

Cryptography: Birthday Paradox

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Theorem 1.1. Let $S = \{1, 2, \dots, N\}$. For n times, uniformly randomly draw one element from set S with replacement. Let x_t be the element we draw at time t . Then $\forall p > 0$, there exists a constant C_1 such that when $n \geq C_1\sqrt{N}$, we have

$$Pr[\exists 1 \leq i, j \leq n, i \neq j \text{ such that } x_i = x_j] > p$$

.

Proof. Let X denote the event that $\exists i, j \leq n, i \neq j$ such that $x_i = x_j$, then we have

$$\begin{aligned} Pr[\bar{X}] &= \frac{N(N-1) \cdots (N-n+1)}{N^n} \\ &= \prod_{i=1}^{n-1} \left(1 - \frac{i}{N}\right) \\ &\leq \prod_{i=1}^{n-1} \exp\left(-\frac{i}{N}\right) \\ &= \exp\left(-\frac{n(n-1)}{2N}\right) \end{aligned} \tag{1}$$

Let $C_1 = \sqrt{-2\ln(1-p)} + 1$. Then when $n \geq C_1\sqrt{N}$, we have

$$n(n-1) > (1 + \sqrt{-2\ln(1-p) \cdot N})(\sqrt{-2\ln(1-p) \cdot N}) > 2\ln(2) \cdot N \tag{2}$$

Which is equivalent to $-\frac{n(n-1)}{2N} < \ln(1-p)$. Thus use (1) we have

$$Pr[\bar{X}] \leq \exp\left(-\frac{n(n-1)}{2N}\right) < 1-p$$

So we have

$$Pr[X] > p$$

□

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Lemma 1.1. For positive integer $n < N$, we have

$$\sum_{i=1}^{n-1} \ln\left(1 - \frac{i}{N}\right) > -\frac{n^2}{N}$$

Proof. Notice that $\forall x \in [1 - \frac{i+1}{N}, 1 - \frac{i}{N}]$ ($0 \leq i \leq n$), we have $\ln(x) < \ln(1 - \frac{i}{N})$. Thus

$$\frac{1}{N} \ln(1 - \frac{i}{N}) \geq \int_{1 - \frac{i+1}{N}}^{1 - \frac{i}{N}} \ln(x) dx$$

Thus we have

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^{n-1} \ln(1 - \frac{i}{N}) &\geq \int_{1 - \frac{n}{N}}^1 \ln(x) dx \\ &= (x \ln(x) - x) \Big|_{1 - \frac{n}{N}}^1 \\ &= -\frac{n}{N} - (1 - \frac{n}{N}) \ln(1 - \frac{n}{N}) \\ &> -\frac{n}{N} - (1 - \frac{n}{N}) (-\frac{n}{N}) \\ &= -\frac{n^2}{N^2} \end{aligned} \tag{3}$$

Thus

$$\sum_{i=1}^{n-1} \ln(1 - \frac{i}{N}) > -\frac{n^2}{N}$$

□

Theorem 1.2. Let $S = \{1, 2, \dots, N\}$. For n times, uniformly randomly draw one element from set S with replacement. Let x_t be the element we draw at time t . Then $\forall p > 0$, there exists a constant C_2 such that when $n \leq C_2 \sqrt{N}$, we have

$$Pr[\exists 1 \leq i, j \leq n, i \neq j \text{ such that } x_i = x_j] < p$$

.

Proof. Let X denote the event that $\exists i, j \leq t, i \neq j$ such that $x_i = x_j$.

Use Lemma 1.1, we have

$$\begin{aligned} Pr[\bar{X}] &= \frac{N(N-1) \cdots (N-n+1)}{N^n} \\ &= \prod_{i=1}^{n-1} (1 - \frac{i}{N}) \\ &= \exp(\sum_{i=1}^{n-1} \ln(1 - \frac{i}{N})) \\ &> \exp(-\frac{n^2}{N}) \end{aligned} \tag{4}$$

Let $C_2 = \sqrt{-\ln(1-p)}$. Then when $n \leq C_2 \sqrt{N}$, we have

$$\exp(-\frac{n^2}{N}) \geq 1 - p$$

So $Pr[\bar{X}] > 1 - p$, thus

$$Pr[X] < p$$

□

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Theorem 2.1. Let $S = \{1, 2, \dots, N\}$. Let $D_1 : S \rightarrow R^+ \cup \{0\}$ be a discrete probability distribution over S . For n times, randomly draw one element from set S according to distribution D_1 with replacement. Let x_t be

the element we draw at time t . Let D_0 be the uniform distribution over S , which satisfies $\forall i \in S, D_0(i) = \frac{1}{N}$. Then we have

$$Pr_{D_1^n}[\exists 1 \leq i, j \leq n, i \neq j \text{ such that } x_i = x_j] \geq Pr_{D_0^n}[\exists 1 \leq i, j \leq n, i \neq j \text{ such that } x_i = x_j]$$

.

