STAT 758, Spring 2012 Key solution for Home Work 2 Prepared by Tracy Backes ACF, stationarity

2.1 Let $\{X_t\}$ be a sequence of uncorrelated random variables, each with mean 0 and variance σ^2 . For each of the following processes, find its representation in terms of lagged X-values (i.e. $X_{t\pm k}$, k=0,1,2,...), in terms of powers of the backshift operator B, find weights of a corresponding linear filter $W_t = \sum w_j X_{t-j}$, find the mean and variance of W_t

a)
$$W_t = \nabla X_t$$
:
 $W_t = (1 - B)X_t = X_t - X_{t-1}$
 $W_t = \sum_{j=0}^1 w_j X_{t-j}$, where $\{w_k\} = \{1, -1\}$.
b) $W_t = \nabla^2 X_t$:
 $W_t = (1 - B)^2 X_t = (1 - 2B + B^2) X_t$
 $W_t = X_t - 2X_{t-1} + X_{t-2}$
 $W_t = \sum_{j=0}^2 w_j X_{t-j}$, where $\{w_k\} = \{1, -2, 1\}$.
c) $W_t = \nabla_3 X_t$:
 $W_t = (1 - B^3)X_t = (1 - B^3)X_t = X_t - X_{t-3}$
 $W_t = \sum_{j=0}^3 w_j X_{t-j}$, where $\{w_k\} = \{1, 0, 0, -1\}$.
d) $W_t = \nabla^2 \nabla_3 X_t$:
 $W_t = (1 - B)^2 (1 - B^3) X_t$
 $W_t = (1 - 2B + B^2 - B^3 + 2B^4 - B^5) X_t$
 $W_t = X_t - 2X_{t-1} + X_{t-2} - X_{t-3} + 2X_{t-4} - X_{t-5}$
 $W_t = \sum_{j=0}^5 w_j X_{t-j}$, where $\{w_k\} = \{1, -2, 1, -1, 2, -1\}$.

2.2 Find all non-trivial (not equal to identity) linear filters of the form $\alpha + \beta B + \gamma B^2$ (*i.e.* find α , β , and γ) that do not change the variance and mean of uncorrelated sequences.

Solution: Let X_t be an uncorrelated sequence of random variables. Applying a linear filter of the form $\alpha + \beta B + \gamma B^2$ yields a new time series Y_t :

$$Y_t = (\alpha + \beta B + \gamma B^2)X_t = \alpha X_t + \beta X_{t-1} + \gamma X_{t-2}$$

We begin by examining the requirement that the mean of Y_t is unchanged

$$E(X_{t}) = E(Y_{t})$$

$$= E(\alpha X_{t} + \beta X_{t-1} + \gamma X_{t-2})$$

$$= E(\alpha X_{t}) + E(\beta X_{t-1}) + E(\gamma X_{t-2})$$

$$= \alpha E(X_{t}) + \beta E(X_{t-1}) + \gamma E(X_{t-2})$$

$$= (\alpha + \beta + \gamma)E(X_{t})$$

In which case, the coefficients of the linear filter must satisfy the following relationship $\alpha + \beta + \gamma = 1$. Next we consider the variance, which also remains unchanged, such that

$$Var(X_t) = Var(Y_t)$$

$$= Var(\alpha X_t + \beta X_{t-1} + \gamma X_{t-2})$$

$$= Var(\alpha X_t) + Var(\beta X_{t-1}) + Var(\gamma X_{t-2})$$

$$= \alpha^2 Var(X_t) + \beta^2 Var(X_{t-1}) + \gamma^2 Var(X_{t-2})$$

$$= (\alpha^2 + \beta^2 + \gamma^2) Var(X_t)$$

Thus the coefficients of the linear filter must also satisfy the condition $\alpha^2 + \beta^2 + \gamma^2 = 1$. Finally, the solution is the set of all triplets (α, β, γ) , excluding $(\alpha = 1, \beta = \gamma = 0)$, obtained as the intersection of a three-dimensional unit sphere and a plane $\alpha + \beta + \gamma = 1$. For example, the following triplet is a valid solutions: $(\alpha = \gamma = 0, \beta = 1)$.

2.3 Let $\{Z_t\}$ be a sequence of independent normal random variables, each with mean 0 and variance σ^2 , and let a, b, c be constants. For each process below decide whether it is stationary or not. For each stationary process find the mean, autocovariance function, and autocorrelation function:

Solution:

1)
$$X_t = a + bZ_t + cZ_{t+1}$$

(i)
$$E(X_t) = \mu$$
:

$$E(X_t) = E(a + bZ_t + cZ_{t+1})$$

$$= a + bE(Z_t) + cE(Z_{t+1})$$

$$= a \qquad \checkmark$$

(ii)
$$\gamma_{X}(h) = f(h)$$
:

$$\begin{split} \gamma_X(h) &= E\{[X_t - \mu][X_{t+h} - \mu]\} \\ &= E\{[a + bZ_t + cZ_{t+1} - a][a + bZ_{t+h} + cZ_{t+1+h} - a]\} \\ &= E\{b^2Z_tZ_{t+h} + bc[Z_tZ_{t+h+1} + Z_{t+1}Z_{t+h}] + c^2Z_{t+1}Z_{t+h+1}\} \\ &= b^2E(Z_tZ_{t+h}) + bcE(Z_{t+1}Z_{t+h}) + c^2E(Z_{t+1}Z_{t+h+1}) \\ &= \begin{cases} (b^2 + c^2)\sigma^2, & h = 0 \\ bc\sigma^2, & h = 1 \\ 0, & h \geq 2 \end{cases} \end{split}$$

