

Fundamentals of Probability

Probability

Basics

- i) Sample space: set of all possible outcomes
Ex: M-ary PAM with $M = 4$ symbols,
 $S = \{-3\alpha, -\alpha, \alpha, 3\alpha\}$
- ii) Event: Is any subset of S
Ex: Event $A = \{-3\alpha, \alpha\}$
Event complement $\bar{A} = \{-\alpha, 3\alpha\}$
- iii) Set theory: Union and Intersection
 $A \cup B$ and $A \cap B$
- iv) NULL Event: $\phi = \{\}$
- v) Mutually Exclusive events: If $A \cap \bar{A} = \{\}$
Contain no common sample points

Axioms of Probability

Ax.1: $P(A) \geq 0$
Ax.2: $P(S) = 1$
Ax.3: $P(A \cup B) = P(A) + P(B)$
Iff $A \cap B = \phi$

Properties

- i) $P(A) \leq 1$
- ii) $P(\phi) = 0$
- iii) $P(A) < P(B)$ if $A \subset B$
- iv) $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

Example

- i) In 4-PAM, let $A = \{-3\alpha, \alpha\}$ and $B = \{\alpha\}$
- ii) $P(B|A) = \frac{P(B \cap A)}{P(A)} = \frac{1/8}{3/8} = \frac{2}{3}$
- iii) $P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{1/4}{1/4} = 1$
- iv) Note $P(A|B) \neq P(B|A)$

Statistically Independent Events

- i) $P(A|B) = P(A)$
i.e., Occurrence of B has no-effect on A
 $\Rightarrow \frac{P(A \cap B)}{P(B)} = P(A)$
 $\Rightarrow P(A \cap B) = P(A)P(B)$
- ii) $\Rightarrow \frac{P(A \cap B)}{P(A)} = P(B)$
 $P(B|A) = P(B)$
Similarly, A has no-effect on B
- iii) Note: If $P(A \cap B) = P(A)P(B)$
then $P(A \cap \bar{B}) = P(A)P(\bar{B})$

Example:

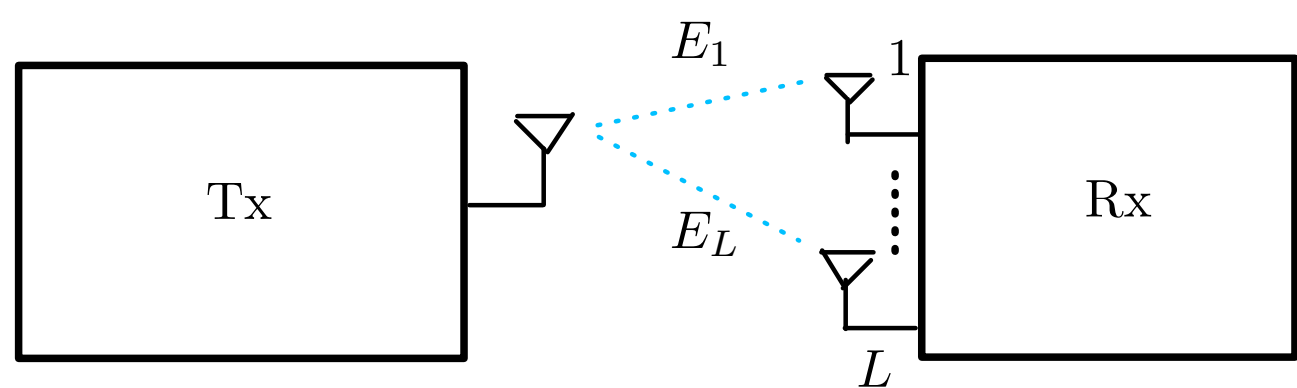
- i) For 4-PAM, let $P = \{\frac{1}{8}, \frac{1}{8}, \frac{1}{4}, \frac{1}{2}\}$ then
- ii) For $A = \{-3\alpha, \alpha\}$, $P(A) = \frac{3}{8}$
and $B = \{\alpha, 3\alpha\}$, $P(B) = \frac{3}{4}$
- iii) We have $P(A \cup B) = P(\{-3\alpha, \alpha, 3\alpha\}) = \frac{7}{8}$
 $= P(A) + P(B) - P(A \cap B) = \frac{7}{8}$

