Statistics 2

Analysis of Variance (ANOVA)

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Overview

Comparing means

Two-sample t test More than two groups — ANOVA Partitioning the variance F-test

Example

Computing CIs for a group's mean

Literature for this lecture

Read:

Agresti, Section 12.3

Comparing two means

Regression with code variables (recall Lecture 8).

Two groups:

- ▶ Group₁ $\sim \mathcal{N}(\mu_1, \sigma)$
- Group₂ $\sim \mathcal{N}(\mu_2, \sigma)$
- \triangleright Same σ assumed
- Sample size n_1 and n_2
- $\vdash \mathcal{H}_0 : \mu_1 = \mu_2$

Coding: 0s for Group 1; 1s for Group 2.

Thus:

$$\mu_1 = \beta_0$$

$$\mu_2 = \beta_0 + \beta_1$$

This implies $\beta_1 = \mu_2 - \mu_1$.

Comparing two means

- $\beta_1 = \mu_2 \mu_1.$
- \blacktriangleright $\mathcal{H}_0: \mu_1 = \mu_2$ is equivalent to $\mathcal{H}_0: \beta_1 = 0$.

Testing \mathcal{H}_0 is done through

$$t = \frac{\bar{y}_2 - \bar{y}_1}{SE_{b_1}}$$

with $n_1 + n_2 - 2$ degrees of freedom.

Conceptually:

$$t = \frac{\text{Distance between groups}}{\text{Variability within groups}}$$

Comparing two means

Another approach to compare two means: The t-test.

Test statistic:

$$t = \frac{\overline{y}_2 - \overline{y}_1}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

Conceptually:

$$t = \frac{\text{Distance between groups}}{\text{Variability within groups}}$$

$$SE_{b_1} = s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

so both t-tests are the same.

t-test and regression with a dummy variable are the same!

Example

- Data from Moore, McCabe, & Craig.
- Reading performance in two groups of pupils: One with and one without 'directed reading activities'.
- ▶ Sample sizes $n_1 = 21$, $n_2 = 23$.

| | Unstandardized | Standard Error | Standardized | t | р |
|-------------|----------------|----------------|--------------|--------|--------|
| (Intercept) | 51.476 | 3.175 | | 16.211 | < .001 |
| Group | -9.954 | 4.392 | -0.330 | -2.267 | 0.029 |
| Огоир | | 4.552 | | 2.201 | |

| | t | df | р |
|-------|-------|-------|-------|
| Group | 2.267 | 42.00 | 0.029 |

Apart from sign issues, due to irrelevant coding choices, both approaches are mathematically equivalent.

Comparing means: More than two groups

What to do when more than two groups need to be compared? Again, there are two approaches:

Regression with multiple dummy variables

ANOVA: ANalysis Of VAriance

Approaches mathematically equivalent.

Both are common within social sciences, thus important to be able to work with both.

Principle:

Study differences in the means of g independent groups.

► Test:

$$\mathcal{H}_0: \mu_1 = \mu_2 = \cdots = \mu_g.$$

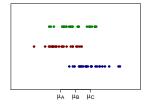
versus

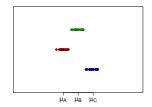
 $\mathcal{H}_{\it a}$: Not all μ 's are equal.

Procedure:

Compare the between and the within group variances using the \emph{F} -test.

Why call it Analysis of Variance when comparing means?





Distance between groups: relative to distance within groups.

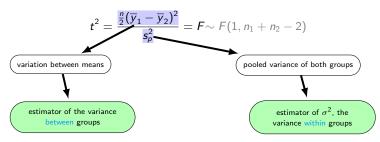
$$F = \frac{\text{Variability between groups}}{\text{Variability within groups}}$$

The *t*-test is a special kind of ANOVA:

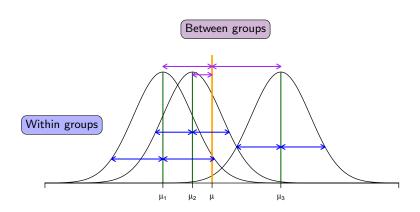
► Two-sample *t* test:

$$t = \frac{\overline{y}_2 - \overline{y}_1}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t(n_1 + n_2 - 2)$$

Look at t^2 (and assume $n_1 = n_2 = n$ for simplicity):



Split total variance: Between groups and within groups



ANOVA: Partitioning the variance

Split the total variance in two parts:

- ▶ A part that can be explained by differences between groups.
- A part that remains unexplained within groups.

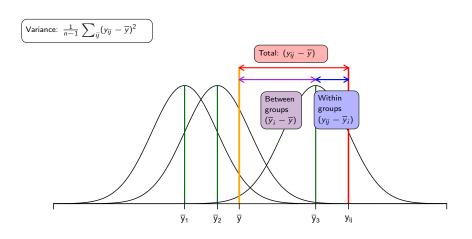
SS = sum of squares

$$\sum_{ij}(y_{ij}-\overline{y})^2=\sum_{ij}(\overline{y}_i-\overline{y})^2+\sum_{ij}(y_{ij}-\overline{y}_i)^2$$

- i indexes the groups.
- j indexes the persons in a group.

- $y_{ij} = \text{observation of person } j \text{ in group } i.$
- \overline{y}_i = mean of the DV y in group i. (i.e., over all persons in group i)
- $\overline{y} = \text{overall, or grand, mean of } y.$ (i.e., over all persons in all groups)

ANOVA: Partitioning the variance



With g groups:

$$\underbrace{\sum_{ij} (y_{ij} - \overline{y})^2}_{TSS} = \underbrace{\sum_{ij} (\overline{y}_i - \overline{y})^2}_{GSS} + \underbrace{\sum_{ij} (y_{ij} - \overline{y}_i)^2}_{RSS}$$

Convert SS's in variances: Divide by degrees of freedom (df)

Mean Squares (MS)

| | | Total | Group | Residual | |
|---|----|----------------|--------------------|---------------------------|-----------|
| S | S | TSS | GSS | RSS | |
| d | lf | df = n - 1 | $df_1 = g-1$ | $df_2 = n - g$ | |
| М | IS | TMS = TSS/df | $GMS {=} GSS/df_1$ | RMS=RSS/df ₂ \ | |
| | | Variance in y! | | $s_p^2 = pooled$ | variance! |

Hypotheses:

$$\mathcal{H}_0: \mu_1 = \mu_2 = \cdots = \mu_g.$$

 $\mathcal{H}_a: \text{Not all } \mu$'s are equal.

► Test statistic:

$$F = \frac{\mathsf{GMS}}{\mathsf{RMS}} = \frac{\mathsf{GSS}/\mathsf{df}_1}{\mathsf{RSS}/\mathsf{df}_2}.$$

If \mathcal{H}_0 holds: $F \approx 1$.

Q: Why?

A: Because:

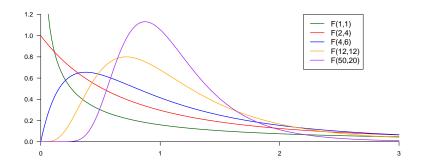
- 1. RMS = s_p^2 always estimates σ^2 , the common group variance.
- 2. Under \mathcal{H}_0 , GMS also estimates σ^2 .
- 3. Hence, under \mathcal{H}_0 , the F ratio is ≈ 1 .
- (If \mathcal{H}_0 does not hold: F > 1.)

Conclusion: Reject \mathcal{H}_0 is F is too large (i.e., $F \gg 1$).

Q: But how large need statistic *F* be?

A: Use the sampling distribution.

$$F \sim F(g-1, n-g)$$



| Source | SS | df | MS | F |
|----------|---|-----|---------------------|---------|
| Group | $\sum_{ij} (\overline{y}_i - \overline{y})^2$ | g-1 | GSS/df ₁ | CMC/DMC |
| Residual | $\sum_{ij}(y_{ij}-\overline{y}_i)^2$ | | RSS/df_2 | GMS/RMS |
| Total | $\sum_{ij}(y_{ij}-\overline{y})^2$ | n-1 | | |

Example: Directed Reading Activities

| Cases | Sum of Squares | df | Mean Square | F | р |
|----------|----------------|----|-------------|-------|-------|
| group | 1088 | 1 | 1087.8 | 5.137 | 0.029 |
| Residual | 8893 | 42 | 211.7 | | |

$$\left(\mathsf{Reject} \,\, \mathcal{H}_0 : \mu_1 = \mu_2
ight)$$

Recall: t-test provided t = 2.267, p = .029.

Indeed, $\sqrt{5.137} = 2.267$: t and F-test equivalent.

Example with more than 2 groups

James et al. (2015) studied whether playing a computer game (Tetris) could prevent intrusive memories (flashbacks) related to a traumatic event from occurring, via a reactivation-reconsolidation mechanism¹.

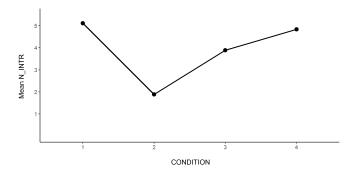
- ► Dependent variable
 - $y = N_INTR = Number of intrusive memories over the next seven days$
- ► Factor

CONDITION, with four levels:

- ▶ 1 = No-task control
- 2 = Reactivation + Tetris
- ▶ 3 = Tetris only
- ► 4 = Reactivation only

¹See Lecture 8

| | | N_II | NTR | |
|----------------|-------|-------|-------|-------|
| | 1 | 2 | 3 | 4 |
| Valid | 18 | 18 | 18 | 18 |
| Mean | 5.111 | 1.889 | 3.889 | 4.833 |
| Std. Deviation | 4.227 | 1.745 | 2.888 | 3.330 |



Question: Is there an effect of CONDITION on N_INTR?
In week 8, we studied this using code variables. Today: ANOVA

Example: Preventing flashbacks

Recall from Lecture 8: Regression with 3 code variables.

Regression output

| | Unstandardized | Standard Error | Standardized | t | р |
|-------------|----------------|----------------|--------------|--------|--------|
| (Intercept) | 4.833 | 0.749 | | 6.457 | < .001 |
| z1 | 0.278 | 1.059 | 0.036 | 0.262 | 0.794 |
| z2 | -2.944 | 1.059 | -0.382 | -2.781 | 0.007 |
| z3 | -0.944 | 1.059 | -0.123 | -0.892 | 0.375 |

- These tests are indicative of specific contrasts
- Overall model fit assessed through R²
- ▶ JASP provided $R^2 = .149$
- F = $(R^2/g)/((1-R^2)/(n-g)) = 3.795$ with p = .014

Example: Preventing flashbacks

ANOVA-approach

 $\mathcal{H}_0: \mu_1 = \mu_2 = \mu_3 = \mu_4.$

| ANOVA table | ΑN | IO | /A | tal | Ы | e |
|-------------|----|----|----|-----|---|---|
|-------------|----|----|----|-----|---|---|

| Model | | Sum of Squares | df | Mean Square | F | р |
|-------|------------|----------------|----|-------------|-------|-------|
| 1 | Regression | 114.8 | 3 | 38.27 | 3.795 | 0.014 |
| | Residual | 685.8 | 68 | 10.09 | | |
| | Total | 800.7 | 71 | | | |

Conclusions:

- ▶ Reject \mathcal{H}_0 .
- ▶ Both the *F*-test in ANOVA and in regression with code variables are equivalent (whatever the choice of dummy coding!).

The Kruskal-Wallis test

In ANOVA one makes three assumptions:

- 1. Independent observations.
- 2. Homogeneity: Variance in each group is equal.
- 3. Normality: Each group is normally distributed.

In Lecture 7 we learned how to check these assumptions and the consequences of violations.

Non-parametric alternative to ANOVA: The Kruskal-Wallis test.

Does not assume normality nor homogeneity.

You don't need to know technical details of KW, just be able to work with it.

 \mathcal{H}_0 : The distribution of observations in each group is identical.

 \mathcal{H}_1 : The distribution of observations in each group is not identical.

| Factor | Statistic | df | р |
|-----------|-----------|----|-------|
| Condition | 13.56 | 3 | 0.004 |

p = .004 thus reject H_0 . Significant differences between groups.

This is not in the textbook but it is important!!

Two ways to compute CIs for group means:

▶ Based on the pooled SD, s_p. (ideal when homoscedasticity is met)

CI for group
$$i = \overline{y}_i \pm t_{n-g}^* \frac{s_p}{\sqrt{n_i}}$$

$$n = \text{total sample size}$$

Based on the groups SD, s_i. (when homoscedasticity is violated)

CI for group
$$i = \overline{y}_i \pm t_{n_i-1}^* \frac{s_i}{\sqrt{n_i}}$$

$$n_i = \text{group sample size}$$

For the next lecture

Contents:

Analysis of Variance (ANOVA): Two-way ANOVA

Read: Section 12.4.