Discrete random variables

Definition 1.1 Let X, Y be two discrete RVs on Ω . We call the function $p : \mathbb{R} \times \mathbb{R} \to \mathbb{R}$ joint probability mass function, short joint PMF, of X, Y, if it holds

$$\begin{array}{lcl} p(x,y) &:= & P(X=x,Y=y) \\ &= & P(\{\omega \in \Omega | X(\omega)=x\} \cap \{\omega \in \Omega | Y(\omega)=y\}) \end{array}$$

Definition 1.3 Let X,Y be two discrete RVs on a common sample space Ω with joint PMF p. The marginal PMF of X is given by

$$p_X(x) = \sum_{y \in R_Y} p(x, y) ,$$

and similarly the marginal PMF of Y is given by

$$p_Y(y) = \sum_{x \in B} p(x, y)$$
.

Example 1.1 (Three coin tosses) We continue our coin toss examples. This time we have a chance experiment with three independent coin tosses, leading to the sample space $\Omega = \{HHH, HHT, HTT, \dots\}$. We introduce two random variables

$$X : \Omega \rightarrow \mathbb{R}$$
, $Y : \Omega \rightarrow \mathbb{R}$.

where the first random variable gives the number of "heads" in the first two tosses and the second RV gives the number of "heads" in the last two tosses.²⁾
We would like to compute their ioint PMF:

$$\begin{array}{ll} p(0,0) &=& P(X=0,Y=0) = P\big(\{\omega\in\Omega|X(\omega)=0\}\cap\{\omega\in\Omega|Y(\omega)=0\}\big) \\ &=& P\big(\{TTH,TTT\}\cap\{HTT,TTT\}\big) = P\big(\{TTT\}\big) = \frac{1}{\circ} \end{array}$$

Theorem 1.1 Let (X,Y) be discrete RVs with joint PMF p, and let further $g:R_X\times R_Y\to\mathbb{R}$ be given. We can compute the expectation value / expectation / mean of g(X,Y), if it exists, by

$$\mathrm{E}\left(g(X,Y)\right) = \sum_{x} \sum_{y} g(x,y) p(x,y) \,.$$

Existence is guaranteed, if $\sum_x \sum_y |g(x,y)| p(x,y) < \infty.$

Definition 1.4 Let (X,Y) be discrete RVs with joint PMF p. We define the conditional distribution of X given Y=y by

$$p(x|y) = P(X = x|Y = y) = \frac{p(x, y)}{p_Y(y)}$$
.

Furthermore, we define the conditional expectation of X given Y = y by

$$\mathbf{E}\left(X|y\right) = \mathbf{E}\left(X|Y=y\right) = \sum_{x} x \, p(x|y) = \frac{\sum_{x} x \, p(x,y)}{p_{Y}(y)} = \frac{\sum_{x} x \, p(x,y)}{\sum_{x} p(x,y)} \; .$$

= 25 c = 1 =) (= 1=

Px(x) = = p(x,y)

3) E(X) = Zx p(x,1)

2) What to show: X, Y are dependent

Py(1) = 5 p(x,1) = 6(1,1) = 25

 $b^{\times}(\sqrt{3}) = \frac{3}{5}b(\sqrt{3}) = b(\sqrt{3}) \cdot b(\sqrt{3}) \cdot b(\sqrt{3})$

 $\frac{1}{25} + \frac{3}{25} + \frac{3}{35} = \frac{0}{25}$

Px(1) · Py(1) = 6 25 25 7 25 = p(1,1) >dependent

= 1. p(1,1) + 1. f(1,2) + 1. p(1,3) - 2. p(2,2) - 2. p(2,3)

 $E(7) = \underbrace{2}_{3,7} + \underbrace{\rho(x, 7)}_{3,7} = \underbrace{2}_{2,5} = \underbrace{3}_{2,5} + \underbrace{2}_{2,7} \underbrace{\rho(x, 2) + 2}_{3,7} \underbrace{\rho(x, 2) + 3}_{4,2} \underbrace{\rho(x, 2) + 3}_{5,7} \underbrace{\rho(x, 2) + 3}_{6,3} \underbrace{\rho(x, 2) + 3}_$

