Random Intervals and Confidence Intervals (Ch. 6.1)

Yifan 7hu

Iowa State University

Random Intervals and Confidence Intervals (Ch. 6.1)

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Motivation

Random Intervals

Outline

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Random Interval

Confidence Intervals $(n \ge 25, \sigma)$

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Confidence Intervals ($n \ge 25$, σ known)

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Confidence ntervals $n \geq 25, \ \sigma$ known)

- Statistical inference: using data from the sample to draw formal conclusions about the population
 - Point estimation (confidence intervals): estimating population parameters and specifying the degree of precision of the estimate.
 - Hypothesis testing: testing the validity of statements about the population that are framed in terms of parameters.

- We want information on a population. For example:
 - ► True mean breaking strength of a kind of wire rope.
 - True mean fill weight of food jars.
 - ▶ True mean instrumental drift of a kind of scale.
 - Average number of cycles to failure of a kind of spring.
- ▶ We can use point estimates:
 - For example: if we measure breaking strengths (in tons) of 6 wire ropes as 5, 3, 7, 3,10, and 1, we might estimate the true mean breaking strength $\mu \approx \overline{x} = \frac{5+3+7+3+10+1}{6} = 4.83$ tons.
- Or, we can use interval estimates:
 - μ is likely to be inside the interval (4.83 2, 4.83 + 2) = (2.83, 6.83).
 - We are confident that the true mean breaking strength, μ , is somewhere in (2.83, 6.83). But how confident can we be?

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- ► A **random interval** is an interval on the real line with a random variable at one or both of the endpoints.
- Examples:
 - $(Z-2,Z+2), Z \sim N(0,1)$
 - \triangleright (Z, ∞)
 - ▶ $(-\infty, X)$, $X \sim N(-2, 9)$
 - $(T s \cdot t_{7,0.975}, T + s \cdot t_{7,0.975}), T \sim t_7$
 - $(X \sigma \cdot z_{1-\alpha}, \infty)$, $X \sim N(5, \sigma^2)$, $0 < \alpha < 1$.
- ightharpoonup Random intervals take into account the uncertainty in the measurement of a true mean, μ .

- Let Z be a measure of instrumental drift of a random voltmeter that comes out of a certain factory. Say $Z \sim N(0,1)$.
- Define a random interval:

$$(Z-2, Z+2)$$

- ▶ What is the probability that -1 is inside the interval?
 - ▶ Equivalent to asking how likely it is that the drift of the next instrument is within 2 units of -1.

Intervals $(n \ge 25, \sigma)$

$$P(-1 \text{ in } (Z-2, Z+2)) = P(Z-2 < -1 < Z+2)$$

$$= P(Z-1 < 0 < Z+3)$$

$$= P(-1 < -Z < 3)$$

$$= P(-3 < Z < 1)$$

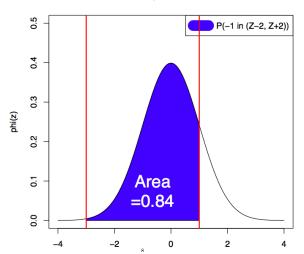
$$= P(Z \le 1) - P(Z \le -3)$$

$$= \Phi(1) - \Phi(-3)$$

$$= 0.84$$

Example: instrumental drift: the range of Z values for which -1 is in (Z-2,Z+2)

pdf of Z



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Confidence ntervals $n \geq 25, \ \sigma$

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Confidence Intervals $(n \ge 25, \sigma \text{known})$

Calculate:

- 1. $P(2 \text{ in } (X-1,X+1)), X \sim N(2,4)$
- 2. $P(6.6 \text{ in } (X-2,X+1)), X \sim N(7,2)$

Here, $0 < \alpha < 1$.

1.
$$X \sim N(2,4)$$

$$P(2 \in (X - 1, X + 1)) = P(X - 1 < 2 < X + 1)$$

$$= P(-1 < 2 - X < 1)$$

$$= P(-1 < X - 2 < 1)$$

$$= P\left(\frac{-1}{2} < \frac{X - 2}{2} < \frac{1}{2}\right)$$

$$= P(-0.5 < Z < 0.5)$$

$$= \Phi(0.5) - \Phi(-0.5)$$

$$= 0.69 - 0.31$$

$$= 0.38$$

Answers: random intervals

2. $X \sim N(7,2)$

$$P(6.6 \in (X - 2, X + 1)) = P(X - 2 < 6.6 < X + 1)$$

$$= P(-2 < 6.6 - X < 1)$$

$$= P(-1 < X - 6.6 < 2)$$

$$= P(-1.4 < X - 7 < 1.6)$$

$$= P\left(\frac{-1.4}{\sqrt{2}} < \frac{X - 7}{\sqrt{2}} < \frac{1.6}{\sqrt{2}}\right)$$

$$= P(-0.99 < Z < 1.13)$$

$$= \Phi(1.13) - \Phi(-0.99)$$

$$= 0.87 - 0.16$$

$$= 0.71$$

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Confidence Intervals $(n \geq 25, \sigma)$

