# Untitled

# Fabian Abrego

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### ##Part 1

For this part use the prostate dataset from the faraway package. Use ?prosate to learn about the dataset. The goal of this exercise is to find a model that is useful for explaining the response lpsa.

Fit a total of five models.

One must use all possible predictors. One must use only lcavol as a predictor. The remaining three you must choose. The models you choose must be picked in a way such that for any two of the five models, one is nested inside the other. Argue that one of the five models is the best among them for explaining the response. Use appropriate methods covered and justify your answer.

```
prostate = faraway::prostate
model_all = lm(lpsa ~ ., data = prostate)
model_lcavol = lm(lpsa ~ lcavol, data = prostate)
model_5 = lm(lpsa ~ lcavol + lweight + age + lbph + lcp, data = prostate)
model_4 = lm(lpsa ~ lcavol + lweight + age + lbph, data = prostate)
model_2 = lm(lpsa ~ lcavol + lweight, data = prostate)
summary(model_all)
##
## Call:
## lm(formula = lpsa ~ ., data = prostate)
##
## Residuals:
##
       Min
                1Q Median
                                3Q
                                       Max
## -1.7331 -0.3713 -0.0170 0.4141
                                   1.6381
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                0.669337
                           1.296387
                                      0.516 0.60693
## lcavol
                0.587022
                           0.087920
                                      6.677 2.11e-09 ***
## lweight
                0.454467
                           0.170012
                                      2.673
                                            0.00896 **
## age
               -0.019637
                           0.011173
                                     -1.758
                                             0.08229
## lbph
                           0.058449
                                      1.832
                                             0.07040
                0.107054
## svi
                0.766157
                           0.244309
                                      3.136
                                             0.00233 **
## 1cp
               -0.105474
                           0.091013
                                     -1.159
                                             0.24964
                0.045142
                           0.157465
                                      0.287
                                             0.77503
## gleason
                0.004525
                           0.004421
                                      1.024 0.30886
## pgg45
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.7084 on 88 degrees of freedom
## Multiple R-squared: 0.6548, Adjusted R-squared: 0.6234
## F-statistic: 20.86 on 8 and 88 DF, p-value: < 2.2e-16
```

```
summary(model_lcavol)
##
## Call:
## lm(formula = lpsa ~ lcavol, data = prostate)
## Residuals:
##
       Min
                 1Q
                     Median
                                   3Q
                                           Max
## -1.67625 -0.41648 0.09859 0.50709 1.89673
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.50730
                          0.12194
                                    12.36
                                            <2e-16 ***
## lcavol
               0.71932
                          0.06819
                                    10.55
                                            <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.7875 on 95 degrees of freedom
## Multiple R-squared: 0.5394, Adjusted R-squared: 0.5346
## F-statistic: 111.3 on 1 and 95 DF, p-value: < 2.2e-16
summary(model_5)
##
## Call:
## lm(formula = lpsa ~ lcavol + lweight + age + lbph + lcp, data = prostate)
##
## Residuals:
##
       Min
                 1Q
                     Median
                                   3Q
                                           Max
## -1.51922 -0.39020 0.00317 0.47268 1.75943
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
                          0.87692 0.921 0.35957
## (Intercept) 0.80750
## lcavol
               0.62519
                          0.09017
                                    6.933 5.78e-10 ***
## lweight
               0.46363
                          0.17576
                                    2.638 0.00981 **
## age
              -0.01345
                          0.01134 -1.186 0.23870
## lbph
               0.08493
