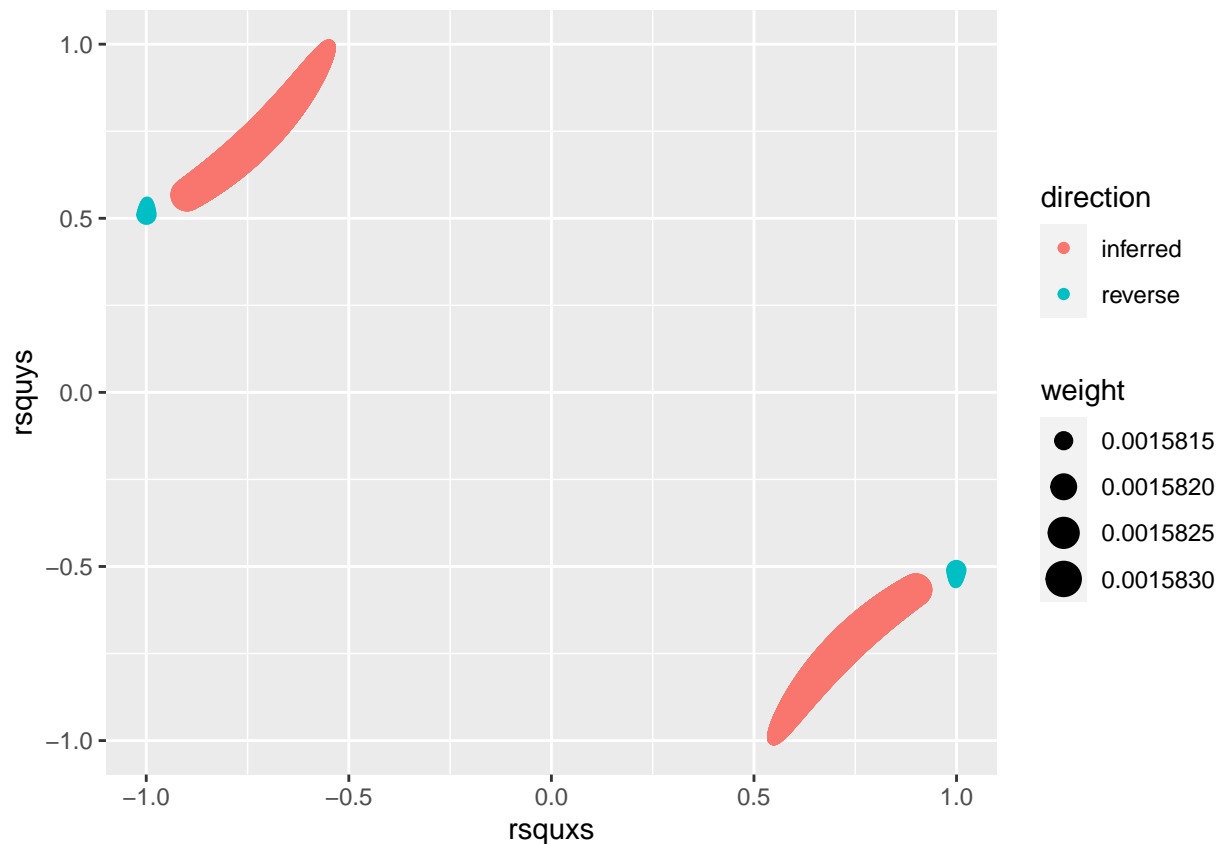


Example 2 - bxy and bxyo are very different

```
r <- u_sensitivity(bxy=0.1, bxyo=-1, vg=1, bgx=1, vx=10, vy=20, scaling=10)
r
```

```
## $result
## # A tibble: 2,212 x 25
## # Groups:   direction [2]
##   bxy  bgx  bux  buy   vg   vu   vex  vey   vx   vy  bxyo  conf
##   <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1  0.1    1   -3    3.67    1    1    0      8.66   10   20   -1   -1.1
## 2  0.1    1  -3.00   3.67    1    1 0.00360   8.65   10   20   -1   -1.1
## 3  0.1    1  -3.00   3.67    1    1 0.00720   8.64   10   20   -1   -1.1
## 4  0.1    1  -3.00   3.67    1    1 0.0108    8.64   10   20   -1   -1.1
## 5  0.1    1  -3.00   3.67    1    1 0.0144    8.63   10   20   -1   -1.1
## 6  0.1    1  -3.00   3.67    1    1 0.0180    8.63   10   20   -1   -1.1
```

