

# Transformation of Two or More Random Variables

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# Transformation of Two or More Random Variables

Suppose we know the joint probability density function for random variable  $X_1$  and  $X_2$  is  $f_{X_1, X_2}(x_1, x_2)$ .

If other random variables are defined,  $Y_1$  and  $Y_2$ , where  $Y_1 = g_1(x_1, x_2)$  and  $Y_2 = g_2(x_1, x_2)$ , then we want to know the joint probability density function for the random variables  $Y_1$  and  $Y_2$ ,  $f_{Y_1, Y_2}(y_1, y_2)$ .

# Bivariate Discrete Random Variable Transformation

Let  $X = (x_1, x_2, \dots, x_n)$  be a discrete random variable with dimension  $n$ . Defined,  $Y = (y_1, y_2, \dots, y_k) = (g_1(x), g_2(x), \dots, g_k(x))$ . Then, the probability mass function (pmf) of  $Y$  is:

$$f_{Y_1, Y_2}(y_1, y_2) = P_{Y_1, Y_2}(y_1, y_2) = \sum_{x_1, x_2, \dots, x_n} P_{x_1, x_2, \dots, x_n}(x_1, x_2, \dots, x_n)$$

# Example 1

$X_1$  and  $X_2$  are 2 random variables with pmf below in the table. Find the pmf for random variable  $Y = |X_1 - X_2|$ .

|       |     | $X_1$ |     |
|-------|-----|-------|-----|
|       |     | -1    | 1   |
| $X_2$ | 0   | 1/6   | 1/6 |
|       | 0.5 | 1/3   | 1/3 |

| $X$       | $f_{X_1, X_2}(x_1, x_2)$ | $Y =  X_1 - X_2 $ |
|-----------|--------------------------|-------------------|
| (-1, 0)   | 1/6                      | 1                 |
| (1, 0)    | 1/6                      | 1                 |
| (-1, 0.5) | 1/3                      | 1.5               |
| (1, 0.5)  | 1/3                      | 0.5               |

Solution:

Define the part of Y.

$$f_{Y_1, Y_2}(y_1, y_2) = \begin{cases} \frac{2}{6} & ; y = 1 \\ \frac{1}{3} & ; y = 1.5 \\ \frac{1}{3} & ; y = 0.5 \end{cases}$$

# Bivariate Discrete Random Variable Transformation

Suppose  $X_1$  and  $X_2$  are 2 discrete random variables with density  $f_{X_1, X_2}(x_1, x_2)$ . If  $Y_1 = g_1(X_1, X_2)$  and  $Y_2 = g_2(X_1, X_2)$  are one on one transformation, then the joint pmf of  $Y_1, Y_2$ :

$$f_{Y_1, Y_2}(y_1, y_2) = \begin{cases} f_{X_1, X_2}(g_1^{-1}(y_1, y_2), g_2^{-1}(y_1, y_2)) & ; \quad y_1, y_2 \in \mathcal{B} \\ 0 & ; \quad otherwise \end{cases}$$

# Theorem

Suppose we know the joint probability density function for the random variable  $X_1$  and  $X_2$  is  $f_{X_1, X_2}(x_1, x_2)$  which is positive and continuous on the group  $\mathcal{A} \subseteq R^2$ , and defined function  $g_1, g_2: \mathcal{A} \rightarrow R$  and  $\mathcal{B}$  is an image of  $\mathcal{A}$  as a one-to-one transformation of  $(g_1, g_2)$ .

Therefore, if  $y_1 = g_1(x_1, x_2)$  and  $y_2 = g_2(x_1, x_2)$  then the inverse,  $x_1 = g_1^{-1}(y_1, y_2)$  and  $x_2 = g_2^{-1}(y_1, y_2) \in \mathcal{B}$ .

Assume that for  $(y_1, y_2) \in \mathcal{B}$ ,  $dx_1/dy_1$  and  $dx_2/dy_2$  exist, continue, and not equal to 0.

Then the joint probability density function for the random variables  $Y_1$  and  $Y_2$  is:

$$f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}\{g_1^{-1}(y_1, y_2), g_2^{-1}(y_1, y_2)\} \cdot |J|, \quad (y_1, y_2) \in \mathcal{B}$$

# Joint probability density function for the random variables $Y_1$ and $Y_2$

$$f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}\{g_1^{-1}(y_1, y_2), g_2^{-1}(y_1, y_2)\} \cdot |J|, (y_1, y_2) \in \mathcal{B}$$

$$J = \text{Jacobian} = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix}$$

# Example 2

Suppose the continuous random variables  $X_1$  and  $X_2$  has a distribution Exponential with *mean* = 1 and its probability density function:  $f(x) = e^{-x}, I_{(0,\infty)}(x)$ .

