Comparison of weighted regression and additive models for trend evaluation of water quality in tidal waters

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Abstract

Key words:

1 Introduction

Needs

- Quantitative tools that describe trends in water quality time series are needed to identify factors that influence ecosystem condition and to evaluate the effects of management activities in the context of multiple drivers

- Recent adaptation of statistical models for evaluating water quality time series have shown promise for application in tidal waters, specifically generalized additive models (GAM) and weighted regression on time, discharge, and season (WRTDS) {acro:GAM} {acro:WRTI}

- These similar techniques can be used to quantify relationships between response measures and different drivers that may vary over time, in addition to an evaluation of trends independent of variation in freshwater inputs

- The relative merits of each approach have not been evaluated, particularly related to accuracy of the empirical description and the desired products for trend evaluation

- Such a comparison could inform the use of each model for addressing management or restoration needs or for developing more robust descriptions of long-term changes in ecosystem characteristics

Goal: Provide a description of the relative abilities of GAMs and WRTDS to describe long-term changes in time series of response endpoints in tidal waters

Objectives:

- Provide a narrative comparison of the statistical foundation of each technique, both as a general description and as a means to evaluate water quality time series

- Use each technique to develop an empirical description of water quality changes in a common dataset with known historical changes in water quality drivers

- Apply the models to simulated data to evaluate ability of the models to describe true changes

- Compare each technique's ability to describe changes, as well as the differences in the information provided by each

- Provide recommendations on the most appropriate context for using each method

2   Methods

2.1   Study site

The Patuxent River Estuary... background, history

Observed trends over time

longitudinal gradient from watershed to mainstem influences, LE1.2, TF1.6

Show plots of trends over time in observed data

2.2   Model descriptions

How, Similarities, differences, optimal smoothing

The selection of optimal model parameters is a challenge that represents a tradeoff between model precision and ability to generalize to novel datasets. Weighted regression requires identifying optimal half-window widths, whereas GAMs requires identifying the optimal degrees of freedom for the smoothing parameter. Overfitting a model with excessively small window widths or excessive degrees of freedom will minimize prediction error but prevent extrapolation of results to different datasets. Similarly, underfitting a model with large window widths or very few degrees of freedom will reduce precision but will improve the ability to generalize results to different datasets. From a statistical perspective, the optimal model parameters provide a balance between over- and under-fitting. Both models use a form of cross-validation to identify model parameters that maximize the precision of model predictions with a novel dataset.

The basic premise of cross-validation is to identify the optimal set of model parameters that minimize prediction error on a dataset that was not used to develop the model. For GAMs (Hastie and Tibshirani 1990, Zuur 2012)...[insert GAMs methods]. Similarly, the tidal adaptation of WRTDS used k-fold cross-validation to identify the optimal half-window widths. For a given set of half-window widhts, the dataset was separated into ten disjoint sets, such that ten models were evaluated for every combination of k - 1 training and remaining test datasets. That is, the training dataset for each fold was all k - 1 folds and the test dataset was the remaining fold,

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repeated k times. The average prediction error of the test datasets across k folds provided an indication of model performance for the given combination of half-window widths. The optimum window widths were those that provided minimum errors on the test data. Evaluating multiple combinations of window-widths can be computationally intensive. An optimization function was implemented in R (Byrd et al. 1995, RDCT (R Development Core Team) 2015) to more efficiently evaluate model parameters using a search algorithm. Window widths were searched using the limited-memory modification of the BFGS quasi-Newton method that imposes upper and lower bounds for each parameter. The chosen parameters were based on a selected convergence tolerance for the error minimization of the search algorithm.

2.3   Comparison of modelled trends

Explanatory power of each method - explained variance/fit in the response, histograms of errors (see page 14 in Moyer) - we can test for significant differences in the errors using a two-sided t-test. Also see page 24/25 in Moyer for average difference comparisons between methods.

Similarity of predictions - observed data, simple scatterplots, similarity coefficients, similarity by time periods, etc.

