

Multiple Regression and ANOVA (Ch. 9.2)

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Multiple Regression and ANOVA

Sums of squares

Advanced inference for multiple regression

The F test statistic and R^2

- ▶ **Analysis of variance (ANOVA):** the use of sums of squares to construct a test statistic for comparing nested models.
- ▶ **Nested models:** a pair of models such that one contains all the parameters of the other.
 - ▶ Examples:
 - ▶ Full model: $Y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \varepsilon_i$ with the reduced model: $Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$.
 - ▶ Full model: $Y_i = \beta_0 + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \varepsilon_i$ with the reduced model: $Y_i = \beta_0 + \varepsilon_i$

Sums of Squares

- ▶ **Total sum of squares (SST)**: the total amount of variation in the response.

$$SST = \sum_i (y_i - \bar{y})^2$$

- ▶ **Regression sum of squares (SSR)**: the amount of variation in response explained by the model.

$$SSR = \sum_i (\hat{y}_i - \bar{y})^2$$

- ▶ **Error sum of squares (SSE)**: the amount of variation in the response *not* explained by the model.

$$SSE = \sum_i (y_i - \hat{y}_i)^2$$

Properties of Sums of Squares

- ▶ They add up:

$$SST = SSR + SSE$$

- ▶ We can use them to calculate R^2 :

$$R^2 = \frac{SST - SSE}{SST} = \frac{SSR}{SST}$$

- ▶ We can calculate the **mean squared error (MSE)**:

$$MSE = \frac{1}{n - p} SSE$$

which satisfies:

$$E(MSE) = \sigma^2$$

$MSE = s_{LF}^2$ for simple linear regression and s_{SF}^2 for multiple regression.

- ▶ The **regression mean square (MSR)** is:

$$MSR = \frac{1}{p - 1} SSR$$

Inference: deciding between nested models

- ▶ Suppose I have the full model:

$$Y_i = \beta_0 + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \cdots + \beta_{p-1} x_{p-1,i} + \varepsilon_i$$

- ▶ And an intercept-only reduced model:

$$Y_i = \beta_0 + \varepsilon_i$$

- ▶ I want to do a hypothesis test to decide if the full model works better than the reduced model.
 - ▶ Does the full model explain significantly more variation in the response than the reduced model?
 - ▶ This is a job for the sums of squares.

The hypothesis test: intercept-only model vs. full model

1.
 - ▶ $H_0 : \beta_1 = \beta_2 = \dots = \beta_{p-1} = 0$
 - ▶ $H_a : \text{not all of the } \beta_i\text{'s} = 0 \ (i = 1, 2, \dots, p - 1)$
2. α is some sensible value (< 0.1).
3. The test statistic is:

$$F = \frac{SSR/(p-1)}{SSE/(n-p)} = \frac{MSR}{MSE} \sim F_{p-1, n-p}$$

Assume:

- ▶ H_0 is true.
 - ▶ The full model is valid with the ε_i 's iid $N(0, \sigma^2)$
4. Reject H_0 if observed $F > F_{p-1, n-p, 1-\alpha}$. Or use the p-value: $P(F_{p-1, n-p} > \text{observed} F)$; reject H_0 when p-value is small.

Example: stack loss

1. Consider a chemical plant that makes nitric acid from ammonia.
2. We want to predict stack loss (y , 10 times the % ammonia that escapes from the absorption column) using:
 - ▶ x_1 : air flow, the rate of operation of the plant
 - ▶ x_2 , inlet temperature of the cooling water
 - ▶ x_3 : (% circulating acid - 50%) $\times 10$

Example: stack loss

i , Observation Number	x_{1i} , Air Flow	x_{2i} , Cooling Water Inlet Temperature	x_{3i} , Acid Concentration	y_i , Stack Loss
1	80	27	88	37
2	62	22	87	18
3	62	23	87	18
4	62	24	93	19
5	62	24	93	20
6	58	23	87	15
7	58	18	80	14
8	58	18	89	14
9	58	17	88	13
10	58	18	82	11
11	58	19	93	12
12	50	18	89	8
13	50	18	86	7
14	50	19	72	8
15	50	19	79	8
16	50	20	80	9
17	56	20	82	15

Example: stack loss

- ▶ Given:
 - ▶ $n = 17$
 - ▶ y : stack loss of nitrogen from the chemical plant.
 - ▶ x_1 : air flow, the rate of operation of the plant
 - ▶ x_2 , inlet temperature of the cooling water
 - ▶ x_3 : (% circulating acid - 50%) $\times 10$
- ▶ We'll test the full model:

$$Y_i = \beta_0 + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + \varepsilon_i$$

against the reduced model:

$$Y_i = \beta_0 + \varepsilon_i$$

at $\alpha = 0.05$.

