## STAT 499: Undergraduate Research

## Week 4: Estimation in noisy settings

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In this lecture note, we consider such a linear model  $y = \mathbf{X}\theta^* + w$ , and mainly study the consistency of the Lagrangian Lasso solution:

$$\min_{\theta \in \mathbb{R}^d} \left\{ \frac{1}{2n} \| y - \mathbf{X}\theta \|_2^2 + \lambda_n \| \theta \|_1 \right\}. \tag{4.1}$$

Extension of restricted nullspace. Define the set

$$\mathbb{C}_{\alpha}(S) := \{ \Delta \in \mathbb{R} : \|\Delta_{S^c}\|_1 \le \alpha \|\Delta_S\|_1 \}.$$

This definition generalizes the set  $\mathbb{C}(S)$  used in our definition of the restricted nullspace property, which corresponds to the special case  $\alpha = 1$ .

**RE condition.** The matrix **X** satisfies the restricted eigenvalue (RE) condition over S with parameters  $(k, \alpha)$  if

$$\frac{1}{n} \|\mathbf{X}\Delta\|_2^2 \ge k \|\Delta\|_2^2 \text{ for all } \Delta \in \mathbb{C}_{\alpha}(S). \tag{4.2}$$

**Assumption 4.1** In the case of hard sparse, we assume that

- (A<sub>1</sub>) The vector  $\theta^*$  is supported on a subset  $S \subset \{1, \ldots, d\}$  with |S| = s.
- $(\mathbf{A_2})$  The design matrix **X** satisfies the RE condition with parameter (k,3):

$$\frac{1}{n} \|\mathbf{X}\Delta\|_2^2 \ge k \|\Delta\|_2^2 \text{ for all } \Delta \in \mathbb{C}_3(S).$$

Theorem 4.2 (Theorem 7.13 in Wainwright's book) Under above assumption, for any solution  $\widehat{\theta}$  of the Lagrangian Lasso (4.1) with  $\lambda_n \geq 2 \|\frac{\mathbf{X}^T w}{n}\|_{\infty}$ , we have

$$\|\widehat{\theta} - \theta^*\|_2 \le \frac{3}{k} \sqrt{s} \lambda_n.$$

**Proof:** Our first step is to show that, under the condition  $\lambda_n \geq 2 \left\| \frac{\mathbf{X}^T w}{n} \right\|_{\infty}$ , the error vector  $\widehat{\Delta} = \widehat{\theta} - \theta^*$  belongs to  $\mathbb{C}_3(S)$ .

Define the Lagrangian  $L(\theta; \lambda_n) = \frac{1}{2n} ||y - \mathbf{X}\theta||_2^2 + \lambda_n ||\theta||_1$ . Since  $\widehat{\theta}$  is optimal, we have

$$L\left(\widehat{\theta};\lambda_n\right) \leq L\left(\theta^*;\lambda_n\right).$$

That is,

$$\frac{1}{2n} \|y - \mathbf{X}\widehat{\theta}\|_{2}^{2} + \lambda_{n} \|\widehat{\theta}\|_{1} \leq \frac{1}{2n} \|y - \mathbf{X}\theta^{*}\|_{2}^{2} + \lambda_{n} \|\theta^{*}\|_{1}$$

$$\iff \frac{1}{2n} \|\mathbf{X}\theta^{*} + w - \mathbf{X}\widehat{\theta}\|_{2}^{2} + \lambda_{n} \|\widehat{\theta}\|_{1} \leq \frac{1}{2n} \|w\|_{2}^{2} + \lambda_{n} \|\theta^{*}\|_{1}$$

$$\iff \frac{1}{2n} (\mathbf{X}\widehat{\Delta} - w)^{T} (\mathbf{X}\widehat{\Delta} - w) + \lambda_{n} \|\widehat{\theta}\|_{1} \leq \frac{1}{2n} w^{T} w + \lambda_{n} \|\theta^{*}\|_{1}.$$

Rearranging it, we have

$$0 \le \frac{1}{2n} \|\mathbf{X}\widehat{\Delta}\|_{2}^{2} \le \frac{w^{\mathrm{T}}\mathbf{X}\widehat{\Delta}}{n} + \lambda_{n} \left\{ \|\boldsymbol{\theta}^{*}\|_{1} - \|\widehat{\boldsymbol{\theta}}\|_{1} \right\}.$$

