Inference for Simple Linear Regression (Ch. 9.1)

Yifan Zhu

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Iowa State University

A Review of Simple Linear Regression (Ch. 4)

Formalizing the Simple Linear Regression Model

Estimating σ^2

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Inference for the slope parameter

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Pressing pressures and specimen densities for a ceramic compound

A mixture of ${\rm Al}_2{\rm O}_3$, polyvinyl alcohol, and water was prepared, dried overnight, crushed, and sieved to obtain 100 mesh size grains. These were pressed into cylinders at pressures from 2,000 psi to 10,000 psi, and cylinder densities were calculated.

x (pressure in psi)	y (density in g/cc)
2000.00	2.49
2000.00	2.48
2000.00	2.47
4000.00	2.56
4000.00	2.57
4000.00	2.58
6000.00	2.65
6000.00	2.66
6000.00	2.65
8000.00	2.72
8000.00	2.77
8000.00	2.81
10000.00	2.86
10000.00	2.88
10000.00	2.86

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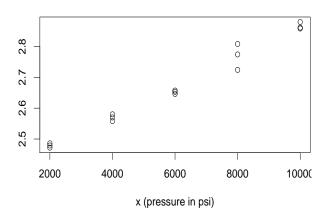
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Scatterplot: ceramics data



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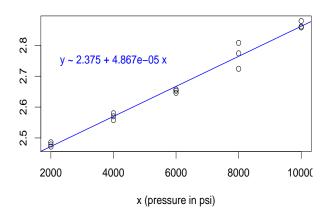
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► The line, $y \approx 2.375 + 4.867 \times 10^{-5} x$, is the **regression** line fit to the data.

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- 1. To predict future values of y based on x.
 - ▶ I.e., a new ceramic under pressure x = 5000 psi should have a density of $2.375 + 4.867 \times 10^{-5} \cdot 5000 = 2.618$ g/cc.
- 2. To characterize the relationship between *x* and *y* in terms of strength, direction, and shape.
 - ► In the ceramics data, density has a strong, positive, linear association with *x*.
 - ▶ On average, the density increases by 4.867×10^{-5} g/cc for every increase in pressure of 1 psi.

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► For a response variable *y* and a predictor variable *x*, we declare:

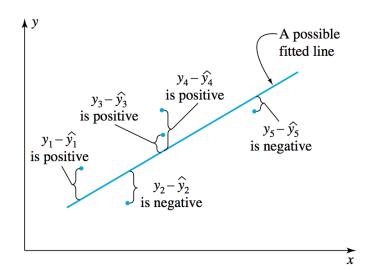
$$y \approx b_0 + b_1 x$$

- ▶ and then calculate the intercept b₀ and slope b₁ using least squares.
 - ▶ We apply the **principle of least squares**: that is, the best-fit line is given by minimizing the **loss function** in terms of b₀ and b₁:

$$S(b_0, b_1) = \sum_{i=1}^n (y_i - \widehat{y}_i)^2$$

 $\blacktriangleright \text{ Here, } \widehat{y}_i = b_0 + b_1 x_i$

Minimize $\sum_{i=1}^{n} (y_i - \hat{y}_i)^2$ to get the line as close as possible to the points.



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$$nb_0 + b_1 \sum_{i=1}^{n} x_i = \sum_{i=1}^{n} y_i$$
$$b_0 \sum_{i=1}^{n} x_i + b_1 \sum_{i=1}^{n} x_i^2 = \sum_{i=1}^{n} x_i y_i$$

▶ and then solve for b_0 and b_1 :

$$b_1 = \frac{\sum (x_i - \overline{x})(y_i - \overline{y})}{\sum (x_i - \overline{x})^2}$$
 $b_0 = \overline{y} - b_1 \overline{x}$

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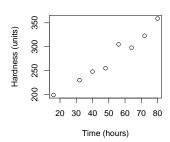
F-test and ANOVA

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Example: plastics hardness data

Eight batches of plastic are made. From each batch one test item is molded. At a given time (in hours), it hardness is measured in units (assume freshly-melted plastic has a hardness of 0 units). The following are the 8 measurements and times.

time	hardness
32.00	230.00
72.00	323.00
64.00	298.00
48.00	255.00
16.00	199.00
40.00	248.00
80.00	359.00
56.00	305.00