Example: stack loss

1.
 - ▶ $H_0 : \beta_1 = \beta_2 = \beta_3 = 0$
 - ▶ Not all of the β_i 's are 0, $i = 1, 2, 3$.
2. $\alpha = 0.05$
3. The test statistic is:

$$F = \frac{SSR/(p-1)}{SSE/(n-p)} = \frac{MSR}{MSE} \sim F_{p-1, n-p}$$

Assume:

- ▶ H_0 is true.
- ▶ The full model is valid with the ε_i 's iid $N(0, \sigma^2)$

Reject H_0 if $F > F_{p-1, n-p, 1-\alpha} = F_{4-1, 17-4, 1-0.05} = F_{3,13,0.95} = 3.41$.

Example: stack loss

4. In JMP, fit the full model and look at the **ANOVA table**:

Analysis of Variance				
Source	DF	Sum of Squares	Mean Square	F Ratio
Model	3	795.83449	265.278	169.0432
Error	13	20.40080	1.569	Prob > F
C. Total	16	816.23529		<.0001*

by reading directly from the table, we can see:

- ▶ $p - 1 = 3$, $n - p = 13$, $n - 1 = 16$
 - ▶ $SSR = 795.83$, $SSE = 20.4$, $SST = 816.24$
 - ▶ $MSR = SSR/(p - 1) = 795.83/3 = 265.28$
 - ▶ $MSE = SSE/(n - p) = 20.4/13 = 1.57$
 - ▶ $observedF = MSR/MSE = 265.78/1.57 = 169.04$
 - ▶ Prob>F gives the p-value,
 $P(F_{3,13} > observedF) < 0.0001$.
5. With $F = 169.04 > 3.41$, we reject H_0 and conclude H_a .
6. There is overwhelming evidence that at least one of air flow, inlet temperature, and % circulating acid is important in explaining the variation in stack loss.

What if I want to compare different nested models?

1.
 - ▶ $H_0 : \beta_{l_1} = \beta_{l_2} = \cdots = \beta_{l_k} = 0$
 - ▶ $H_a : \text{not all of } \beta_{l_1}, \beta_{l_2}, \cdots, \beta_{l_k} \text{ are } 0.$
 - ▶ (For example, $H_0 : \beta_2 = \beta_3 = 0$ vs $H_a : \text{either } \beta_2 \text{ or } \beta_3 \neq 0 \text{ or both. The model is } Y_i = \beta_0 + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \beta_3 x_{i,3} + \beta_4 x_{i,4} + \varepsilon_i, \text{ and } k = 2)$
2. α is some sensible value.
3. The test statistic is:

$$F = \frac{(SSR_f - SSR_r)/k}{SSE_f/(n-p)} \sim F_{k, n-p}$$

- ▶ SSR_r is for the reduced model and SSR_f is for the full model.
- ▶ Of course, we assume H_0 is true and the full model is valid with the ε_i 's iid $N(0, \sigma^2)$.

What if I want to compare different nested models?

4. We can construct a combined ANOVA table:

Source	SS	df	MS	F
Reg (full)	SSR_f	$p - 1$		
Reg (reduced)	SSR_r	$p - k - 1$		
Reg (full red)	$SSR_f - SSR_r$	k	$\frac{SSR_f - SSR_r}{k}$	$\frac{MSR_{f r}}{MSE_f}$
Error	SSE_f	$n - p$	$\frac{SSE_f}{n - p}$	
Total	SST	$n - 1$		

5. Reject H_0 if observed $F > F_{p-1, n-p, 1-\alpha}$. Or use the p-value:
 $P(F_{p-1, n-p} > \text{observed } F)$; reject H_0 when p-value is small.

