

# An introduction to PCA

## Weekly AI pills

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Fabio Brau.

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SSSA, Emerging Digital Technologies, Pisa.

ISTITUTO  
DI TECNOLOGIE DELLA  
COMUNICAZIONE,  
DELL'INFORMAZIONE  
E DELLA  
PERCEZIONE



Scuola Superiore  
Sant'Anna



- The aim of Principal Component Analysis
- Derivation
  1. A Geometrical idea
  2. A statistical Derivation
  3. Singular Value Decomposition
- PCA from Encoder Decoder NN
- Dummy examples



# Geometrical Introduction

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# Geometrical Introduction

Let  $X \in \mathbb{R}^{N \times n}$  be a dataset of  $N$  **observation** within  $n$  **variables**.

$$X = \begin{bmatrix} x_1^T \\ \vdots \\ x_N^T \end{bmatrix} = \begin{bmatrix} x^{(1)} & | & \dots & | & x^{(n)} \end{bmatrix} \quad (1)$$

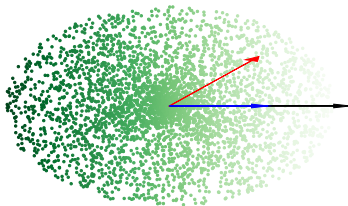
## Notations:

- $x_i \in \mathbb{R}^n$  represents a single **observation**, i.e a **sample** in the feature space.
- $x^{(i)} \in \mathbb{R}^N$  represents the single **variable**, i.e a **column** of the dataset.
- The object  $\mathbb{1}_n \in \mathbb{R}^n$  is the unitary columnar vector of length  $n$   
 $\mathbb{1}_n = [1, \dots, 1]^T$ .
- $X$  is centered if  $X^T \mathbb{1}_N = 0$

# Geometrical Introduction: Finding a principal direction.

1. Scalar product measures the projection of  $x_j$  along the direction  $w$ .
2. We are only interested on module.
3. Summation over samples to get the global projection's contribute.
4. Searching for  $w$  which maximizes projection.
5. Adding constraint to avoid  $w \rightarrow \infty$  solution.

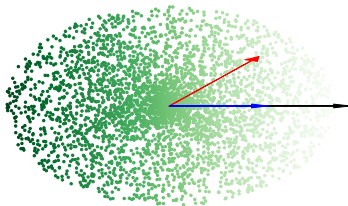
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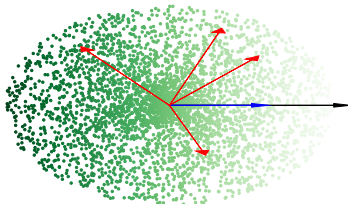
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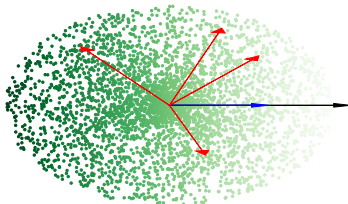
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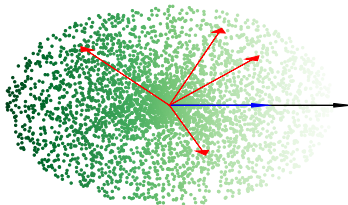




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# Geometrical Introduction: Finding other directions

We search for other orthogonal directions which maximize projections.

$$w_1 \in \operatorname{argmax}_{\|w\|_2=1} \sum_{j=1}^N (w \cdot x)^2$$

$$w_2 \in \operatorname{argmax}_{\|w\|_2=1} \sum_{j=1}^N (w \cdot x)^2 \quad \text{and} \quad w_2 \perp w_1$$

$$\vdots$$

$$w_n \in \operatorname{argmax}_{\|w\|_2=1} \sum_{j=1}^N (w \cdot x)^2 \quad \text{and} \quad w_n \perp \{w_1, \dots, w_{n-1}\}$$

Example

# Geometrical Introduction: Direction Selection

$$V(w) = \sum_j (w \cdot x_j)^2 \quad \text{momentum along } w$$

If  $w_1, w_2, w_3$  orthogonal that maximizes  $V$  in the 3D example, then

1.  $V(w_1) = 3181.20$   $\approx 82.5\%$
2.  $V(w_2) = 646.25$   $\approx 17.0\%$
3.  $V(w_3) = 19.23$   $\approx 0.5\%$

What if we forget the last direction?