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Fitting the line

- $\overline{x} = 51$
- $\overline{v} = 277.125$

X	у	$x_i - \overline{x}$	$y_i - \overline{y}$	$(x_i - \overline{x})(y_i - \overline{y})$	$(x_i - \overline{x})^2$
32.00	230.00	-19.00	-47.12	895.38	361.00
72.00	323.00	21.00	45.88	963.38	441.00
64.00	298.00	13.00	20.88	271.38	169.00
48.00	255.00	-3.00	-22.12	66.38	9.00
16.00	199.00	-35.00	-78.12	2734.38	1225.00
40.00	248.00	-11.00	-29.12	320.38	121.00
80.00	359.00	29.00	81.88	2374.38	841.00
56.00	305.00	5.00	27.88	139.38	25.00

$$\sum (x_i - \overline{x})(y_i - \overline{y}) = 895.38 + 963.38 + \cdots 139.38 = 7765$$

$$b_1 = \frac{7765}{3192} = 2.43$$

$$b_0 = \overline{y} - b_1 \overline{x} = 277.125 - 2.43 \cdot 51 = 153.19$$

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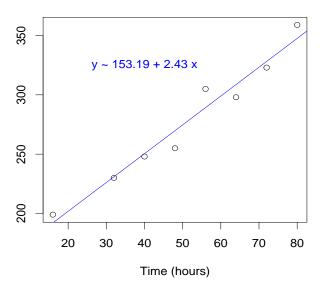
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Plot the line to check the fit.



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- ▶ b₁ = 2.43 means that on average, the plastic hardens 2.43 more units for every additional hour it is allowed to harden.
- ▶ $b_0 = 153.19$ means that if the model is completely true, at the very beginning of the hardening process (time = 0 hours), the plastics had a hardness of 153.19 on average.
 - ▶ But we know that the plastics were completely molten at the very beginning, with a hardness of 0.
 - ▶ Don't **extrapolate**: i.e., predict *y* values beyond the range of the *x* data.

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► Linear correlation: a measure of usefulness of a fitted line, defined by:

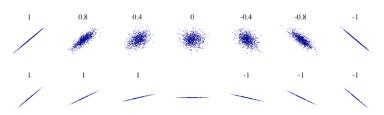
$$r = \frac{\sum (x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum (x_i - \overline{x})^2 \sum (y_i - \overline{y})^2}}$$

As it turns out:

$$r=b_1\frac{s_x}{s_y}$$

where s_x is the standard deviation of the x_i 's and x_y is the standard deviation of the y_i 's.

- ▶ $-1 \le r \le 1$
- ightharpoonup r < 0 means a negative slope, r > 0 means a positive slope
- ▶ High |r| means x and y have a strong linear relationship (high correlation), and low |r| implies a weak linear relationship (low correlation).



$$R^{2} = \frac{\sum (y_{i} - \overline{y})^{2} - \sum (y_{i} - \widehat{y}_{i})^{2}}{\sum (y_{i} - \overline{y})^{2}}$$

where $\hat{y}_i = b_0 + b_1 x_i$.

Fortunately,

$$R^2 = r^2$$

- ▶ Interpretation: R^2 is the fraction of variation in the response variable (y) explained by the fitted line.
- ▶ Ceramics data: $R^2 = r^2 = 0.9911^2 = 0.9822792$, so 98.227921% of the variation in density is explained by pressure. Hence, the line is useful for predicting density from pressure.
- ▶ Plastics data: $R^2 = r^2 = 0.9796^2 = 0.9596162$, so 95.961616% of the variation in hardness is explained by time. Hence, so the line is useful for predicting hardness from time.

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-test and ANOVA able ▶ Up until now, we have looked at fitted lines of the form:

$$y_i = b_0 + b_1 x_i + e_i$$

where:

- ▶ $y_1, y_2, ..., y_n$ are the fixed, observed values of the response variable.
- \triangleright x_1, x_2, \dots, x_n are the fixed, observed values of the predictor variable.
- b₀ is the estimated slope of the line based on sample data.
- b₁ is the estimated intercept of the line based on sample data.
- $ightharpoonup e_i$ is the residual of the *i*'th unit of the sample.

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$$Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

- Y_1, Y_2, \dots, Y_n are random variables that describe the response.
- x₁, x₂,...,x_n are still fixed, observed values of the predictor variable.
- \triangleright β_0 is a parameter denoting the *true* intercept of the line if we fit it to the population.
- \triangleright β_1 is a parameter denoting the *true* slope of the line if we fit it to the population.
- \triangleright $\varepsilon_1, \ \varepsilon_2, \dots, \varepsilon_n$ are random variables called **error terms**.