Proof. Let X denote the event that $\exists i, j \leq t, i \neq j$ such that $x_i = x_j$. Let X_m denote the event that $\exists 1 \leq i, j \leq n, i \neq j$ such that $x_i = x_j = m$.

First, to change D_1 to D_0 , we can apply the following algorithm:

1. $t := 1$
2. While $D_t \neq D_0$:
3. find $i, j \in S$ such that $D_t[i] < \frac{1}{N} < D_t[j]$
4. let $D_{t+1}[j] := D_t[i] + D_t[j] - \frac{1}{N}, D_{t+1}[i] := \frac{1}{N}, \forall k \neq i, j, D_{t+1}[k] := D_t[k]$
5. $t++$
6. End While

Since the number of $\frac{1}{N}$ in D increases at each iteration, this algorithm will terminate in N steps.

We only need to prove that

$$\forall t, Pr_{D_t^n}[X] \geq Pr_{D_{t+1}^n}[x]$$

Without losing generality, suppose when generate D_{t+1} from D_t , we choose $i = 1, j = 2$.

Let Y be the number of times that the element we draw is in $\{1, 2\}$.

$$\begin{aligned} Pr_{D^n}[X] &= Pr_{D^n}[\bigcup_{k=3}^N X_k] + (1 - Pr_{D^n}[\bigcup_{k=3}^N X_k]) Pr_{D^n}(X_1 \cup X_2 | \bigcap_{k=3}^N \bar{X}_k) \\ &= Pr_{D^n}[\bigcup_{k=3}^N X_k] + (1 - Pr_{D^n}[\bigcup_{k=3}^N X_k]) \sum_{i=0}^{\infty} Pr_{D^n}(Y = i | \bigcap_{k=3}^N \bar{X}_k) Pr_{D^n}(X_1 \cup X_2 | Y = i) \end{aligned} \quad (5)$$

The last equation holds because $\forall i, Pr_{D^n}(X_1 \cup X_2 | Y = i) = Pr_{D^n}(X_1 \cup X_2 | Y = i, \bigcap_{k=3}^N \bar{X}_k)$.

Notice that

$$\begin{aligned} &Pr_{D_t^n}(X_1 \cup X_2 | Y = 2) - Pr_{D_{t+1}^n}(X_1 \cup X_2 | Y = 2) \\ &= \frac{1}{(D_t[1] + D_t[2])^2} (D_t[1]^2 + D_t[2]^2 - (\frac{1}{N})^2 - (D_t[1] + D_t[2] - \frac{1}{N})^2) \\ &= -\frac{2}{(D_t[1] + D_t[2])^2} (D_t[1] - \frac{1}{N})(D_t[2] - \frac{1}{N}) \\ &> 0 \end{aligned} \quad (6)$$

And for any distribution D over S we have

$$Pr_{D^n}(X_1 \cup X_2 | Y = i) = \begin{cases} 0 & i = 0, 1 \\ 1 & i \geq 3 \end{cases} \quad (7)$$

Thus

$$\forall i, Pr_{D_t^n}(X_1 \cup X_2 | Y = i) \geq Pr_{D_{t+1}^n}(X_1 \cup X_2 | Y = i)$$

Since we only adjust $D_t[0], D_t[1]$,

$$Pr_{D_t^n}[\bigcup_{k=3}^N X_k] = Pr_{D_{t+1}^n}[\bigcup_{k=3}^N X_k]$$

$$\forall i, Pr_{D_t^n}(Y = i | \bigcap_{k=3}^N \bar{X}_k) = Pr_{D_{t+1}^n}(Y = i | \bigcap_{k=3}^N \bar{X}_k)$$

So consider equation (6) and we get

$$Pr_{D_t^n}[X] \geq Pr_{D_{t+1}^n}[x]$$

□

Acknowledgement: