Yifan Zhu

Bernolli Distributior

Distribution

Geometric Distribution

Poisson Distribution

# Special Discrete Random Variables (Ch. 5.1)

Yifan 7hu

Iowa State University

### Outline

Special Discrete Random Variables (Ch. 5.1)

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 $ightharpoonup X \sim \text{Bernolli}(p) - \text{i.e.}, X \text{ is distributed as a bernolli}$ random variable with parameter p (0 < p < 1) if:

$$f_X(x) = \begin{cases} p^x (1-p)^{1-x} & x = 0, 1 \\ 0 & \text{otherwise} \end{cases}$$

- E(X) = p Var(X) = p(1-p)
  - A bernolli random variable indicate success (encoded as 1) or failure (encoded as 0) in one success-failure trial.
  - **Examples**:
    - A hexamine pellet made from a pelletizing machine conforms the specification (success, 1) or not (failure, 0).
    - A run of the chemical process has a percent yield above 80% (success, 1) or not (failure, 0).

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fcx):

x=(. f()=p

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 $\triangleright$  X  $\sim$  Binomial(n, p) - i.e., X is distributed as a binomial random variable with parameters n and p(0 if:

$$f_X(x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & x = 0, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

where:

- mere.  $\binom{n}{x} = \underbrace{\binom{n!}{x!(n-x)!}}_{read} \text{ "$n$ choose $x$"}$   $n! = n \cdot (n-1) \cdot \dots \cdot 2 \cdot 1, \text{ the factorial function.}$
- E(X) = np Var(X) = np(1-p)

## The Binomial Distribution

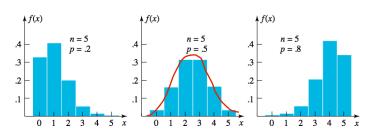


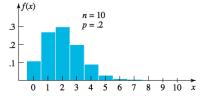
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Geometric Distribution





X: # of Successes. our of n trials.

[] [] [] ... (] = n Success for line trials.

independent.

for each one:

Success prob: 1.

Support: {0,1,..., n}.

 $f(x) = P(X=x) = \binom{n}{x} p^{x} (1-p)^{n-x}$ 

$$= \sum_{x=1}^{n} \frac{n!}{(x-1)! (n-x)!} \cdot p^{x} (-p)^{n-x}, \quad x-1 \to x$$

$$= \frac{n-1}{2} \frac{n!}{x! (n-1-x)!} \cdot p^{x+1} (1-p)^{n-x-1}$$

$$= \sum_{x=0}^{n-1} \frac{(n-1)!}{x! (n-1-x)!} \cdot p^{x} \cdot p^{x} (1-p)^{n-1-x}$$

 $= NP \sum_{x=0}^{N-1} {N-1 \choose x} p^{x} (1-p)^{N-1-x} = NP$ 

 $E(x) = \sum_{k=0}^{n} x \cdot f(x) = \sum_{k=0}^{n} x \cdot {n \choose k} p^{k} (1-p)^{n-k}$ 

 $= \frac{n}{\sum_{k=0}^{N} x} \cdot \frac{n!}{x! (n-x)!} p^{x} (1-p)^{n-x}$ 

$$\frac{E(x^2)}{E(x(x-1))} = E((x-1)x) + E(x)$$

$$= \sum_{x=0}^{n} x(x-1) {n \choose x} p^x (1-p)^{n-x}$$

$$= \sum_{x=2}^{n} x(x-1) {n \choose x} p^x (1-p)^{n-x}$$

$$= \sum_{x=2}^{n} x(x-1) {n \choose x} p^x (1-p)^{n-x}$$

$$= \sum_{x=5}^{\infty} \frac{(x-5)!(n-x)!}{x!(n-b)!} b_{x}(1-b)_{x-x}$$

$$= \sum_{y=5}^{\infty} \frac{(x-5)!(n-x)!}{x!(n-b)!} b_{x}(1-b)_{x-x}$$

$$E(x(x-1)) = n(n-1)p^{2}$$

$$E(x^{2}) = n(n-1)p^{2} + np = n^{2}p^{2} - np^{2} + np$$

$$V_{nr}(x) = E(x^{2}) - (E(x))^{2}$$

- np-np= np(1-p)