Example: stack loss

1.
 - ▶ $H_0 : \beta_2 = \beta_3 = 0$
 - ▶ $H_a : \text{either } \beta_2 \neq 0 \text{ or } \beta_3 \neq 0$
2. $\alpha = 0.05$
3. The test statistic is:

$$\begin{aligned} F &= \frac{(SSR_f - SSR_r)/k}{SSE_f/(n-p)} = \frac{(SSR_f - SSR_r)/2}{SSE_f/(17-4)} \\ &= \frac{(SSR_f - SSR_r)/2}{SSE_f/13} \end{aligned}$$

- ▶ Assume H_0 is true and the full model is valid with the ε_i 's iid $N(0, \sigma^2)$.
- ▶ Then, $F \sim F_{k, n-p} = F_{2,13}$.
- ▶ I will reject H_0 if $F > F_{2,13,0.95} = 3.81$.

Example: stack loss

4. Look at the ANOVA tables in JMP for both the full model ($Y_i = \beta_0 + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + \varepsilon_i$):

▼ Analysis of Variance				
Source	DF	Sum of Squares	Mean Square	F Ratio
Model	3	795.83449	265.278	169.0432
Error	13	20.40080	1.569	Prob > F
C. Total	16	816.23529		<.0001*

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and the reduced model ($Y_i = \beta_0 + \beta_1 x_{1,i} + \varepsilon_i$):

▼ Analysis of Variance				
Source	DF	Sum of Squares	Mean Square	F Ratio
Model	1	775.48219	775.482	285.4318
Error	15	40.75311	2.717	Prob > F
C. Total	16	816.23529		<.0001*

Example: stack loss

I construct a different ANOVA table for this test:

Source	SS	df	MS	F
Reg (full)	795.83	4		
Reg (reduced)	775.48	2		
Reg (full red)	20.35	2	10.18	6.48
Error	20.4	13	1.57	
Total	SST	16		

5. With $observedF = 6.48 > 3.81$, I reject H_0 and conclude H_a .
6. There is enough evidence to conclude that at least one of inlet temperature and % circulating acid is associated with stack loss.

Example: stack loss

- Attempt to eliminate inlet temperature (x_2) from the model at $\alpha = 0.05$. Here is the ANOVA table for the full model:

▼ Analysis of Variance				
Source	DF	Sum of Squares	Mean Square	F Ratio
Model	3	795.83449	265.278	169.0432
Error	13	20.40080	1.569	Prob > F
C. Total	16	816.23529		<.0001*

and for the reduced model:

▼ Analysis of Variance				
Source	DF	Sum of Squares	Mean Square	F Ratio
Model	2	776.84496	388.422	138.0520
Error	14	39.39033	2.814	Prob > F
C. Total	16	816.23529		<.0001*

Example: stack loss

1. $H_0 : \beta_2 = 0, H_a : \beta_2 \neq 0$
2. $\alpha = 0.05$
3. The test statistic is:

$$\begin{aligned} F &= \frac{(SSR_f - SSR_r)/k}{SSE_f/(n-p)} = \frac{SSR_f - SSR_r}{SSE_f/(17-4)} \\ &= \frac{SSR_f - SSR_r}{SSE_f/13} \end{aligned}$$

- ▶ Assume H_0 is true and the full model is valid with the ε_i 's iid $N(0, \sigma^2)$.
- ▶ Then, $F \sim F_k, n-p = F_{1,13}$.
- ▶ I will reject H_0 if $F > F_{1,13,0.95} = 4.67$.

Example: stack loss

4. I construct a different ANOVA table for this test:

Source	SS	df	MS	F
Reg (full)	795.83	4		
Reg (reduced)	776.84	3		
Reg (full red)	18.99	1	18.99	12.10
Error	20.4	13	1.57	
Total	SST	16		

5. With $observedF = 12.10 > 4.67$, we reject H_0 .
6. There is enough evidence to conclude that stack loss varies with inlet temperature.