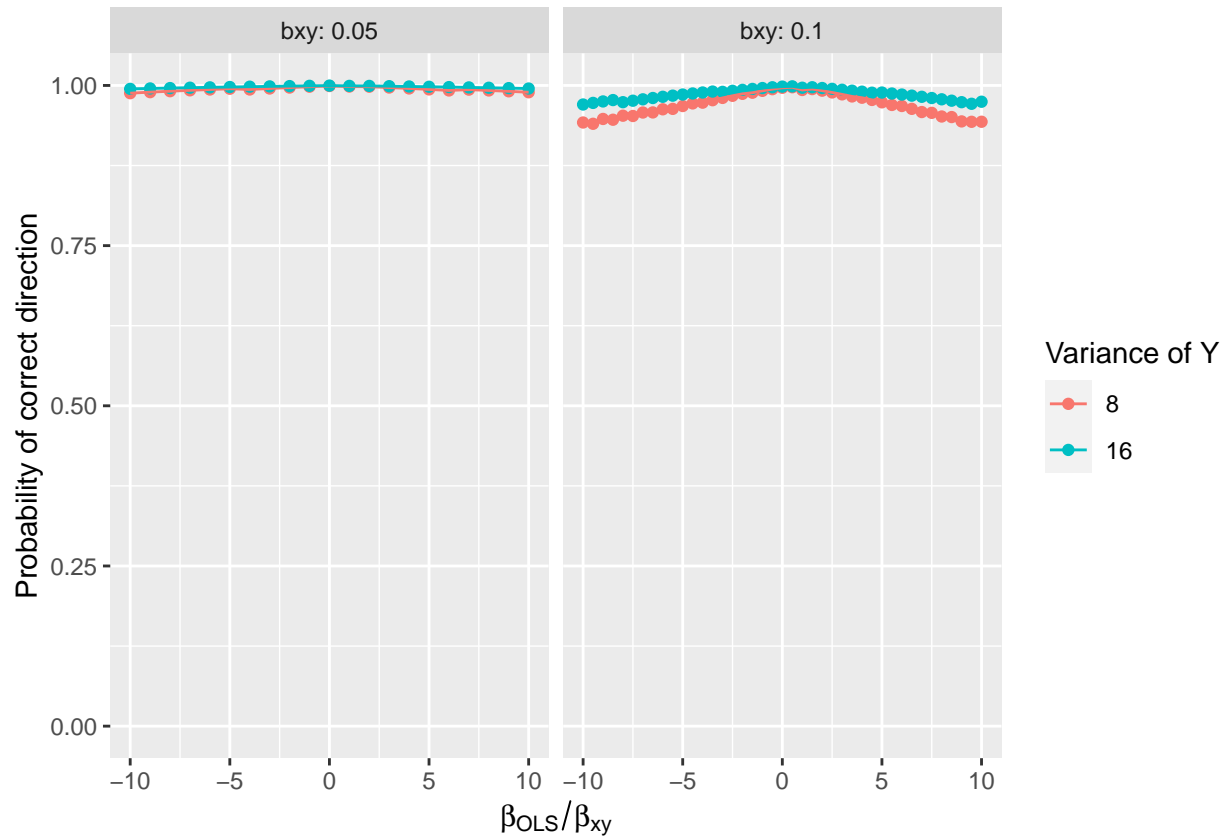
```
## 7  0.1    1 -3.00  3.67    1    1 0.0216  8.62   10   20   -1  -1.1
## 8  0.1    1 -3.00  3.67    1    1 0.0252  8.62   10   20   -1  -1.1
## 9  0.1    1 -3.00  3.67    1    1 0.0288  8.61   10   20   -1  -1.1
## 10 0.1    1 -2.99  3.67    1    1 0.0324  8.61   10   20   -1  -1.1
## # ... with 2,202 more rows, and 13 more variables: rsqgx <dbl>, rsqgy <dbl>,
## #   rsqux <dbl>, rsquy <dbl>, rsqxy <dbl>, rsqxyo <dbl>, rsqxyos <dbl>,
## #   rsquxs <dbl>, rsquys <dbl>, rsqxys <dbl>, direction <chr>, d <dbl>,
## #   weight <dbl>
##
## $prop
## [1] 0.9434405
##
## $pl
```