- np+np-np

- ► A Bin(n, p) random variable counts the number of successes in n success-failure trials that:
  - are independent of one another.
  - each succeed with probability p.
- Examples:
  - Number of conforming hexamine pellets in a batch of n = 50 total pellets made from a pelletizing machine.
  - Number of runs of the same chemical process with percent yield above 80%, given that you run the process a total of n = 1000 times.
  - Number of rivets that fail in a boiler of n=25 rivets within 3 years of operation. (Note; "success" doesn't always have to be good.)

Distribution

- Bernolli(p) = Binomial(1, p).
   Let X<sub>1</sub>, X<sub>2</sub>,..., X<sub>n</sub> be i.i.d. (independent and identically distributed) Bernolli(p) random variables, then

$$\sum_{i=1}^{n} X_i \sim \mathsf{Binomial}(n, p)$$

Informally,  $X_1, X_2, \dots, X_n$  are independent means their outcomes do not affect each other. Formal definition of independence will be given in the lecture of "joint distribution and independence".

$$eg \left( \sum_{i=1}^{n} \chi_{i} \right) = \sum_{i=1}^{n} E(\chi_{i}) \quad (\text{true for non independent to } ).$$

$$eg Var \left( \sum_{i=1}^{n} \chi_{i} \right) = \sum_{i=1}^{n} Var \left( \chi_{i} \right) \quad (\text{only true for independent variables})$$

$$E(X) = \sum_{i=1}^{\infty} E(X_i) = NP(i-p).$$

$$Var(X) = \sum_{i=1}^{\infty} Var(X_i) = NP(i-p).$$

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Binomial Distribution

Distribution

- 1 2 3 4 5 6 7 8 9 10
  - Suppose you have a machine with 10 independent components in series. The machine only works if all the components work.
  - ► Each component succeeds with probability p = 0.95 and fails with probability 1 p = 0.05.
  - ▶ Let *Y* be the number of components that succeed in a given run of the machine. Then:

$$Y \sim \mathsf{Binomial}(n = 10, p = 0.95)$$

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$$P(\text{machine succeeds}) = P(Y = 10)$$

$$= \binom{10}{10} p^{10} (1-p)^{10-10}$$

$$= p^{10}$$

$$= 0.95^{10}$$

$$= 0.5987$$

This machine isn't very reliable.

## Example: machine with 10 components

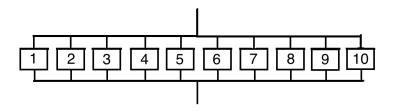


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- What if I arrange these 10 components in parallel? This machine succeeds if at least 9 of the components succeed.
- What is the probability that the new machine succeeds?

$$P(\text{improved machine succeeds})$$

$$= P(Y \ge 9)$$

$$= P(Y = 9) + P(Y = 10)$$

$$= {10 \choose 9} p^9 (1 - p) + {10 \choose 10} p^{10} (1 - p)^{10 - 10}$$

$$= (10) \cdot 0.95^9 \cdot 0.05 + (1) \cdot 0.95^{10}$$

By allowing just one component to fail, we made this machine far more reliable.