Example: Block Transmission

Symbols generated at different time instances are independent

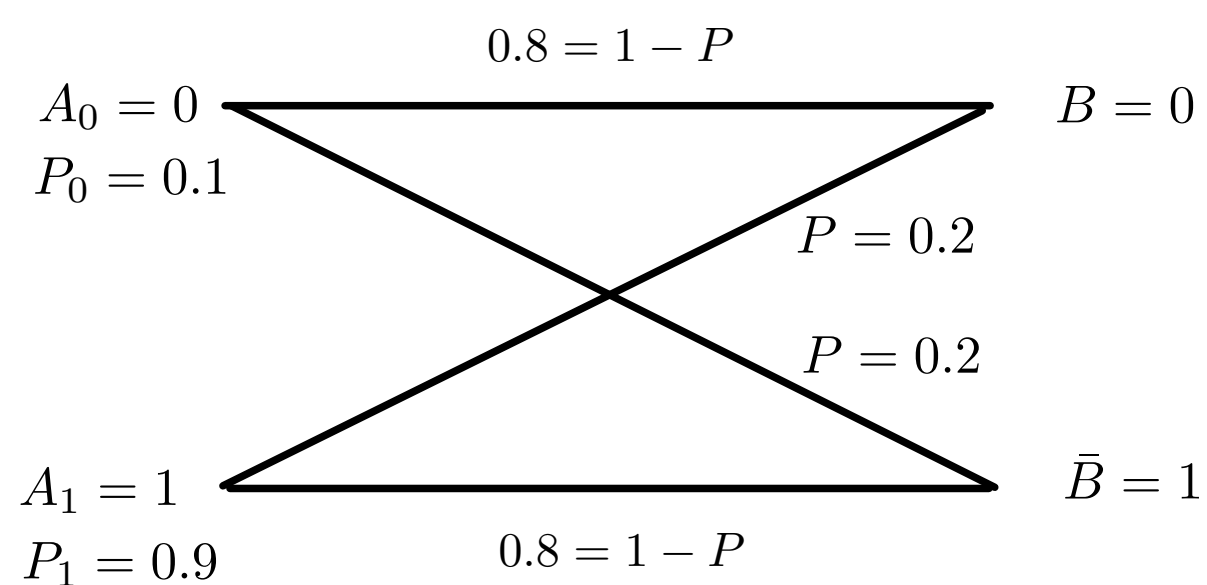
- i) $[x_1 x_2 \dots x_i]$ where $x_i \in S$ and is IID
Each x_i is generated at different time instance
- ii) IID - Independent Identically
Distributed from 4-PAM constellation S
- iii) $P(x_1 = 3\alpha, x_2 = \alpha) = \frac{1}{4} \frac{1}{2} = \frac{1}{8}$

Example: Wireless Systems (Deep Fade)



- i) SIMO: Single Input (Tx) - Multiple Output (Rx) System
- ii) E_i denotes the event that link i is in deep-fade
We have, $P(E_i) = \frac{1}{SNR} = \frac{1}{P/\sigma^2}$
where P is Tx Power and σ^2 is Noise Power
- iii) System is in deep-fade when all links are in deep-fade
 $P(\text{Deep-fade}) = P(E_1 \cap E_2 \cap \dots \cap E_L)$
 $P_{DF} = P(E_1)P(E_2) \dots P(E_L) = \frac{1}{SNR^L}$
This **Diversity** leads to significant decrease in BER

Example: MAP Receiver (Maximum A Posteriori Probability)

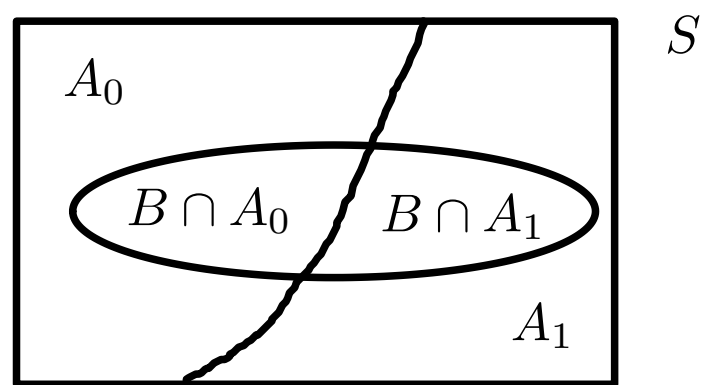


- i) Prior prob.: $P(A_0) = 0.1$ and $P(A_1) = 0.9$
- ii) Likelihood: $P(B|A_1) = P(\bar{B}|A_0) = 0.2$
 $P(B|A_0) = P(\bar{B}|A_1) = 0.8$
- iii) Compute Aposteriori probs.
 $P(A_0|B) = \frac{0.8 \times 0.1}{0.8 \times 0.1 + 0.2 \times 0.9} = \frac{8}{26}$
 $P(A_1|B) = \frac{0.2 \times 0.9}{0.8 \times 0.1 + 0.2 \times 0.9} = \frac{18}{26}$
- iv) \therefore Given B , MAP receiver decides A_1 was transmitted
- v) Whereas the ML receiver decides A_0 was transmitted
- vi) MAP is optimal and minimizes the prob. of error

Conditional Probability

- i) $P(A|B) \triangleq \frac{P(A \cap B)}{P(B)}$
 $\Rightarrow P(A \cap B) = P(A|B)P(B)$
- ii) $P(B|A) \triangleq \frac{P(B \cap A)}{P(A)}$
 $\Rightarrow P(B \cap A) = P(B|A)P(A)$
- iii) **Bayes Rule:**
 $P(A|B)P(B) = P(B|A)P(A)$
Note $P(A|B) \neq P(B|A)$

