

Probability, 2

Frank Edwards

Three kinds of probability

The joint probability of two events (A and B) occurring is expressed as

$$P(A \text{ and } B)$$

The marginal probability of an event B is

$$P(B)$$

The conditional probability of event A occurring given that event B occurred is the ratio of the joint probability of A and B divided by the marginal probability of B

$$P(A|B) = \frac{P(A \text{ and } B)}{P(B)}$$

Working with some real data

Voter files

```
data("FLVoters")  
voters<-na.omit(FLVoters)  
head(voters)
```

##	surname	county	VTD	age	gender	race
## 1	PIEDRA	115	66	58	f	white
## 2	LYNCH	115	13	51	m	white
## 4	LATHROP	115	80	54	m	white
## 5	HUMMEL	115	8	77	f	white
## 6	CHRISTISON	115	55	49	m	white
## 7	HOMAN	115	84	77	f	white

Marginal probability

What is the probability that a randomly sampled voter in the population is Black: $P(\text{Black}) = ?$

```
voters %>%  
  count(race, name = "voters") %>%  
  mutate(p = voters/sum(voters))
```

```
##      race voters      p  
## 1   asian    175 0.01920336  
## 2   black   1194 0.131021617  
## 3 hispanic  1192 0.130802151  
## 4   native    29 0.003182267  
## 5   other    310 0.034017338  
## 6   white   6213 0.681773291
```

Is a woman: $P(\text{Woman}) = ?$

```
voters %>%  
  count(gender) %>%  
  mutate(n = n/sum(n))
```

```
##   gender      n  
## 1      f 0.5358279  
## 2      m 0.4641721
```


Joint probability

What is the probability that a voter is a Black woman:

$P(\text{Black and woman}) = ?$

```
voters %>%  
  count(gender, race) %>%  
  mutate(n = n/sum(n)) %>%  
  pivot_wider(names_from = gender, values_from = n)
```

```
## # A tibble: 6 x 3  
##   race      f      m  
##   <chr>    <dbl> <dbl>  
## 1 asian    0.00911 0.0101  
## 2 black    0.0744  0.0566  
## 3 hispanic 0.0731  0.0577  
## 4 native   0.00187 0.00132  
## 5 other    0.0173  0.0167  
## 6 white    0.360   0.322
```

What is the probability that a voter is a woman?

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##   race          f      m
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## 1 asian    0.00911 0.0101
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```

Use the law of total probability:

$$P(A) = P(A \text{ and } B) + P(A \text{ and not } B)$$

put differently, for all categories of B i :

$$P(A) = \sum_{i=1}^n P(A \text{ and } B_i)$$

Conditional probability

If a voter is a man, what is the probability that he is Asian:

$$P(\text{Asian}|\text{man}) = ?$$

```
voters %>%  
  filter(gender=="m") %>%  
  count(race) %>%  
  mutate(n=n/sum(n))
```

```
##      race      n  
## 1   asian 0.021749409  
## 2   black 0.121985816  
## 3 hispanic 0.124349882  
## 4   native 0.002836879  
## 5    other 0.035933806  
## 6    white 0.693144208
```

Conditional probability

Alternatively, we can use the definition of conditional probability as the ratio of the joint probability to the marginal probability:

$$P(\text{Asian}|\text{man}) = \frac{P(\text{Asian and man})}{P(\text{man})}$$

```
voters %>%  
  count(gender, race) %>%  
  mutate(n = n/sum(n))%>%  
  pivot_wider(names_from = gender, values_from = n)
```

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```

Conditioning on more than one variable

What is the probability that a male voter over age 60 is white?

$$P(\text{white} | \text{male and over 60})$$

```
voters %>%  
  mutate(over60=age>60) %>%  
  count(over60, gender, race) %>%  
  mutate(n=n/sum(n)) %>%  
  pivot_wider(names_from = gender, values_from = n)
```

```
## # A tibble: 12 x 4  
##   over60 race      f      m  
##   <lg1> <chr>    <dbl>  <dbl>  
## 1 FALSE asian    0.00691 0.00823  
## 2 FALSE black    0.0555  0.0435  
## 3 FALSE hispanic 0.0549  0.0436  
## 4 FALSE native  0.00121 0.000768  
## 5 FALSE other   0.0124  0.0129  
## 6 FALSE white   0.212   0.198  
## 7 TRUE  asian    0.00219 0.00187  
## 8 TRUE  black    0.0189  0.0132  
## 9 TRUE  hispanic 0.0182  0.0142  
## 10 TRUE native  0.000658 0.000549  
## 11 TRUE other   0.00494 0.00373  
## 12 TRUE white   0.148   0.124
```

Conditioning on more than one variable

In general:

$$P(A \text{ and } B|C) = \frac{P(A \text{ and } B \text{ and } C)}{P(C)}$$

and

$$P(A|B \text{ and } C) = \frac{P(A \text{ and } B \text{ and } C)}{P(B \text{ and } C)}$$

Independence

Two events are independent if knowledge of one event gives us no information about the other event.

$$P(A|B) = P(A) \text{ and } P(B|A) = P(B)$$

$$A \perp B$$

if and only if

$$P(A \text{ and } B) = P(A)P(B)$$

Recall that a Bayesian perspective treats probability as a subjective opinion about how likely an event is. How should we change our beliefs after we make observations about the world?