## Observation

- $x_j = \alpha_{1j}w_1 + \alpha_{2j}w_2 + \alpha_{3j}w_3$  (where  $\alpha_{ij} = w_i \cdot x_j$ ).
- $\tilde{x}_j = \alpha_{1j}w_1 + \alpha_{2j}w_2$ .

$$\frac{1}{N} \sum_j \|x_j - \tilde{x}_j\|^2 = \frac{V(w_3)}{N} \approx 4.8 \cdot 10^{-3} \quad (\text{MSE})$$



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# Geometrical Introduction: Conclusion

- Given a set of data  $X \in \mathbb{R}^{N \times n}$
- We can find  $w_1, \dots, w_n$  principal (orthonormal) directions that maximize their momentum.
- $V(w_1) > V(w_2) > \dots > V(w_n)$
- Approximating  $X$  with  $\tilde{X}$  by taking only the first  $k$  directions we are getting an error that is  $V(w_{k+1})/N$

What's the catch?

$$\begin{aligned} \max_{w \in \mathbb{R}^n} \quad & \sum_{j=1}^N (w \cdot x_j)^2 \\ \text{s.t.} \quad & w_i \cdot w = 0, \forall i < k \\ & w \cdot w = 1 \end{aligned} \tag{MP}$$





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# Classical Derivation

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# Classical Derivation: An Eigenvalue Problem

$$\max_{\|w\|=1} V(w) = \max_{\|w\|=1} \sum_j (w^T x_j)^2 = \max_{w^T w=1} w^T (X^T X) w \quad (\text{MP})$$

## Lagrange Multipliers Technique

Let consider the Lagrangian Function of MP

$$\mathcal{L}(w, \lambda) = V(w) - \lambda(w^T w - 1), \quad \forall w \in \mathbb{R}^n, \lambda \in \mathbb{R}$$

### Claim

If  $w^*$  is a solution of MP then there exists  $\lambda^*$  such that

$$\nabla \mathcal{L}(w^*, \lambda^*) = 0, \quad \text{i.e.} \quad (X^T X)w^* - \lambda^* w^* = 0 \quad (2)$$

$w$  Principal Direction  $\implies w$  Eigenvector of  $X^T X$



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# Classical Derivation: An Eigenvalue Problem

Why switching to an eigen-pair problem?<sup>1</sup>

$$X^T X \quad + \quad \longrightarrow \quad \begin{aligned} & \bullet w_1, \dots, w_n \quad \text{eigenvectors} \\ & \bullet w_i^T X^T X w_i = V(w_i) \quad \text{eigenvalues} \\ & \bullet V(w_1) > \dots > V(w_n) \geq 0 \end{aligned}$$

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<sup>1</sup>Appendix for further details.

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$X^T X$

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- $w_1, \dots, w_n$  eigenvectors
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# Dimensionality Reduction

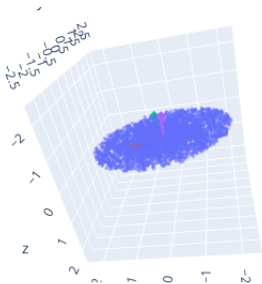
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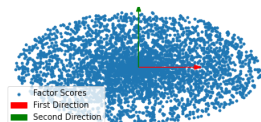
# Dimensionality Reduction

The matrix  $W = [w_1 | \dots | w_n]$  can be used to reduce the dimensionality

$$F = \begin{bmatrix} f^{(1)} & | & \dots & | & f^{(n)} \end{bmatrix} = X W \quad (\text{factors scores})$$

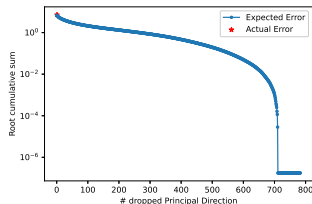
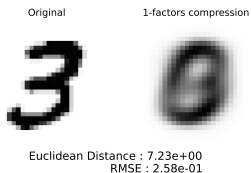


Feature space



Factor scores restricted to the first two principal directions.

# Dimensionality Reduction: A concrete example

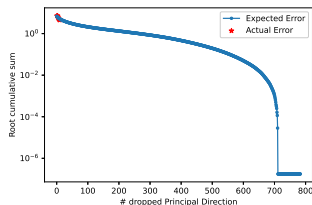
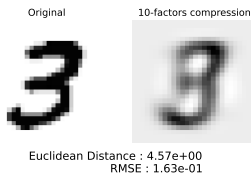


## Compression Error Estimation

1.  $x \in \mathbb{R}^n$  original sample.
2.  $f = W^T x \in \mathbb{R}^n$  coordinates in factor-scores space.
3.  $\tilde{f} = [f_1, \dots, f_{n-k}] \in \mathbb{R}^{n-k}$  dropping last  $k$  coordinates.
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$$\|x - \tilde{x}\| \approx \frac{1}{\sqrt{N}} \|X - X_k\| = \frac{1}{\sqrt{N}} \sqrt{V(n-k+1) + \dots + V(n)} \quad (\text{EB})$$