 $E(XY) = \sum_{x,y} x \cdot y \, p(x,y) = \sum_{x=1}^{4} \frac{1}{25} = 1 \cdot \frac{4}{35} + 1 \cdot \frac{2}{25} + 3 \cdot \frac{3}{25} + 2 \cdot \frac{2}{25}$

1. 1. p(1,1) +1.2. p(1,2) + 1.3. p(1,3) +2.2. p(2,2) + 2.3. p(2,3) -33. p(3.3)

Continuous random variables

Definition 1.11 Let $X = (X_1, \dots, X_n)$ be an n-dimensional random vector of continuous RVs $X_i : \Omega \to \mathbb{R}$. We call the function $\rho : \mathbb{R}^n \to \mathbb{R}$ **joint density**, short **density**, of X, if it holds

$$P(a_1 \le X_1 \le b_1, \dots, a_n \le X_n \le b_n)$$

$$= \int_{a_n}^{b_n} \dots \int_{a_1}^{b_1} \rho(x_1, \dots, x_n) dx_1 \dots dx_n$$

Theorem 1.3 Let $\mathbf{X} = (X_1, \dots, X_n)$ be an n-dimensional random vector of continuous RVs with joint density ρ . We are further given a function $g: \mathbb{R}^n \to \mathbb{R}$. We can compute the expectation value f expectation f mean of $g(X_1, \dots, X_n)$, if it exists, by

$$E\left(g(X_1,\ldots,X_n)\right) = \int_{-\infty}^{\infty} g(x_1,\ldots,x_n)\rho(x_1,\ldots,x_n)dx_1\cdots dx_n.$$

Existence is guaranteed, if $\int_{\mathbb{R}^n} |g(x_1, \dots, x_n)| p(x_1, \dots, x_n) dx_1 \cdots dx_n < \infty$.

Theorem 1.4 (Bayes) Let (X,Y) be continuous RVs with joint density ρ . For all $x,y\in\mathbb{R}$, such that $\rho_X(x),\rho_Y(y)>0$, it holds

$$\rho(y|x) = \frac{\rho(x|y)\rho_Y(y)}{\rho_X(x)}.$$





Definition 1.13 Let $X = (X_1, \ldots, X_n)$ be an n-dimensional random vector of continuous RVs with joint density ρ . We are further given $1 \le p < n$. The marginal joint density of (X_1, \ldots, X_n) is defined by

$$\rho_{1,2,\ldots,p}(x_1,\ldots,x_p) = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} \rho(x_1,\ldots,x_n) dx_{p+1} \cdots dx_n.$$

Definition 1.14 Let (X,Y) be continuous RVs with joint density ρ . We define the conditional density of X given Y=y by

$$\rho(x|y) = \rho(x|Y=y) = \frac{\rho(x,y)}{\rho_Y(y)}\,,$$

for all y such that $\rho_Y(y)>0$. Furthermore, we define the **conditional expectation of** X **given** Y=y by

$$\mathrm{E}\left(X|y\right) = \mathrm{E}\left(X|Y=y\right) = \int_{-\infty}^{\infty} x \, \rho(x|y) dx = \frac{\int_{-\infty}^{\infty} x \, \rho(x,y) dx}{\int_{-\infty}^{\infty} \rho(x,y) dx} \, ,$$

for all y such that $\rho_Y(y) > 0$.

Let (XY) be two continuous RVs with joint densit

$$y) = \begin{cases} cxy & \text{if } x, y \in [0, 1] \\ 0 & \text{otherwise} \end{cases}$$

 $\rho(x, y) = \begin{cases} cxy & \text{if } x, y \in [0, 1] \\ 0 & \text{else} \end{cases}$ 2. Find the marginal densities and expectations of X.Y Find the conditional expectation of X given Y = v.] [[(x,y) dxdy =] [(xy dxdy =] [xy dxdy = cfy fxdxdy = cfy [2x4] dy 3) Px(N= Je(myldy = Ju. xydy = [4x== 2]0=2x Py (V= [P(x,y)dx =]+ xyd ~= [+ y = x] ~= Ly

 $\text{Proof} \ \text{Proof} \ \text{Proof}$

F[11 X=] P(x,y)

$$| \frac{1}{3} | \frac{$$

J 2 2 6 (1) 4x91 = 1