                          0.06064
                                   1.400 0.16477
## lcp
               0.09044
                          0.07392
                                    1.223 0.22432
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.746 on 91 degrees of freedom
## Multiple R-squared: 0.6041, Adjusted R-squared: 0.5823
## F-statistic: 27.77 on 5 and 91 DF, p-value: < 2.2e-16
summary(model 4)
##
## Call:
## lm(formula = lpsa ~ lcavol + lweight + age + lbph, data = prostate)
## Residuals:
      Min
               1Q Median
                               3Q
                                      Max
## -1.4885 -0.4241 -0.0001 0.4031 1.8073
```

```
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.73074 0.87703 0.833 0.4069
                        0.06754 10.343
## lcavol
              0.69854
                                          <2e-16 ***
## lweight
              0.45770 0.17617 2.598 0.0109 *
              -0.01371
                         0.01137 -1.206 0.2309
## age
                         0.06080 1.382 0.1702
## lbph
              0.08404
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.748 on 92 degrees of freedom
## Multiple R-squared: 0.5976, Adjusted R-squared: 0.5801
## F-statistic: 34.15 on 4 and 92 DF, p-value: < 2.2e-16
summary(model_2)
##
## Call:
## lm(formula = lpsa ~ lcavol + lweight, data = prostate)
## Residuals:
       Min
                 1Q
                     Median
                                  30
## -1.61965 -0.50778 -0.02095 0.52291 1.89885
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.30262  0.56904 -0.532  0.59612
## lcavol
             0.67753
                         0.06626 10.225 < 2e-16 ***
## lweight
               0.51095
                         0.15726
                                  3.249 0.00161 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.7506 on 94 degrees of freedom
## Multiple R-squared: 0.5859, Adjusted R-squared: 0.5771
## F-statistic: 66.51 on 2 and 94 DF, p-value: < 2.2e-16
#RMSE
RMSE = function(test)
{sqrt((summary(test)$sigma^2)*test$df.residual/length(test$fitted.values))}
RMSE(model_all)
## [1] 0.674751
RMSE(model_lcavol)
## [1] 0.7793386
RMSE(model_5)
## [1] 0.7225684
RMSE(model 4)
## [1] 0.7284869
```

```
RMSE(model_2)
## [1] 0.7389478
#Testing Residuals
#Normality Assumption
shapiro.test(resid(model_all))
##
##
   Shapiro-Wilk normality test
##
## data: resid(model_all)
## W = 0.99113, p-value = 0.7721
shapiro.test(resid(model_lcavol))
##
##
   Shapiro-Wilk normality test
##
## data: resid(model_lcavol)
## W = 0.97985, p-value = 0.1419
shapiro.test(resid(model_5))
##
    Shapiro-Wilk normality test
##
## data: resid(model_5)
## W = 0.98824, p-value = 0.5486
shapiro.test(resid(model_4))
##
## Shapiro-Wilk normality test
## data: resid(model_4)
## W = 0.98684, p-value = 0.4491
shapiro.test(resid(model_2))
##
##
   Shapiro-Wilk normality test
## data: resid(model_2)
## W = 0.99043, p-value = 0.718
#Constant Variance Assumption
library(lmtest)
## Loading required package: zoo
##
## Attaching package: 'zoo'
## The following objects are masked from 'package:base':
##
##
       as.Date, as.Date.numeric
bptest(model_all)
```

```
##
## studentized Breusch-Pagan test
##
## data: model_all
## BP = 10.08, df = 8, p-value = 0.2594
bptest(model_lcavol)
##
## studentized Breusch-Pagan test
##
## data: model_lcavol
## BP = 0.12623, df = 1, p-value = 0.7224
bptest(model_5)
##
## studentized Breusch-Pagan test
##
## data: model_5
## BP = 4.2837, df = 5, p-value = 0.5093
bptest(model_4)
##
## studentized Breusch-Pagan test
##
## data: model_4
## BP = 2.0293, df = 4, p-value = 0.7304
bptest(model_2)
##
## studentized Breusch-Pagan test
##
## data: model_2
## BP = 3.3046, df = 2, p-value = 0.1916
All models meet assumptions.
#AIC similar to Mallows CP in comparing models - the smaller the better
AIC(model_all)
## [1] 218.9522
AIC(model_lcavol)
## [1] 232.908
AIC(model_5)
## [1] 226.2351
AIC(model_4)
## [1] 225.8177
AIC(model_2)
## [1] 224.5837
```

```
#BIC the smaller the better
BIC(model_all)
## [1] 244.6993
BIC(model_lcavol)
## [1] 240.6321
BIC(model_5)
## [1] 244.2581
BIC(model_4)
## [1] 241.2659
BIC(model_2)
## [1] 234.8825
#Anova Testing
#Model All - at 5 percent significance- reject the null for all- suggesting linear relationship between
anova(model_2,model_all)
## Analysis of Variance Table
##
## Model 1: lpsa ~ lcavol + lweight
## Model 2: lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
##
      pgg45
              RSS Df Sum of Sq
##
    Res.Df
                                   F Pr(>F)
## 1
       94 52.966
## 2
        88 44.163 6
                      8.8032 2.9236 0.01199 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
anova(model_4,model_all)
## Analysis of Variance Table
##
## Model 1: lpsa ~ lcavol + lweight + age + lbph
## Model 2: lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
      pgg45
##
   Res.Df
              RSS Df Sum of Sq
                                   F Pr(>F)
       92 51.477
## 1
        88 44.163 4
                      7.3142 3.6436 0.00855 **
## 2
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
anova(model_5,model_all)
## Analysis of Variance Table
## Model 1: lpsa ~ lcavol + lweight + age + lbph + lcp
## Model 2: lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
##
      pgg45
##
    Res.Df
              RSS Df Sum of Sq
                                   F Pr(>F)
## 1
        91 50.644
## 2
        88 44.163 3
                      6.4812 4.3048 0.00699 **
```

```
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
anova(model_lcavol,model_all)
## Analysis of Variance Table
## Model 1: lpsa ~ lcavol
## Model 2: lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
      pgg45
              RSS Df Sum of Sq
##
   Res.Df
                                  F
                                        Pr(>F)
        95 58.915
        88 44.163 7 14.752 4.1992 0.0004916 ***
## 2
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
#Model 5
anova(model_4, model_5) # no significant difference between models - the smaller the better
## Analysis of Variance Table
## Model 1: lpsa ~ lcavol + lweight + age + lbph
## Model 2: lpsa ~ lcavol + lweight + age + lbph + lcp
   Res.Df RSS Df Sum of Sq
                                  F Pr(>F)
## 1
        92 51.477
## 2
        91 50.644 1
                       0.83304 1.4968 0.2243
anova(model_2,model_5) # no significant difference between models - the smaller the better
## Analysis of Variance Table
## Model 1: lpsa ~ lcavol + lweight
## Model 2: lpsa ~ lcavol + lweight + age + lbph + lcp
   Res.Df
            RSS Df Sum of Sq
                                  F Pr(>F)
## 1
       94 52.966
        91 50.644 3
                       2.3221 1.3908 0.2507
anova(model_lcavol,model_5) # sigificant - model 5 preferred to lcavol
## Analysis of Variance Table
##
## Model 1: lpsa ~ lcavol
## Model 2: lpsa ~ lcavol + lweight + age + lbph + lcp
## Res.Df RSS Df Sum of Sq
                                  F Pr(>F)
## 1
        95 58.915
        91 50.644 4 8.2706 3.7152 0.007575 **
## 2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#Model 4
anova(model_lcavol,model_4) # significant - Model 4 preferred to Lcavol
## Analysis of Variance Table
## Model 1: lpsa ~ lcavol
## Model 2: lpsa ~ lcavol + lweight + age + lbph
```

```
Res.Df
                RSS Df Sum of Sq
                                            Pr(>F)
##
## 1
         95 58.915
                          7.4375 4.4308 0.005902 **
## 2
         92 51.477 3
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
anova(model_2,model_4) # not significant - the smaller the better
## Analysis of Variance Table
##
## Model 1: lpsa ~ lcavol + lweight
## Model 2: lpsa ~ lcavol + lweight + age + lbph
     Res.Df
                RSS Df Sum of Sq
## 1
         94 52.966
## 2
         92 51.477 2
                            1.489 1.3306 0.2694
#Model 2
anova(model_lcavol,model_2) #significant- Model 2 preferred against lcavol.
## Analysis of Variance Table
##
## Model 1: lpsa ~ lcavol
## Model 2: lpsa ~ lcavol + lweight
                RSS Df Sum of Sq
##
     Res.Df
                                       F
                                            Pr(>F)
## 1
         95 58.915
         94 52.966
## 2
                    1
                          5.9485 10.557 0.001606 **
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Model All:
RSE = .7084 RMSE = .674751 R^2 = .6548 Adjusted R^2 = .6234 Beta (fail to reject): 5 predictors out of 8
AIC = 218.9522 BIC = 244.6993
Model lcavol:
RSE = .7875 \text{ RMSE} = .7793386 \text{ R}^2 = .5394 \text{ Adjusted R}^2 = .5346 \text{ Beta fails} = 0 \text{ out of 1 AIC} = 232.9522
BIC = 240.6321
Model 5:
RSE = .746
RMSE = .7225684 R^2 = .6041 Adjusted R^2 = .5823 Beta fails = 3 out of 5 AIC = 226.2351 BIC = 244.2581
Model 4:
RSE = .748
RMSE = .7284869 R^2 = .5976
Adjusted R^2 = .5801 beta fails = 2 out of 4 AIC = 225.8177 BIC = 241.2659
Model 2:
RSE = .7506
RMSE = .7389478 R^2 = .5859 Adjusted R^2 = .5771 beta fails = 0 out of 2 AIC = 224.5837 BIC = 234.8825
#Explaining Test Results
```