Indications of change - direction/magnitude of trends by different time periods

2.4   Comparison of flow-normalized trends

The relative abilities of each model to characterize flow or salinity-normalized trends in chlorophyll were evaluated using simulated datasets with known components. This approach was used because the flow-independent component of chlorophyll is typically not observed in raw data such that the true signal must be empirically estimated. Accordingly, the ability of each model to isolate the flow-normalized trend cannot be evaluated with reasonable certainty unless the true signal is known. Simulated time series of observed chlrophyll (Chl_{obs}) were created as additive components related to flow (Chl_{flo}, analogous to salinity) and a flow-independent biological component of chlorophyll (Chl_{bio}):

$$Chl_{obs} = Chl_{flo} + Chl_{bio} \qquad\qquad (1) \quad \{\text{chlobs}\}$$

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A distinction between Chl_{flo} and Chl_{bio} is that the former describes variation in the observed time series with changes in discharge (e.g., concentration dilution with increased flow) and the latter describes a true, desired measure of chlorophyll in the water column that is directly linked to primary production. The biological component of chlorophyll is comparable to an observation in a closed system that is not affected by flow and is the time series that is estimated by flow-normalization with WRTDS and GAMs.

The simulated time series was based on a stochastic model derived from actual flow and water quality measurements to ensure the statistical properties were comparable to existing datasets. This approach allowed us to evaluate GAMs and WRTDS under different sampling regimes (e.g., monthly rather than daily), while ensuring the simulated datasets had statistical properties that were consistent with known time series. Daily flow observations were obtained from the US Geological Survey (USGS) stream gage station 01594440 near Bowie, Maryland {acro:USGS (38°57′21.3″N, 76°41′37.3″W) from 1985 to 2014. Daily chlorophyll records were obtained from the Jug Bay station (38°46′50.6″N, 76°42′29.1″W) of the Chesapeake Bay Maryland National Estuarine Research Reserve. Daily chlorophyll concentrations were estimated from fluorescence values that did not include blue-green algae blooms. Our primary concern was simulating chlorophyll concentrations at monthly or bimonthly timesteps such that taxa-specific concentrations on a daily time step were not relevant.

Four time series were estimated or simulated from the actual datasets to create the complete, simulated time series: 1) estimated discharge as a stationary seasonal component ($\hat{Q}_{seas}$), 2) simulated error structure from the residuals of the seasonal discharge model ($\varepsilon_{Q,sim}$), 3) estimated chlorophyll independent of discharge as a stationary seasonal component ($\hat{Chl}_{seas}$), and 4) simulated error structure from the residuals of the seasonal chlorophyll model ($\varepsilon_{Chl,sim}$). Unless otherwise noted, chlorophyll and discharge are in ln-transformed units. Each of the four components was used to simulate the components in eq. (1):

$$Chl_{flo} = I\left(\hat{Q}_{seas} + \sigma \cdot \varepsilon_{Q,sim}\right) \tag{2} \quad \{\text{chlflo}\}$$

$$Chl_{bio} = \hat{Chl}_{seas} + \sigma \cdot \varepsilon_{Chl,sim} \tag{3} \quad \{\text{chlbio}\}$$

The estimated flow time series within the parentheses, $\hat{Q}_{seas} + \sigma \cdot \varepsilon_{Q,sim}$, is floored at zero to

simulate an additive effect of increasing flow on Chl_{obs}. Although the actual relationship of water quality measurements with flow is more complex, we assumed that a simple addition was sufficient for the simulations where the primary objective was to create an empirical and linear link between flow and chlorophyll. Moreover, the vector I (where $0 \leq I \leq 1$) can be manually changed to represent an independent effect of flow based on the desired simulation. For example, a flow effect that changes from non-existent to positive throughout the period of observation can be simulated by creating a vector ranging from zero to one. For the simulated Chl_{bio} time series, the seasonal and error components were characterized using the daily time series at Jug Bay that likely included an effect of flow in the observed data. For the simulated models, we assumed that the actual flow effect was part of the seasonal component to obtain an accurate estimate of the error component that was independent of both flow and season. Methods for estimating each of the components in eqs. (2) and (3) are described in detail below.