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Bayes' rule formalizes how we should update our beliefs based on evidence:

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If we have a *prior* belief that event A has $P(A)$ chance of occurring, then we observe some data, represented as event B , we update our beliefs and obtain a *posterior probability* $P(A|B)$.

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

Example: Detecting breast cancer

How good is a mammogram at detecting breast cancer?

What we know: One percent of women have breast cancer. 80 percent of people who have cancer and take a mammogram test positive. 9.6 percent of people who take a mammogram get a positive result when they do not have breast cancer.

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What we know: One percent of women have breast cancer. 80 percent of people who have cancer and take a mammogram test positive. 9.6 percent of people who take a mammogram get a positive result when they do not have breast cancer.

If you take a mammogram and get a positive result, what is the probability that you have breast cancer?

Rewriting as probabilities

One percent of women have breast cancer

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$$P(\text{Test positive}|\text{Cancer}) = 0.8$$

9.6 percent of people who take a mammogram get a positive result when they do not have breast cancer

$$P(\text{Test positive}|\text{No cancer}) = 0.096$$

The prior probability of having cancer is 0.01. How should we update our belief that someone has cancer based on a positive test?

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

Using the law of total probability, we can rewrite the denominator as:

$$P(B) = P(B|A)P(A) + P(B|\text{not } A)P(\text{not } A)$$

Using Bayes' rule

We can apply Bayes' rule for A = Cancer, B = positive test:

$$P(\text{Cancer}|\text{Test positive}) =$$

$$\frac{P(\text{Test positive}|\text{Cancer})P(\text{Cancer})}{P(\text{Test positive})}$$

$$P(\text{Cancer}|\text{Test positive}) = \frac{0.8 \times 0.01}{0.8 \times 0.01 + 0.096 \times 0.99}$$

```
(0.8 * 0.01)/(0.8 * 0.01 + 0.096 * 0.99)
```

```
## [1] 0.07763975
```

The probability that someone has cancer given a prior probability of one percent and a positive test is about 0.078. What would the probability of a true positive be if the test were more sensitive? Say 0.95?

Random Variables

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- The set of all values is exhaustive (the sample space Ω)
- Discrete random variables take a finite number of values (e.g. TRUE, FALSE)
- Continuous random variables are real numbers, and take on an infinite number of values

The simplest random variable: the binary Bernoulli

Any random variable with two values is called a Bernoulli random variable.

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Bernoulli (binary) variables are typically represented as $[0, 1]$ or $[T, F]$.

They can also be two-level character variables, like $[\text{pass}, \text{fail}]$ or $[\text{plaid}, \text{stripes}]$.

A coin flip as a Bernoulli random variable

A coin flip can be defined as a discrete random variable X

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- If the coin lands on heads, $X = 1$
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The probability of a Bernoulli variable is the probability of success, or $X = 1$

$$P(X = 1) = p$$

$$X \sim \text{Bernoulli}(p)$$

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Reads: X is a Bernoulli distributed random variable with probability p

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In this notation, we name the variable X , note that it is randomly distributed \sim , name the distribution it follows *Bernoulli*, and list the parameters for that distribution p .

Let's flip some coins

```
set.seed(12345)  
sample_of_flips<-rbinom(5, 1, 0.5)  
table(sample_of_flips)
```

```
## sample_of_flips  
## 0 1  
## 1 4
```

