

Bayesian inference for stochastic differential random effects models

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- ▷ Consider an Itô process $\{\mathbf{X}_t, t \geq 0\}$ satisfying

$$d\mathbf{X}_t = \boldsymbol{\alpha}(\mathbf{X}_t, \boldsymbol{\theta})dt + \sqrt{\boldsymbol{\beta}(\mathbf{X}_t, \boldsymbol{\theta})}d\mathbf{W}_t$$

- $\boldsymbol{\alpha}(\mathbf{X}_t, \boldsymbol{\theta})$ is the drift
 - $\boldsymbol{\beta}(\mathbf{X}_t, \boldsymbol{\theta})$ is the diffusion coefficient
 - \mathbf{W}_t is standard Brownian motion
- ▷ Seek a numerical solution via the Euler-Maruyama approximation

$$\Delta\mathbf{X}_t \equiv \mathbf{X}_{t+\Delta t} - \mathbf{X}_t = \boldsymbol{\alpha}(\mathbf{X}_t, \boldsymbol{\theta})\Delta t + \sqrt{\boldsymbol{\beta}(\mathbf{X}_t, \boldsymbol{\theta})}\Delta\mathbf{W}_t$$

where $\Delta\mathbf{W}_t \sim N(\mathbf{0}, \mathbf{I}\Delta t)$

- ▶ Consider the case where we have ℓ subjects and that each individual can be represented by the same SDE
- ▶ Common parameters θ
- ▶ Different parameters $\psi^{(i)}$, $i = 1, \dots, \ell$
- ▶ This gives us a stochastic differential random effects model:

$$d\mathbf{X}_t^{(i)} = \alpha\left(\mathbf{X}_t^{(i)}, \theta, \psi^{(i)}\right) dt + \sqrt{\beta\left(\mathbf{X}_t^{(i)}, \theta, \psi^{(i)}\right)} d\mathbf{W}_t^{(i)}$$

for $i = 1, \dots, \ell$

- ▶ Suppose we have data available at times $t_0^{(i)}, t_1^{(i)}, \dots, t_{n_i}^{(i)}$ for each individual i
- ▶ We implement a data augmentation approach

- ▷ Consider $[t_j^{(i)}, t_{j+1}^{(i)}]$ and introduce a partition

$$t_j^{(i)} = \tau_{j,0}^{(i)} < \underbrace{\tau_{j,1}^{(i)} < \dots < \tau_{j,m_j^{(i)}-1}^{(i)}}_{\text{latent times}} < \tau_{j,m_j^{(i)}}^{(i)} = t_{j+1}^{(i)}$$

- ▷ Time step between observations

$$\Delta_{\tau_j}^{(i)} = \frac{t_{j+1}^{(i)} - t_j^{(i)}}{m_j^{(i)}}$$

- ▷ Allows for irregularly spaced data for each individual

- ▷ For individual i

$$\mathbf{d}^{(i)} = \left(\mathbf{x}_{t_0}^{(i)}, \mathbf{x}_{t_1}^{(i)}, \dots, \mathbf{x}_{t_{n_i}}^{(i)} \right)$$

$$\mathbf{x}^{(i)} = \left(\mathbf{x}_{\tau_{0,1}}^{(i)}, \mathbf{x}_{\tau_{0,2}}^{(i)}, \dots, \mathbf{x}_{\tau_{0, m_0^{(i)} - 1}}^{(i)}, \mathbf{x}_{\tau_{1,1}}^{(i)}, \dots, \mathbf{x}_{\tau_{n_i - 1, m_{n_i - 1}^{(i)} - 1}}^{(i)} \right)$$

- ▷ $\mathbf{x}^{(i)}$ is the values of the skeleton path at times $\tau_{0,1}^{(i)}, \tau_{0,2}^{(i)}$ etc
- ▷ Putting these together for ℓ individuals

$$\mathbf{d} = \left(\mathbf{d}^{(1)}, \mathbf{d}^{(2)}, \dots, \mathbf{d}^{(\ell)} \right)$$

$$\mathbf{x} = \left(\mathbf{x}^{(1)}, \mathbf{x}^{(2)}, \dots, \mathbf{x}^{(\ell)} \right)$$