Bayes Theorem



- i) Let $A_0 \cup A_1 = S$ and $A_0 \cap A_1 = \phi$
Mutually exclusive and exhaustive
- ii) Note, for any event B,
 $(B \cap A_0)$ and $(B \cap A_1)$ are mutually exclusive
and $(B \cap A_0) \cup (B \cap A_1) = B$
 $\Rightarrow P(B) = P(B \cap A_0) + P(B \cap A_1)$
 $= P(B|A_0)P(A_0) + P(B|A_1)P(A_1)$
- iii) $\Rightarrow P(B) = P(B|A_0)P(A_0) + P(B|A_1)P(A_1)$
- iv) Using **Bayes rule** we get **Bayes Theorem** as,
 $P(A_0|B) = \frac{P(B|A_0)P(A_0)}{P(B)}$
 $= \frac{P(B|A_0)P(A_0)}{P(B|A_0)P(A_0) + P(B|A_1)P(A_1)}$
 $P(A_1|B) = \frac{P(B|A_1)P(A_1)}{P(B|A_0)P(A_0) + P(B|A_1)P(A_1)}$

Key Terms

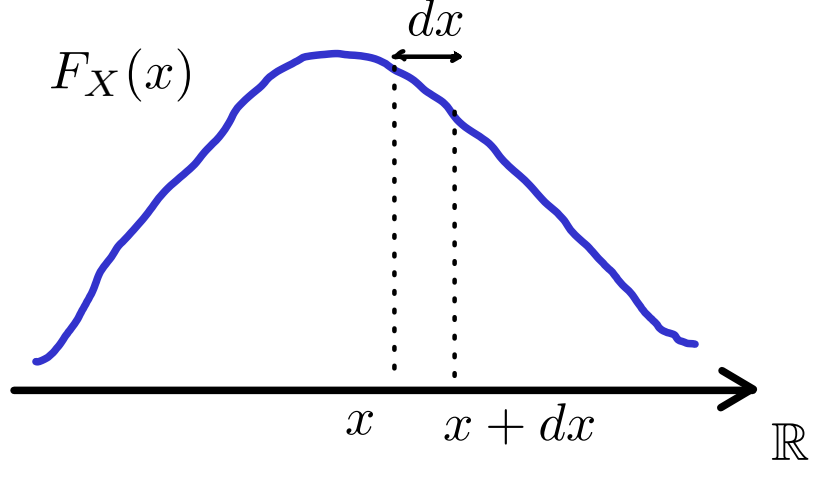
- i) $P(A_0|B)$ and $P(A_1|B)$ are called **Aposteriori** probabilities
- ii) $P(A_0)$ and $P(A_1)$ are called **Prior** probabilities
- iii) $P(B|A_0)$ and $P(B|A_1)$ are called **Likelihoods**

Generalized Bayes Theorem

For N mutually exclusive and exhaustive events
 $P(A_i|B) = \frac{P(B|A_i)P(A_i)}{\sum_{j=0}^{N-1} P(B|A_j)P(A_j)}$
Also note, $\sum_{i=0}^{N-1} P(A_i|B) = 1$

Random Variables

Characteristics



$$F_X(x)$$

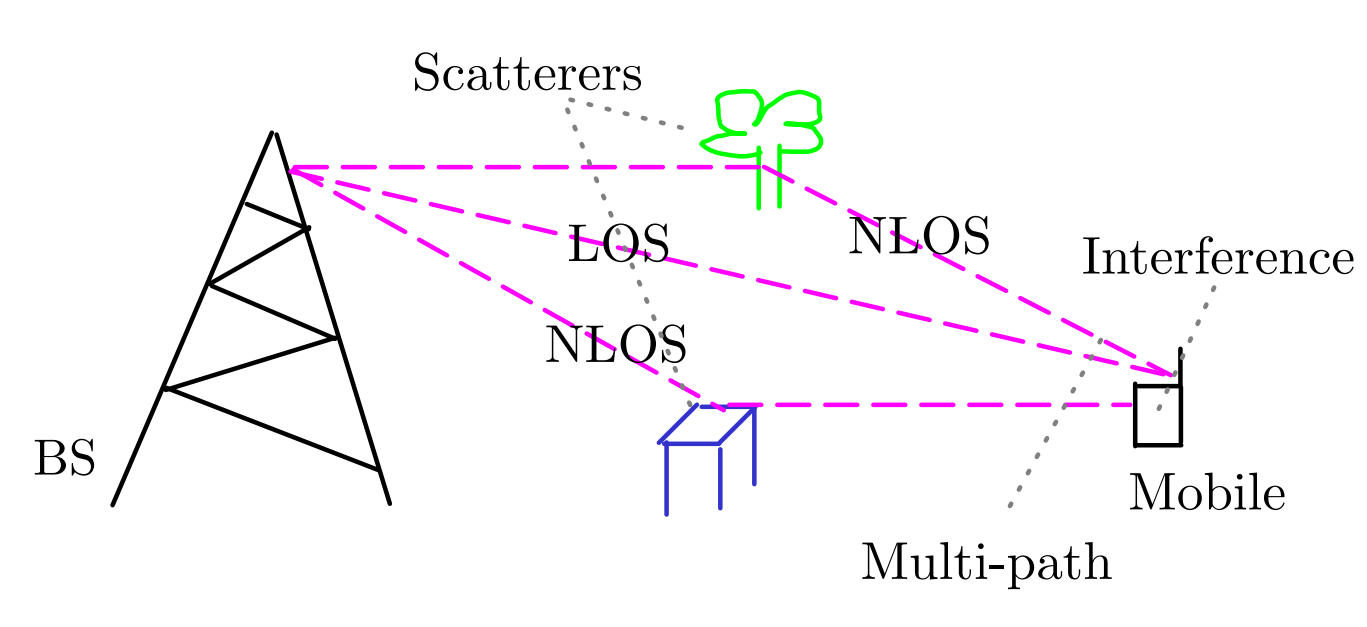
$$x \quad x+dx \quad \mathbb{R}$$

- Characterized by PDF $F_X(x)$
- Quantity $F_X(x)dx$ represents the probability that X takes value in the infinitesimal interval $[x, x+dx]$
- $F_X(x) \geq 0$ for all values of x
- $P(X \in [a, b]) = \int_a^b F_X(x)dx$
- Naturally, $\int_{-\infty}^{\infty} F_X(x)dx = 1$

Example: PDF

- Consider $F_X(x) = Ke^{-ax}$
- $\Rightarrow \int_{-\infty}^{\infty} Ke^{-ax} = 1$
 $-\frac{K}{a}e^{-ax} \Big|_0^{\infty} = 1$
 $\Rightarrow K = a$
- $F_X(x) = ae^{-ax}$, is known as the Exponential PDF

Application: Wireless Channel



- Multipath propagation leads to 'Fading' Signal
- Power of Fading channel g , is random
- $F_G(g) \triangleq e^{-g}$ where $g > 0$
- This is known as **Rayleigh** Fading channel where power g is distributed as Exponential and the amplitude of the channel coefficient h is distributed as Rayleigh

Example:

Probability that the attenuation of the channel is worse than 20dB.

$$\Rightarrow 10 \log_{10} g \leq -20 \text{ dB}$$

$$\Rightarrow P(g \in [0, 0.01]) = \int_0^{0.01} e^{-g} dg$$

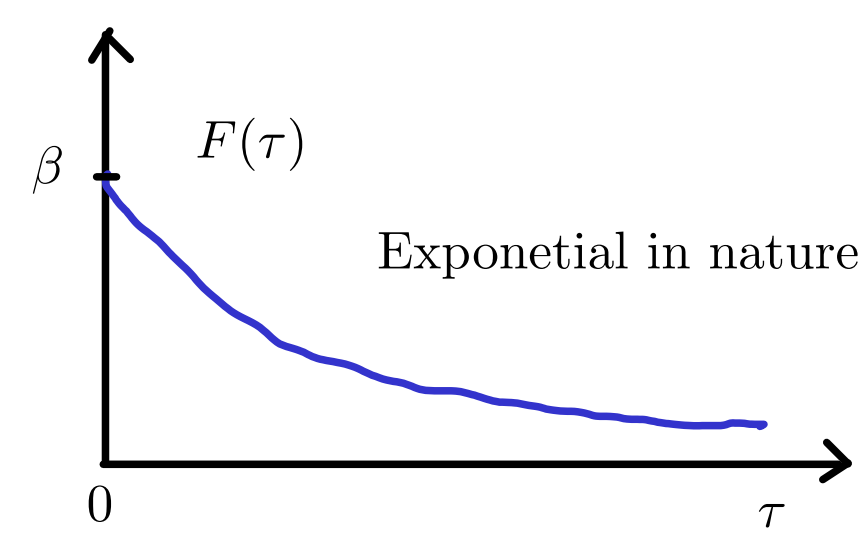
$$= 1 - e^{-0.01}, \text{ where } e^{-0.01} \approx (1 - 0.01)$$

$$\Rightarrow P(g \in [0, 0.01]) = 0.01$$

Statistical Parameters (Mean μ and Variance σ^2)

- Mean $\mu \triangleq$ Average value of X
- $\mu = E\{X\}$, the expected value of X
- $\mu = \int_{-\infty}^{\infty} x F_X(x)dx$
i.e., integration of the values that X takes weighed with the corresponding probability
- Variance $\sigma_X^2 \triangleq$ Average square deviation of X about mean μ
- $\sigma_X^2 = E\{(X - \mu)^2\}$
 $= \int_{-\infty}^{\infty} (x - \mu)^2 F_X(x)dx$
 $= \int_{-\infty}^{\infty} x^2 F_X(x)dx - \mu^2$
 $= E\{X^2\} - E\{X\}^2$

Application: (Power Profile of wireless channel)



$$\beta \quad F(\tau) \quad \tau$$

Exponential in nature

- At receiver, multiple signal copies with different delays
- Power arriving with different delays
- Hence **power profile** of the channel
- Used to compute **Average delay** (μ_τ) and **variance delay** (σ_τ^2)
- Typical power profile is $F(\tau) = \beta e^{-\beta\tau}$

Average Delay (μ_τ)

- $\mu_\tau = \int_0^{\infty} \tau e^{-\beta\tau} d\tau$
 $= -\tau e^{-\beta\tau} \Big|_0^{\infty} + \int_0^{\infty} e^{-\beta\tau} d\tau$
 $= -\frac{1}{\beta} e^{-\beta\tau} \Big|_0^{\infty} = \frac{1}{\beta}$

Variance Delay (σ_τ^2)

- $\sigma_\tau^2 = E\{(\tau - \mu_\tau)^2\}$
 $= E\{\tau^2\} - \mu_\tau^2$
 $= \frac{1}{\beta^2}$
- Root Mean square** delay spread
 $\sigma_\tau = \frac{1}{\beta}$

Random Variables

- RV X , takes values randomly from \mathbb{R} or a subset of \mathbb{R}

Transformation of RVs (Function of RV)

- Consider RV $Y = \Psi(X)$ where $\Psi(\cdot)$ is 1-to-1
 $\Rightarrow \Psi(\cdot)$ is invertible and $x = \Psi^{-1}(y)$
- If $X \in [x, x+dx] \Rightarrow Y \in [y, y+dy]$
 $\therefore P(X \in [x, x+dx]) = P(Y \in [y, y+dy])$
 $F_X(x)dx = F_Y(y)dy$
 $F_Y(y) = \frac{F_X(x)|_{x=\Psi^{-1}(y)}}{\left|\frac{dy}{dx}\right|_{x=\Psi^{-1}(y)}}$

Example:

- For a Rayleigh fading channel, we have $F_X(x) = e^{-x}$, where X denotes the power
- Amplitude Y is then given by $Y = \sqrt{X}$
 $y = \Psi(x) = \sqrt{x}$, where $x \geq 0$ and $y \geq 0$
 $\therefore \Psi(\cdot)$ is 1-to-1 and an invertible fn we have, $x = \Psi^{-1}(y) = y^2$
 $\Rightarrow F_Y(y) = \frac{e^{-x}|_{x=y^2}}{\left|\frac{1}{2\sqrt{x}}\right|_{x=y^2}} = \frac{e^{-y^2}}{\frac{1}{2\sqrt{y^2}}}$
 $= 2ye^{-y^2}$
which is known as the **Rayleigh** Fading channel.

Applications:

- Noise** at an instant in time in comm. system is modeled as Gaussian RV
- Wireless **fading channel coefficient** is modeled as complex Gaussian RV which leads to Rayleigh channel (i.e., amplitude is Rayleigh)

Property of Gaussian RV:

- Linear combination of Gaussian RV results in another Gaussian RV
- Consider X_1, X_2, \dots, X_L Gaussian RVs
 $E\{X_i\} = \mu_i$ and $E\{(X_i - \mu_i)^2\} = \sigma_i^2$, for $i \in \mathbb{N}_L^1$
Covariance of X_i, X_j is $E\{(X_i - \mu_i)(X_j - \mu_j)\} = \sigma_{ij}$
- Let linear combination of Gaussian RVs be
 $X \triangleq a_1X_1 + a_2X_2 + \dots + a_LX_L$
 $\Rightarrow X = [a_1 \ a_2 \ \dots \ a_L] \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_L \end{bmatrix} = \mathbf{a}^T \mathbf{X}$
 X is another Gaussian RV

Mean and Variance of X

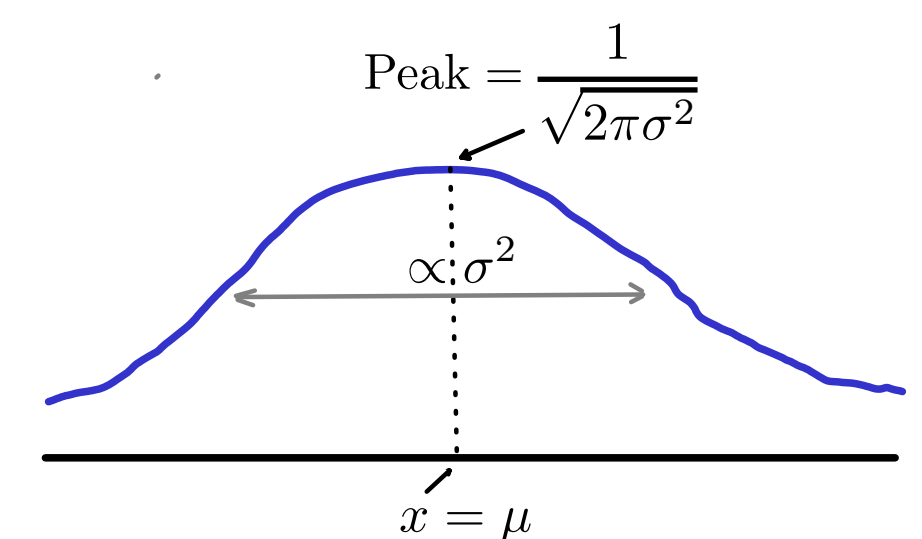
- $\mu = E\{X\} = E\{\sum_i a_i X_i\}$
 $= \sum_i a_i E\{X_i\} = \sum_i a_i \mu_i = \mathbf{a}^T \boldsymbol{\mu}$
- $\sigma^2 = E\{(X - \mu)^2\}$
 $= E\{(\sum_i a_i X_i - \mu)^2\} = E\{(\sum_i a_i X_i - \sum_i a_i \mu_i)^2\}$
 $= E\{(\sum_i a_i (X_i - \mu_i))^2\}$
 $= E\{\sum_i \sum_j a_i a_j (X_i - \mu_i)(X_j - \mu_j)\}$
 $= \sum_i a_i^2 \sigma_i^2 + \sum_i \sum_{j \neq i} a_i a_j \sigma_{ij}$
- $X \sim \mathcal{N}(\sum_i a_i \mu_i, \sum_i a_i^2 \sigma_i^2 + \sum_i \sum_{j \neq i} a_i a_j \sigma_{ij})$

Special Case:

Consider $X_i, \forall i$ are **zero-mean, equal variance and uncorrelated**

- Zero Mean : $E\{X_i\} = \mu_i = 0$
- Same Variance: $E\{(X_i - \mu_i)^2\} = \sigma_i^2 = \sigma^2$
- Uncorrelated: $E\{(X_i - \mu_i)(X_j - \mu_j)\} = \sigma_{ij} = 0$
- Note that **uncorrelated** Gaussian RVs are **independent** as well
- \Rightarrow Considered group of X_i 's are **Independent and Identically Distributed (IID)** RVs
- $X = \mathbf{a}^T \mathbf{X}$ is Gaussian
- $E\{X\} = \mathbf{a}^T \boldsymbol{\mu} = 0$
- $E\{(X - \mu)^2\} = E\{X^2\} = \sum_i a_i^2 \sigma^2 = \sigma^2 \|\mathbf{a}\|^2$
- $X \sim \mathcal{N}(0, \sigma^2 \|\mathbf{a}\|^2)$

Gaussian Random Variable X , (Normal RV)



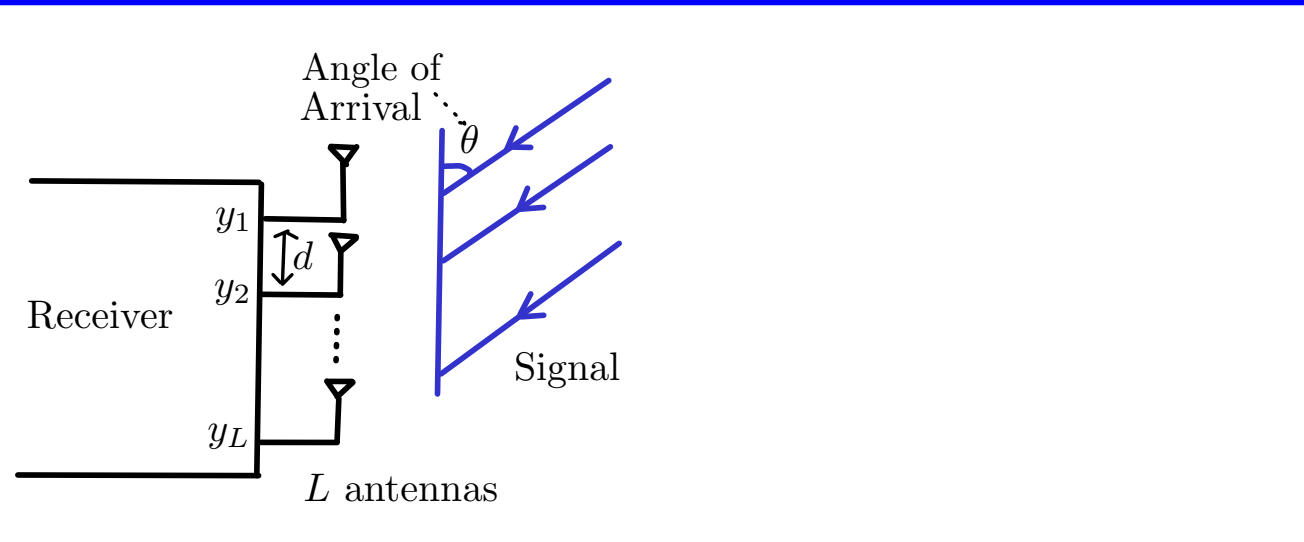
$$\text{Peak} = \frac{1}{\sqrt{2\pi\sigma^2}}$$

$$\sigma, \sigma^2$$

$$x = \mu$$

- $F_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$ where,
 $E\{X\} = \mu$ is the Mean
 $E\{(X - \mu)^2\} = \sigma^2$ is the Variance
- Notation $X \sim \mathcal{N}(\mu, \sigma^2)$

Application: Uniform Linear Antenna Array



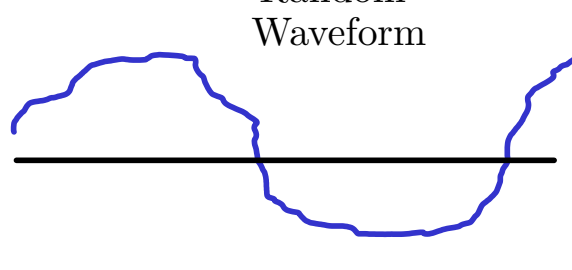
- $y_i = x e^{-j(i-1)\phi} + w_i, \forall i$, where x is the transmitted signal, y_i is the received signal at i^{th} receiver antenna and w_i is the Gaussian noise.
 $\Rightarrow \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_L \end{bmatrix} = \begin{bmatrix} 1 \\ e^{-j\phi} \\ \vdots \\ e^{-j(L-1)\phi} \end{bmatrix} + \begin{bmatrix} w_1 \\ w_2 \\ \vdots \\ w_L \end{bmatrix}, \mathbf{y} = \mathbf{h}(\phi)\mathbf{x} + \mathbf{w}$
where $\mathbf{h}(\phi)$ is known as Array Steering Vector.
- Received signal at each successive antenna is delayed by an additional phase of ϕ , hence named Linear Phased array where $\phi = 2\pi f_c \frac{d \cos(\theta)}{c} = \frac{2\pi d}{\lambda} \cos(\theta)$

Application: (continuation ...)