= 0.9139

► If we allow up to 2 components to fail:

P(improved machine succeeds)

$$= P(Y \ge 8)$$

$$= P(Y = 8) + P(Y = 9) + P(Y = 10)$$

$$= {10 \choose 8} p^8 (1 - p)^{10 - 8} + {10 \choose 9} p^9 (1 - p) + {10 \choose 10} p^{10} (1 - p)^{10 - 10}$$

$$= {10! \over (10 - 8)!8!} \cdot 0.95^8 \cdot 0.05^2 + (10) \cdot 0.95^9 \cdot 0.05 + (1) \cdot 0.95^{10}$$

$$= 0.9885$$

Geometric Distribution

- ►  $E(Y) = np = 10 \cdot 0.95 = 9.5$ . So the number of components to fail per run on average is 9.5.
- $Var(Y) = np(1-p) = 10 \cdot 0.95 \cdot (1-0.95) = 0.475.$
- $ightharpoonup SD(Y) = \sqrt{Var(Y)} = \sqrt{np(1-p)} = 0.689.$

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▶  $X \sim \text{Geometric}(p) - \text{that is, } X \text{ has a geometric}$  distribution with parameter  $p \ (0 is:$ 

$$f_X(x) = \begin{cases} p(1-p)^{x-1} & x = 1, 2, 3, \dots \\ 0 & \text{otherwise} \end{cases}$$

and its cdf is:

$$F_X(x) = egin{cases} 1 - (1-p)^x & x = 1, 2, 3, \dots \\ 0 & \text{otherwise} \end{cases}$$

$$E(X) = \frac{1}{p}$$

$$Var(X) = \frac{1-p}{p^2}$$

$$X: \# \text{ if frials taken to get a first success.}$$

$$\{(, z_1, ..., 3).$$

$$f(x) = P(X = x).$$

$$y = 1 \quad P(y = 0) = 0$$

$$x=1.$$
  $P(x=1)=p$ 

$$x=2: FS. \rightarrow P(x=2) = (1-p)P$$

3, EES 
$$\rightarrow P(X=3)=(1-6)^{k_1}$$

$$E(x) = \sum_{k=1}^{\infty} x \left( (-p)^{k-1} \cdot p \right)$$

$$\sum_{k=0}^{\infty} \left( (-p)^{k} \right) = \frac{1}{p} ,$$

$$\sum_{x=0}^{\infty} (1-p)^{x} = \frac{1}{p}-1,$$

$$(1-p)^{n} = \frac{1}{p-1},$$

$$\sum_{p} + x \left(1 - p\right)^{x-1} = r \frac{1}{p^2}$$

$$\sum_{k=1}^{\infty} + x \left( 1 - p \right)^{x-1} = + \frac{1}{p^2}$$

$$\langle (1-p)^{\lambda-1} = + \frac{1}{p^2}$$

 $\sum_{i=1}^{\infty} x_i p(1-p)^{x-1} = \frac{1}{p} = \overline{t}(x)$ 

$$\frac{\sum_{X=1}^{\infty} (1-p)^{X-1} |p| = 1$$

$$\sum_{X=1}^{N} x (1-p)^{X-1} = \frac{1}{p^2}$$

$$\sum_{X=1}^{N} x (1-p)^{X} = \frac{1-p}{p^2} = \frac{1}{p^2} - \frac{1}{p}$$

$$\sum_{X=1}^{N} x x (1-p)^{X-1} (-1) = -\frac{2}{p^3} + \frac{1}{p^2}$$

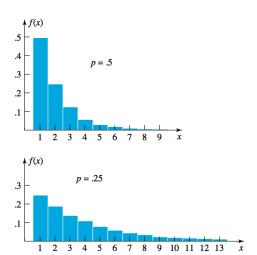
$$\sum_{X=1}^{N} x^2 (1-p)^{X-1} = \frac{2}{p^3} - \frac{1}{p^2} = E(x^2)$$

$$\sum_{X=1}^{N} x^2 p (1-p)^{X-1} = \frac{2}{p^3} - \frac{1}{p} = E(x^2)$$

$$Var(x) = E(x^2) - (E(x))^2$$

$$= \frac{2}{p^3} - \frac{1}{p} - (\frac{1}{p})^2 = \frac{1}{p^2} - \frac{1}{p} = \frac{1-p}{p^2}$$

# A look at the Geom(p) distribution



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- ► For an indefinitely-long sequence of independent, success-failure trials, each with P(success) = p, X is the number of trials it takes to get a success.
- Examples:
  - Number of rolls of a fair die until you land a 5
  - Number of shipments of raw material you get until you get a defective one.
  - The number of enemy aircraft that fly close before one flies into friendly airspace.
  - Number hexamine pellets you make before you make one that does not conform.
  - Number of buses that come defore yours.

