

07_bootstrap

December 18, 2025

1 Bootstrapping

Due to some issues we noticed when replicating Galvao et al's asymptotic variance estimator and test statistic, we turn to bootstrapping to try to come up with bootstrap bias estimate, bootstrap standard errors, and an alternative Wald-type test.

1.1 Notebook setup

```
[63]: import pandas as pd
import numpy as np
import matplotlib.pyplot as plt
from scipy.stats import norm
from scipy.stats import chi2
from scipy import stats
import torch
from tqdm import tqdm
import sys
sys.path.append('../')
from utils import utils
sys.executable
```

```
[63]: '/Users/fanghema/Desktop/aaSTAT_5200/STAT_5200_final_project/env/bin/python'
```

```
[2]: data = pd.read_csv(
    '../data/processed/data_extended.csv',
    index_col=0,
    parse_dates=True
)

factors = ['Mkt-RF', 'SMB', 'HML', 'RMW', 'CMA']
assets = [col for col in data.columns if col != 'RF' and col not in factors]
data['Quarter'] = data.index.to_period("Q")

beta_loading, returns_df, realized_covariance, residuals = utils.
    ↪calculate_factor_loading(
    data,
    factors=factors,
    assets=assets
```

```

)

excess_returns = (
    returns_df
    .groupby("Quarter")
    .sum()
    [assets]
    .T
    .values
)

industries = beta_loading.index.get_level_values(0).unique().tolist()
factors = beta_loading.index.get_level_values(1).unique().tolist()

N = len(industries)
K = len(factors)
T = beta_loading.shape[1]
R = 3

beta_hat_np = np.zeros((N, K, T))

for i, asset in enumerate(industries):
    for j, factor in enumerate(factors):
        beta_hat_np[i, j, :] = beta_loading.loc[(asset, factor)].values

beta_hat_np.shape

```

[2]: (47, 5, 250)

1.2 Using Jackknife

```

[ ]: p = N * (K + 1)
jackknife_eta_estimates = np.zeros((T, p))

for t in range(T):
    if (t % 10) == 0:
        print(f"Processing {t} out of {T}")

    time_mask = np.ones(T, dtype=bool)
    time_mask[t] = False

    beta_hat_subset = beta_hat_np[:, :, time_mask]          # (N, K, T-1)
    excess_returns_subset = excess_returns[:, time_mask]    # (N, T-1)

    eta_jk, G_jk, beta_star_jk, objective_jk = utils.penalty_based_minimization(
        beta_hat=beta_hat_subset,
        excess_returns=excess_returns_subset,
        N=N,

```

```

        K=K,
        R=R,
        T=T-1,          # updated sample size
        n_iter=500,
    )

    jackknife_eta_estimates[t, :] = eta_jk.reshape(p)

```

Now, we do a full sample estimation.

```

[7]: eta, G, beta_star, objective = utils.iterative_convergence(
    beta_hat=beta_hat_np,
    excess_returns=excess_returns,
    N=N,
    K=K,
    R=R,
    T=T,
    n_iter=2000,
    verbose=False
)

p = N * (K + 1)
eta_hat_full = eta.reshape(p)

```

```

[42]: # Jackknife mean
eta_jk_mean = jackknife_eta_estimates.mean(axis=0)

# Jackknife bias
jk_bias = (T - 1) * (eta_jk_mean - eta_hat_full)

# Jackknife covariance
eta_centered = jackknife_eta_estimates - eta_jk_mean
jk_cov = ((T - 1) / T) * (eta_centered.T @ eta_centered)

```

```

[43]: eta_centered.shape

```

```

[43]: (250, 282)

```

```

[44]: print(np.linalg.inv(jk_cov).min())
      print(np.linalg.inv(jk_cov).max())

```

```

-1.863547211797388e+17
1.638447607959742e+17

```

```

[45]: p = N * (K + 1)

eta_mean = eta.mean(axis=0)
eta_centered = eta - eta_mean

```

```

d_vec = eta_centered.reshape(-1)
print("Typical |d|:", np.median(np.abs(d_vec)))

print("Diag(jk_cov) head:", np.diag(jk_cov)[:10])

rough_var = jackknife_eta_estimates.var(axis=0)[:10]
print("Naive var across jk replicates (first 10):", rough_var)

```

```

Typical |d|: 0.050935989443553834
Diag(jk_cov) head: [1.26641957e+01 9.11005233e-03 2.70422503e-02 5.87296361e-02
7.69373618e-02 1.26184372e-01 5.24237287e+01 7.07177949e-02
2.01213236e-01 4.07243362e-02]
Naive var across jk replicates (first 10): [5.08602237e-02 3.65865555e-05
1.08603415e-04 2.35861993e-04
3.08985389e-04 5.06764545e-04 2.10537063e-01 2.84007209e-04
8.08085283e-04 1.63551551e-04]

```

```

[40]: def _flatten_eta(eta: np.ndarray) -> np.ndarray:
    """
    vec(eta) stacking assets one after another.
    eta: (N, K+1)
    returns: (N*(K+1),)
    """
    return eta.reshape(-1)

def wald_full_homogeneity_jackknife(
    eta: np.ndarray,
    jk_cov: np.ndarray,
    N: int,
    K: int
):
    """
    Wald-type test for joint homogeneity:
        H0: all intercepts equal AND all slopes equal across assets.
        Here we use centered eta_i (deviation from cross-sectional mean)
        and jackknife covariance of vec(eta).

    Returns:
        gamma_ad: standardized test statistic ~ N(0,1) under H0
        W: chi-square statistic ~  $\chi^2_q$  under H0
        p_value: p-value based on N(0,1) approximation for gamma_ad
    """
    p = N * (K + 1)
    assert eta.shape == (N, K + 1)
    assert jk_cov.shape == (p, p)

```

```

# Center eta across assets
eta_mean = eta.mean(axis=0)          # (K+1,)
eta_centered = eta - eta_mean        # (N, K+1)
d_vec = _flatten_eta(eta_centered)   # (p,)

V = jk_cov
V_inv = np.linalg.pinv(V)

W = float(d_vec.T @ V_inv @ d_vec)

q = (N - 1) * (K + 1)

gamma_ad = (W - q) / np.sqrt(2 * q)

p_val = 2 * (1 - norm.cdf(abs(gamma_ad)))

return gamma_ad, W, p_val

def wald_intercept_homogeneity_jackknife(
    eta: np.ndarray,
    jk_cov: np.ndarray,
    N: int,
    K: int
):
    """
    Wald-type test for:
         $H_0: \alpha_i = 0$  for all  $i$  (intercepts jointly zero).

    Uses jackknife covariance submatrix corresponding to intercepts.
    """
    p = N * (K + 1)
    assert eta.shape == (N, K + 1)
    assert jk_cov.shape == (p, p)

    alpha = eta[:, 0]          # (N,)

    # indices of intercepts in vec(eta): 0, K+1, 2(K+1), ...
    alpha_idx = np.arange(0, p, K + 1)

    V_eta = jk_cov
    V_alpha = V_eta[np.ix_(alpha_idx, alpha_idx)]   # (N, N)
    V_alpha_inv = np.linalg.pinv(V_alpha)

    W = float(alpha.T @ V_alpha_inv @ alpha)

    q = N

```

```

gamma_a = (W - q) / np.sqrt(2 * q)
p_val = 2 * (1 - norm.cdf(abs(gamma_a)))

return gamma_a, W, p_val

def wald_slope_homogeneity_jackknife(
    eta: np.ndarray,
    jk_cov: np.ndarray,
    N: int,
    K: int
):
    """
    Wald-type test for:
         $H_0: \beta_i = \beta_j$  for all  $i, j$  (all slope vectors equal across assets).

    We work with cross-sectionally centered slopes, so under  $H_0$ 
    the mean of  $\text{vec}(\text{centered slopes})$  is approximately 0, and we
    use the jackknife covariance of those components.
    """
    p = N * (K + 1)
    assert eta.shape == (N, K + 1)
    assert jk_cov.shape == (p, p)

    slopes = eta[:, 1:] # (N, K)
    slopes_centered = slopes - slopes.mean(axis=0) # (N, K)
    slopes_vec = slopes_centered.reshape(N * K) # (NK,)

    slope_idx = []
    for i in range(N):
        for j in range(K):
            slope_idx.append(i * (K + 1) + 1 + j)
    slope_idx = np.array(slope_idx)

    V_eta = jk_cov
    V_lambda = V_eta[np.ix_(slope_idx, slope_idx)] # (NK, NK)
    V_lambda_inv = np.linalg.pinv(V_lambda)

    W = float(slopes_vec.T @ V_lambda_inv @ slopes_vec)

    q = (N - 1) * K

    gamma_lambda = (W - q) / np.sqrt(2 * q)
    p_val = 2 * (1 - norm.cdf(abs(gamma_lambda)))

    return gamma_lambda, W, p_val

```

```
[41]: gamma_a_lam_jk, W_a_lam_jk, _ = wald_full_homogeneity_jackknife(
        eta=eta,
        jk_cov = jk_cov,
        N = N,
        K = K
    )

    gamma_a_jk, W_a_jk, _ = wald_intercept_homogeneity_jackknife(
        eta=eta,
        jk_cov = jk_cov,
        N = N,
        K = K
    )

    gamma_lam_jk, W_lam_jk, _ = wald_slope_homogeneity_jackknife(
        eta=eta,
        jk_cov = jk_cov,
        N = N,
        K = K
    )

    print(gamma_a_lam_jk)
    print(gamma_a_jk)
    print(gamma_lam_jk)
```

```
1829.0568303859861
7.00146195513254
3411.160745721216
```

1.2.1 Using jackknife for full empirical test

```
[ ]: factor_options = [
    ['Mkt-RF'],
    ['Mkt-RF', 'SMB', 'HML'],
    ['Mkt-RF', 'SMB', 'HML', 'RMW', 'CMA'],
]
R_options = [1, 2, 5]
sample_period_options = [
    ('1963-01-01', '2025-12-31'),
    ('1963-01-01', '1983-01-01'),
    ('1973-01-01', '1993-01-01'),
    ('1983-01-01', '2003-01-01'),
    ('1993-01-01', '2013-01-01'),
    ('2003-01-01', '2023-01-01'),
]

results = pd.DataFrame(
    index=pd.MultiIndex.from_product([
```

```

        list(map(tuple, factor_options)),    # convert lists → tuples
        R_options,
        sample_period_options
    ]),
    columns=['gamma_a_lam', 'gamma_a', 'gamma_lam']
)

print(f"Total combinations: {results.shape[0]}")
counter = 0

for factors in factor_options:
    K = len(factors)
    for R in R_options:
        for sample_period in sample_period_options:
            print(f"Processing {counter}/{results.shape[0]}: {factors} - {R} - {sample_period}")
            data_slice = data.loc[
                (data.index > sample_period[0]) &
                (data.index < sample_period[1])
            ]
            beta_loading, returns_df, realized_covariance, residuals = utils.
↪calculate_factor_loading(
                data_slice,
                factors=factors,
                assets=assets
            )

            excess_returns = returns_df.groupby("Quarter").sum()[assets].T.
↪values
            industries = beta_loading.index.get_level_values(0).unique().
↪tolist()
            factors_names = beta_loading.index.get_level_values(1).unique().
↪tolist()

            N = len(industries)
            K = len(factors)
            T = beta_loading.shape[1]

            beta_hat_np = np.zeros((N, K, T))

            for i, asset in enumerate(industries):
                for j, factor in enumerate(factors):
                    beta_hat_np[i, j, :] = beta_loading.loc[(asset, factor)].
↪values

```



```

eta, G, beta_star, objective = utils.iterative_convergence(
    beta_hat_np,
    excess_returns,
    N = N,
    K = K,
    R = R,
    T = T,
    n_iter=2000
)

avar = utils.estimate_avar(
    beta_hat=beta_hat_np,
    excess_returns=excess_returns,
    eta=eta,
    G=G,
    beta_star=beta_star,
    realized_covariance=realized_covariance,
    residuals=residuals,
    N = N,
    K = K,
    R = R,
    T = T,
)

gamma_a_lambda = utils.full_homogeneity_test(
    eta = eta,
    avar = avar,
    N = N,
    K = K,
    T = T
)

gamma_a = utils.intercept_homogeneity_test(
    eta = eta,
    avar = avar,
    N = N,
    K = K,
    T = T
)

gamma_lambda = utils.slope_homogeneity_test(
    eta = eta,
    avar = avar,
    N = N,
    K = K,
    T = T
)

```

```

    )
    print(f"Test statistics")
    print(f"gamma_a_lam: {gamma_a_lambda}")
    print(f"gamma_a: {gamma_a}")
    print(f"gamma_lam: {gamma_lambda}")

    results.loc[(
        tuple(factors), R, sample_period
    )] = np.asarray([
        gamma_a_lambda,
        gamma_a,
        gamma_lambda
    ])
    counter += 1
    print(f"=====")

```

1.3 Testing for Type I error rate

```

[85]: def run_mc_jackknife(
    N=20, K=3, R=1, T=200,
    MC_REPS=100,
    n_iter_est=400,
    seed=12345,
    verbose=False,
    heterogeneity_strength = 0.0
):
    rng = np.random.default_rng(seed)

    # store statistics
    gamma_full_list = []
    gamma_a_list = []
    gamma_lam_list = []

    for rep in range(MC_REPS):
        rep_seed = rng.integers(1_000_000)

        # heterogeneity_strength = 0
        beta_true, r, realized_cov, residuals, G_true, beta_star_true,
        lambda_true = utils.simulate_dgp(
            N=N, K=K, R=R, T=T,
            heterogeneity_strength=heterogeneity_strength,
            seed=rep_seed
        )

        eta_hat, G_hat, beta_star_hat, obj = utils.iterative_convergence(
            beta_true,
            r,

```

```

        N=N, K=K, R=R, T=T,
        n_iter=n_iter_est,
        verbose=False
    )

    p = N * (K + 1)
    jk_eta = np.zeros((T, p))

    for t in range(T):
        mask = np.ones(T, dtype=bool)
        mask[t] = False

        beta_sub = beta_true[:, :, mask]
        r_sub = r[:, mask]

        eta_jk, _, _ = utils.iterative_convergence(
            beta_sub, r_sub,
            N=N, K=K, R=R, T=T-1,
            n_iter=n_iter_est,
            verbose=False
        )

        jk_eta[t, :] = eta_jk.reshape(p)

    jk_mean = jk_eta.mean(axis=0)

    diffs = jk_eta - jk_mean
    jk_cov = (T - 1) / T * (diffs.T @ diffs)

    gamma_full, _, _ = wald_full_homogeneity_jackknife(
        eta=eta_hat,
        jk_cov=jk_cov,
        N=N,
        K=K
    )

    gamma_a, _, _ = wald_intercept_homogeneity_jackknife(
        eta=eta_hat,
        jk_cov=jk_cov,
        N=N,
        K=K
    )

    gamma_lam, _, _ = wald_slope_homogeneity_jackknife(
        eta=eta_hat,
        jk_cov=jk_cov,
        N=N,
        K=K
    )

```

```

gamma_full_list.append(gamma_full)
gamma_a_list.append(gamma_a)
gamma_lam_list.append(gamma_lam)

if (rep+1) % 2 == 0 and verbose:
    print(f"MC rep {rep+1}/{MC_REPS} complete.")

return np.array(gamma_full_list), np.array(gamma_a_list), np.
↪array(gamma_lam_list)

```

```

[ ]: gamma_full, gamma_a, gamma_lam = run_mc_jackknife(
    N=20, K=3, R=1, T=200,
    MC_REPS=100,
    verbose=True
)

def type1_error(g):
    return np.mean(np.abs(g) > 1.96)

```

```

[79]: print("Type-I error (full test):", type1_error(gamma_full))
      print("Type-I error (alpha test):", type1_error(gamma_a))
      print("Type-I error (lambda test):", type1_error(gamma_lam))

```

```

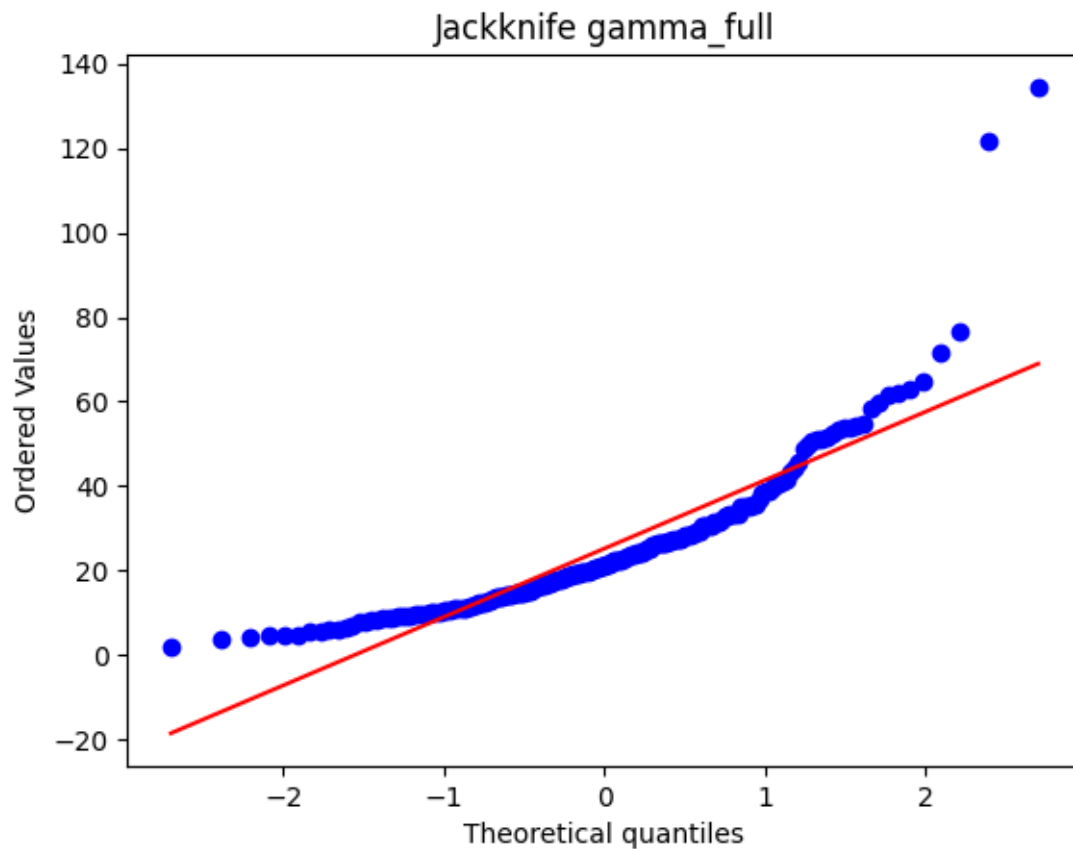
Type-I error (full test): 0.995
Type-I error (alpha test): 0.105
Type-I error (lambda test): 0.61

```

```

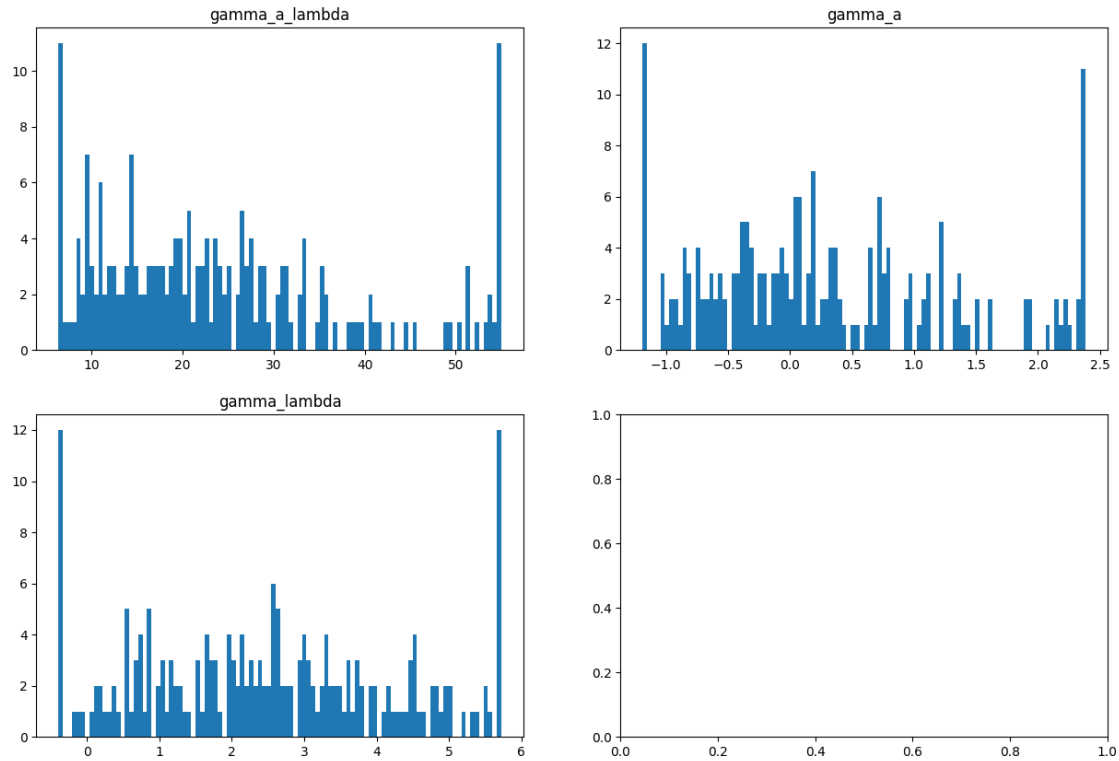
[60]: stats.probplot(gamma_full, dist="norm", plot=plt)
      plt.title("Jackknife gamma_full")
      plt.show()

```



```
[58]: fig, axes = plt.subplots(2, 2, figsize = (15, 10))
      axes = np.ravel(axes)
      labels = ['gamma_a_lambda', 'gamma_a', 'gamma_lambda']

      for i, test_statistic in enumerate([gamma_full, gamma_a, gamma_lam]):
          axes[i].hist(
              utils.clean(test_statistic),
              bins = 100,
          )
          axes[i].set_title(labels[i])
```



```
[87]: gamma_h1_full, gamma_h1_a, gamma_h1_lam = run_mc_jackknife(
      N=20, K=3, R=1, T=200,
      MC_REPS=200,
      verbose=True,
      heterogeneity_strength=1.0
    )

    def reject_rate(g):
        return np.mean(np.abs(g) > 1.96)

    print("Power (full test):", reject_rate(gamma_h1_full))
    print("Power (alpha test):", reject_rate(gamma_h1_a))
    print("Power (lambda test):", reject_rate(gamma_h1_lam))
```

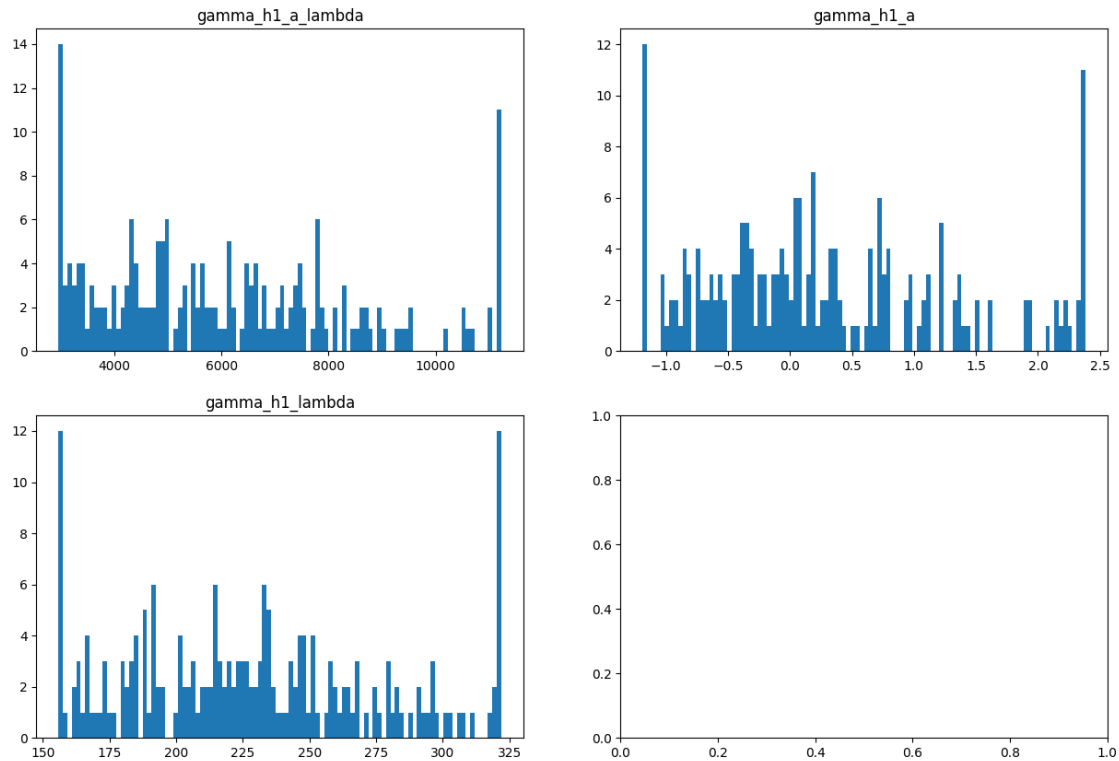
```
MC rep 2/200 complete.
MC rep 4/200 complete.
MC rep 6/200 complete.
MC rep 8/200 complete.
MC rep 10/200 complete.
MC rep 12/200 complete.
MC rep 14/200 complete.
MC rep 16/200 complete.
```

MC rep 18/200 complete.
MC rep 20/200 complete.
MC rep 22/200 complete.
MC rep 24/200 complete.
MC rep 26/200 complete.
MC rep 28/200 complete.
MC rep 30/200 complete.
MC rep 32/200 complete.
MC rep 34/200 complete.
MC rep 36/200 complete.
MC rep 38/200 complete.
MC rep 40/200 complete.
MC rep 42/200 complete.
MC rep 44/200 complete.
MC rep 46/200 complete.
MC rep 48/200 complete.
MC rep 50/200 complete.
MC rep 52/200 complete.
MC rep 54/200 complete.
MC rep 56/200 complete.
MC rep 58/200 complete.
MC rep 60/200 complete.
MC rep 62/200 complete.
MC rep 64/200 complete.
MC rep 66/200 complete.
MC rep 68/200 complete.
MC rep 70/200 complete.
MC rep 72/200 complete.
MC rep 74/200 complete.
MC rep 76/200 complete.
MC rep 78/200 complete.
MC rep 80/200 complete.
MC rep 82/200 complete.
MC rep 84/200 complete.
MC rep 86/200 complete.
MC rep 88/200 complete.
MC rep 90/200 complete.
MC rep 92/200 complete.
MC rep 94/200 complete.
MC rep 96/200 complete.
MC rep 98/200 complete.
MC rep 100/200 complete.
MC rep 102/200 complete.
MC rep 104/200 complete.
MC rep 106/200 complete.
MC rep 108/200 complete.
MC rep 110/200 complete.
MC rep 112/200 complete.

MC rep 114/200 complete.
MC rep 116/200 complete.
MC rep 118/200 complete.
MC rep 120/200 complete.
MC rep 122/200 complete.
MC rep 124/200 complete.
MC rep 126/200 complete.
MC rep 128/200 complete.
MC rep 130/200 complete.
MC rep 132/200 complete.
MC rep 134/200 complete.
MC rep 136/200 complete.
MC rep 138/200 complete.
MC rep 140/200 complete.
MC rep 142/200 complete.
MC rep 144/200 complete.
MC rep 146/200 complete.
MC rep 148/200 complete.
MC rep 150/200 complete.
MC rep 152/200 complete.
MC rep 154/200 complete.
MC rep 156/200 complete.
MC rep 158/200 complete.
MC rep 160/200 complete.
MC rep 162/200 complete.
MC rep 164/200 complete.
MC rep 166/200 complete.
MC rep 168/200 complete.
MC rep 170/200 complete.
MC rep 172/200 complete.
MC rep 174/200 complete.
MC rep 176/200 complete.
MC rep 178/200 complete.
MC rep 180/200 complete.
MC rep 182/200 complete.
MC rep 184/200 complete.
MC rep 186/200 complete.
MC rep 188/200 complete.
MC rep 190/200 complete.
MC rep 192/200 complete.
MC rep 194/200 complete.
MC rep 196/200 complete.
MC rep 198/200 complete.
MC rep 200/200 complete.
Power (full test): 1.0
Power (alpha test): 0.105
Power (lambda test): 1.0


```
[88]: fig, axes = plt.subplots(2, 2, figsize = (15, 10))
      axes = np.ravel(axes)
      labels = ['gamma_h1_a_lambda', 'gamma_h1_a', 'gamma_h1_lambda']

      for i, test_statistic in enumerate([gamma_h1_full, gamma_h1_a, gamma_h1_lam]):
          axes[i].hist(
              utils.clean(test_statistic),
              bins = 100,
          )
          axes[i].set_title(labels[i])
```



1.4 Jackknife Attempt #2

```
[76]: def mc_null_jackknife_theoretical_avar_tests(
      N=20, K=3, R=1, T=200,
      MC_REPS=50,
      n_iter_est=400,
      seed=1234,
      heterogeneity_strength=0.0
      ):
      rng = np.random.default_rng(seed)
```

```

gamma_full_obs      = np.zeros(MC_REPS)
gamma_alpha_obs     = np.zeros(MC_REPS)
gamma_lambda_obs    = np.zeros(MC_REPS)

gamma_full_resample  = np.zeros((MC_REPS, T))
gamma_alpha_resample = np.zeros((MC_REPS, T))
gamma_lambda_resample = np.zeros((MC_REPS, T))

for mc in range(MC_REPS):
    print(f"Processing {mc}", end = '\r')
    rep_seed = rng.integers(1_000_000_000)

    # Simulate under the NULL
    beta_true, r, realized_cov, residuals, G_true, beta_star_true,
    ↪ lambda_true = utils.simulate_dgp(
        N=N, K=K, R=R, T=T,
        heterogeneity_strength=heterogeneity_strength,
        seed=rep_seed
    )

    # FULL-SAMPLE ESTIMATE
    eta_full, G_full, beta_star_full, obj_full = utils.
    ↪ iterative_convergence(
        beta_hat=beta_true,
        excess_returns=r,
        N=N, K=K, R=R, T=T,
        n_iter=n_iter_est,
        verbose=False
    )

    avar_full = utils.estimate_avar(
        beta_hat=beta_true,
        excess_returns=r,
        eta=eta_full,
        G=G_full,
        beta_star=beta_star_full,
        realized_covariance=realized_cov,
        residuals=residuals,
        N=N, K=K, R=R, T=T
    )

    # Full-sample test stats
    gamma_full_obs[mc] = utils.full_homogeneity_test(eta_full, avar_full,
    ↪ N, K, T)
    gamma_alpha_obs[mc] = utils.intercept_homogeneity_test(eta_full,
    ↪ avar_full, N, K, T)

```

```

        gamma_lambda_obs[mc] = utils.slope_homogeneity_test(eta_full,
↪avar_full, N, K, T)

    # jackknife resampling
    for t in range(T):

        mask = np.ones(T, dtype=bool)
        mask[t] = False

        beta_sub = beta_true[:, :, mask]    # (N, K, T-1)
        r_sub = r[:, mask]                  # (N, T-1)
        realized_cov_sub = realized_cov[:, :, mask]
        residuals_sub = residuals[:, :, mask]

        # Re-estimate parameters on the resampled dataset
        eta_jk, G_jk, beta_star_jk, obj_jk = utils.iterative_convergence(
            beta_hat=beta_sub,
            excess_returns=r_sub,
            N=N, K=K, R=R, T=T-1,
            n_iter=n_iter_est,
            verbose=False
        )

        avar_jk = utils.estimate_avar(
            beta_hat=beta_sub,
            excess_returns=r_sub,
            eta=eta_jk,
            G=G_jk,
            beta_star=beta_star_jk,
            realized_covariance=realized_cov_sub,
            residuals=residuals_sub,
            N=N, K=K, R=R, T=T-1
        )

        # Test statistics for leave-one-out
        gamma_full_resample[mc, t] = utils.full_homogeneity_test(eta_jk,
↪avar_jk, N, K, T-1)
        gamma_alpha_resample[mc, t] = utils.
↪intercept_homogeneity_test(eta_jk, avar_jk, N, K, T-1)
        gamma_lambda_resample[mc, t] = utils.slope_homogeneity_test(eta_jk,
↪avar_jk, N, K, T-1)

    return (
        gamma_full_obs, gamma_alpha_obs, gamma_lambda_obs,
        gamma_full_resample, gamma_alpha_resample, gamma_lambda_resample
    )

```

```
[69]: gamma_full_obs, gamma_alpha_obs, gamma_lambda_obs, \
      gamma_full_resample, gamma_alpha_resample, gamma_lambda_resample \
      = mc_null_jackknife_theoretical_avar_tests()
```

Size (full test): 0.1
 Size (alpha test): 0.08
 Size (lambda test): 0.06

```
[70]: def size_from_resamples_two_sided(g_obs, g_res):
      """
      g_obs: shape (MC_REPS,)
      g_res: shape (MC_REPS, T) - each row is T resampled statistics
      """

      lower = np.percentile(g_res, 2.5, axis=1) # per MC rep
      upper = np.percentile(g_res, 97.5, axis=1)

      reject = (g_obs < lower) | (g_obs > upper)

      return np.mean(reject)

      print("Size (full test):", size_from_resamples_two_sided(gamma_full_obs, \
      ↪gamma_full_resample))
      print("Size (alpha test):", size_from_resamples_two_sided(gamma_alpha_obs, \
      ↪gamma_alpha_resample))
      print("Size (lambda test):", size_from_resamples_two_sided(gamma_lambda_obs, \
      ↪gamma_lambda_resample))
```

Size (full test): 0.08
 Size (alpha test): 0.06
 Size (lambda test): 0.04

```
[81]: gamma_h1_full_obs, gamma_h1_alpha_obs, gamma_h1_lambda_obs, \
      gamma_h1_full_resample, gamma_h1_alpha_resample, gamma_h1_lambda_resample \
      = mc_null_jackknife_theoretical_avar_tests(
          heterogeneity_strength=10
      )
```

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```
[82]: print("When heterogeneity is 1.0")
      print("Power (full test):", size_from_resamples_two_sided(gamma_h1_full_obs, \
      ↪gamma_h1_full_resample))
      print("Power (alpha test):", size_from_resamples_two_sided(gamma_h1_alpha_obs, \
      ↪gamma_h1_alpha_resample))
      print("Power (lambda test):", \
      ↪size_from_resamples_two_sided(gamma_h1_lambda_obs, gamma_h1_lambda_resample))
```

When heterogeneity is 1.0
 Power (full test): 0.02

Power (alpha test): 0.02
Power (lambda test): 0.02