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```
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## 0 1
## 1 4
```

This is the result of taking 5 draws from a Bernoulli random variable with probability 0.5.

Describing a probability distribution: probability mass

We use a probability mass function to show how likely each value is in a random variable

The probability mass function (PMF) of a variable X is defined as the probability that a variable takes on a particular value x .

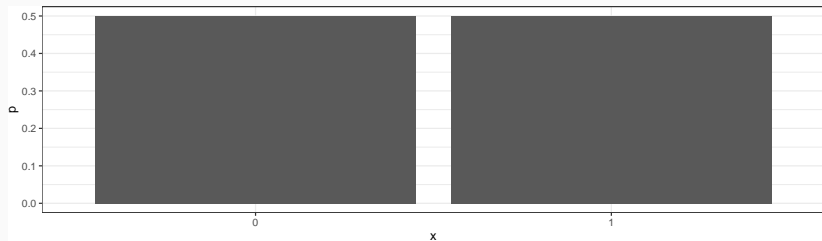
$$PMF(x) = P(X = x)$$

For a Bernoulli variable, $PMF(X = 1) = p$ and $PMF(X = 0) = 1 - p$

The probability mass function for our coin flip

$$PMF(X = 1) = p = 0.5$$

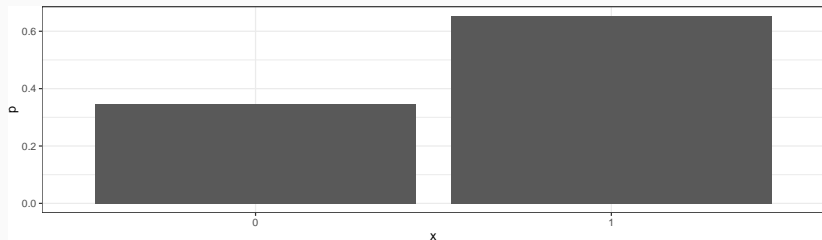
$$PMF(X = 0) = 1 - p = 0.5$$



The probability mass function for passing the bar in NJ ($p=0.653$)

$$PMF(X = 1) = p = 0.653$$

$$PMF(X = 0) = 1 - p = 0.347$$



Describing a probability distribution: cumulative probability

How likely is a variable to take a value less than or equal to a specified value?

We define the cumulative distribution function as the sum of all probabilities up to a value x

$$CDF(X) = P(X \leq x) = \sum_{k \leq x}^x PMF(k)$$

The CDF always ranges from 0 to 1, and never decreases as x increases.

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Uniform random variables

Let's simulate it! 10 draws

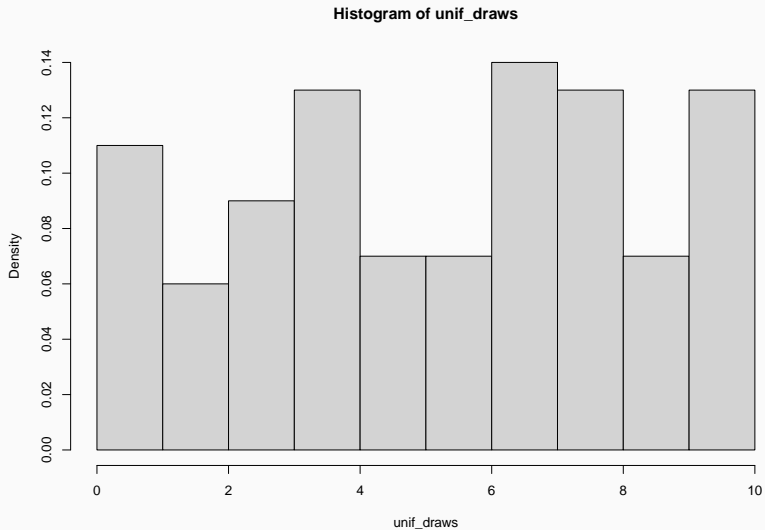
```
unif_draws<-runif(10, min=0, max=10)  
unif_draws
```

