

STAT 510 Homework 4

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1. (a)

$$\begin{bmatrix} \hat{\mathbf{y}} \\ \mathbf{y} - \hat{\mathbf{y}} \end{bmatrix} = \begin{bmatrix} \mathbf{P}_X \mathbf{y} \\ (\mathbf{I} - \mathbf{P}_X) \mathbf{y} \end{bmatrix} = \begin{bmatrix} \mathbf{P}_X \\ \mathbf{I} - \mathbf{P}_X \end{bmatrix} \mathbf{y}$$

We know $\mathbf{y} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I})$, thus $\begin{bmatrix} \hat{\mathbf{y}} \\ \mathbf{y} - \hat{\mathbf{y}} \end{bmatrix} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, where

$$\begin{aligned} \boldsymbol{\mu} &= \begin{bmatrix} \mathbf{P}_X \\ \mathbf{I} - \mathbf{P}_X \end{bmatrix} \mathbf{X}\boldsymbol{\beta} = \begin{bmatrix} \mathbf{P}_X \mathbf{X}\boldsymbol{\beta} \\ (\mathbf{I} - \mathbf{P}_X) \mathbf{X}\boldsymbol{\beta} \end{bmatrix} = \begin{bmatrix} \mathbf{X}\boldsymbol{\beta} \\ \mathbf{0} \end{bmatrix} \\ \boldsymbol{\Sigma} &= \begin{bmatrix} \mathbf{P}_X \\ \mathbf{I} - \mathbf{P}_X \end{bmatrix} \sigma^2 \mathbf{I} \begin{bmatrix} \mathbf{P}_X^T & (\mathbf{I} - \mathbf{P}_X)^T \end{bmatrix} \\ &= \begin{bmatrix} \mathbf{P}_X \mathbf{P}_X^T & \mathbf{P}_X (\mathbf{I} - \mathbf{P}_X)^T \\ (\mathbf{I} - \mathbf{P}_X) \mathbf{P}_X^T & (\mathbf{I} - \mathbf{P}_X) (\mathbf{I} - \mathbf{P}_X)^T \end{bmatrix} \\ &= \begin{bmatrix} \mathbf{P}_X & \mathbf{0} \\ \mathbf{0} & \mathbf{I} - \mathbf{P}_X \end{bmatrix} \end{aligned}$$

Hence

$$\begin{bmatrix} \hat{\mathbf{y}} \\ \mathbf{y} - \hat{\mathbf{y}} \end{bmatrix} \sim N \left(\begin{bmatrix} \mathbf{X}\boldsymbol{\beta} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{P}_X & \mathbf{0} \\ \mathbf{0} & \mathbf{I} - \mathbf{P}_X \end{bmatrix} \right)$$

(b) We have

$$\hat{\mathbf{y}}^T \hat{\mathbf{y}} = \mathbf{y}^T \mathbf{P}_X^T \mathbf{P}_X \mathbf{y} = \mathbf{y}^T \mathbf{P}_X \mathbf{y}$$

So let $\mathbf{A} = \frac{1}{\sigma^2} \mathbf{P}_X$, we have $\text{rank}(\mathbf{A}) = \text{rank}(\mathbf{P}_X) = \text{rank}(\mathbf{X}) = r$. Also, $\boldsymbol{\Sigma} = \sigma^2 \mathbf{I}$ is positive definite, and

$$\mathbf{A} \boldsymbol{\Sigma} \mathbf{A} \boldsymbol{\Sigma} = \frac{1}{\sigma^2} \mathbf{P}_X \sigma^2 \mathbf{I} \frac{1}{\sigma^2} \mathbf{P}_X \sigma^2 \mathbf{I} = \frac{1}{\sigma^2} \mathbf{P}_X \sigma^2 \mathbf{I} = \mathbf{A} \boldsymbol{\Sigma}$$

From $\mathbf{y} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I})$ we have

$$\mathbf{y}^T \mathbf{A} \mathbf{y} = \frac{1}{\sigma^2} \mathbf{y}^T \mathbf{P}_X \mathbf{y} \sim \chi_r^2 \left(\boldsymbol{\beta}^T \mathbf{X}^T \frac{1}{\sigma^2} \mathbf{P}_X \mathbf{X} \boldsymbol{\beta} / 2 \right) = \chi_r^2 \left(\frac{1}{\sigma^2} \boldsymbol{\beta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\beta} / 2 \right)$$

Thus

$$\hat{\mathbf{y}}^T \hat{\mathbf{y}} \sim \sigma^2 \chi_r^2 \left(\frac{1}{\sigma^2} \boldsymbol{\beta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\beta} / 2 \right)$$