If  $Y_1 = X_1$  and  $Y_2 = X_1 + X_2$  are defined, determine:

- a) Joint density function for random variables  $X_1$  and  $X_2$  ,  
 $f_{X_1, X_2}(x_1, x_2)$ .
- b) Joint density function for random variables  $Y_1$  and  $Y_2$  ,  
 $f_{Y_1, Y_2}(y_1, y_2)$ .
- c) Marginal density function for random variables  $Y_1$ ,  $f_{Y_1}(y_1)$ .

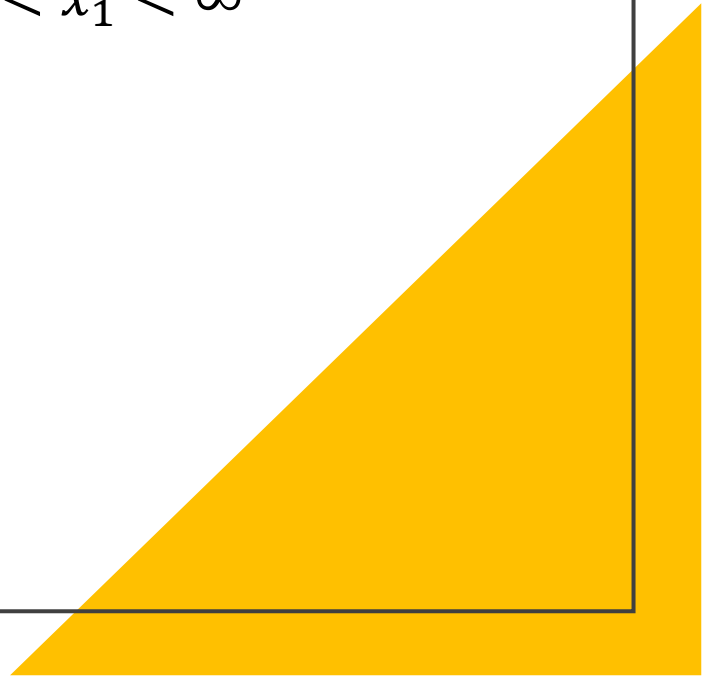


## Example 2 solution (a)

$$f_1(x_1) = e^{-x_1}; 0 < x_1 < \infty$$

$$f_2(x_2) = e^{-x_2}; 0 < x_2 < \infty$$

$$f_{X_1, X_2} = f_1(x_1)f_2(x_2) = e^{-x_1}e^{-x_2} = e^{-(x_1+x_2)}; 0 < x_1 < \infty, 0 < x_2 < \infty$$



# Example 2 solution (b)

Transformation one-to-one:

$$\begin{aligned}y_1 &= x_1 \rightarrow x_1 = y_1 \\y_2 &= x_1 + x_2 \rightarrow x_2 = y_2 - x_1 = y_2 - y_1\end{aligned}$$

Limit:

$$\begin{aligned}\mathcal{A} &= \{(x_1, x_2); 0 < x_1 < \infty, 0 < x_2 < \infty\} \\ \mathcal{B} &= \{(y_1, y_2); 0 < y_1 < y_2 < \infty\}\end{aligned}$$

$$J = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix} = \begin{vmatrix} 1 & 0 \\ -1 & 1 \end{vmatrix} = 1$$

$$\begin{aligned}f_{Y_1, Y_2}(y_1, y_2) &= f_{X_1, X_2}\{g_1^{-1}(y_1, y_2), g_2^{-1}(y_1, y_2)\} \cdot |J| \\ &= e^{-y_1} e^{-(y_2 - y_1)} = e^{-y_2}; (y_1, y_2) \in \mathcal{B}\end{aligned}$$

## Example 2 solution (c)

$$\begin{aligned} f_{Y_1}(y_1) &= \int_{y_1}^{\infty} f_{Y_1, Y_2}(y_1, y_2) dy_2 \\ &= \int_{y_1}^{\infty} e^{-y_2} dy_2 \\ &= -e^{-y_2} \Big|_{y_1}^{\infty} \\ &= 0 + e^{-y_1} \\ &= e^{-y_1}; 0 < y_1 < \infty \\ &\quad y_1 \sim \text{exponential}(1) \end{aligned}$$

# Example 3

Suppose the continuous random variable  $X$  has a distribution  $U(0,1)$ , while  $X_1$  and  $X_2$  are independent random variable examples of this distribution. If  $Y_1 = X_1 + X_2$  and  $Y_2 = X_1 - X_2$  are defined, determine:

- a) Joint density function for random variables  $Y_1$  and  $Y_2$  ,  
 $f_{Y_1, Y_2}(y_1, y_2)$ .
- b) Marginal density function for random variables  $Y_1$  and  $Y_2$  ,  
 $f_{Y_1}(y_1)$  and  $f_{Y_2}(y_2)$ .