```
## [1] 1.66371785 3.25095387 5.09224336 7.27705254 9.89736938 0.34535435  
## [7] 1.52373490 7.35684952 0.01136587 3.91203335
```

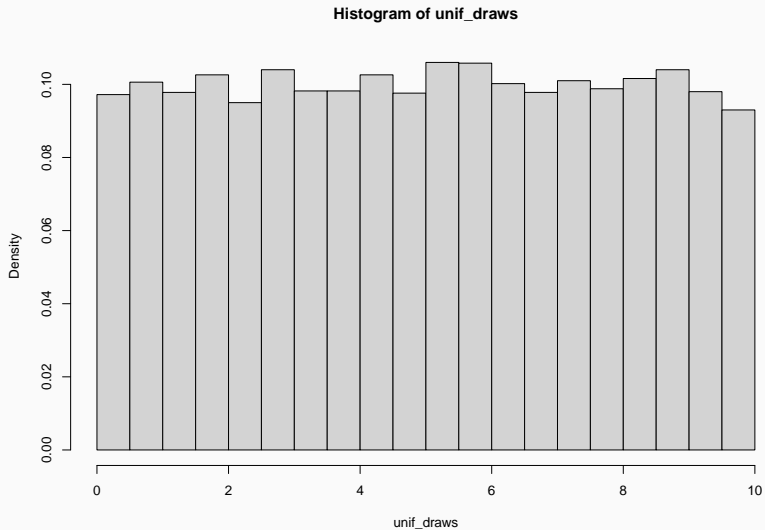
```
hist(unif_draws, freq=F)
```



Uniform random variable: 100 draws



Uniform random variable: 10000 draws



Properties of uniform random variables

For a uniform random variable on the interval $[a, b]$, the probability of drawing any value between a and b is

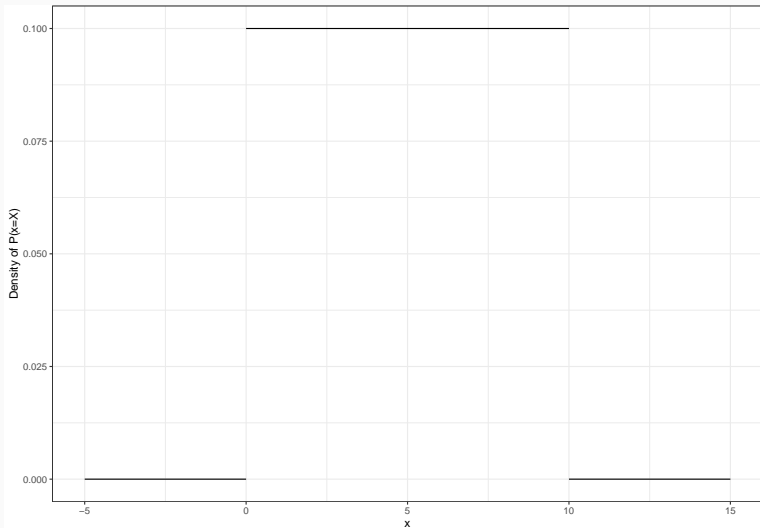
$$\frac{1}{b - a}$$

Formally, the PDF (density, not mass for continuous) and CDF are defined as

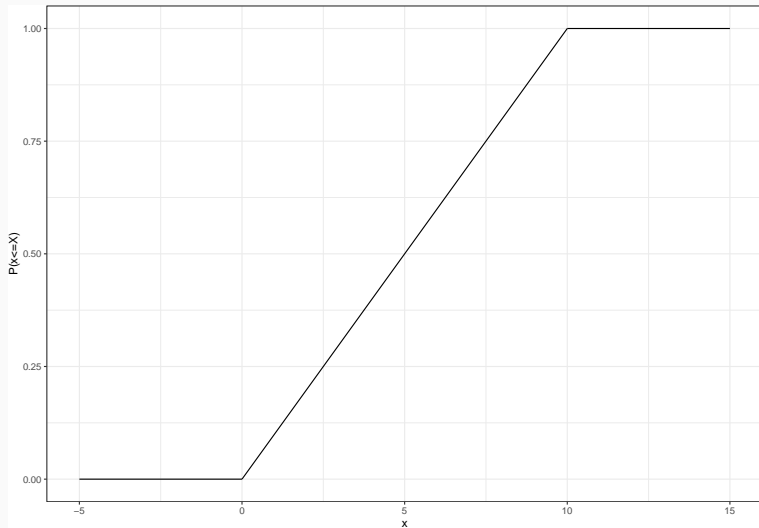
$$\text{PDF: } \begin{cases} \frac{1}{b-a} & \text{for } x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$$

$$\text{CDF: } \begin{cases} 0 & \text{for } x < a \\ \frac{x-a}{b-a} & \text{for } x \in [a, b] \\ 1 & \text{for } x \geq b \end{cases}$$

Probability Density function for $X \sim \text{Uniform}(0, 10)$



Cumulative Distribution Function for $X \sim \text{Uniform}(0, 10)$



A note on CDF for continuous variables

Recall that a CDF for a discrete variable is the sum of all probabilities for values $x \leq X$

We can't sum over each value when X is continuous. Instead, we'll take the integral

$$CDF(x) = P(x \leq X) = \int_{-\infty}^x PDF(x)dx$$

The binomial distribution

When we repeat Bernoulli trials many times, we get a binomial random variable.

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Binomial random variables represent the count of successes in a fixed number of trials of a Bernoulli experiment.

Formally:

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Binomial variables take on integer values between 0 and n

Back to flipping coins

Imagine we flipped a coin 5 times, and then repeated the exercise twice more

```
## [1] 0 1 0 0 0
```

```
## [1] 0 0 0 0 1
```

```
## [1] 1 1 1 1 0
```

Each of these trials is a sample from $X \sim \text{Binomial}(n, p)$ where $n = 5$ and $p = 0.5$

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```
## [1] 0 1 0 0 0
```