Formulate joint posterior for parameters and latent data as

$$\begin{aligned} \pi(\boldsymbol{\theta}, \boldsymbol{\psi}, \mathbf{x} | \mathbf{d}) &\propto \pi(\boldsymbol{\theta}) \pi(\boldsymbol{\psi} | \boldsymbol{\theta}) \pi(\mathbf{x}, \mathbf{d} | \boldsymbol{\theta}, \boldsymbol{\psi}) \\ &\propto \underbrace{\pi(\boldsymbol{\theta}) \pi(\boldsymbol{\psi} | \boldsymbol{\theta})}_{\text{prior}} \times \prod_{i=1}^{\ell} \prod_{j=0}^{n_i-1} \prod_{k=0}^{m_j^{(i)}-1} \underbrace{\pi\left(\mathbf{x}_{\tau_{j,(k+1)}}^{(i)} \middle| \mathbf{x}_{\tau_{j,k}}^{(i)}, \boldsymbol{\theta}, \boldsymbol{\psi}^{(i)}\right)}_{\text{Euler density}} \end{aligned}$$

where

$$\mathbf{x}_{\tau_{j,(k+1)}}^{(i)} \middle| \mathbf{x}_{\tau_{j,k}}^{(i)}, \boldsymbol{\theta}, \boldsymbol{\psi}^{(i)} \sim N\left(\mathbf{x}_{\tau_{j,k}}^{(i)} + \boldsymbol{\alpha}_{j,k}^{(i)} \Delta_{t_j}^{(i)}, \boldsymbol{\beta}_{j,k}^{(i)} \Delta_{t_j}^{(i)}\right)$$

Note $\boldsymbol{\alpha}_{j,k}^{(i)} = \boldsymbol{\alpha}\left(\mathbf{x}_{\tau_{j,k}}^{(i)}, \boldsymbol{\theta}, \boldsymbol{\psi}^{(i)}\right)$ and $\boldsymbol{\beta}_{j,k}^{(i)} = \boldsymbol{\beta}\left(\mathbf{x}_{\tau_{j,k}}^{(i)}, \boldsymbol{\theta}, \boldsymbol{\psi}^{(i)}\right)$

- ▷ The posterior distribution is typically analytically intractable
- ▷ Use a Gibbs sampler, alternating between draws of
 - $\theta|x, d, \psi$
 - $\psi|x, d, \theta$
 - $x|\theta, d, \psi$
- ▷ We can update ψ componentwise, for each individual

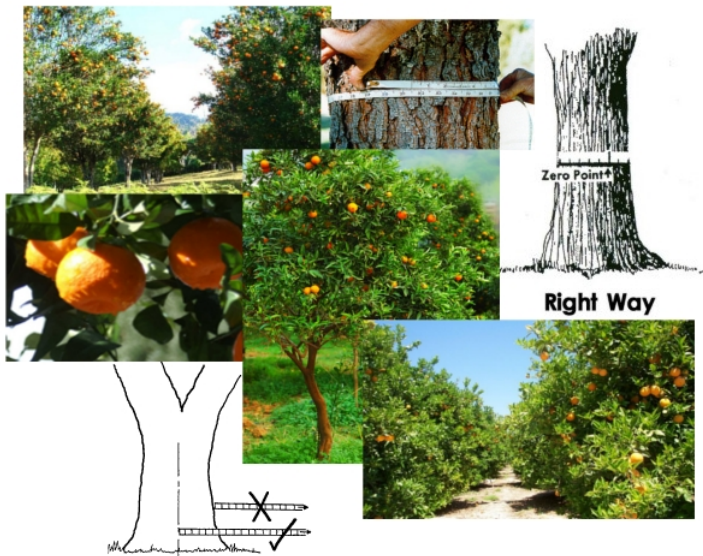
$$\pi(\psi|x, d, \theta) = \prod_{i=1}^{\ell} \pi\left(\psi^{(i)}|x^{(i)}, d^{(i)}, \theta\right)$$

- ▷ Similarly

$$\pi(x|\theta, d, \psi) = \prod_{i=1}^{\ell} \pi\left(x^{(i)}|\theta, d^{(i)}, \psi^{(i)}\right)$$

- ▷ Typically Metropolis within Gibbs updates are needed

Example: orange tree growth



- ▷ Picchini and Ditlevsen (2011) discuss a model for the growth of orange trees incorporating random effects. We use an equivalent reparameterisation, written as

$$dX_t^{(i)} = \frac{1}{\phi_1^{(i)} \phi_2^{(i)}} X_t^{(i)} \left(\phi_1^{(i)} - X_t^{(i)} \right) dt + \sigma \sqrt{X_t^{(i)}} dW_t^{(i)}, \quad i = 1, \dots, 100$$