First, a model for simulating flow-related chlorophyll (eq. (2)) was estimated from the stream gage data as the additive combination of a stationary seasonal component and serially-correlated errors:

$$Q_{seas} = \beta_0 + \beta_1 \sin\left(2\pi T\right) + \beta_2 \cos\left(2\pi T\right) \qquad (4) \quad \{\texttt{qseas}\}$$

$$\varepsilon_Q = Q_{seas} - \hat{Q}_{seas} \qquad (5) \quad \{\texttt{qerr}\}$$

A seasonal model of flow was estimated using linear regression for time, T, on an annual sinusoidal period (eq. (4)). The residuals from this regression, ε_Q (eq. (5)), were used to estimate the structure of the error distribution for simulating the stochastic component of flow. The error distribution was characterized using an Autoregressive Moving Average (ARMA) model to {acro:ARMA} identify appropriate p and q coefficients (Hyndman and Khandakar 2008). The parameters were chosen using stepwise estimation for nonseasonal univariate time series that minimized Akaike {acro:AIC} Information Criterion (AIC). The resulting coefficients were used to generate random errors from a standard normal distribution for the length of the original time series, $\varepsilon_{Q,sim}$. These stochastic errors were multiplied by the standard deviation of the residuals in eq. (5) and added to the seasonal component in eq. (4) to create a simulated, daily time series of the flow-component for chlorophyll, Chl_{flo} (eq. (2)).

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The chlorophyll time series was created using a similar approach. The first step estimated the stationary seasonal component of the chlorophyll time series by fitting a WRTDS model (Hirsch et al. 2010) that explicitly included discharge from the gaged station using one year of data from the whole time series:

$$Chl_{seas} = \beta_0 + \beta_1 T + \beta_2 Q + \beta_3 \sin(2\pi T) + \beta_4 \cos(2\pi T) \qquad (6) \quad \{\texttt{chlseas}\}$$

$$\varepsilon_{Chl} = Chl_{seas} - \hat{Chl}_{seas} \qquad (7) \quad \{\texttt{chlerr}\}$$

This approach was used to isolate an error structure for simulation that was independent of flow and biology, where the seasonal component (as time T on a sinusoidal annual period) was assumed to be related to biological processes. The error distribution was then estimated from the residuals (eq. (7)) as before using an ARMA estimate of the residual parameters, p and q. Standard error estimates from the regression used at each point in the one-year time series were also retained for each residual. Errors were simulated ($\varepsilon_{Chl,\,sim}$, eq. (3)) for the entire year using the estimated auto-regressive structure and multiplied by the corresponding standard error estimate from the regression. The entire year was repeated for every year in the observed time series. All simulated errors were rescaled to the range of the original residuals that were used to estimate the distribution. Finally, the simulated flow-component, Chl_{flo}, was added to the simulated bilogical model, Chl_{bio}, to create the final chlorophyll-flow time series, Chl_{obs}, in eq. (1).

A daily time series for the entire period of record was simulated using the above methods and then used to compare the relative abilities of WRTDS and GAMs to characterize flow-normalized trends. Multiple time series with a monthly sampling frequencies and varying contributions of the flow component, (Chl_{flo} in eq. (1)) were created from the daily time series. One day in each month for each year was randomly sampled to create a monthly time series. Varying contributions of the flow component on observed chlorophyll were creating by multiplying Chl_{flo} by different indicator vectors (I in eq. (2)). The contribution of the flow component varied from constant, non-existent, steadily increasing, and steadily decreasing. Respectively, the vector of coefficients applied to each flow component was a constant vector of ones, a constant vector of zeroes, a linear increase starting at zero and ending at one, and a linear

decrease starting at one and ending at zero. This created four monthly time series that were used to evaluate each model. The flow-normalized results of WRTDS and GAMs for each simulated time series were compared to each other and to the original biological chlorophyll component of each time series (Chl_{bio}, eqs. (1) and (3)).

3   Results

Predictions with actual data

Simulations

4   Discussion

Qualitative comparison

- Computational requirements and potential limitations

- Data needs or transferability of each technique to novel datasets

- Products, e.g., conditional quantiles of WRTDS, confidence intervals for GAMs, handling censored data, hypothesis testing vs description

- Appropriate context for using each approach

4.1   Conclusions

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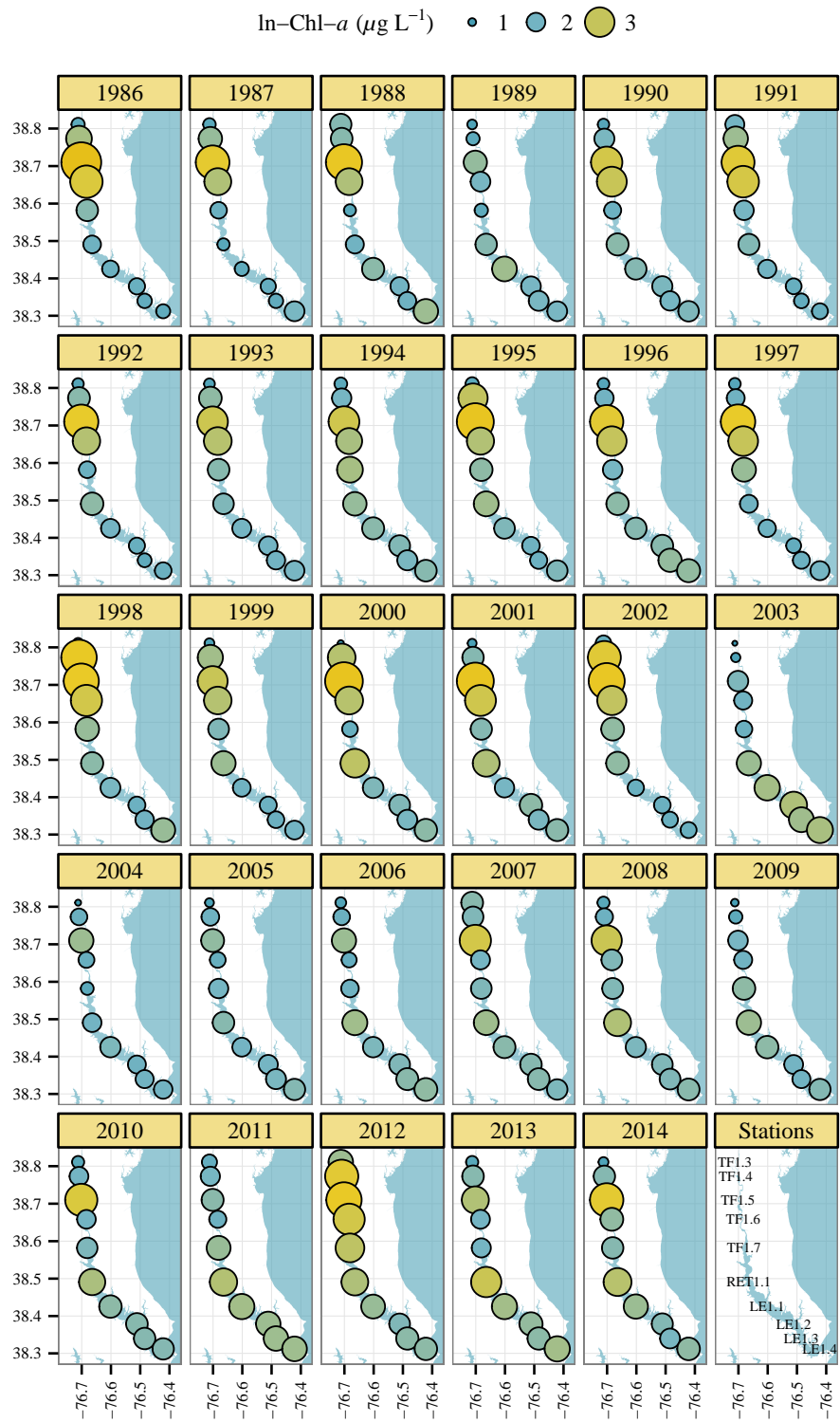