Explore general performance

```
param <- expand.grid(
  bxy = c(0.1, 0.05),
  bxyo = seq(-1, 1, by=0.05),
  vg=0.5,
  bgx = 0.01,
  vx = 4,
  vy = c(8,16),
  plot=FALSE,
  scaling=10
)
param$prop <- sapply(1:nrow(param), function(i) do.call(u_sensitivity, param[i,])$prop)
```

```
ggplot(param %>% filter(abs(bxyo / bxy) <=10), aes(bxyo / bxy, prop)) +
  geom_point(aes(colour=as.factor(vy))) +
  geom_line(aes(colour=as.factor(vy))) +
  ylim(c(0,1)) +
  facet_grid(. ~ bxy, labeller=label_both) +
  labs(y="Probability of correct direction", colour="Variance of Y", x=TeX(r'($\beta_{OLS}/\beta_{xy})$'))
```



For a particular analysis, the observational association needs to be substantially larger than the causal effect in order for there to be some chance of unmeasured confounding inferring the wrong causal direction.

Empirical analysis

BMI on SBP

```
library(TwoSampleMR)
a <- extract_instruments("ieu-a-2")
b <- extract_outcome_data(a$SNP, "ukb-b-19953") %>% convert_outcome_to_exposure()
```

```
## Extracting data for 79 SNP(s) from 1 GWAS(s)
```

```
c <- extract_outcome_data(a$SNP, "ukb-b-20175")
```

```
## Extracting data for 79 SNP(s) from 1 GWAS(s)
```

```
d <- harmonise_data(b,c)
```

```
## Harmonising Body mass index (BMI) || id:ukb-b-19953 (ukb-b-19953) and Systolic blood pressure, autom
```

```

d <- add_metadata(d)
d <- add_rsqr(d)

bgx <- sqrt(sum(d$rsqr.exposure))
bxy <- mr(d, method="mr_ivw")$b

## Analysing 'ukb-b-19953' on 'ukb-b-20175'

# From https://www.ncbi.nlm.nih.gov/pmc/articles/PMC6324286/
ols <- (0.8 + 1.7)/2 / 20.4 * 3.5
u_sensitivity(bxy=bxy, bxyo=ols, bgx=sqrt(sum(d$rsqr.exposure)), vx=1, vy=1, vg=1, vu=1, scaling=0.5)

## $result
## # A tibble: 8,812 x 25
## # Groups:   direction [2]
##       bxy    bgx    bux    buy    vg    vu    vex    vey    vx    vy    bxyo    conf
##   <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1 0.0942 0.134 -0.991 -0.121 1 1 1.11e-16 0.954 1 1 0.214 0.120
## 2 0.0942 0.134 -0.991 -0.121 1 1 3.93e- 4 0.954 1 1 0.214 0.120
## 3 0.0942 0.134 -0.991 -0.121 1 1 7.86e- 4 0.954 1 1 0.214 0.120
## 4 0.0942 0.134 -0.990 -0.121 1 1 1.18e- 3 0.954 1 1 0.214 0.120
## 5 0.0942 0.134 -0.990 -0.121 1 1 1.57e- 3 0.954 1 1 0.214 0.120
## 6 0.0942 0.134 -0.990 -0.121 1 1 1.96e- 3 0.954 1 1 0.214 0.120
## 7 0.0942 0.134 -0.990 -0.121 1 1 2.36e- 3 0.954 1 1 0.214 0.120
## 8 0.0942 0.134 -0.990 -0.122 1 1 2.75e- 3 0.954 1 1 0.214 0.120
## 9 0.0942 0.134 -0.989 -0.122 1 1 3.14e- 3 0.954 1 1 0.214 0.120
## 10 0.0942 0.134 -0.989 -0.122 1 1 3.53e- 3 0.954 1 1 0.214 0.120
## # ... with 8,802 more rows, and 13 more variables: rsqgx <dbl>, rsqgy <dbl>,
## # rsqux <dbl>, rsquy <dbl>, rsqxy <dbl>, rsqxyo <dbl>, rsqxyos <dbl>,
## # rsquxs <dbl>, rsquys <dbl>, rsqxys <dbl>, direction <chr>, d <dbl>,
## # weight <dbl>
##
## $prop
## [1] 0.9955696
##
## $pl

```

