

# Variance of a random variable

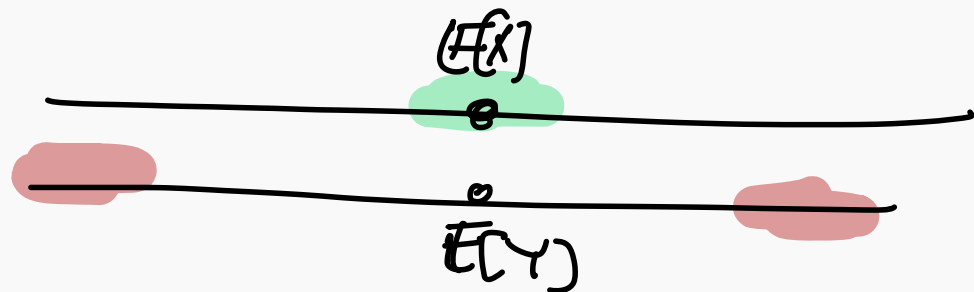
## Definition

The variance of a random variable  $X$  is defined by

$$\text{Var}[X] = \mathbb{E} \left[ (X - \mathbb{E}[X])^2 \right].$$

$$\mathbb{E}[X] = \sum p_i x_i$$

$$\mathbb{E}[X - \mathbb{E}[X]] = 0.$$



## Alternative formula for the variance

### Theorem

The *variance* of a random variable  $X$  is given by

$$\text{Var}[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2.$$

Set  $\mu = \mathbb{E}[X]$

definition  $\text{Var}[X] = \mathbb{E}[(X - \mu)^2] = \mathbb{E}[X^2 - 2\mu X + \mu^2]$

$$= \mathbb{E}[X^2] - 2\mu \underbrace{\mathbb{E}[X]}_{\mathbb{E}[\mu^2]} + \mu^2 \underbrace{\mathbb{E}[1]}_{\mathbb{E}[1] = 1 \cdot 1 = 1}$$

C  $= \mathbb{E}[X^2] - 2\mu^2 + \mu^2 \cdot 1$

$$= \mathbb{E}[X^2] - 2\mu^2 + \mu^2 = \mathbb{E}[X^2] - \mu^2$$

C constant  $\mathbb{E}[C] = C$   $= \mathbb{E}[X^2] - \mathbb{E}[X]^2.$

# Properties of the variance

## Theorem

Let  $X, Y$  be random variables and  $c$  a constant. We have

1.  $\text{Var}[cX] = c^2 \text{Var}[X]$ ,
2.  $\text{Var}[X] \geq 0$ , with equality when  $X = c$  for a constant  $c$ .
3.  $\text{Var}[X + a] = \text{Var}[X]$ , translation invariance.

$$\begin{aligned}\text{Var}[cX] &= \mathbb{E}[(cX)^2] - \mathbb{E}[cX]^2 \\ &= \mathbb{E}[c^2 X^2] - c^2 \mathbb{E}[X]^2 \\ \mu = \mathbb{E}[X] \quad &= c^2 \mathbb{E}[X^2] - c^2 \mathbb{E}[X]^2 = c^2 (\mathbb{E}[X^2] - \mathbb{E}[X]^2) \\ &= c^2 \text{Var}[X].\end{aligned}$$

$$\text{Var}[X] = \mathbb{E}[(X - \mu)^2] \geq 0. \quad \mathbb{E}[X^2] \geq \mathbb{E}[X]^2$$

$$\mathbb{E}[X + a] = \mathbb{E}[X] + \mathbb{E}[a] = a + \mathbb{E}[X]$$

$$\text{Var}[X + a] = \mathbb{E}[(X + a - \mathbb{E}[X + a])^2] = \mathbb{E}[(X - \mathbb{E}[X])^2] = \text{Var}[X].$$