# Example 3 solution (a)

$X \sim U(0,1)$  and  $X_1, X_2$  are independent and identical random sample of this distribution. Then the joint probability density function for  $X_1$  and  $X_2$  is:

$$f_{X_1, X_2}(x_1, x_2) = f_{X_1}(x_1) \cdot f_{X_2}(x_2) = 1; 0 < x_1 < 1 \text{ and } 0 < x_2 < 1$$

then it is defined that:

$$\begin{aligned} y_1 &= g_1(x_1, x_2) = x_1 + x_2 \\ y_2 &= g_2(x_1, x_2) = x_1 - x_2 \end{aligned}$$

# Example 3 solution (a)

Through the substitution or elimination method from the above equation, the following equation will be obtained:

$$x_1 = g_1^{-1}(y_1, y_2) = (y_1 + y_2)/2$$

$$x_2 = g_2^{-1}(y_1, y_2) = (y_1 - y_2)/2$$

$$J = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix} = \begin{vmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{vmatrix}$$

# Example 3 solution (a)

So, the common density for the random variables  $Y_1$  and  $Y_2$  is

$$f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}\{g_1^{-1}(y_1, y_2), g_2^{-1}(y_1, y_2)\} \cdot |J|$$

$$f_{Y_1, Y_2}(y_1, y_2) = \{(y_1 + y_2)/2, (y_1 - y_2)/2\} \cdot \left| -\frac{1}{2} \right|$$

$$f_{Y_1, Y_2}(y_1, y_2) = (1) \cdot \frac{1}{2}$$

$$f_{Y_1, Y_2}(y_1, y_2) = \frac{1}{2}, (y_1, y_2) \in T$$

# Example 3 solution (a)

determine the limit value for  $y_1$  and  $y_2$  that is  $T$ ,

For  $0 < x_1 < 1$

$$0 < x_1 < 1 \rightarrow 0 < (y_1 + y_2)/2 < 1 \rightarrow 0 < y_1 + y_2 < 2$$

$$0 < y_1 + y_2 \text{ and } y_1 + y_2 < 2$$

$$y_2 > y_1 \text{ and } y_2 < 2 - y_1$$

For  $0 < x_2 < 1$

$$0 < x_2 < 1 \rightarrow 0 < (y_1 - y_2)/2 < 1 \rightarrow 0 < y_1 - y_2 < 2$$

$$0 < y_1 - y_2 \text{ and } y_1 - y_2 < 2$$

$$y_2 < y_1 \text{ and } y_2 > y_1 - 2$$



# Example 3 solution (b)

The marginal distribution for  $y_1$  is

For  $0 < y_1 \leq 1$

$$f_{Y_1}(y_1) = \int_{-y_1}^{y_1} f_{Y_1 Y_2}(y_1, y_2) dy_2 = \int_{-y_1}^{y_1} \frac{1}{2} dy_2 = y_1$$

For  $1 < y_1 < 2$

$$f_{Y_1}(y_1) = \int_{y_1-2}^{2-y_1} f_{Y_1 Y_2}(y_1, y_2) dy_2 = \int_{y_1-2}^{2-y_1} \frac{1}{2} dy_2 = 2 - y_1$$

Therefore,

$$f_{Y_1}(y_1) = \begin{cases} y_1 & ; \quad 0 < y_1 \leq 1 \\ 2 - y_1 & ; \quad 1 < y_1 < 2 \\ 0 & ; \quad \textit{otherwise} \end{cases}$$

# Example 3 solution (b)

The marginal distribution for  $y_2$  is

For  $-1 < y_2 \leq 0$

$$f_{Y_2}(y_2) = \int_{-y_2}^{y_2+2} f_{Y_1Y_2}(y_1, y_2) dy_1 = \int_{-y_2}^{y_2+2} \frac{1}{2} dy_1 = y_2 + 1$$