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▶ We assume:

$$\varepsilon_1, \ \varepsilon_2, \ldots, \varepsilon_n \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$$

▶ Which means that for all *i*:

$$Y_i \stackrel{\text{iid}}{\sim} N(\beta_0 + \beta_1 x_i, \sigma^2)$$

► We often say:

$$\mu_{v|x} = \beta_0 + \beta_1 x$$

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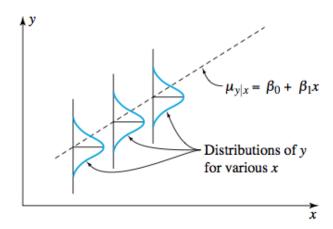
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The line-fitting sample variance

- ► Recall:
 - $\hat{\mathbf{y}}_i = b_0 + b_1 x_i$
 - $ightharpoonup e_i = y_i \widehat{y}_i$
- ► The line-fitting sample variance, also called mean squared error (MSE) is:

$$s_{LF}^2 = \frac{1}{n-2} \sum_i (y_i - \hat{y}_i)^2 = \frac{1}{n-2} \sum_i e_i^2$$

and it satisfies:

$$E(s_{LF}^2) = \sigma^2$$

The line-fitting sample standard deviation is just $s_{LF} = \sqrt{s_{LF}^2}$

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▶ A mixture of Al₂O₃, polyvinyl alcohol, and water was prepared, dried overnight, crushed, and sieved to obtain 100 mesh size grains. These were pressed into cylinders at pressures from 2,000 psi to 10,000 psi, and cylinder densities were calculated.

х,	у,
Pressure (psi)	Density (g/cc)
2,000	2.486
2,000	2.479
2,000	2.472
4,000	2.558
4,000	2.570
4,000	2.580
6,000	2.646
6,000	2.657
6,000	2.653
8,000	2.724
8,000	2.774
8,000	2.808
10,000	2.861
10,000	2.879
10,000	2.858

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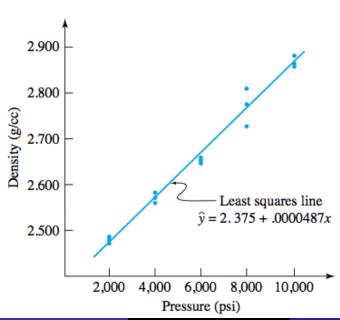
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-test and ANOVA able

- ▶ The fitted least squares line is $\hat{y}_i = 2.375 + 0.0000487x_i$.
- ▶ The fitted values \hat{y}_i are:

x, Pressure	\hat{y} , Fitted Density
2,000	2.4723
4,000	2.5697
6,000	2.6670
8,000	2.7643
10,000	2.8617

And $\sum (y_i - \hat{y}_i)^2$ is:

$$\sum (y_i - \hat{y}_i)^2 = (2.486 - 2.4723)^2 + (2.479 - 2.4723)^2 + (2.472 - 2.4723)^2 + (2.558 - 2.5697)^2 + \dots + (2.879 - 2.8617)^2 + (2.858 - 2.8617)^2$$

$$= .005153$$

► Thus,
$$s_{LF}^2 = \frac{1}{n-2} \sum (y_i - \hat{y}_i)^2 = \frac{1}{15-2} \cdot 0.005153 = 0.00396(g/cc)^2$$

$$s_{LF} = \sqrt{s_{LF}^2} = 0.0199g/cc$$

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- ▶ Recall that we assume $\varepsilon_1, \ldots, \varepsilon_n \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$.
- We also have $E(e_i) = 0$, but because we're estimating the slope and intercept instead of using the true slope and intercept,

$$Var(e_j) = \sigma^2 \left(1 - \frac{1}{n} - \frac{(x_j - \overline{x})^2}{\sum_i (x_i - \overline{x})^2} \right)$$

▶ We don't want $Var(e_j)$ to vary with j, so we define the j'th **standardized residual**:

$$e_j^* = rac{e_j}{s_{LF}\sqrt{1-rac{1}{n}-rac{(x_j-\overline{x})^2}{\sum_i(x_i-\overline{x})^2}}}$$

which, under our model assumptions, is $\approx N(0,1)$.

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▶ Since $\overline{x} = 6000$, we can calculate $\sum (x_i - \overline{x})^2 = 1.2 \times 10^8$.