# Expected value of continuous random variable

## Definition

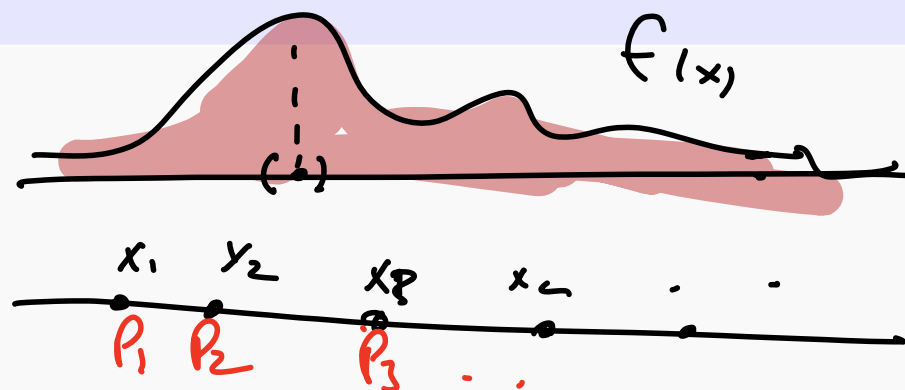
Let  $X$  be a continuous random variable with the probability density function  $f(x)$ . The **expected value of  $X$**  is defined by

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} xf(x)dx$$

why is this a good definition?

$$\mathbb{E}[x] = \sum x_i \cdot p_i$$

physics



$m(x)$  line

$$\int_{-\infty}^{\infty} x \cdot m(x) dx = \text{Center of mass}$$


mass density

## Example

### Example

Let  $X$  have uniform distribution over the interval  $[a, b]$ . Compute  $\mathbb{E}[X]$ .

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \mathbb{E}[X] &= \int_{-\infty}^{+\infty} x \cdot f(x) dx = \int_a^b x \cdot \frac{1}{b-a} dx \\ &= \frac{1}{b-a} \left. \frac{x^2}{2} \right|_a^b = \frac{1}{b-a} \cdot \frac{b^2 - a^2}{2} = \frac{b+a}{2} \end{aligned}$$


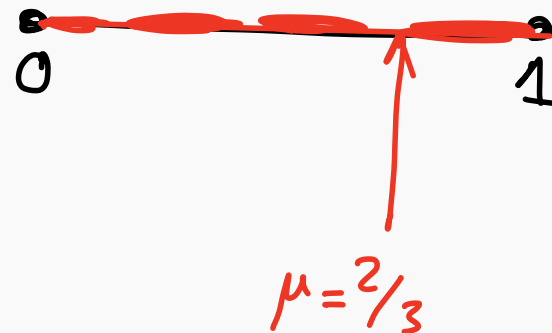
## Example

Suppose  $X$  is a continuous with the density function

$$f_X(x) = \begin{cases} 2x & \text{if } 0 \leq x \leq 1 \\ 0 & \text{if otherwise} \end{cases}$$

Find  $\mathbb{E}[X]$ .

$$\begin{aligned} \mathbb{E}[X] &= \int_{-\infty}^{\infty} x f(x) dx = \int_0^1 2x^2 dx \\ &= \left. \frac{2x^3}{3} \right|_0^1 = \frac{2}{3} \end{aligned}$$



## Example

Let  $X$  have the Cauchy distribution given by

$$f_X(t) = \frac{1}{\pi(1 + x^2)}.$$

Find  $\mathbb{E}[X]$ .

# Alternative formula for the expected value of non-negative random variables

## Theorem

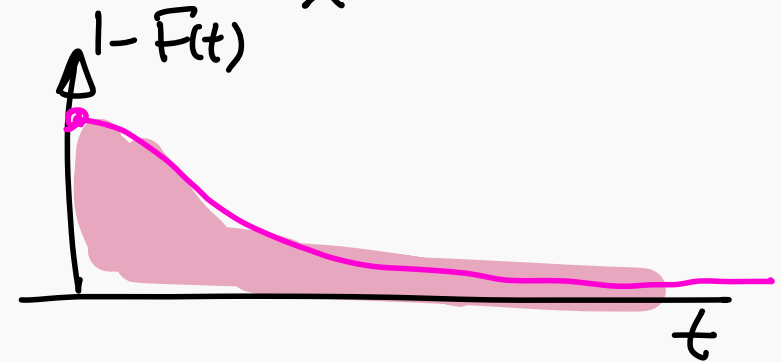
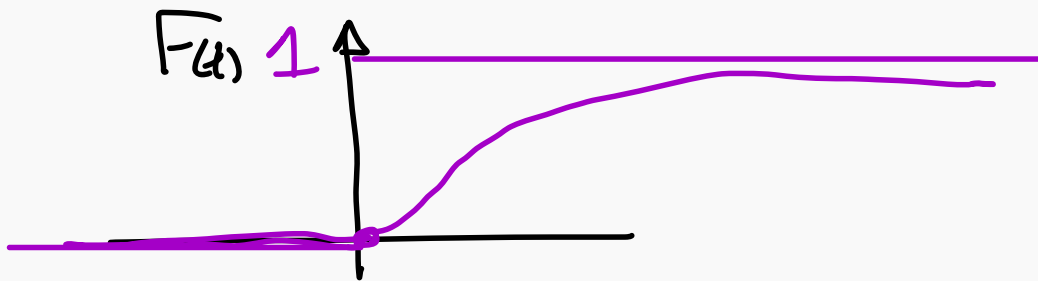
Let  $X$  be a **non-negative** continuous random variable with the distribution function  $F_X(t)$ . Then