## Example: shorts in NiCad batteries

- An experimental program was successful in reducing the percentage of manufactured NiCad cells with internal shorts to around 1%.
- ▶ Let T be the test number at which the first short is discovered. Then, T ~ Geom(p).

P(1st or 2nd cell tested is has the 1st short) = P(T=1 or T=2) = f(1) + f(2) = p + p(1-p) = 0.01 + 0.01(1-0.01) = 0.02

 $P(\text{at least 50 cells tested w/o finding a short}) = \underbrace{P(T > 50)}_{} = 1 - \underbrace{P(T \le 50)}_{} = 1 - F(50)$   $= 1 - (1 - (1 - \rho)^{x})$   $= (1 - \rho)^{x}$   $= (1 - 0.01)^{50}$  = 0.61

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$$E(T) = \frac{1}{p} = \frac{1}{0.01}$$

 $=100\ \text{tests}$  for the first short to appear, on avg.

$$SD(T) = \sqrt{Var(T)} = \sqrt{\frac{1-p}{p^2}}$$

$$= \sqrt{\frac{1-0.01}{0.01^2}} = 99.5 \text{ tested batteries}$$

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▶  $X \sim \text{Poisson}(\lambda)$  — that is, X has a geometric distribution with parameter  $\lambda > 0$  — if its pmf is:

$$f_X(x) = egin{cases} rac{\mathrm{e}^{-\lambda}\lambda^x}{x!} & x = 0, 1, 2, 3, \dots \\ 0 & ext{otherwise} \end{cases}$$

- $ightharpoonup E(X) = \lambda$
- $Var(X) = \lambda$

$$f(x) = \frac{e^{-\lambda} \lambda^{x}}{x!}, x = 0, 1, \dots$$

$$f(x) \gg 0.$$

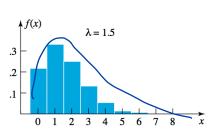
$$\sum_{x=0}^{\infty} f(x) = 1.$$

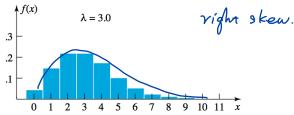
$$e^{x} = \sum_{n=0}^{\infty} \frac{x^{n}}{n!}$$

$$e^{\lambda} = \sum_{x=0}^{\infty} \frac{\lambda^x}{x!}$$

$$e^{\lambda} = \sum_{x=0}^{\infty} \frac{\lambda}{x}$$

## A look at the Poisson distribution





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- A Poisson ( $\lambda$ ) random variable counts the number of occurrences that happen over a fixed interval of time or space.
- These occurrences must:
  - be independent
  - be sequential in time (no two occurrences at once)
  - ightharpoonup occur at the same constant rate,  $\lambda$ .
- $\lambda$ , the **rate parameter**, is the expected number of occurrences in the specified interval of time or space.

- Y is the number of shark attacks off the coast of CA <u>next</u> year.  $\lambda = 100$  attacks per year.
- $\blacktriangleright$  Z is the number of shark attacks off the coast of CA next month.  $\lambda=100/12=8.3333333$  attacks per month
- N is the number of  $\beta$  particles emitted from a small bar of plutonium, registered by a Geiger counter, in a minute.  $\lambda = 459.21$  particles/minute.
- ▶ *J* is the number of particles per three minutes.  $\lambda = ?$

$$\begin{split} \lambda &= \frac{\text{459.21 (units particle)}}{1 \text{ (unit minute)}} \cdot \frac{\text{3 (units minute)}}{1 \text{ (unit of 3 minutes)}} \\ &= \frac{1377.63 \text{ (units particle)}}{1 \text{ (unit of 3 minutes)}} = 1377.62 \text{ particles per 3 minutes} \end{split}$$