Calculations for Standardized Residuals in the Pressure/Density Study

x	$\sqrt{1 - \frac{1}{15} - \frac{(x - 6,000)^2}{120,000,000}}$
2,000	.894
4,000	.949
6,000	.966
8,000	.949
10,000	.894

Residuals and Standardized Residuals for the Pressure/Density Study

x	e	Standardized Residual
2,000	.0137, .0067,0003	.77, .38,02
4,000	0117, .0003, .0103	62, .02, .55
6,000	0210,0100,0140	-1.09,52,73
8,000	0403, .0097, .0437	-2.13, .51, 2.31
10,000	0007, .0173,0037	04, .97,21

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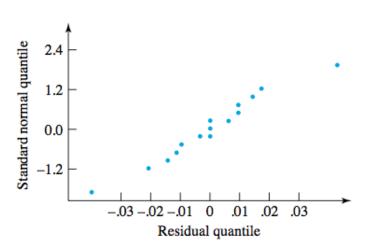
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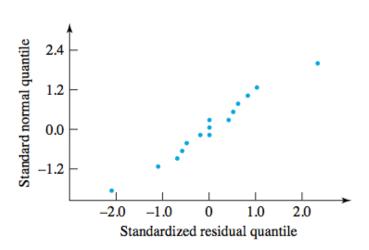
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- ► Since *b*₁ was estimated from the data, we can treat it as a random variable.
- Under the assumptions of the simple linear regression model,

$$b_1 \sim N\left(\beta_1, \ \frac{\sigma^2}{\sum_i (x_i - \overline{x})^2}\right)$$

► Thus:

$$Z = rac{b_1 - eta_1}{\sqrt{\sum_i (x_i - \overline{x})^2}} \sim N(0, 1)$$

and

$$T = rac{b_1 - eta_1}{\sqrt{\sum_i (x_i - \overline{x})^2}} \sim t_{n-2}$$

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▶ If we want to test H_0 : $\beta_1 = \#$, we can use the test statistic:

$$T=rac{b_1-\#}{rac{s_{LF}}{\sqrt{\sum_i(x_i-\overline{x})^2}}}\sim t_{n-2}$$

which has a t_{n-2} distribution if H_0 is true and the model assumptions are true.

• We can write a two-sided $1-\alpha$ confidence interval as:

$$\left(b_{1}-t_{n-2,\ 1-\alpha/2}\cdot\frac{s_{LF}}{\sqrt{\sum_{i}(x_{i}-\overline{x})^{2}}},b_{1}+t_{n-2,1-\alpha/2}\cdot\frac{s_{LF}}{\sqrt{\sum_{i}(x_{i}-\overline{x})^{2}}}\right)$$

▶ The one-sided confidence intervals are analogous.

- ► From before, $b_1 = 0.0000487$ g/cc/psi, $\sum_i (x_i \overline{x})^2 = 1.2 \times 10^8$, and $s_{LF} = 0.0199$.
- $t_{n-2, 1-\alpha/2} = t_{13, 0.975} = 2.16.$
- ► The confidence interval is then:

$$\left(0.0000487 - 2.16 \frac{0.0199}{\sqrt{1.2 \times 10^8}}, \ 0.0000487 + 2.16 \frac{0.0199}{\sqrt{1.2 \times 10^8}}\right)$$

$$\left(0.0000448, \ 0.0000526\right)$$

▶ We're 95% confident that for every unit increase in psi, the density of the next ceramic increases by anywhere between 0.0000448 g/cc and 0.0000526 g/cc.

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- ► In JMP:
 - Open the data in a spreadsheet with:
 - ▶ 1 column for x
 - ▶ 1 column for y
 - ► For simple linear regression
 - ightharpoonup Click Analyze ightharpoonup Fit Y by X
 - ▶ Y variable in Y, Response
 - X variable in X, Factor
 - ► Click red triangle Fit line

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▼ Analysis of Variance

 Source
 DF
 Squares
 Mean

 Model
 1
 0.28421333
 0

 Error
 13
 0.00515267
 0

 C. Total
 14
 0.28936600
 0

 Squares
 Mean Square
 F Ratio

 28421333
 0.284213
 717.0604

 00515267
 0.000396
 Prob > F

 28936600
 < .0001*</td>

▼ Parameter Estimates

 Term
 Estimate
 Std Error
 t Ratio
 Prob>ltl

 Intercept
 2.375
 0.012055
 197.01
 <.0001*</td>

 pressure
 4.8667e-5
 1.817e-6
 26.78
 <.0001*</td>

Bivariate Fit of density By pressure 2.85-2.8-2.75-2.7-2.65 2.6-2.55-2.5-2 45 2000 4000 6000 8000 10000 pressure Linear Fit ▼ Linear Fit

density = 2.375 + 4.8667e-5*pressure

Summary of Fit

0.982193

0.980824

0.019909

2.667

15

RSquare

RSquare Adi

Root Mean Square Error

Observations (or Sum Wats)

Mean of Response

▼ Parameter Estimates

Term	Estimate	Std Error	t Ratio	Prob>ltl
Intercept	2.375	0.012055	197.01	<.0001*
pressure	4.8667e-5	1.817e-6	26.78	<.0001*