On the basis of selecting a model to *explain*, it is important to keep such models small as it is easier to derive and explain relationships between the predictor variables and the explanatory variable. On that basis the smaller models considered are favored over the larger ones. Firstly, all models meet the constant variance and normality assumption of errors which is an important factor in considering models for explaining. The next important aspect to consider is the significance of regression which is measured through the F test of all predictors in a model, the T test of the individual predictors in a model, and anova tests between models at

a significance level of .05. In terms of the F tests, all models would fail to accept the null suggesting a linear relationship between the model predictors and the explanatory variable. In terms of the individual T tests, models model\_2 and model\_lcavol were the only models with betas that would fail to reject suggesting individual linear relationships between the predictors and the explanatory variable. This is an important aspect in explaining an output as we can prove the existence of a linear relationship between variables. In terms of anova testing, the model model\_all was preferred to all other models meaning there may be a linear relationship with additional variables in the model that were not captured in the other models. Using anova on the latter models, we see no significant difference between such models except when they are compared to model\_lcavol - all models are preferred to such. The last measures used were AIC and BIC, which are typically used in model selection. In terms of AIC the smallest value was attributed to model\_all making it the best selection and the second smallest was with attributed to model\_all. Typically the AIC and BIC agree in picking the best model but since they are on two separate spectrums, we can conclude that model\_2 is the best selection based on these measures.

### #Conclusion

The best model for *explaining* is model\_2. This is due to it being a smaller model (the second smallest), the proven linear relationship between each predictor and the explanatory variable (T-test/F-test), the outcome of the AIC (second best) and BIC (best) measures. Lastly, in terms of the anova testing we saw this model as being not significantly different from model\_5 and model\_4 suggesting the model\_2 is the better option. Stacked up against model\_lcavol we find that the additional beta in model\_2 is significant, making model\_2 the preferred choice. When using anova to compare model\_all to all other models, we see it is a better option but given the model's performance in other areas of testing and the preference of smaller models, we dismiss the model\_all.

In terms of prediction, RSE, RMSE, R^2 and Adjusted R^2 are better measures to consider but because we are trying to *explain*, those aspects were not as heavily considered as much as the results of testing for linearity and thus proving a relationship that can be explained.