Now since  $\theta^*$  is S-sparse, we can write

$$\|\theta^*\|_1 - \|\widehat{\theta}\|_1 = \|\theta_S^*\|_1 - \|\theta_S^* + \widehat{\Delta}_S\|_1 - \|\widehat{\Delta}_S\|_1.$$

Thus,

$$0 \leq \frac{1}{n} \|\mathbf{X}\widehat{\Delta}\|_{2}^{2} \leq 2 \frac{w^{\mathsf{T}} \mathbf{X}\widehat{\Delta}}{n} + 2\lambda_{n} \left\{ \|\boldsymbol{\theta}_{S}^{*}\|_{1} - \left\|\boldsymbol{\theta}_{S}^{*} + \widehat{\Delta}_{S}\right\|_{1} - \left\|\widehat{\Delta}_{S^{c}}\right\|_{1} \right\}$$

$$\stackrel{(i)}{\leq} 2 \|\mathbf{X}^{\mathsf{T}} w/n\|_{\infty} \|\widehat{\Delta}\|_{1} + 2\lambda_{n} \left\{ \|\widehat{\Delta}_{S}\|_{1} - \|\widehat{\Delta}_{S^{c}}\|_{1} \right\}$$

$$\stackrel{(ii)}{\leq} \lambda_{n} \left\{ 3 \|\widehat{\Delta}_{S}\|_{1} - \|\widehat{\Delta}_{S^{c}}\|_{1} \right\}, \tag{4.3}$$

where step (i) follows from a combination of Hölder's inequality and the triangle inequality, whereas step (ii) follows from the choice of  $\lambda_n$ .

Inequality (4.3) shows that  $\widehat{\Delta} \in \mathbb{C}_3(S)$ , so that the RE condition may be applied. Doing so, we obtain

$$\kappa \|\widehat{\Delta}\|_2^2 \le 3\lambda_n \sqrt{s} \|\widehat{\Delta}\|_2,$$

which implies the consistency.

**Example 7.14 in Wainwright's book.** Consider the classical linear Gaussian model, where  $w \in \mathbb{R}^n$  has i.i.d.  $\mathcal{N}(0, \sigma^2)$  entries.

Consider  $\mathbf{X} \in \mathbb{R}^{n \times d}$  is fixed. Suppose that  $\mathbf{X}$  satisfies the RE condition, and that it is C-column normalized, i.e.,

$$\max_{j=1,\dots,d} \frac{\|\mathbf{X}(\cdot,j)\|_2}{\sqrt{n}} \le C.$$

Thus, the random variable  $\|\frac{\mathbf{X}^T w}{n}\|_{\infty}$  corresponds to the absolute maximum of d zero-mean Gaussian variables, each with variance at most  $\frac{C^2 \sigma^2}{n}$ , since

$$var(\frac{\mathbf{X}(\cdot,j)^Tw}{n}) = \frac{\sigma^2}{n^2}\sum_{i=1}^n\mathbf{X}(i,j)^2 = \sigma^2\frac{\|\mathbf{X}(\cdot,j)\|_2^2}{n^2} \leq \frac{C^2\sigma^2}{n}.$$

The standard Gaussian tail bounds states that, for any  $j \in \{1, ..., d\}$ ,

$$P\left(\left|\frac{\mathbf{X}(\cdot,j)^Tw}{n}\right| \ge t\right) \le 2\exp\{-\frac{nt^2}{2C^2\sigma^2}\} \text{ for all } t > 0.$$

Thus, for all  $\delta > 0$ ,

$$\begin{split} P\left(\left\|\frac{\mathbf{X}^T w}{n}\right\|_{\infty} &\geq C\sigma\left(\sqrt{\frac{2\log d}{n}} + \delta\right)\right) \leq \sum_{j=1}^d P\left(\frac{\mathbf{X}(\cdot,j)^T w}{n} \geq C\sigma\left(\sqrt{\frac{2\log d}{n}} + \delta\right)\right) \\ &\leq 2d \exp\left\{-\frac{nC^2\sigma^2(\sqrt{\frac{2\log d}{n}} + \delta)^2}{2C^2\sigma^2}\right\} \\ &\leq 2e^{-n\delta^2/2}. \end{split}$$

If we set  $\lambda_n = 2C\sigma(\frac{2\log d}{n} + \delta)$ , then the theorem implies that

$$\|\widehat{\theta} - \theta^*\|_2 \le \frac{6C\sigma}{k} \sqrt{s} \left\{ \frac{2\log d}{n} + \delta \right\}.$$