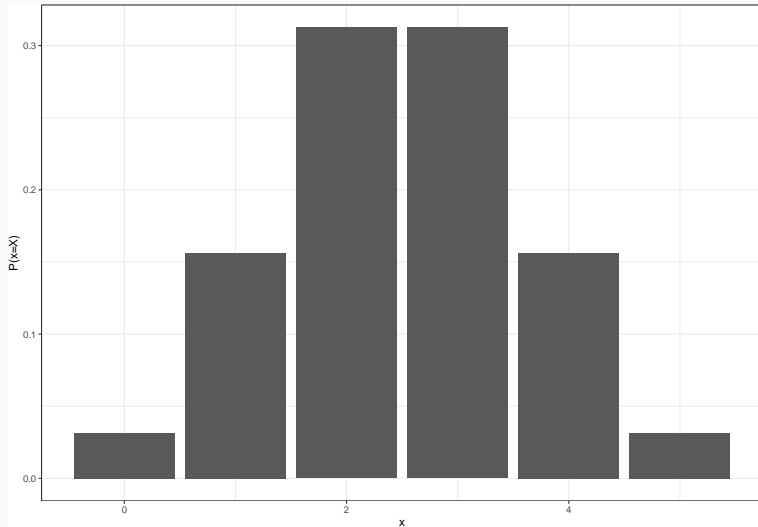
```
## [1] 0 0 0 0 1
```

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```

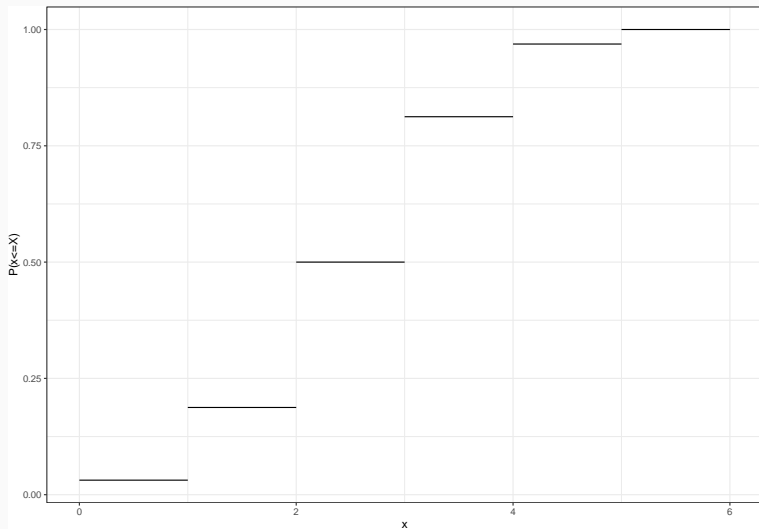
Each of these trials is a sample from $X \sim \text{Binomial}(n, p)$ where $n = 5$ and $p = 0.5$

What is x for each trial?