2. (a)

$$\mathbf{X} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 \\ 1 & 1 & 1 & 1 \\ 1 & 1 & 1 & 1 \\ 1 & 1 & 1 & 1 \end{bmatrix}$$

(b) By Gauss-Markov Theorem, BLUE here is OLSE, thus

$$\text{BLUE}(\beta_4) = \text{OLSE}(\beta_4) = \text{OLSE}(\beta_1 + \beta_2 + \beta_3 + \beta_4) - \text{OLSE}(\beta_1 + \beta_2 + \beta_3) = 26.3 - 22.8 = 3.5$$

(c)

$$\hat{\text{Var}}(\beta_4) = \hat{\sigma}^2 \mathbf{C}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{C}^T = \mathbf{A} \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{A}^T = \mathbf{A} \mathbf{P}_{\mathbf{X}} \mathbf{A}^T$$

The matrix \mathbf{A} we use here is $\begin{bmatrix} 0 & \cdots & 0 & -1 & 0 & 0 & 0 & 1 \end{bmatrix}$, which is literally take the last row of a matrix and minus the last 5th row. And in this model, the estimated response variable is just sample means of each treatment. Then we have

$$\mathbf{A} \mathbf{P}_{\mathbf{X}} = \begin{bmatrix} 0 & \cdots & 0 & -1/2 & -1/2 & 1/4 & 1/4 & 1/4 & 1/4 \end{bmatrix}$$

Then we have $\mathbf{A} \mathbf{P}_{\mathbf{X}} \mathbf{A}^T = \frac{1}{4} + \frac{1}{2} = \frac{3}{4}$.

And

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^4 s_i^2 (n_i - 1)}{\sum_{i=1}^4 n_i - 4} = 3.428571$$

Then

$$se(\hat{\beta}_4) = \sqrt{\frac{3}{4} \hat{\sigma}^2} = 1.603576$$

3. (a)

$$\begin{aligned} \mathbf{X}^T \mathbf{X} &= \begin{bmatrix} \mathbf{1}^T \\ \mathbf{x}^T \end{bmatrix} \begin{bmatrix} \mathbf{1} & \mathbf{x} \end{bmatrix} \\ &= \begin{bmatrix} \mathbf{1}^T \mathbf{1} & \mathbf{1}^T \mathbf{x} \\ \mathbf{x}^T \mathbf{1} & \mathbf{x}^T \mathbf{x} \end{bmatrix} \\ &= \begin{bmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{bmatrix} \end{aligned}$$

Take the inverse, we have

$$(\mathbf{X}^T \mathbf{X})^{-1} = \begin{bmatrix} \frac{\sum_{i=1}^n x_i^2}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} & \frac{-\sum_{i=1}^n x_i}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} \\ \frac{-\sum_{i=1}^n x_i}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} & \frac{n}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} \end{bmatrix}$$

And

$$\mathbf{X}^T \mathbf{y} = \begin{bmatrix} \mathbf{1}^T \\ \mathbf{x}^T \end{bmatrix} \mathbf{y} = \begin{bmatrix} \mathbf{1}^T \mathbf{y} \\ \mathbf{x}^T \mathbf{y} \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix}$$

Then

$$\begin{aligned}
\hat{\beta} &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \\
&= \begin{bmatrix} \frac{\sum_{i=1}^n x_i^2}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} & \frac{-\sum_{i=1}^n x_i}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} \\ \frac{-\sum_{i=1}^n x_i}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} & \frac{n}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} \end{bmatrix} \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix} \\
&= \begin{bmatrix} \frac{(\sum_{i=1}^n x_i^2)(\sum_{i=1}^n y_i) - (\sum_{i=1}^n x_i)(\sum_{i=1}^n x_i y_i)}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} \\ \frac{n \sum_{i=1}^n x_i y_i - (\sum_{i=1}^n x_i)(\sum_{i=1}^n y_i)}{n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2} \end{bmatrix} \\
&= \begin{bmatrix} \frac{x^2 \bar{y} - \bar{x} \bar{xy}}{x^2 - (\bar{x})^2} \\ \frac{\bar{xy} - \bar{x} \bar{y}}{x^2 - (\bar{x})^2} \end{bmatrix}
\end{aligned}$$

(b) $[\mathbf{1} \quad \mathbf{x} - \bar{x}\mathbf{1}] = [\mathbf{1} \quad \mathbf{x}] \begin{bmatrix} 1 & -\bar{x} \\ 0 & 1 \end{bmatrix}$. Thus we choose $\mathbf{B}^{-1} = \begin{bmatrix} 1 & -\bar{x} \\ 0 & 1 \end{bmatrix}$