# $\#\#\operatorname{Part} 2$

```
boston = MASS::Boston
library(MASS)
set.seed(42)
train_index = sample(1:nrow(Boston), 400)
train_data = boston[train_index,]
test_data = boston[-train_index,]
Model_ALL = lm(medv ~ ., data = train_data)
Model_crim = lm(medv ~ crim, data = train_data)
Model 6 = lm(medv ~ crim + indus + nox + rm + age + dis, data = train data)
Model 5 = lm(medv ~ crim + indus + nox + rm + age, data = train data)
Model_3 = lm(medv ~ crim + nox + dis, data = train_data)
Y = test_data[,14]
#Model All
RSE_ALL = summary(Model_ALL)$sigma
Train_RMSE_ALL = sqrt((RSE_ALL^2)*Model_ALL$df.residual/length(Model_ALL$fitted.values))
beta_ALL = as.vector(Model_ALL$coefficients)
X_ALL = cbind(1,test_data[,-14])
Y_hat_ALL = as.matrix(X_ALL) %*% beta_ALL
SSE_ALL = sum((Y - Y_hat_ALL)^2)
Test_RMSE_ALL = sqrt(SSE_ALL/length(Y))
```

```
#Train RMSE = 4.675465 Test RMSE = 4.767746
#Model crim
RSE_crim = summary(Model_crim)$sigma
Train_RMSE_crim = sqrt((RSE_crim^2)*Model_crim$df.residual/length(Model_crim$fitted.values))
beta_crim = as.vector(Model_crim$coefficients)
X_crim = cbind(1,test_data[,1])
Y_hat_crim = as.matrix(X_crim) %*% beta_crim
SSE_crim = sum((Y - Y_hat_crim)^2)
Test_RMSE_crim = sqrt(SSE_crim/length(Y))
#Train RMSE = 8.238496  Test RMSE = 9.318085
#Model 6
RSE_6 = summary(Model_6)$sigma
Train_RMSE_6 = sqrt((RSE_6^2)*Model_6$df.residual/length(Model_6$fitted.values))
beta_6 = as.vector(Model_6$coefficients)
X_6 = cbind(1,test_data[,c("crim","indus","nox","rm","age","dis")])
Y_hat_6 = as.matrix(X_6) %*% beta_6
SSE_6 = sum((Y - Y_hat_6)^2)
Test_RMSE_6 = sqrt(SSE_6/length(Y))
#Train RMSE = 5.758958 Test RMSE = 5.95507
#Model 5
RSE_5 = summary(Model_5)$sigma
Train_RMSE_5 = sqrt((RSE_5^2)*Model_5$df.residual/length(Model_5$fitted.values))
beta_5 = as.vector(Model_5$coefficients)
X_5 = cbind(1,test_data[,c("crim","indus","nox","rm","age")])
Y_{hat_5} = as.matrix(X_5) %*% beta_5
SSE_5 = sum((Y - Y_hat_5)^2)
Test_RMSE_5 = sqrt(SSE_5/length(Y))
#Train RMSE = 5.995325   Test RMSE = 6.148281
#Model 3
RSE_3 = summary(Model_3)$sigma
Train_RMSE_3 = sqrt((RSE_3^2)*Model_3$df.residual/length(Model_3$fitted.values))
beta_3 = as.vector(Model_3$coefficients)
X_3 = cbind(1,test_data[,c("crim","nox","dis")])
Y_hat_3 = as.matrix(X_3) %*% beta_3
```

```
SSE_3 = sum((Y - Y_hat_3)^2)
Test_RMSE_3 = sqrt(SSE_3/length(Y))
#Train RMSE = 7.72839 Test RMSE = 8.643548
ALL = c(Test_RMSE_ALL, Train_RMSE_ALL) # Variance = .004258
crim = c(Test_RMSE_crim, Train_RMSE_crim) #Variance = .58276
Six = c(Test RMSE 6, Train RMSE 6) #Variance = .01923
Five = c(Test_RMSE_5, Train_RMSE_5) #Variance = .01170
Three = c(Test_RMSE_3, Train_RMSE_3) #Variance = .41876
summary(Model_ALL)
##
## Call:
## lm(formula = medv ~ ., data = train_data)
## Residuals:
      Min
               1Q Median
                               3Q
                                      Max
## -14.3126 -2.7134 -0.5522
                           1.5431 25.5431
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 40.211363 5.823305 6.905 2.07e-11 ***
## crim
             ## zn
              0.037754 0.016166
                                 2.335 0.020038 *
## indus
              ## chas
              1.918167 0.999327 1.919 0.055663 .
## nox
            -17.987178 4.304668 -4.179 3.63e-05 ***
## rm
              3.478935
                        0.457299
                                 7.608 2.16e-13 ***
                        0.014798 -0.209 0.834880
## age
             -0.003087
## dis
             0.074539 4.167 3.81e-05 ***
## rad
              0.310637
             -0.011081
                       0.004234 -2.617 0.009212 **
## tax
## ptratio
             ## black
             0.007692 0.003214 2.393 0.017194 *
## lstat
             -0.533910
                       0.055318 -9.652 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 4.759 on 386 degrees of freedom
## Multiple R-squared: 0.7262, Adjusted R-squared: 0.7169
## F-statistic: 78.73 on 13 and 386 DF, p-value: < 2.2e-16
summary(Model_crim)
##
## Call:
## lm(formula = medv ~ crim, data = train_data)
## Residuals:
      Min
              1Q Median
                            3Q
                                  Max