$$\mathbb{E}[X] = \int_0^{\infty} \mathbb{P}[X \geq t] dt = \int_0^{\infty} (1 - F_X(t)) dt.$$

$X$  continuous

$$F_X(t) = \mathbb{P}[X \leq t]$$

$$\mathbb{P}[X \geq t] = 1 - F_X(t)$$



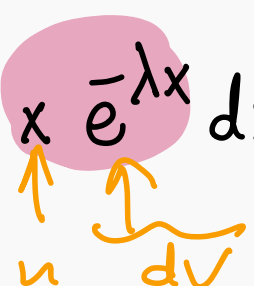


## Example

Find the expected value of an exponential random variable.

$$X \quad f(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0 \\ 0 & x \leq 0 \end{cases} \quad F(t) = \begin{cases} 1 - e^{-\lambda t} & t > 0 \\ 0 & t \leq 0 \end{cases}$$

$$E[X] = \int_0^{\infty} \lambda x e^{-\lambda x} dx$$

 integration by parts.

$$\begin{aligned} E[X] &= \int_0^{\infty} (1 - F(t)) dt = \int_0^{\infty} (1 - (1 - e^{-\lambda t})) dt \\ &= \int_0^{\infty} e^{-\lambda t} dt = \left. \frac{e^{-\lambda t}}{-\lambda} \right|_0^{\infty} = \frac{1}{\lambda}. \end{aligned}$$

## Expected value of $h(X)$ .

### Theorem

Let  $X$  be a continuous random variable with the density function  $f(x)$ . For any continuous function  $h : \mathbb{R} \rightarrow \mathbb{R}$ , we have

$$\mathbb{E}[h(X)] = \int_{-\infty}^{\infty} h(x)f(x) dx.$$

$$\mathbb{E}[X] = \int_{-\infty}^{+\infty} x f(x) dx$$

$X$  random variable with density  $f(x)$ ,

$\downarrow$

$$X^2 = h(X) \quad h(x) = x^2$$

$$\sin X = g(x) \quad g(x) = \sin x$$

## Example

### Example

Let  $X$  have the uniform distribution over the interval  $[0, \pi]$ . Compute  $\mathbb{E}[\sin X]$ .

$$\mathbb{E}[\sin X] = \int_0^{\pi} \sin x \cdot \frac{1}{\pi} dx = \frac{1}{\pi} \left( -\cos x \right) \Big|_0^{\pi}$$

$$f(x) = \begin{cases} \frac{1}{\pi} & 0 \leq x < \pi \\ 0 & \text{elsewhere} \end{cases} \quad = \frac{1}{\pi} (1 + 1) = \frac{2}{\pi}.$$

## The expected value of log-normal distributions

Let  $Y = e^X$ , where  $X$  has the  $N(0, 1)$  distribution. Find  $\mathbb{E}[Y]$ .

## Example

Let  $X$  be a continuous random variable with the uniform distribution on  $[0, 1]$ .  
Find  $\text{Var}[X]$ .

$$\text{Var}[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

$$f(x) = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & \text{elsewhere} \end{cases}$$

$$\mathbb{E}[X] = \int_0^1 x f(x) dx = \int_0^1 x dx = \frac{x^2}{2} \Big|_0^1 = \frac{1}{2}$$

$$\mathbb{E}[X^2] = \int_0^1 x^2 f(x) dx = \int_0^1 x^2 dx = \frac{x^3}{3} \Big|_0^1 = \frac{1}{3}$$

$$\text{Var}[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{1}{3} - \left(\frac{1}{2}\right)^2 = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}.$$

## Expected value as the value of a game

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Consider a lottery with the possible outcomes

$$x_1, \dots, x_n$$

attained with probabilities  $p_1, \dots, p_n$ .

Ex.