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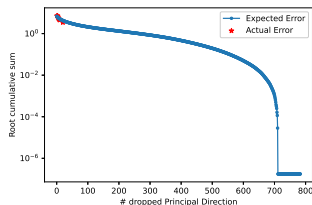
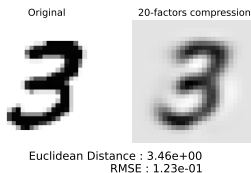


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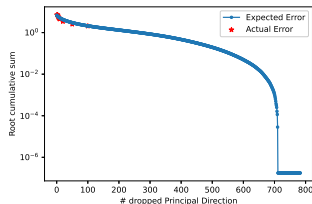
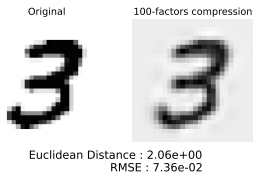


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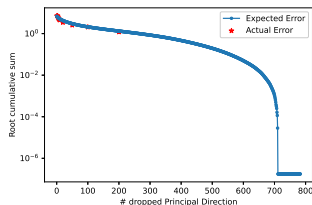
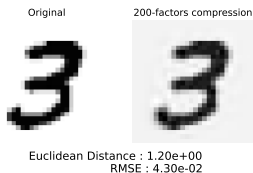


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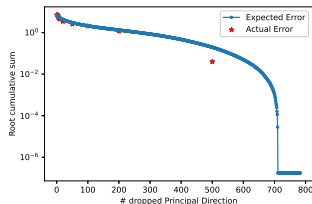
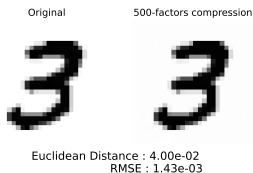


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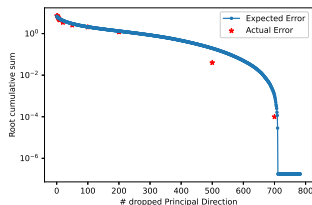
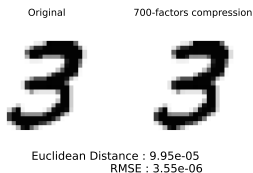
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Where is statistic?

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# Statistical Point of View: Notations

$\mathcal{V}$  random variable,  $V = (v_1, \dots, v_N)$   $N$  observations of the variable.

- Expected Value

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## Compression Error Estimation

$$\|x_j - \tilde{x}_j\| \approx \sqrt{\mathbb{E}[\|x_j - \tilde{x}_j\|^2]} = \sqrt{\text{Var}(w_{k+1} \cdot X) + \dots + \text{Var}(w_n \cdot X)}$$

a.k.a

$w_1, \dots, w_k$  explain  $100 * \left( \frac{\sum_{i=1}^k \text{Var}(w_i \cdot X)}{\sum_i \text{Var}(w_i \cdot X)} \right) \%$  of the variance.

## appendix

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## Eigen-pairs of Symmetric def.positive matrices

A matrix  $A \in M(n)$  is symmetric and def.positive if respectively

$$A^T A = A A^T, \quad v^T A v > 0 \quad \forall v \in \mathbb{R}^n \quad (3)$$

From spectral theorem it's exists an isometry  $V = [v_1 | \cdots | v_n]$  such that

$$V^T A V = D$$

where  $D = \text{diag}(\lambda_1, \dots, \lambda_n)$  is a diagonal matrix.

Because of  $V^T V = Id$  then

$$A V = \begin{bmatrix} A v_1 & | \cdots | & A v_n \end{bmatrix} = V D = \begin{bmatrix} \lambda_1 v_1 & | \cdots | & \lambda_n v_n \end{bmatrix} \quad (4)$$

This shows that **there exists an orthonormal bases of eigenvectors for A**. Because of A is def.positive then

$$\lambda_i = v_i^T A v_i > 0$$

and so A has only positive eigenvalues.



# Approximation Error

For each  $j = 1, \dots, N$  we can write  $x_j = f_{j1}v_1 + \dots + f_{jn}v_n$  where  $f_{ij} = w_i \cdot x_j$ . The approximated samples can be written as  $\tilde{x}_j = f_{j1}v_1 + \dots + f_{j,n-k}v_{n-k}$ . The main idea is to write the **expected value of the square euclidean distance** between the two samples (i.e. original and compressed).

$$\begin{aligned}\|x - \tilde{x}\|^2 &\approx \frac{1}{N} \sum_{j=1}^N \|f_{j,n-k+1}v_{n-k+1} + \dots + f_{jn}v_n\|^2 \\ &= \frac{1}{N} \sum_{j=1}^N f_{j,n-k+1}^2 + \dots + f_{j,n}^2 \\ &= \frac{1}{N} (V(w_1) + \dots + V(w_n))\end{aligned}\tag{5}$$

By taking the root we obtain the approximation in EB. Moreover we can compute also the **Variance of the squared euclidean distance** to increase the accuracy of the error approximation.