- Let's say X_1, X_2, \ldots, X_n are iid with:
 - ▶ n > 25
 - ightharpoonup mean μ
 - variance σ^2
- The random interval, $(\overline{X} z_{1-\alpha} \frac{\sigma}{\sqrt{n}}, \infty)$, is useful for estimating μ $(0 < \alpha < 1)$.
- ▶ The interval contains μ with probability 1α .

$$\begin{split} P(\mu \in (\overline{X} - z_{1-\alpha} \frac{\sigma}{\sqrt{n}}, \, \infty)) \\ &= P\left(\overline{X} - z_{1-\alpha} \frac{\sigma}{\sqrt{n}} < \mu\right) \\ &= P\left(\overline{X} - \mu < z_{1-\alpha} \frac{\sigma}{\sqrt{n}}\right) \\ &= P\left(\frac{\overline{X} - \mu}{\sigma/\sqrt{n}} < z_{1-\alpha}\right) \\ &\approx P(Z < z_{1-\alpha}) \quad \text{(Central Limit Theorem)} \\ &= \Phi(z_{1-\alpha}) \\ &= 1 - \alpha \quad \text{(by the definition of } z_p) \end{split}$$

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Calculate:

- 1. $P(\mu \in (-\infty, \overline{X} + z_{1-\alpha} \frac{\sigma}{\sqrt{n}})), \overline{X} \sim N(\mu, \sigma^2)$
- 2. $P(\mu \in (\overline{X} z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}, \ \overline{X} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}})), \ X \sim N(\mu, \sigma^2)$

Remember the Central Limit Theorem:

$$rac{\overline{X} - \mu}{\sigma / \sqrt{n}} pprox \mathit{N}(0, 1)$$

Answers: abstract random intervals

1.

$$\begin{split} P\big(\mu \in & (-\infty, \ \overline{X} + z_{1-\alpha} \frac{\sigma}{\sqrt{n}})\big) \\ &= P\left(\mu < \overline{X} + z_{1-\alpha} \frac{\sigma}{\sqrt{n}}\right) \\ &= P\left(-z_{1-\alpha} \frac{\sigma}{\sqrt{n}} < \overline{X} - \mu\right) \\ &= P\left(-z_{1-\alpha} < \frac{\overline{X} - \mu}{\sigma/\sqrt{n}}\right) \\ &\approx P\left(-z_{1-\alpha} < Z\right) \quad \text{(Central Limit Theorem)} \\ &= 1 - P(Z \le -z_{1-\alpha}) \\ &= 1 - \Phi(-z_{1-\alpha}) \\ &= 1 - \Phi(z_{\alpha}) \quad \text{(by symmetry: N(0,1) pdf)} \\ &= 1 - \alpha \quad \text{(by the definition of } z_p) \end{split}$$

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2.

$$\begin{split} P(\mu \in (X - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}, \ \overline{X} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}})) \\ &= P\left(\overline{X} - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}} < \mu < \overline{X} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right) \\ &= P\left(-z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} < \mu - \overline{X} < z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right) \\ &= P\left(-z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} < \overline{X} - \mu < z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right) \\ &= P\left(-z_{1-\alpha/2} < \frac{\overline{X} - \mu}{\sigma/\sqrt{n}} < z_{1-\alpha/2}\right) \\ &\approx P(-z_{1-\alpha/2} < Z < z_{1-\alpha/2}) \quad \text{(Central Limit Theorem)} \\ &= \Phi(z_{1-\alpha/2}) - \Phi(-z_{1-\alpha/2}) \\ &= \Phi(z_{1-\alpha/2}) - \Phi(z_{\alpha/2}) \quad \text{(by symmetry: N(0,1) pdf)} \\ &= (1 - \frac{\alpha}{2}) - \frac{\alpha}{2} = 1 - \alpha \end{split}$$

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Confidence Intervals ($n \ge 25$, σ known)

- ▶ A $1-\alpha$ confidence interval for an unknown parameter is the finite realization of a random interval that contains that parameter with probability $1-\alpha$.
- ▶ 1α is called the **confidence level** of the interval.
- Example: for observations $x_1, x_2, \ldots x_n$ from random variables X_1, X_2, \ldots, X_n iid with $E(X_1) = \mu$, $Var(X_1) = \sigma^2$, a 1α confidence interval for μ is:

$$\left(\overline{x}-z_{1-\alpha/2}\frac{\sigma}{\sqrt{n}},\overline{x}+z_{1-\alpha/2}\frac{\sigma}{\sqrt{n}}\right)$$

which is a random draw from the random interval:

$$\left(\overline{X}-z_{1-\alpha/2}\frac{\sigma}{\sqrt{n}},\overline{X}+z_{1-\alpha/2}\frac{\sigma}{\sqrt{n}}\right)$$