► I can construct the same confidence interval using the JMP output:

▶
$$b_1 = 4.87 \times 10^{-5}$$
, $t_{n-1,1-\alpha/2} = 2.16$, $\widehat{SD}(b_1) = 1.817 \times 10^{-6}$

•

$$(4.87 \times 10^{-5} - 2.16 \cdot 1.817 \times 10^{-6},$$

$$4.87 \times 10^{-5} + 2.16 \cdot 1.817 \times 10^{-6})$$

= (0.0000448, 0.0000526)

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▼ Parameter Estimates

Term	Estimate	Std Error	t Ratio	Prob>ltl
Intercept	2.375	0.012055	197.01	<.0001*
pressure	4.8667e-5	1.817e-6	26.78	<.0001*

At $\alpha = 0.05$, conduct a two-sided hypothesis test of $H_0: \beta_1 = 0$ using the method of p-values.

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- 1. $H_0: \beta_1 = 0, H_a: \beta_1 \neq 0.$
- 2. $\alpha = 0.05$
- 3. Use the test statistic:

$$T=rac{b_1-0}{rac{s_{LF}}{\sqrt{\sum(x_i-ar{x})^2}}}=rac{b_1}{\widehat{SD}(b_1)}$$

Lassume:

- $ightharpoonup H_0$ is true.
- ▶ The model, $Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$ with errors $\varepsilon_i \sim \text{iid}$ $N(0, \sigma^2)$, is correct.

Under these assumptions, $T \sim t_{n-2} = t_{15-2} = t_{13}$

4. Observed test statistic:

$$\begin{split} t &= \frac{4.87 \times 10^{-5}}{1.817 \times 10^{-6}} = 26.80 \quad \text{("t Ratio" in JMP output)} \\ \text{p-value} &= P(|t_{13}| > |26.8|) = P(t_{13} > 26.8) + P(t_{13} < -26.8) \\ &< 0.0001 \quad \text{("Prob>} |t|" in JMP output)} \end{split}$$

- 5. With a p-value $< 0.0001 < 0.05 = \alpha$, we reject H_0 and conclude H_a .
- 6. There is overwhelming evidence that the true slope of the line is different from 0.

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$$H_0: \beta_1 = 0, H_a: \beta_1 \neq 0$$

- ▶ Another method for testing $H_0: \beta_1 = 0, H_a: \beta_0 \neq 0$ is the "ANalysi Of VAriance" (ANOVA) method.
- ► Fact: the Total Sum of Squares can be decomposed into Error Sum of Squares and Regression Sum of Squares.

$$\sum_{i=1}^{n} (y_i - \bar{y})^2 = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 + \sum_{i=1}^{n} (\hat{y}_i - \bar{y})$$
SST ol SSE

• Under the assumptions of SLR model, and assuming $H_0: \beta_1 = 0$ is true, the test statistic

$$F = \frac{SSR/1}{SSE/(n-2)}$$

has a $F_{1,n-2}$ distribution. (Reiview F distribution in ch5part5_Mar_3.pdf.)

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$$H_0: \beta_1 = 0, H_a: \beta_1 \neq 0$$

• We reject $H_0: \beta_1 = 0$ in favor of $H_a: \beta_1 \neq 0$ when the test statistic F is large. So the p-value is

$$P(F_{1,n-2} > \text{observed } F)$$

▶ In fact, the square of the t statistic for testing $H_0: \beta_1 = 0$ is

$$T^2 = \left(rac{b_1 - 0}{rac{s_{LF}}{\sqrt{\sum (x_i - ar{x})^2}}}
ight) = rac{SSR/1}{SSE/(n-2)} = F$$

which has an $F_{1,n-2}$ distribution if H_0 is true and tends to be large if H_0 is false. So counting large F as evidence against H_0 in favor of $H_a:\beta_1\neq 0$ is a sensible significance testing method.

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ANOVA table

Calculations in the ANOVA method can be summarized in the ANOVA table:

Source	SS	df	MS	F
Regression	SSR	1	MSR = SSR/1	F = MSR/MSE
Error	SSE	n-2	MSE = SSE/(n-2)	
Total	SSTot	n — 1		

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Source

Model

C. Total

Error

△ Analysis of Variance

DF

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F-test and ANOVA

The n-value	in the	F_test	is very	small	So we	reject Ho	

► There is significant evidence that the true slope is different from 0.

Sum of

Squares

0.28421333

0.00515267

0.28936600

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Mean Square

0.284213

F Ratio

<.0001*

717.0604

0.000396 Prob > F