- ▷ $\phi_1^{(i)} \sim N(\phi_1, \sigma_{\phi_1}^2)$ and $\phi_2^{(i)} \sim N(\phi_2, \sigma_{\phi_2}^2)$
- ▷ X_t is the circumference (mm)
- ▷ t is the number of days since December 31st 1968
- ▷ $\phi_1^{(i)}$ is the asymptotic circumference
- ▷ $\phi_2^{(i)}$ is the rate of change parameter

$$dX_t^{(i)} = \frac{1}{\phi_1^{(i)} \phi_2^{(i)}} X_t^{(i)} \left(\phi_1^{(i)} - X_t^{(i)} \right) dt + \sigma \sqrt{X_t^{(i)}} dW_t^{(i)}, \quad i = 1, \dots, 100$$

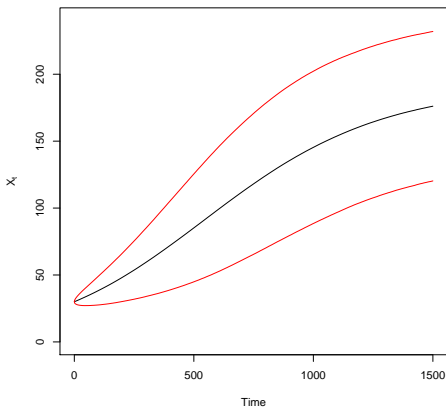
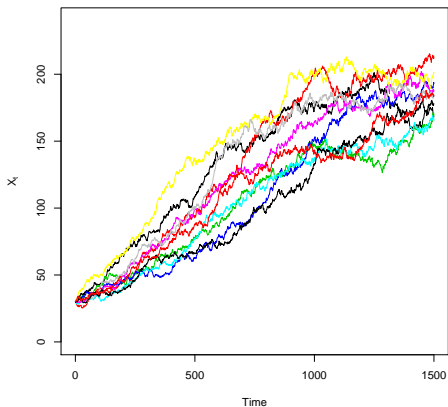
- ▶ Common parameters: ϕ_1 , ϕ_2 , σ_{ϕ_1} , σ_{ϕ_2} and σ
- ▶ Tree specific parameters: $\phi_1^{(i)}$ and $\phi_2^{(i)}$
- ▶ We repeat Picchini and Ditlevsen's simulation study for 100 trees with $x_0^{(i)} = 30$,

$$(\phi_1, \phi_2, \sigma_{\phi_1}, \sigma_{\phi_2}, \sigma) = (195, 350, 25, 52.5, 0.08),$$

$$\phi_1^{(i)} \sim N(195, 25^2) \text{ and } \phi_2^{(i)} \sim N(350, 52.5^2)$$

- ▶ This gives us 16 observations at intervals of 100 days on 100 trees

Figure: **left:** 10 simulated skeleton paths
right: mean, upper and lower 2.5 percentiles for 1k realisations



▷ Recall we wish to sample

- $\theta|x, d, \psi$
- $\psi|x, d, \theta$
- $x|\theta, d, \psi$

▷ We can formulate the joint posterior as

$$\begin{aligned}\pi(\theta, \psi, x|d) &\propto \pi(\phi_1)\pi(\phi_2)\pi(\sigma_{\phi_1})\pi(\sigma_{\phi_2})\pi(\sigma) \\ &\quad \times \pi(\phi_1|\phi_1, \sigma_{\phi_1})\pi(\phi_2|\phi_2, \sigma_{\phi_2}) \\ &\quad \times \prod_{i=1}^{100} \underbrace{\pi\left(x^{(i)}|\phi_1^{(i)}, \phi_2^{(i)}, \sigma, d^{(i)}\right)}_{\text{Euler density}}\end{aligned}$$

- ▷ Sample common parameters θ via Gibbs updates
- ▷ Sample each $\psi^{(i)}$ using random normal proposals
- ▷ Use fairly uninformative priors
 - ϕ_1 and ϕ_2 are Normally distributed
 - $\sigma_{\phi_1}^{-2}$, $\sigma_{\phi_2}^{-2}$ and σ^{-2} follow Gamma distributions
- ▷ Recall

$$dX_t^{(i)} = \frac{1}{\phi_1^{(i)} \phi_2^{(i)}} X_t^{(i)} \left(\phi_1^{(i)} - X_t^{(i)} \right) dt + \sigma \sqrt{X_t^{(i)}} dW_t^{(i)}, \quad i = 1, \dots, 100$$

with $\phi_1^{(i)}$ and $\phi_2^{(i)}$ Normally distributed

- ▷ Sample x by updating each path x^i separately, conditional on $\phi_1^{(i)}$, $\phi_2^{(i)}$ and σ