Confidence intervals for μ : σ known, $n \ge 25$

▶ Two-sided $1 - \alpha$ confidence interval:

$$\left(\overline{x} - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}, \overline{x} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right)$$

▶ One-sided $1 - \alpha$ upper confidence interval:

$$\left(-\infty, \ \overline{x} + z_{1-\alpha} \frac{\sigma}{\sqrt{n}}\right)$$

▶ One-sided $1 - \alpha$ lower confidence interval:

$$\left(\overline{x}-z_{1-\alpha}\frac{\sigma}{\sqrt{n}}, \infty\right)$$

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Example: fill weight of jars

- Suppose a manufacturer fills jars of food using a stable filling process with a known standard deviation of $\sigma = 1.6g$.
- ▶ We take a sample of n = 47 jars and measure the sample mean weight $\overline{x} = 138.2$ g.
- A two-sided 90% confidence interval ($\alpha=0.1$) for the true mean weight μ is:

$$\left(\overline{x} - z_{1-0.1/2} \frac{\sigma}{\sqrt{n}}, \ \overline{x} + z_{1-0.1/2} \frac{\sigma}{\sqrt{n}}\right)$$

$$= \left(138.2 - z_{0.95} \frac{1.6}{\sqrt{47}}, \ 138.2 + z_{0.95} \frac{1.6}{\sqrt{47}}\right)$$

$$= (138.2 - 1.64 \cdot 0.23, \ 138.2 + 1.64 \cdot 0.23)$$

$$= (137.82, 138.58)$$

I could have also written the interval as:

$$138.2 \pm 0.38 \ g$$

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Interpreting the confidence interval: fill weight of jars

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- ▶ We are 90% confident that the true mean fill weight is between 137.82g and 138.58g.
- If we took 100 more samples of 47 jars each, roughly 90 of those samples would yield confidence intervals containing the true mean fill weight.
- These methods of interpretation generalize to all confidence intervals.

Example: fill weight of jars.

- What if we just want to be sure that the true mean fill weight is high enough?
- ► Then, we would use a one-side lower 90% confidence interval:

$$\left(\overline{x} - z_{1-\alpha} \frac{\sigma}{\sqrt{n}}, \infty\right) \\
= \left(138.2 - z_{1-\alpha} \frac{\sigma}{\sqrt{n}}, \infty\right) \\
= \left(138.2 - z_{0.9} \frac{1.6}{\sqrt{47}}, \infty\right) \\
= (138.2 - 1.28 \cdot 0.23, \infty) \\
= (137.91, \infty)$$

▶ We're 90% confident that the true mean fill weight is above 137.91 g.

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known)

- Consider a grinding process used to rebuild car engines, which involves grinding rod journals for engine crankshafts.
- Of interest is the deviation of the true mean rod journal diameter from the target diameter.
- ▶ Suppose the standard deviation of the individual differences from the target diameter is 0.7×10^{-4} in.
- ▶ 32 consecutive rod journals are ground, with a sample mean deviation of -0.16×10^{-4} in from the target diameter.
- ► Calculate and interpret a two-sided 95% confidence interval for the true mean deviation from the target diameter. Is there enough evidence that we're missing the target on average?

- $\alpha = 1 0.95 = 0.05$, n = 32, $\sigma = 0.7 \times 10^{-4}$, and $\overline{x} = -0.16 \times 10^{-4}$.
- ► Interval:

$$\begin{split} &\left(\overline{x}-z_{1-0.05/2}\frac{\sigma}{\sqrt{n}},\ \overline{x}+z_{1-0.05/2}\frac{\sigma}{\sqrt{n}}\right)\\ &=\left(-0.16\times10^{-4}-z_{0.975}\frac{0.7\times10^{-4}}{\sqrt{32}},\ -0.16\times10^{-4}+z_{0.975}\frac{0.7\times10^{-4}}{\sqrt{32}}\right)\\ &=\left(-0.16\times10^{-4}-1.96\cdot1.2\times10^{-5},\ -0.16\times10^{-4}+1.96\cdot1.2\times10^{-5}\right)\\ &=\left(-4.0\times10^{-5},7.5\times10^{-6}\right) \end{split}$$

- ▶ We are 95% confident that the true mean deviation from the target diameter of the rod journals is between -4.0×10^{-5} in and 7.5×10^{-6} in
- Since 0 is in the confidence interval, there is not enough evidence to conclude that the rod journal grinding process is off target.