## -16.734 -5.147 -1.788
                         2.329
                               29.619
## Coefficients:
```

```
Estimate Std. Error t value Pr(>|t|)
                      0.44704 53.243 < 2e-16 ***
## (Intercept) 23.80157
             -0.37045
## crim
                        0.04425 -8.372 9.75e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 8.259 on 398 degrees of freedom
## Multiple R-squared: 0.1497, Adjusted R-squared: 0.1476
## F-statistic: 70.09 on 1 and 398 DF, p-value: 9.753e-16
summary(Model 6)
##
## Call:
## lm(formula = medv ~ crim + indus + nox + rm + age + dis, data = train_data)
## Residuals:
      Min
              1Q Median
                             3Q
                                   Max
## -16.219 -3.166 -0.600
                          2.097 37.638
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -0.93105
                        4.52606 -0.206 0.837125
## crim
             -0.19154
                         0.03456 -5.543 5.47e-08 ***
## indus
              4.66224 -2.612 0.009334 **
## nox
             -12.17993
## rm
              ## age
              -1.45760 0.25404 -5.738 1.92e-08 ***
## dis
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 5.81 on 393 degrees of freedom
## Multiple R-squared: 0.5845, Adjusted R-squared: 0.5782
## F-statistic: 92.15 on 6 and 393 DF, p-value: < 2.2e-16
summary(Model_5)
##
## lm(formula = medv ~ crim + indus + nox + rm + age, data = train_data)
##
## Residuals:
      Min
              1Q Median
                             3Q
                                   Max
## -17.127 -3.120 -0.847
                          2.174 39.005
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -17.64406 3.60183 -4.899 1.41e-06 ***
                         0.03589 -5.067 6.23e-07 ***
## crim
              -0.18183
                         0.07177 -1.646
## indus
              -0.11816
                                          0.101
## nox
              -3.78538
                         4.60257 -0.822
                                          0.411
              7.28048
                         0.46594 15.625 < 2e-16 ***
## rm
## age
              -0.02127 0.01586 -1.341
                                          0.181
## ---
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 6.041 on 394 degrees of freedom
## Multiple R-squared: 0.5497, Adjusted R-squared: 0.544
## F-statistic: 96.2 on 5 and 394 DF, p-value: < 2.2e-16
summary(Model_3)
##
## Call:
## lm(formula = medv ~ crim + nox + dis, data = train_data)
## Residuals:
##
      Min
               1Q Median
                               3Q
                                      Max
## -16.488 -4.947 -2.063
                            2.625 29.138
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 48.09033
                           3.85312 12.481 < 2e-16 ***
               -0.26589
                           0.04577 -5.809 1.29e-08 ***
## crim
## nox
              -37.35806
                           5.35091 -6.982 1.24e-11 ***
                           0.29019 -3.537 0.000452 ***
               -1.02654
## dis
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 7.767 on 396 degrees of freedom
## Multiple R-squared: 0.2518, Adjusted R-squared: 0.2461
## F-statistic: 44.42 on 3 and 396 DF, p-value: < 2.2e-16
#LOOCV RMSE
calc_loocv_rmse = function(model) {sqrt(mean((resid(model) / (1 - hatvalues(model))) ^ 2))}
calc_loocv_rmse(Model_ALL)
## [1] 4.908037
calc_loocv_rmse(Model_crim)
## [1] 8.310566
calc_loocv_rmse(Model_6)
## [1] 5.875081
calc_loocv_rmse(Model_5)
## [1] 6.107382
calc_loocv_rmse(Model_3)
## [1] 7.805229
#Train RMSE
Train_RMSE_ALL
## [1] 4.675465
Train_RMSE_crim
## [1] 8.238496
```

```
{\tt Train\_RMSE\_6}
## [1] 5.758958
{\tt Train\_RMSE\_5}
## [1] 5.995325
{\tt Train\_RMSE\_3}
## [1] 7.72839
#Test RMSE
Test_RMSE_ALL
## [1] 4.767746
{\tt Test\_RMSE\_crim}
## [1] 9.318085
Test_RMSE_6
## [1] 5.95507
Test_RMSE_5
## [1] 6.148281
Test_RMSE_3
## [1] 8.643548
#AIC and BIC
AIC(Model_ALL)
## [1] 2399.014
AIC(Model_crim)
## [1] 2828.205
AIC(Model_6)
## [1] 2551.756
AIC(Model_5)
## [1] 2581.935
AIC(Model_3)
## [1] 2781.071
BIC(Model_ALL)
## [1] 2458.886
BIC(Model_crim)
## [1] 2840.179
BIC(Model_6)
## [1] 2583.688
```