(c) We have

$$\mathbf{W}^T \mathbf{W} = \begin{bmatrix} \mathbf{1}^T \\ (\mathbf{x} - \bar{x}\mathbf{1})^T \end{bmatrix} [\mathbf{1} \quad \mathbf{x} - \bar{x}\mathbf{1}] = \begin{bmatrix} n & 0 \\ 0 & \sum_{i=1}^n (x_i - \bar{x})^2 \end{bmatrix}$$

Also,

$$\mathbf{W}^T \mathbf{y} = \begin{bmatrix} \mathbf{1}^T \mathbf{y} \\ (\mathbf{x} - \bar{x}\mathbf{1})^T \mathbf{y} \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n (x_i - \bar{x}) y_i \end{bmatrix}$$

Hence

$$\begin{aligned}
\hat{\alpha} &= (\mathbf{W}^T \mathbf{W})^{-1} \mathbf{W}^T \mathbf{y} \\
&= \begin{bmatrix} 1/n & 0 \\ 0 & 1/\sum_{i=1}^n (x_i - \bar{x})^2 \end{bmatrix} \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n (x_i - \bar{x}) y_i \end{bmatrix} \\
&= \begin{bmatrix} \frac{\sum_{i=1}^n y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \end{bmatrix} \\
&= \begin{bmatrix} \bar{y} \\ \frac{x\bar{y} - \bar{x}\bar{y}}{x^2 - (\bar{x})^2} \end{bmatrix}
\end{aligned}$$

(d)

$$\begin{aligned}
\hat{\beta} &= \begin{bmatrix} 1 & -\bar{x} \\ 0 & 1 \end{bmatrix} \begin{bmatrix} \frac{\sum_{i=1}^n y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \end{bmatrix} \\
&= \begin{bmatrix} \frac{\sum_{i=1}^n y_i / n - \bar{x} \sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \end{bmatrix}
\end{aligned}$$

(e)

$$\begin{aligned}
\hat{\beta} &= \begin{bmatrix} 1 & -\bar{x} \\ 0 & 1 \end{bmatrix} \begin{bmatrix} \bar{y} \\ \frac{x\bar{y} - \bar{x}\bar{y}}{x^2 - (\bar{x})^2} \end{bmatrix} \\
&= \begin{bmatrix} \frac{x^2 \bar{y} - (\bar{x})^2 \bar{y} - \bar{x} \bar{xy} + (\bar{x})^2 \bar{y}}{x^2 - (\bar{x})^2} \\ \frac{x\bar{y} - \bar{x}\bar{y}}{x^2 - (\bar{x})^2} \end{bmatrix} \\
&= \begin{bmatrix} \frac{x^2 \bar{y} - \bar{x} \bar{xy}}{x^2 - (\bar{x})^2} \\ \frac{x\bar{y} - \bar{x}\bar{y}}{x^2 - (\bar{x})^2} \end{bmatrix}
\end{aligned}$$

It matches (a).

4. (a) The null hypothesis tested is the difference in score students gave for different perceived instructor when ruling out any possible impact the actually instructor gave is zero.
- (b) The matrix $(\mathbf{X}_R^T \mathbf{X}_R)^-$ we got is

$$\begin{bmatrix} 0.1 & -0.1 & -0.1 & 0.1 \\ -0.1 & 0.2 & 0.1 & -0.2 \\ -0.1 & 0.1 & 0.2 & -0.2 \\ 0.1 & -0.2 & -0.2 & 0.377 \end{bmatrix}$$

With we can have

$$\hat{\sigma}^2 = (se(\text{Intercept}))^2 = 0.0507$$

To calculate the variance of main effect of perceived instructor, we use $\mathbf{C} = [0 \ 0 \ 1 \ 1/2]$

Hence

$$\hat{\text{Var}}(\text{main effect of perceived instructor}) = \hat{\sigma}^2 \mathbf{C} (\mathbf{X}_R^T \mathbf{X}_R)^- \mathbf{C}^T = 0.0047775$$

Thus the standard error is $\sqrt{0.0047775} = 0.0691$.

Also, we have the estimate of main effect of perceived instructor to be $[0 \ 0 \ 1 \ 1/2] \hat{\boldsymbol{\beta}} = 0.8700 - 0.1831/2 = 0.77845$. And we know $t_{33-4, 0.975} = t_{29, 0.975} = 2.04523$. Thus the 95% confidence interval

$$[0.77845 - 0.0691 \times 2.04523, 0.77845 + 0.0691 \times 2.04523] = [0.6371, 0.9198]$$