(Modified) Innovation scheme for random effects SDEs

- ▶ This type of scheme suffers from intolerably poor mixing as $m \rightarrow \infty$
- ▶ Caused by relation between the parameters and the path in the quadratic variation \Rightarrow acceptance probability $\rightarrow 0$ as $m \rightarrow \infty$
- ▶ The (Modified) Innovation scheme (Golightly and Wilkinson (2008)) conditions on the Brownian increments

$$\boldsymbol{w} = \left(\boldsymbol{w}^{(1)}, \boldsymbol{w}^{(2)}, \dots, \boldsymbol{w}^{(\ell)} \right)$$

(which drive a tractable conditioned diffusion) to overcome the dependence between the parameters and the path

- ▶ **Insight:** the quadratic variation of \boldsymbol{w} does not itself determine any of the model parameters and should therefore be effective in decoupling the problematic dependence

\Rightarrow the scheme should not become degenerate as $m \rightarrow \infty$

Application: orange tree growth

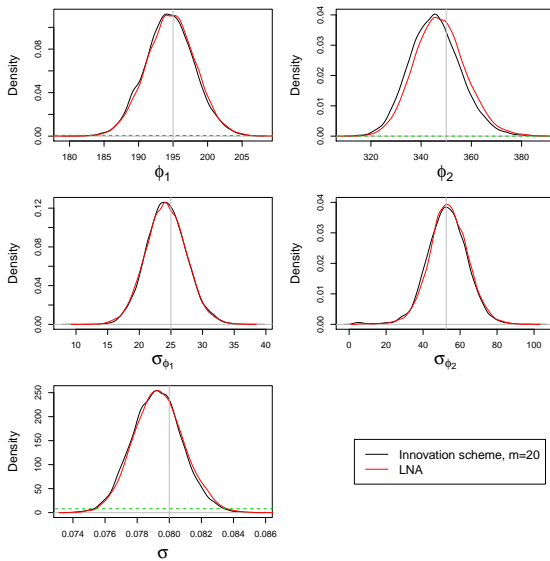
- ▷ Recall we have 16 observations on 100 trees at intervals of 100 days
- ▷ Parameter choice:

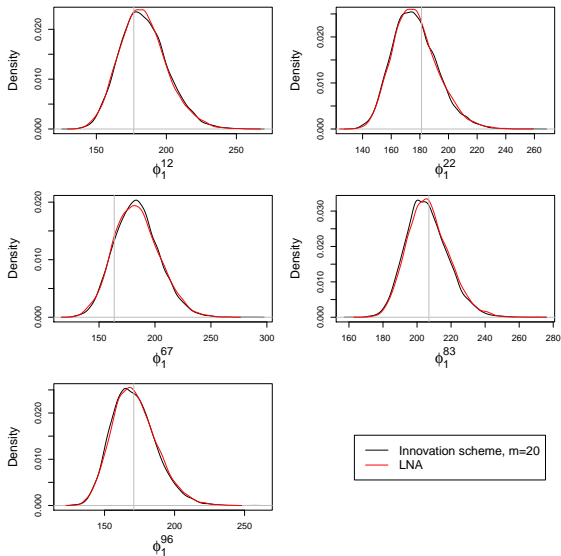
$$(\phi_1, \phi_2, \sigma_{\phi_1}, \sigma_{\phi_2}, \sigma) = (195, 350, 25, 52.5, 0.08)$$