# ## [1] 2609.875 BIC(Model\_3) ## [1] 2801.029 Model All Train RMSE = 4.675465 Test RMSE = 4.767746 LOOCV RMSE = 4.908037 Variance = .004258 R^2= .7262 Adjusted R^2 = .7169 AIC = 2399.014 BIC = 2458.886 Model crim Train RMSE = 8.238496 Test RMSE = 9.318085 LOOCV RMSE = 8.310566 Variance = .58276 R^2 = .1497 Adjusted R^2 = .1476 AIC = 2828.205 BIC = 2840.179 Model 6 Train RMSE = 5.758958 Test RMSE = 5.95507 LOOCV RMSE = 5.875081 Variance = .01923 R^2 = .5845 Adjusted R^2 = .5782 AIC = 2551.756 BIC = 2583.688 Model 5 Train RMSE = 5.995325 Test RMSE = 6.148281 LOOCV RMSE = 6.107382 Variance = .01170 R^2 = .5497 Adjusted R^2 = .544 AIC = 2581.935 BIC = 2609.875

 $Model\ 3$  Train RMSE = 7.72839 Test RMSE = 8.643548

 $R^2 = .2518 \text{ Adjusted } R^2 = .2461 \text{ AIC} = 2781.071 \text{ BIC} = 2801.029$ 

LOOCV RMSE = 7.805229 Variance = .41876

The model that is the best for *predicting* the response variable medv is the Model\_ALL which contains all predictors. This is based on the fact that it had the smallest RMSE for train and test data and the smallest LOOCV RMSE. This is important because the smaller the RMSE, the less error is attributed to the model. It is also important in regard to the LOOCV RMSE because this RMSE measure implicity penalizes models for having more predictors. Secondly, this is due to this model having the smallest variance, or spread, between the RMSE calculated from the test data and the train data. Lastly, this choice is based on the fact that such model has the highest R^2 and Adjusted R^2 values. This is important because these values both essentially give the percentage of variation in the response variable that is described by the model. These factors are weighed heavily as they provide measures of error that are attributed to a model. The less error, the better for predicting. It is also important to note, in terms of AIC and BIC statistics, Model\_ALL is considered the best model.