	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$
$x$	1	2	3	4	5	6
Payoff	1	4	9	16	25	-15

$$\begin{aligned} E[x] &= \frac{1}{6} \cdot 1 + \frac{1}{6} \cdot 4 + \frac{1}{6} \cdot 9 + \frac{1}{6} \cdot 16 + \frac{1}{6} \cdot 25 - \frac{1}{6} \cdot 15 > 0 \\ &= \frac{1}{6}(40) = \frac{40}{6} = \frac{20}{3} = 6.66. \end{aligned}$$

## Expected value as the value of a game

Consider a lottery with the possible outcomes

$$x_1, \dots, x_n$$

attained with probabilities  $p_1, \dots, p_n$ .

The expected value of the lottery is given by

$$E := p_1 x_1 + \dots + p_n x_n.$$



# Decision making using expected payoff

# The idea of utility and its history

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Suppose a coin is thrown until a head occurs. If this happens at round  $n$  for the first time then the player gets  $2^n$  dollars.

	$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{8}$	$\frac{1}{16}$	$\frac{1}{32}$	$\dots$
Round	1	2	3	4	5	$\dots$
Payoff	2	4	8	16	32	$\dots$

$$\begin{aligned} E[\text{Payoff}] &= 2 \cdot \frac{1}{2} + 4 \cdot \frac{1}{4} + 8 \cdot \frac{1}{8} + 16 \cdot \frac{1}{16} + \dots \\ &= 1 + 1 + 1 + 1 + \dots = \infty. \end{aligned}$$

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If tails occurs in the first round, and heads in the second round then the player get 4 dollars.

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How much are you willing to pay to play this game?

The value of the game is

$$2 \times \frac{1}{2} + 4 \times \frac{1}{4} + 8 \times \frac{1}{8} + \cdots = \infty$$



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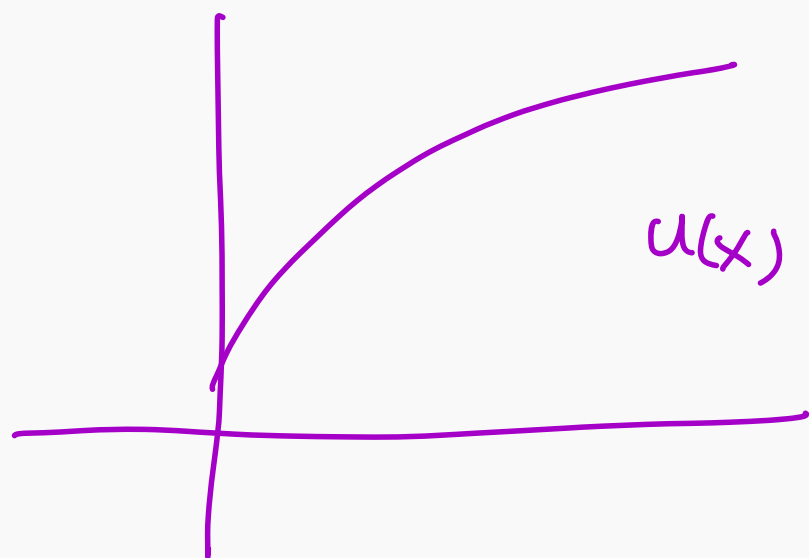
In reality people only pay around 3-4 dollars to play this game.