- ▶ F. Willett, in the article *The Case of the Derailed Disk Drives* (Mechanical Engineering, 1988), discusses a study done to isolate the cause of *blink code A failure* in a model of Winchester hard disk drive.
- For each disk, the investigator measured the breakaway torque (in. oz.) required to loosen the drive's interrupter flag on the stepper motor shaft.
- Breakaway torques for 26 disk drives were recorded, with a sample mean of 11.5 in. oz.
- Suppose you know the true standard deviation of the breakaway torques is 5.1 in. oz.
- Calculate and interpret:
 - 1. A two-sided 90% confidence interval for the true mean breakaway torque of the relevant type of Winchester drive.
 - 2. An analogous two-sided 95% confidence interval.
 - 3. An analogous two-sided 99% confidence interval.
- Is there enough evidence to conclude that the mean breakaway torque is different from the factory's standard of 33.5 in. oz.?

- $\sigma = 5.1, \overline{x} = 11.5, n = 26.$
- ▶ All three confidence intervals have the form:

$$\begin{split} &\left(\overline{x} - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}, \ \overline{x} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right) \\ &= \left(11.5 - z_{1-\alpha/2} \frac{5.1}{\sqrt{26}}, \ 11.5 + z_{1-\alpha/2} \frac{5.1}{\sqrt{26}}\right) \\ &= \left(11.5 - 1.0002 \cdot z_{1-\alpha/2}, \ 11.5 + 1.0002 \cdot z_{1-\alpha/2}\right) \end{split}$$

- The confidence intervals are thus:
 - 1. 90% CI means $\alpha = 0.1$

$$\begin{aligned} &(11.5 - 1.0002 \cdot z_{1-0.1/2}, \ 11.5 + 1.0002 \cdot z_{1-0.1/2}) \\ &= (11.5 - 1.0002 \cdot z_{0.95}, \ 11.5 + 1.0002 \cdot z_{0.95}) \\ &= (11.5 - 1.0002 \cdot 1.64, \ 11.5 + 1.0002 \cdot 1.64) \\ &= (9.86, 13.14) \end{aligned}$$

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2. 95% CI means $\alpha = 0.05$

$$(11.5 - 1.0002 \cdot z_{1-0.05/2}, \ 11.5 + 1.0002 \cdot z_{1-0.05/2})$$

$$= (11.5 - 1.0002 \cdot z_{0.975}, \ 11.5 + 1.0002 \cdot z_{0.975})$$

$$= (11.5 - 1.0002 \cdot 1.96, \ 11.5 + 1.0002 \cdot 1.96)$$

$$= (9.54, 13.46)$$

3. 99% CI means $\alpha = 0.01$

$$(11.5 - 1.0002 \cdot z_{1-0.01/2}, 11.5 + 1.0002 \cdot z_{1-0.01/2})$$

$$= (11.5 - 1.0002 \cdot z_{0.995}, 11.5 + 1.0002 \cdot z_{0.995})$$

$$= (11.5 - 1.0002 \cdot 2.33, 11.5 + 1.0002 \cdot 2.33)$$

$$= (9.17, 13.83)$$

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- Notice: the confidence intervals get wider as the confidence level 1α increases.
- ▶ None of these confidence intervals contains the manufacturer's target of 33.5 in. oz., so there is significant evidence that the process misses this target.
- Hence, there is a design flaw in the manufacturing process of the disk drives that must be corrected.

- If you want to estimate the breakaway torque with a 2-sided, 95% confidence interval with ± 2.0 in. oz. of precision, what sample size would you need?
- ► The confidence interval is:

$$\begin{split} &\left(\overline{x} - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}, \ \overline{x} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right) \\ &= \left(11.5 - z_{1-0.05/2} \cdot \frac{5.1}{\sqrt{n}}, \ 11.5 + z_{1-0.05/2} \cdot \frac{5.1}{\sqrt{n}}\right) \\ &= \left(11.5 - z_{0.975} \cdot \frac{5.1}{\sqrt{n}}, \ 11.5 + z_{0.975} \cdot \frac{5.1}{\sqrt{n}}\right) \\ &= \left(11.5 - 1.96 \cdot 5.1 \cdot n^{-1/2}, 11.5 + 1.96 \cdot 5.1 \cdot n^{-1/2}\right) \\ &= \left(11.5 - 9.996 \cdot n^{-1/2}, 11.5 + 9.996 \cdot n^{-1/2}\right) \end{split}$$

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The interval precision (half-width) δ is:

$$\delta = \frac{1}{2} \left((11.5 + 9.996 \cdot n^{-1/2}) - (11.5 - 9.996 \cdot n^{-1/2}) \right)$$

= 9.996 \cdot n^{-1/2}

We require δ to be at most 2:

$$2.0 \le 9.996 \cdot n^{-1/2}$$
$$n > 25$$

▶ We would need a sample of 25 disk drives to meet a precision of ± 2.0 .