This time series meets the criteria for a weakly stationary process, with mean and autocovariance as shown above and the following ACF:

$$\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = \begin{cases} 1, & h = 0\\ \frac{bc}{b^2 + c^2}, & h = 1\\ 0, & h \ge 2 \end{cases}$$

- 2) $X_t = a + bZ_0$
- (i) $E(X_t) = \mu$:

$$E(X_t) = E(a + bZ_0)$$

$$= a + bE(Z_0)$$

$$= a \qquad \checkmark$$

(ii)
$$\gamma_{X}(h) = f(h)$$
:

$$\gamma_X(h) = E\{[X_t - \mu][X_{t+h} - \mu]\}$$

$$= E\{[a + bZ_0 - a][a + bZ_0 - a]\}$$

$$= E(b^2 Z_0^2)$$

$$= b^2 \sigma^2 \qquad \checkmark$$

This time series meets the criteria for a weakly stationary process, with mean a, covariance $b^2\sigma^2$ and an ACF of 1 at lag h=0.

- $3) X_t = Z_t Z_{t-1}$
- (i) $E(X_t) = \mu$:

$$E(X_t) = E(Z_t Z_{t-1})$$
$$= 0 \qquad \checkmark$$

(ii) $\gamma_{x}(h) = f(h)$:

$$\begin{split} \gamma_X(h) &= E\{[X_t - \mu][X_{t+h} - \mu]\} \\ &= E\{[Z_t Z_{t-1} - 0][Z_{t+h} Z_{t+h-1} - 0]\} \\ &= E(Z_t Z_{t-1} Z_{t+h} Z_{t+h-1}) \\ &= \begin{cases} \sigma^4, & h = 0 \\ 0, & h \neq 0 \end{cases} \checkmark \end{split}$$

This time series meets the criteria for a weakly stationary process, with mean and autocovariance as shown above and the following ACF:

$$\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = \begin{cases} 1, & h = 0\\ 0, & h \neq 0 \end{cases}$$

2.4 Let $\{Z_t\}$ be a stationary process with mean zero and let a and b be constants. If $X_t = a + bt + S_t + Z_t$, where S_t is a seasonal component with period d, show that process $\nabla \nabla_d X_t$ is stationary and express its autocovariance function in terms of that of $\{Z_t\}$.

Solution: We can express the process $Y_t = \nabla \nabla_d X_t$ in terms of the backshift operator B:

$$Y_t = \nabla \nabla_d X_t = (1 - B)(1 - B^d) X_t$$

= $(1 - B - B^d + B^{d+1}) X_t$

We can then rewrite this equation in terms of process Z_t

$$Y_{t} = a + bt + S_{t} + Z_{t}$$

$$- a - b(t - 1) - S_{t-1} - Z_{t-1}$$

$$- a - b(t - d) - S_{t-d} - Z_{t-d}$$

$$+ a + b(t - d - 1) + S_{t-d-1} + Z_{t-d-1}$$

$$= Z_{t} - Z_{t-1} - Z_{t-d} + Z_{t-d-1}.$$

We are now ready to consider the two necessary conditions for a weakly stationary process.

(i)
$$E(Y_t) = \mu$$
:

$$E(Y_t) = E(Z_t - Z_{t-1} - Z_{t-d} + Z_{t-d-1})$$

$$= E(Z_t) - E(Z_{t-1}) - E(Z_{t-d}) + E(Z_{t-d-1})$$

$$= 4E(Z_t)$$

$$= 0 \quad \checkmark$$

(ii) $\gamma_{V}(h) = f(h)$:

$$\begin{split} \gamma_Y(h) &= E\{[Y_t - \mu][Y_{t+h} - \mu]\} \\ &= E\{[Y_t - 0][Y_{t+h} - 0]\} \\ &= E(Y_t Y_{t+h}) \\ &= E\{[Z_t - Z_{t-1} - Z_{t-d} + Z_{t-d-1}][Z_{t+h} - Z_{t+h-1} - Z_{t+h-d} + Z_{t+h-d-1}]\} \\ &= E\{Z_t Z_{t+h} - Z_t Z_{t+h-1} - Z_t Z_{t+h-d} + Z_t Z_{t+h-d-1} \\ &- Z_{t-1} Z_{t+h} + Z_{t-1} Z_{t+h-1} + Z_{t-1} Z_{t+h-d} - Z_{t-1} Z_{t+h-d-1} \\ &- Z_{t-d} Z_{t+h} + Z_{t-d} Z_{t+h-1} + Z_{t-d} Z_{t+h-d} - Z_{t-d} Z_{t+h-d-1} \\ &+ Z_{t-d-1} Z_{t+h} - Z_{t-d-1} Z_{t+h-1} - Z_{t-d-1} Z_{t+h-d} + Z_{t-d-1} Z_{t+h-d-1} \}. \end{split}$$

Assuming $d \geq 1$, this simplifies to

$$\gamma_{Y}(h) = \begin{cases} \gamma_{Y}(-h), & h <= -1 \\ 4\gamma_{Z}(0), & h = 0, d \neq 1 \\ 5\gamma_{Z}(0), & h = 0, d = 1 \\ -4\gamma_{Z}(0), & h = d = 1 \\ -2\gamma_{Z}(0), & h = 1 \text{ or } h = d, d \neq 1 \\ \gamma_{Z}(0), & h = d \pm 1 \\ 0, & \text{o.w.} \end{cases}$$