For  $0 < y_2 < 1$

$$f_{Y_2}(y_2) = \int_{y_2}^{2-y_2} f_{Y_1Y_2}(y_1, y_2) dy_1 = \int_{y_2}^{2-y_2} \frac{1}{2} dy_1 = 1 - y_2$$

Therefore,

$$f_{Y_2}(y_2) = \begin{cases} y_2 + 1 & ; \quad -1 < y_2 \leq 0 \\ 1 - y_2 & ; \quad 0 < y_2 < 1 \\ 0 & ; \quad \textit{otherwise} \end{cases}$$

# Example 4

Suppose the continuous random variable  $X$  has the following probability density function

$$f_X(x) = e^{-x}, x \geq 0$$

$X_1$  and  $X_2$  is an independent and identic random variable of this probability density function. Determine the probability density function of random variable  $Y = X_1/(X_1 + X_2)$ .

# Example 4 solution

$X_1$  and  $X_2$  is an independent and identic random variable for this probability density function (pdf), then the joint pdf for  $X_1$  and  $X_2$ :

$$f_{X_1, X_2}(x_1, x_2) = e^{-x_1} e^{-x_2} = e^{-(x_1 + x_2)}; x_1 \geq 0 \text{ and } x_2 \geq 0$$

# Example 4 solution

It is necessary to define another random variable for the transformation. occurs from two-dimensional space to two-dimensional space. Suppose  $Z = X_1 + X_2$ , so that we get a pair of transformations i.e.  $y = x_1/(x_1 + x_2)$  and  $z = x_1 + x_2$ . This transformation is one-to-one for all functional areas.

# Example 4 solution

Through the substitution or elimination method from the above equation, the following equation will be obtained:

$$\begin{aligned}x_1 &= yz \\ x_2 &= (1 - y)z\end{aligned}$$

$$J = \begin{vmatrix} \frac{\partial x_1}{\partial y} & \frac{\partial x_1}{\partial z} \\ \frac{\partial x_2}{\partial y} & \frac{\partial x_2}{\partial z} \end{vmatrix} = \begin{vmatrix} z & y \\ -z & 1 - y \end{vmatrix} = z$$

# Example 4 solution

So, the joint density for the random variables  $Y$  and  $Z$  is

$$\begin{aligned} f_{Y,Z}(y, z) &= f_{X_1, X_2}(x_1, x_2) \cdot |J| \\ &= e^{-(yz + (1-y)z)} \cdot |J| \\ &= ze^{-z}, (y, z) \in T \end{aligned}$$

# Example 4 solution

Next determine the limit value for  $y$  and  $z$ , namely  $T$ .

Pay attention, because  $x_1 \geq 0$  and  $x_2 \geq 0$ , then

$$0 \leq y = x_1/(x_1 + x_2) \leq 1 \rightarrow 0 \leq y \leq 1$$

$$z = x_1 + x_2 \geq 0 \rightarrow z \geq 0$$

Therefore,

$$f_{Y,Z}(y, z) = ze^{-z} \quad , 0 \leq y \leq 1 \quad \text{and } z \geq 0$$



# Example 4 solution

Marginal distribution for random variables  $Y = X_1/(X_1 + X_2)$ :

$$\int_0^{\infty} ze^{-z} dz = 1$$

Hence, the pdf for random variable  $Y = X_1/(X_1 + X_2)$ :

$$f_Y(y) = \begin{cases} 1 & ; \quad 0 \leq y \leq 1 \\ 0 & ; \quad otherwise \end{cases}$$

# Exercise 1

Suppose the random variables  $X$  and  $Y$  are independent and have a Negative Exponential probability density function with  $\lambda = 1$ , and it is defined that the random variable  $U = (X + Y)/2$  and  $V = (X - Y)/2$ . Determine:

- a) Joint pdf  $f_{U,V}(u, v)$ .
- b) Marginal pdf  $f_U(u)$  and  $f_V(v)$ .

Note:

$$f_X(x) = \lambda e^{-\lambda x}, x > 0, \lambda > 0$$

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## Exercise 2

Suppose the random variables  $X$  and  $Y$  are independent and have pdf  $Normal(0, 1)$  and defined  $U = (X + Y)$  and  $V = (X - Y)$ . Determine:

- a) Joint pdf  $f_{U,V}(u, v)$ .
  - b) Marginal pdf  $f_U(u)$  and  $f_V(v)$ .
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A blue ribbon graphic with a 3D effect, featuring a dark blue shadow on the left side. The text "Thank you" is written in white on the main blue surface.

Thank you