Probability Mass Function for $X \sim \text{Binomial}(5, 0.5)$



Cumulative Distribution Function for $X \sim \text{Binomial}(5, 0.5)$

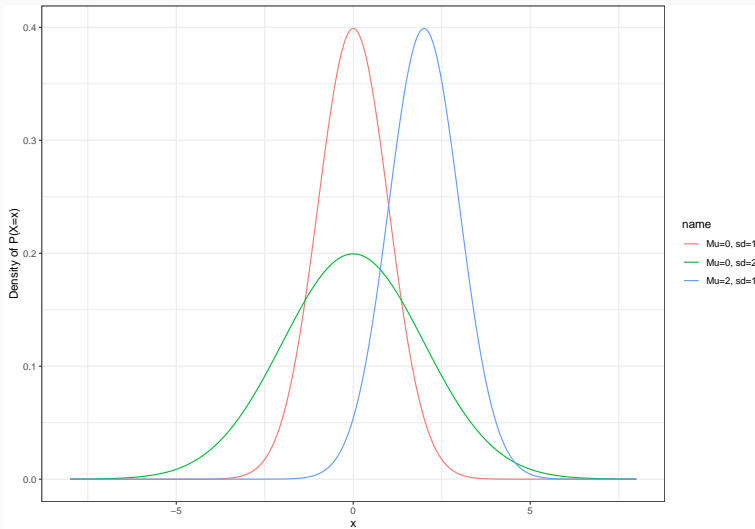


The Normal Distribution

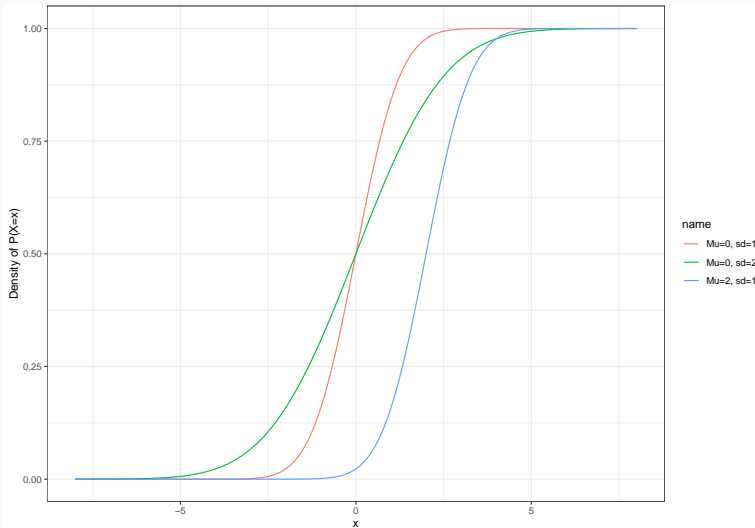
The Normal (Gaussian) distribution is continuous, and takes on values from $[-\infty, \infty]$. It has two parameters, the mean μ and standard deviation σ (or variance σ^2).

- μ determines the location of the distribution
- σ determines the spread of the distribution

The Normal PDF



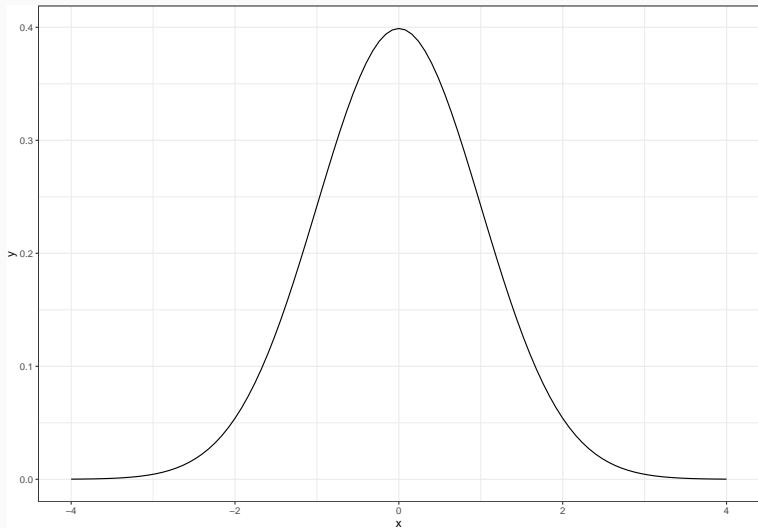
The Normal CDF



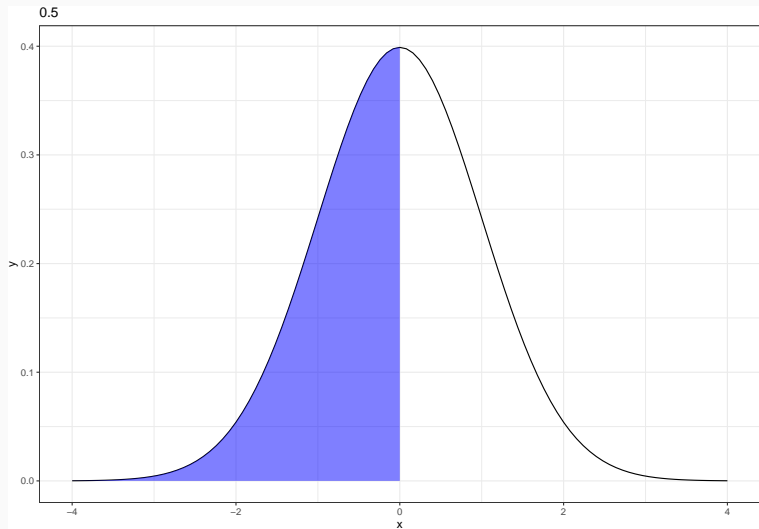
Special features of Normal distributions:

- The sum of many random variables from other distributions are often Normal
- For $X \sim N(\mu, \sigma^2)$, $Z = X + c$ is also Normal: $Z \sim (\mu + c, \sigma^2)$
- $Z = cX$ is distributed $Z \sim N(c\mu, (c\sigma)^2)$
- Z-scores of a Normal random variable are $N(0, 1)$

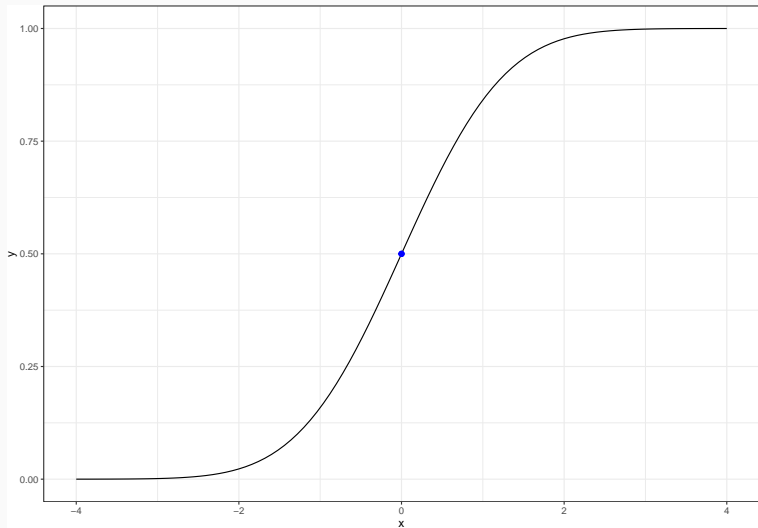
Area under the curve: interpreting the PDF and CDF