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- Rutherford and Geiger measured the number of  $\alpha$  particles detected near a small bar of plutonium for 8-minute periods.
- The average number of particles per 8 minutes was  $\lambda=3.87$  particles / 8 min
- Let  $S \sim \text{Poisson}(\lambda)$ , the number of particles detected in the next 8 minutes.

$$f(s) = \begin{cases} \frac{e^{-3.87}(3.87)^s}{s!} & s = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

P(at least 4 particles recorded)

$$= P(S \ge 4) = I - P(S \le 4).$$

$$= f(4) + f(5) + f(6) + \cdots$$

$$= 1 - f(0) - f(1) - f(2) - f(3)$$

$$= 1 - \frac{e^{-3.87}(3.87)^{0}}{0!} - \frac{e^{-3.87}(3.87)^{1}}{1!}$$

$$- \frac{e^{-3.87}(3.87)^{2}}{2!} - \frac{e^{-3.87}(3.87)^{3}}{3!}$$

$$= 0.54$$

# Example: arrival at a university library

1 lo min

- Some students' data indicate that between 12:00 and 12:10 P.M. on Monday through Wednesday, an average of around 125 students entered a library at lowa State University library.
- ► Let *M* be the number of students entering the ISU library between 12:00 and 12:01 PM next Tuesday.
- ► Model  $M \sim \text{Poisson}(\lambda)$ .
- Having observed 125 students enter between 12:00 and 12:10 PM last Tuesday, we might choose:

```
\begin{split} \lambda &= \frac{125 \text{ (units of student)}}{1 \text{ (unit of 10 minutes)}} \cdot \frac{1 \text{ (unit of 10 minutes)}}{10 \text{ (units of minute)}} \\ &= \frac{12.5 \text{ (units of student)}}{1 \text{ (unit minute)}} = 12.5 \text{ students per minute} \end{split}
```

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▶ Under this model, the probability that between 10 and 15 students arrive at the library between 12:00 and 12:01 PM is:

$$P(10 \le M \le 15) = f(10) + f(11) + f(12) + f(13) + f(14) + f(15)$$

$$= \frac{e^{-12.5}(12.5)^{10}}{10!} + \frac{e^{-12.5}(12.5)^{11}}{11!} + \frac{e^{-12.5}(12.5)^{12}}{12!} + \frac{e^{-12.5}(12.5)^{13}}{13!} + \frac{e^{-12.5}(12.5)^{14}}{14!} + \frac{e^{-12.5}(12.5)^{15}}{15!}$$

$$= 0.60$$

(0)r

- ► Let *X* be the number of unprovoked shark attacks that will occur off the coast of Florida next year.
- ▶ Model  $X \sim \text{Poisson}(\lambda)$ .
- From the shark data at http://www.flmnh.ufl.edu/fish/sharks/statistics/FLactivity.htm, 246 unprovoked shark attacks occurred from 2000 to 2009.
- ► Hence, I calculate:

$$\begin{split} \lambda &= \frac{246 \text{ (units attack)}}{1 \text{ (unit of 10 years)}} \cdot \frac{1 \text{ (unit of 10 years)}}{10 \text{ (units year)}} \\ &= \frac{24.6 \text{ (units attack)}}{1 \text{ (unit year)}} = 24.6 \text{ attacks per year} \end{split}$$

$$P(\text{no attacks next year}) = f(0) = e^{-24.6} \cdot \frac{24.6^{0}}{0!}$$

$$\approx 2.07 \times 10^{-11} \text{ P(x > 5)} = 1 - \text{P(x < 5)}$$

$$P(\text{at least 5 attacks}) = 1 - P(\text{at most 4 attacks})$$

$$= 1 - F(4)$$

$$= 1 - f(0) - f(1) - f(2) - f(3) - f(4)$$

$$= 1 - e^{-24.6} \frac{24.6^{0}}{0!} - e^{-24.6} \frac{24.6^{1}}{1!} - e^{-24.6} \frac{24.6^{2}}{2!}$$

$$- e^{-24.6} \frac{24.6^{3}}{3!} - e^{-24.6} \frac{24.6^{4}}{4!}$$

$$\approx 0.99999996$$

P(more than 30 attacks) = 1 - P(at least 30 attacks)

$$=1-e^{-24.6}\sum_{i=0}^{30}\frac{24.6^{x}}{x!}=1-e^{-24.6}\cdot 4.251\times 10^{10}$$