Fig. 1: Annual chlorophyll trends at each monitoring station in the Patuxent River Estuary. Values are annual medians of ln-chlorophyll-a with size and color proportional between years.
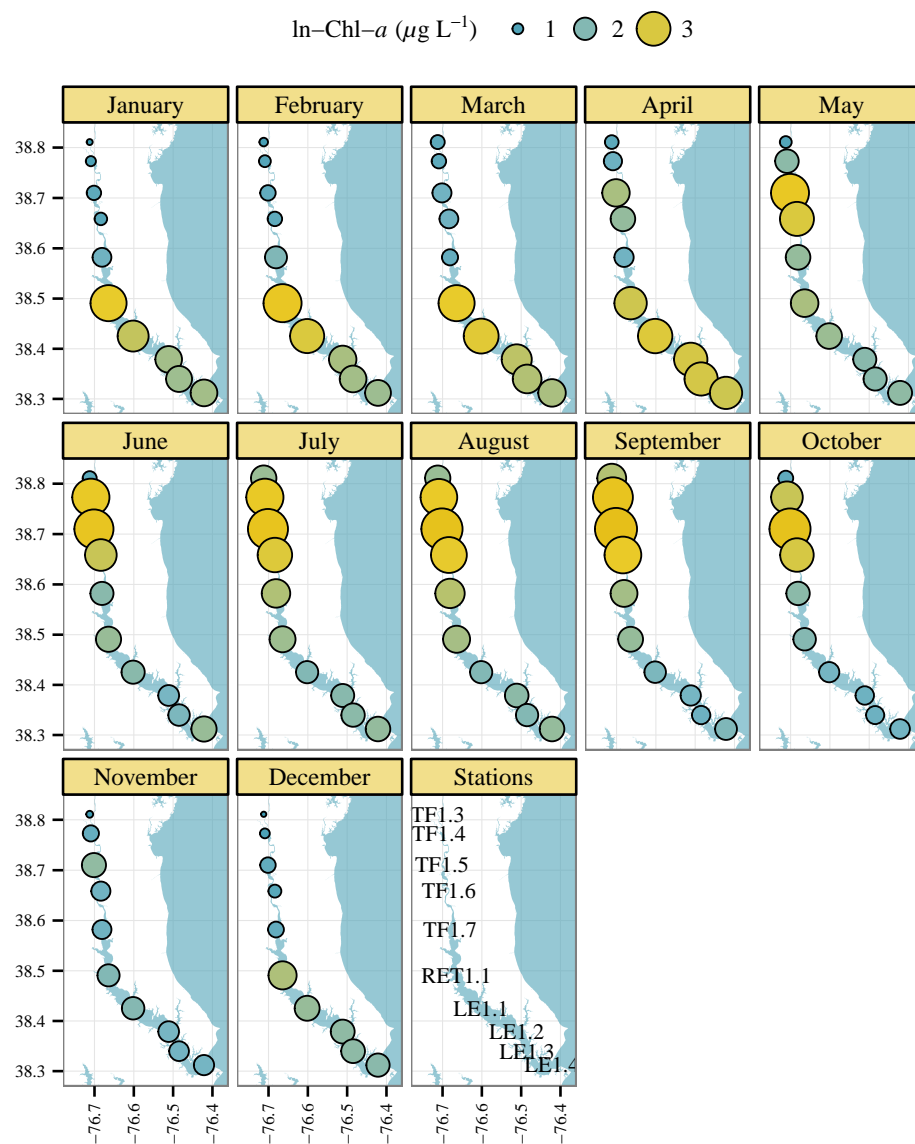
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Fig. 2: Monthly chlorophyll trends at each monitoring station in the Patuxent River Estuary. Values are monthly medians of ln-chlorophyll-a with size and color proportional between months.

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