Example: stack loss

- ▶ The F test for eliminating one parameter is analogous to the t test from before:

▼ Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	-37.65246	4.732051	-7.96	<.0001*
x1	0.7976856	0.067439	11.83	<.0001*
x2	0.5773405	0.165969	3.48	0.0041*
x3	-0.06706	0.061603	-1.09	0.2961

- ▶ The t statistic for $H_0 : \beta_2 = 0$ vs. $H_0 : \beta_2 \neq 0$ is 3.48.
- ▶ But $3.48^2 = 12.1$, which is our F statistic from the ANOVA test!
- ▶ Fun fact:

$$F_{1, \nu} = t_{\nu}^2$$

The F test statistic and R^2

- ▶ If F is the test statistic from a test of $H_0 : \beta_1 = \beta_2 = \cdots = \beta_{p-1} = 0$ vs. H_a : not all of $\beta_1, \beta_2, \dots, \beta_{p-1}$ are 0, then F can be expressed in terms of the coefficient of determination of the full model:

$$F = \frac{R^2/(p-1)}{(1-R^2)/(n-p)}$$

- ▶ For the stack loss example, the full model's $R^2 = 0.975$, and so:

$$F = \frac{0.975/(4-1)}{(1-0.975)/(17-4)} = 169$$

The F test statistic and R^2

▼ Summary of Fit

RSquare	0.975006
RSquare Adj	0.969238
Root Mean Square Error	1.252714
Mean of Response	14.47059
Observations (or Sum Wgts)	17

▼ Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Ratio
Model	3	795.83449	265.278	169.0432
Error	13	20.40080	1.569	Prob > F
C. Total	16	816.23529		<.0001*

The F test statistic and R^2

- For $H_0 : \beta_1 = \beta_2 = \cdots = \beta_{p-1} = 0$ vs. $H_a : \text{not all of } \beta_1, \beta_2, \dots, \beta_{p-1},$

$$\begin{aligned} F &= \frac{SSR \frac{1}{p-1}}{SSE \frac{1}{n-p}} = \frac{\frac{SSR}{SST} \frac{1}{p-1}}{\frac{SSE}{SST} \frac{1}{n-p}} = \frac{\frac{SSR}{SST} \frac{1}{p-1}}{\frac{SST-SSR}{SST} \frac{1}{n-p}} = \frac{\frac{SSR}{SST} \frac{1}{p-1}}{\left(1 - \frac{SSR}{SST}\right) \frac{1}{n-p}} \\ &= \frac{R^2 \frac{1}{p-1}}{(1 - R^2) \frac{1}{n-p}} \end{aligned}$$

The F test statistic and R^2

- ▶ If F is the test statistic from a test of $H_0 : \beta_{l_1} = \beta_{l_2} = \cdots = \beta_{l_k} = 0$ vs. H_a : not all of $\beta_{l_1}, \beta_{l_2}, \dots, \beta_{l_k}$ are 0, then F can be expressed in terms of the coefficient of determination of the full model (R_f^2) and that of the reduced model (R_r^2):

$$F = \frac{(R_f^2 - R_r^2)/k}{(1 - R_f^2)/(n - p)}$$

- ▶ For the stack loss example when we tested $H_0 : \beta_2 = \beta_3 = 0$, $R_f^2 = 0.975$ and $R_r^2 = 0.95$.

$$F = \frac{(0.975 - 0.95)/2}{(1 - 0.975)/(17 - 4)} = 6.50$$

which is close to the test statistic of 6.48 that we calculated before.

The F test statistic and R^2

- ▶ When we tested $H_0 : \beta_2 = 0$, R_r^2 was 0.9517, so:

$$F = \frac{(0.975 - 0.9517)/1}{(1 - 0.975)/(17 - 4)} = 12.117$$

which is close to the test statistic of 12.10 that was calculated directly from the ANOVA table.

The F test statistic and R^2

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$$\begin{aligned} F &= \frac{(SSR_f - SSR_r) \frac{1}{k}}{SSE_f \frac{1}{n-p}} = \frac{\frac{SSR_f - SSR_r}{SST} \frac{1}{k}}{\frac{SSE_f}{SST} \frac{1}{n-p}} = \frac{\left(\frac{SSR_f}{SST} - \frac{SSR_r}{SST} \right) \frac{1}{k}}{\frac{SST - SSR_f}{SST} \frac{1}{n-p}} \\ &= \frac{\left(\frac{SSR_f}{SST} - \frac{SSR_r}{SST} \right) \frac{1}{k}}{\left(1 - \frac{SSR_f}{SST} \right) \frac{1}{n-p}} = \frac{(R_f^2 - R_r^2) \frac{1}{k}}{(1 - R_f^2) \frac{1}{n-p}} \end{aligned}$$