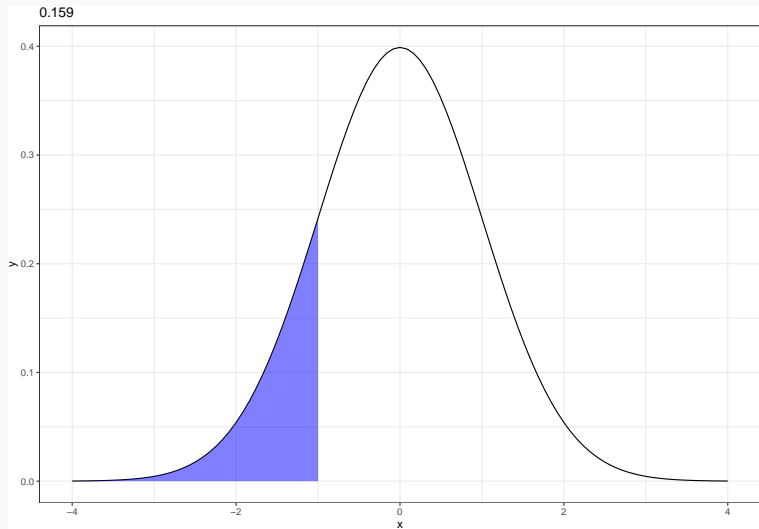
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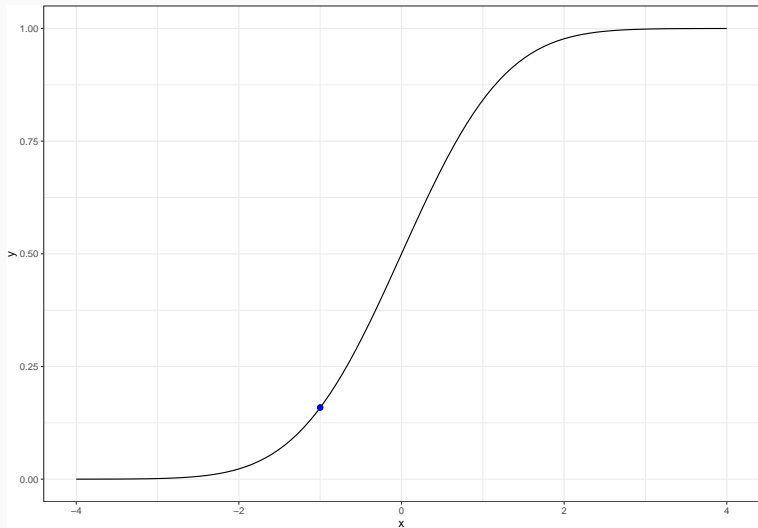
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Area under the curve: interpreting the PDF and CDF



Area under the curve: interpreting the PDF and CDF



Recall that to obtain a z-score, we subtract the mean and divide by the standard deviation:

$$\text{z-score} = \frac{X - \mu}{\sigma}$$

For a Normal variable, z-scores are distributed $z \sim N(0, 1)$

Z-scores and area under the curve

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What does a z-score of 0 indicate?