 $\approx 0.1193$ 

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- Now, let Y be the total number of shark attacks in Florida during the next 4 months.
- ▶ Let  $Y \sim \mathsf{Poisson}(\theta)$ , where  $\theta$  is the true shark attack rate per 4 months:

$$\begin{split} \theta &= \frac{24.6 \text{ (units attack)}}{1 \text{ (unit year)}} \cdot \frac{1/3 \text{ (unit year)}}{1 \text{ (unit of 4 months)}} \\ &= \frac{8.2 \text{ (units attack)}}{1 \text{ (unit of 4 months)}} = 8.2 \text{ attacks per 4 months} \end{split}$$

$$P(\text{no attacks next year}) = f(0) = e^{-8.2} \cdot \frac{8.2^0}{0!}$$

$$\approx 0.000275$$

 $\approx 0.9113$ 

$$P(\text{at least 5 attacks}) = 1 - P(\text{at most 4 attacks})$$

$$= 1 - F(4)$$

$$= 1 - f(0) - f(1) - f(2) - f(3) - f(4)$$

$$= 1 - e^{-8.2} \frac{8.2^{0}}{0!} - e^{-8.2} \frac{8.2^{1}}{1!} - e^{-8.2} \frac{8.2^{2}}{2!}$$

$$- e^{-8.2} \frac{8.2^{3}}{3!} - e^{-8.2} \frac{8.2^{4}}{4!}$$

P(more than 30 attacks) = 1 - P(at least 30 attacks)

$$= 1 - e^{-8.2} \sum_{i=0}^{30} \frac{8.2^{x}}{x!} = 1 - e^{-8.2} \cdot 4.251 \times 10^{10}$$
  
 
$$\approx 9.53 \times 10^{-10}$$

$$f(x) = e^{-\lambda} \cdot \frac{\lambda^{x}}{x!}, \quad x = 0, 1, 2, \dots$$

$$E(x) = \sum_{x=0}^{\infty} \frac{x}{x!} \cdot \frac{\lambda^{x}}{x!} \cdot e^{-\lambda}$$

$$= \sum_{x=1}^{\infty} \frac{\lambda^{x}}{(x-i)!} e^{-\lambda} = \sum_{x=0}^{\infty} \frac{\lambda^{x+i}}{x!} e^{-\lambda}$$

$$= \lambda \left| \frac{\sum_{x=0}^{\infty} \lambda^{x}}{\sum_{x=0}^{\infty} x!} e^{-\lambda} \right| = 1$$

$$\mathcal{E}(\chi(\chi-I)) = \bar{\mathcal{E}}(\chi^2) - \mathcal{E}(\chi).$$

$$\sum_{x=2}^{\infty} \frac{x(x-1)}{x!} e^{-\lambda}$$

$$= \sum_{x=2}^{\infty} \frac{\lambda^{x}}{(x-2)!} e^{-\lambda}$$

$$= \sum_{x=0}^{\infty} \frac{\lambda^{x+2}}{x!} e^{-\lambda} = \lambda^{2} \underbrace{\sum_{x=0}^{\infty} \frac{\lambda^{x}}{x!} e^{\lambda}}_{= \lambda}$$

$$E(x^{2}) = \lambda^{2} + \lambda$$

$$\widehat{E}(x^2) = \lambda^2 + \lambda.$$

$$V_{or}(x) = E(x^2) - (\widehat{E}(x))^2 = \lambda^2 + \lambda - \lambda^2 - \lambda$$