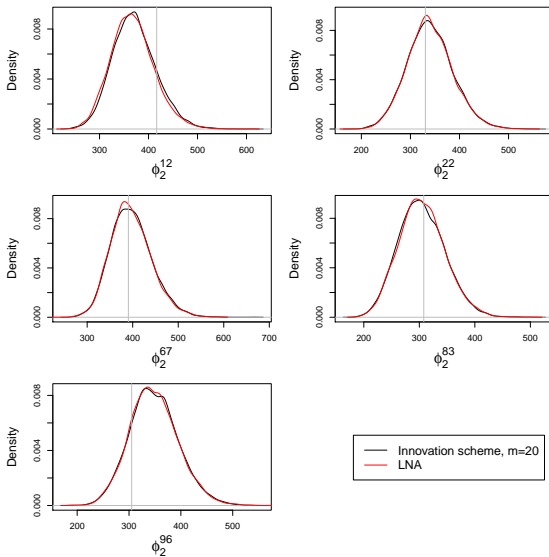
$$\phi_1^{(i)} \sim N(195, 25^2)$$

$$\phi_2^{(i)} \sim N(350, 52.5^2)$$

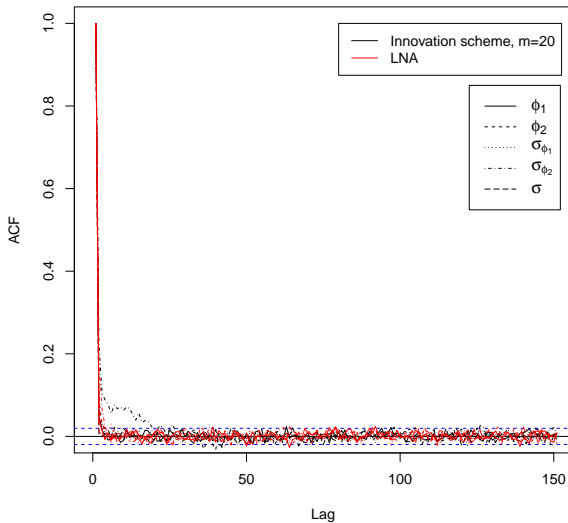
- ▷ We run the (Modified) Innovation scheme with $m = 20$, for 1 million iterations with a thin of 100
- ▷ We compare the output with that obtained under the linear noise approximation (LNA)



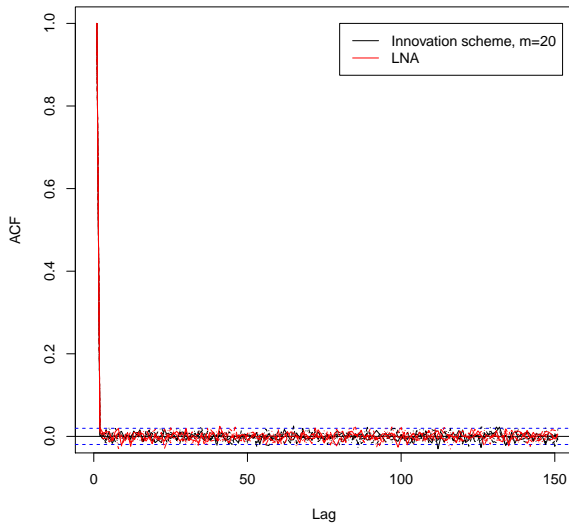




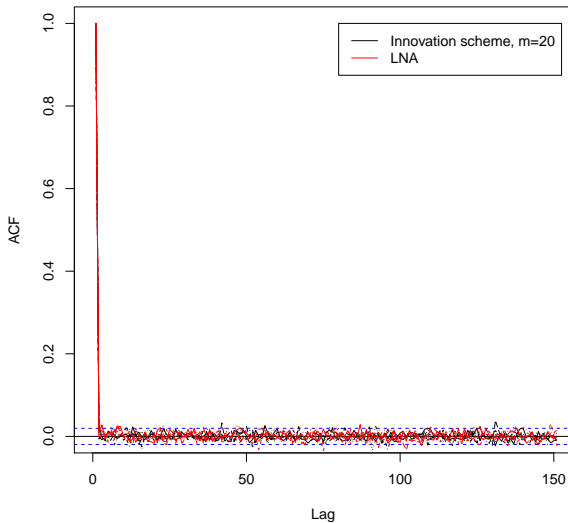
Auto-correlation for θ



Auto-correlation for ϕ_1^i , $i=12,22,67,83,96$



Auto-correlation for ϕ_2^i , $i=12,22,67,83,96$



Comparison of the schemes

- ▶ Comparing the LNA against the (Modified) Innovation scheme
 - LNA: $\text{ESS/sec} = 1.178$
 - (Modified) Innovation scheme: $\text{ESS/sec} = 0.057$

which gives a comparable $\text{ESS/sec} = 20.667$

- ▶ The gain by using the LNA is slightly exaggerated by the fact that the system of ODEs can be solved analytically
- ▶ It should still be advantageous to use the LNA over the (Modified) Innovation scheme on a system that can't be solved analytically
- ▶ However it should be noted that the LNA only gives an approximation to the posterior density

Future work

- ▷ Extend these methods to data where we only have partial observations
- ▷ Examine the case where we have observations observed with error, typically Gaussian error
- ▷ Apply these schemes to a more challenging example e.g. larger model
- ▷ Compare these methods with pMCMC
- ▷ Apply these schemes to lemon trees

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