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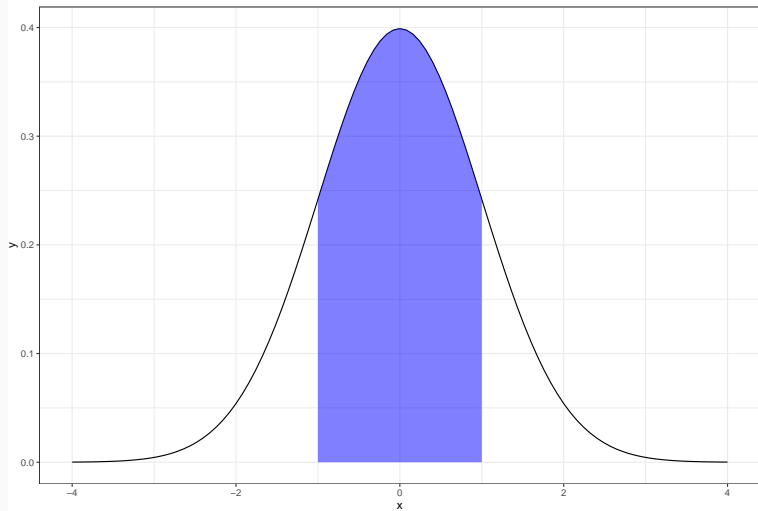
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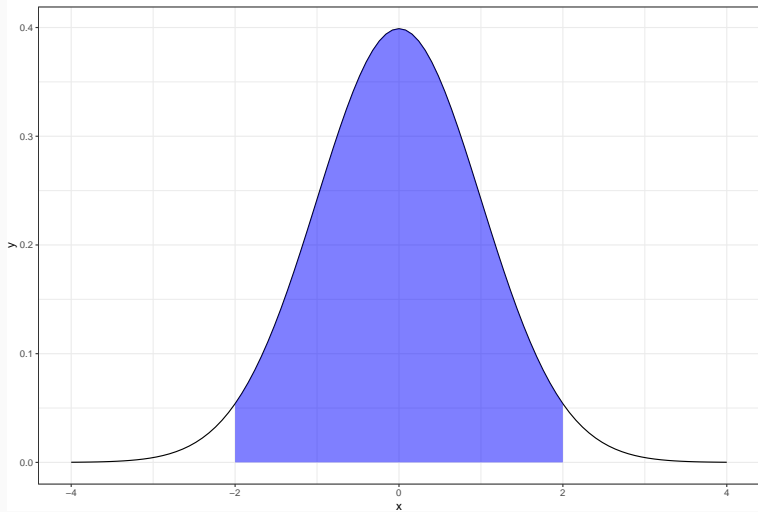
For a Normal variable, z-scores are distributed $z \sim N(0, 1)$

What does a z-score of 0 indicate? -1? 2?

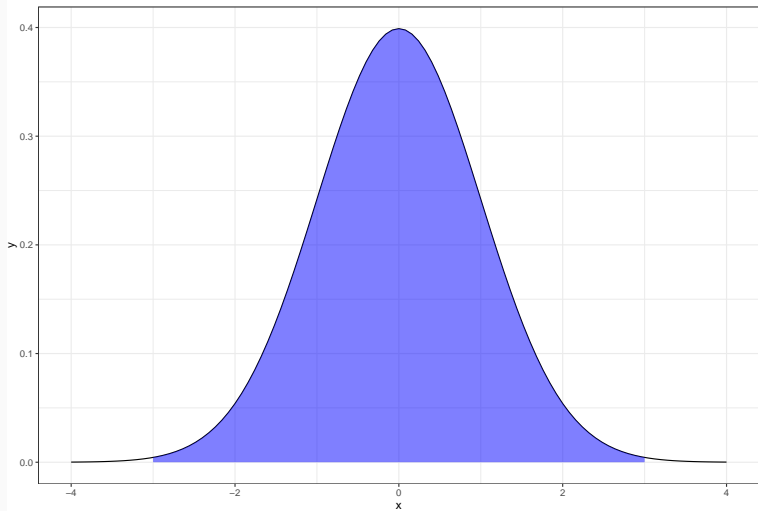
Mean \pm 1 SD = 0.683



Mean \pm 2 SD = 0.954



Mean \pm 3 SD = 0.997



Useful probability distribution functions

```
### Normal(0,1) probability density function
```

```
dnorm(x = 0, mean = 0, sd = 1)
```

```
## [1] 0.3989423
```

```
### Normal(0,1) cumulative distribution function
```

```
pnorm(q = 0, mean = 0, sd = 1)
```

```
## [1] 0.5
```

```
### Random draw from a normal(0,1) distribution
```

```
rnorm(n = 1, mean = 0, sd = 1)
```

```
## [1] 1.231011
```

```
### CDF position for a given probability (quantile)
```

```
qnorm(p = 0.75, mean = 0, sd = 1)
```

```
## [1] 0.6744898
```

```
### You can also use dbinom(), pbinom(), rbinom(), qbinom()
```

The expectation of a random variable

The expectation of a random variable $E(X)$ is the mean of a random variable.

Be careful not to confuse $E(X)$ and \bar{x} .

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For a discrete variable, the expectation is the sum of all values of x weighted by their probability, given by the PDF $f(x)$.

$$E(X) = \sum_x x \times f(x)$$

Because continuous variables take on an infinite number of values, we compute the expectation with an integral

$$\int x \times f(x) dx$$

Variance and standard deviation of a random variable

Recall that for a sample, the standard deviation sd is

$$sd = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$

And the sample variance is sd^2

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For a random variable X , the variance is defined via the expectation instead of sample mean

$$V(X) = E[\{X - E(X)\}^2]$$

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Note the similarities in the two equations