Bernoulli: it is not the absolute value of money but its utility

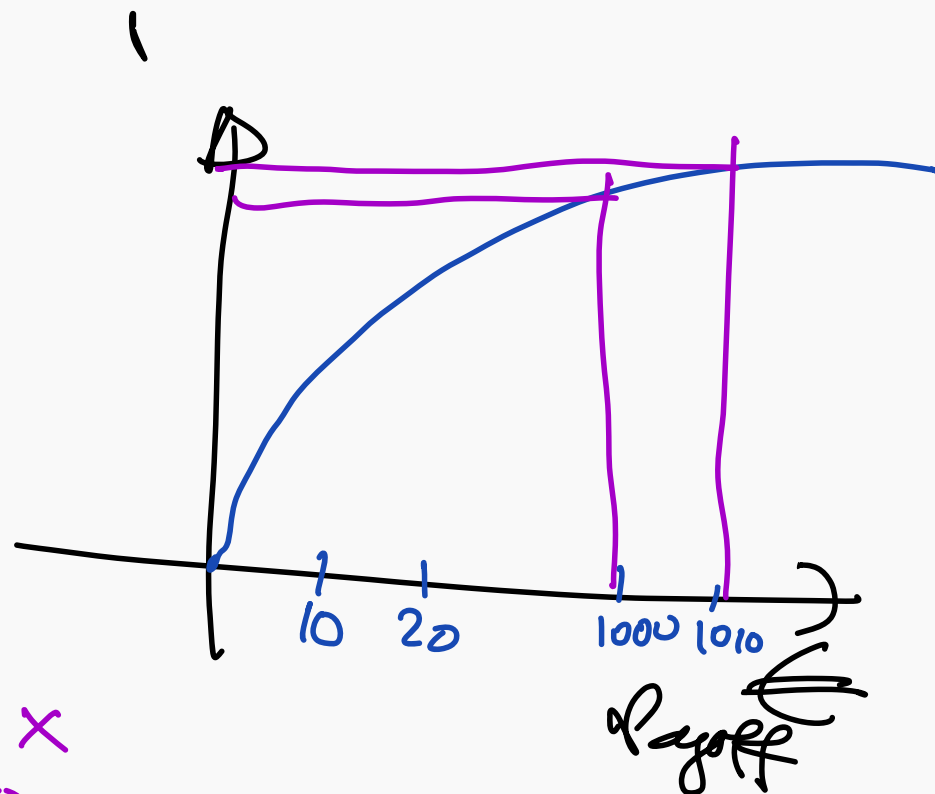
# Solution to the paradox

$E[\text{Payoff}] \rightsquigarrow$

$E[\underline{\text{util}}(\text{Payoff})]$



$$u(x) = \log_{10} x$$



+1  
10  $\rightarrow$  100

100  $\rightarrow$  1000

# Solution to the paradox

If the utility is

$$u(x) = 2 \log_2 x$$

then:

1. With probability  $1/2$ , one gets  $u(2) = \log_2 2 = 2$
2. With probability  $1/4$ , one gets  $u(4) = \log_2 4 = 4$
3. With probability  $1/8$ , one gets  $u(8) = \log_2 8 = 6$ .

So the expected utility is

$$\frac{2}{2} + \frac{4}{4} + \frac{6}{8} + \frac{8}{16} + \dots = 4.$$

Round	$1/2$	$1/4$	$1/8$	$1/16$	$1/32$	$\dots$
	1	2	3	4	5	$\dots$
Payoff	2	4	8	16	32	$\dots$
	2	3	4	5	6	$\dots$

$$\frac{2}{2} + \frac{3}{4} + \frac{4}{8} + \frac{5}{16} + \dots$$

= finite sum



Suppose you are offered the choice between the following alternatives:

1. 1 dollars.
2. first throwing a fair coin. If the outcomes is heads 2 dollars and if the outcome is tails zero.

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- 
1. 100000 dollars.
  2. first throwing a fair coin. If the outcomes is heads 200000 dollars and if the outcome is tails zero.

# Lotteries

Suppose  $A_1, A_2, \dots, A_n$  denotes possible options in a choice decision making. A lottery is a scenario where each outcome will happen with a given probability. We write  $L = (A_1, p_1; \dots, A_n; p_n)$ .

For two lotteries  $L$  and  $M$  we write  $M > L$  if  $M$  is preferable to  $L$  and  $M \sim L$  if the decision maker is indifferent between  $M$  and  $L$ .

Combination of lotteries:

$$A_1 > A_2 > A_3 > \dots > A_{n-1} > A_n$$

$$\begin{array}{ccccc} A_1 & A_5 & \text{vs.} & A_2 & A_3 \\ \frac{1}{2} & \frac{1}{2} & & \frac{1}{3} & \frac{2}{3} \\ \hline 0.9 & 0.10 & & & \end{array}$$



1. For any two lotteries  $L$  and  $M$  we have  $L < M$ ,  $M < L$  or  $M \sim L$ .
- ✓ 2. If  $L < M$  and  $M < N$  then  $L < N$ .
3. If  $L < M < N$  then there exists  $p \in [0, 1]$  such that  $pL + (1 - p)N \sim M$ .
4. For any  $N$  and  $0 < p \leq 1$  we have  $L < M$  iff  
 $pL + (1 - p)N \leq pM + (1 - p)N$ .

# Von Neumann–Morgenstern utility theorem

For any preference satisfying axioms 1-4 above, there exists a function  $u : \{A_1, \dots, A_n\} \rightarrow \mathbb{R}$  such that  $L < M$  iff  $\mathbb{E}[u(L)] \leq \mathbb{E}[u(M)]$ .