- Assuming the Gaussian noise samples to be zero-mean IID, i.e., $E\{w_i\} = 0, E\{|w_i|^2\} = \sigma^2$ and $E\{w_i w_j^*\} = 0$ (uncorrelated)
- Considering beamforming vector $\mathbf{a}^T = [a_1 \ a_2 \ \dots \ a_L]$
 $\tilde{y} = a_1^* y_1 + a_2^* y_2 + \dots + a_L^* y_L = [a_1^* \ a_2^* \ \dots \ a_L^*] \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_L \end{bmatrix} = \mathbf{a}^H \mathbf{y}$
This combining process is called **beam forming**.
- We have, $\tilde{y} = \mathbf{a}^H \mathbf{y} = \mathbf{a}^H \mathbf{h}(\phi)x + \mathbf{a}^H \mathbf{w}$ and considering transmitted signal power $E\{|x|^2\} = P$
We have, $\text{SNR} = \frac{\text{Signal Power}}{\text{Noise Power}} = \frac{|\mathbf{a}^H \mathbf{h}(\phi)|^2 E\{|x|^2\}}{E\{|\mathbf{a}^H \mathbf{w}|^2\}} = \frac{|\mathbf{a}^H \mathbf{h}(\phi)|^2 P}{\sigma^2 \|\mathbf{a}\|^2}$
- To find the maximum SNR possible, we use the Cauchy-schwartz Inequality $|\mathbf{a}^H \mathbf{h}(\phi)|^2 \leq \|\mathbf{a}\|^2 \|\mathbf{h}(\phi)\|^2$,
 $\Rightarrow \text{SNR} \leq \frac{\|\mathbf{a}\|^2 \|\mathbf{h}(\phi)\|^2 P}{\sigma^2 \|\mathbf{a}\|^2} = \frac{\|\mathbf{h}(\phi)\|^2 P}{\sigma^2} = \text{SNR}_{\max}$
Note $\|\mathbf{h}(\phi)\|^2 = 1 + |e^{-j\phi}|^2 + \dots + |e^{-j(L-1)\phi}|^2 = 1 + 1 + \dots + 1 = L$
- $\therefore \text{SNR}_{\max} = \frac{LP}{\sigma^2}$, i.e., L times the initial SNR. (i.e., Array gain = L)
- To obtain $\text{SNR} = \text{SNR}_{\max}$, **spatially matched filter (matched beam former)** is used, i.e., $\mathbf{a} = \mathbf{h}(\phi)$ (also called as **Maximum Ratio Combining** and coherent combining)

Random Processes

Characteristics and Statistical Parameters



- Can be used to model random signal or random noise process
- Is characterized by $F_X(x, t)$, a PDF at each instance of time
- Statistical parameters are also function of time
Mean : $\mu_X(t) = E\{X(t)\} = \int_{-\infty}^{\infty} x F_X(x, t) dx$
- Auto-Correlation (or self-correlation) at time instances t_1 and t_2 , is
 $R_{XX}(t_1, t_2) \triangleq E\{X(t_1) X(t_2)\}$
 $= \int_{-\infty}^{\infty} x_1 x_2 F(x_1, x_2, t_1, t_2) dx_1 dx_2$
where $F(x_1, x_2, t_1, t_2)$ is the joint PDF for t_1 and t_2

Random Processes

- A **random process** $X(t)$, is a random variable X which is a function of time t
- Is a RV at every time instant or is a RV **indexed** by time

Special Class of Random Process

Special Random Process: (Wide Sense Stationary, WSS)

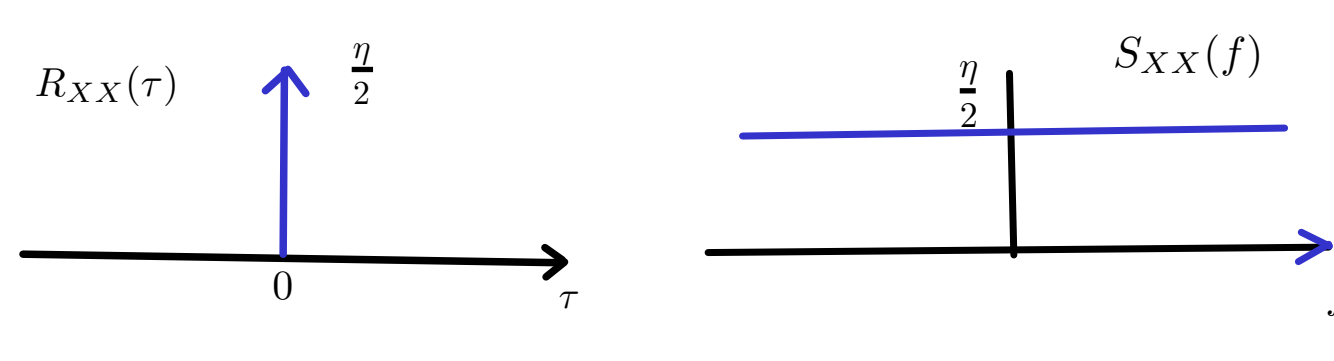
Satisfies two properties:

- $E\{X(t)\} = \mu_X = \text{Constant}$
i.e., mean does **NOT depend** on time (is stationary)
- Let $t_1 = t$ and $t_2 = t + \tau$
 $E\{X(t)X(t + \tau)\} = R_{XX}(t, t + \tau) = R_{XX}(\tau)$
i.e., Auto-correlation **depends only** on the shift τ and NOT on the specific time instance t .
 - Note, when $\tau = 0$
 $E\{X(t)X(t)\} = E\{X(t)^2\} = R_{XX}(0)$
i.e., $R_{XX}(0) = \text{Avg. power in the process } X(t)$

Gaussian Random Process

- Random Process $X(t)$ is a **Gaussian** random process, if $X(t_1), X(t_2), \dots, X(t_n)$ are **jointly Gaussian** at any time instance i.e., Multivariate Density $F_{X(t_1)X(t_2)\dots X(t_n)}(x_1, x_2, \dots, x_n)$ is Gaussian $\forall n$ and $\forall t_1, t_2, \dots, t_n$
- Considering $n = 1$, the PDF of $X(t)$ at any time instance must be Gaussian
 $\Rightarrow F_{X(t)}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$ where μ and σ^2 are function of time
- Considering a **WSS Gaussian Process**, we have
 $E\{X(t)\} = \mu_X$ and $E\{X(t)X(t + \tau)\} = R_{XX}(\tau)$ are stationary. Further with $\tau = 0$ we have $E\{X(t)^2\} = R_{XX}(0)$ and $\sigma^2 = R_{XX}(0) - \mu_X^2$
 $\Rightarrow F_{X(t)}(x) = \frac{1}{\sqrt{2\pi(R_{XX}(0) - \mu_X^2)}} e^{-\frac{(x-\mu)^2}{2(R_{XX}(0) - \mu_X^2)}}$

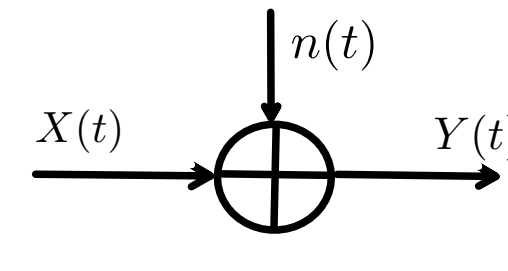
White Noise (White Random Process)



- A Random Process is **white** if it is
 - WSS**
 - and its **autocorrelation** $R_{XX}(\tau) = \frac{\eta}{2}\delta(\tau)$, i.e., $R_{XX}(\tau) = 0$ if $\tau \neq 0$.
For any non-zero shift τ , the correlation $E\{X(t)X(t + \tau)\} = 0$, i.e., $X(t)$ and $X(t + \tau)$ are **un-correlated**.
- PSD** of White Noise is obtained by Fourier Transform of $R_{XX}(\tau)$, i.e., $S_{XX}(f) = \frac{\eta}{2}$ for $-\infty < f < \infty$ i.e., **constant over entire freq. range**. Uniform power distribution over the entire freq. range similar to white light

Example: Additive White Gaussian Noise

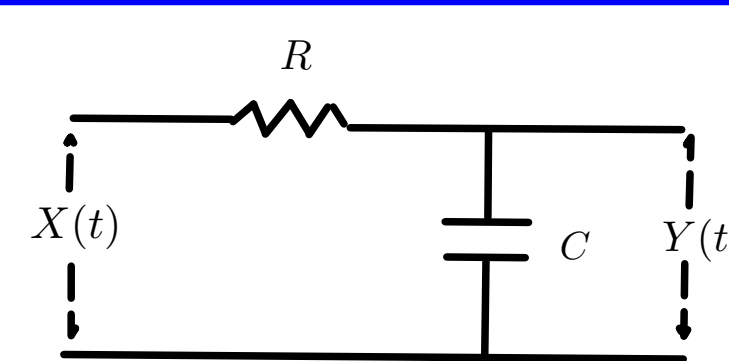
- Random process $n(t)$, is both **Gaussian** and **White**, i.e., $S_{nn}(f) = \frac{\eta}{2}$
- Ex: Additive White Gaussian Noise (**AWGN**) adds to the transmitted signal



$Y(t) = X(t) + n(t)$

Such a comm. channel that adds WGN to transmit signal is termed as **AWGN channel** and is a typical model for **wireline** channel. (Note that there is no fading coefficient h as in wireless channel)

White Gaussian Noise when passed through RC (low-pass) filter



- Consider $X(t)$ is zero-mean White Gaussian noise, i.e., $\mu_X = 0$, $R_{XX}(\tau) = \frac{\eta}{2}\delta(\tau)$ and $S_{XX}(f) = \frac{\eta}{2}$ for entire freq. range
- Characteristics of RP $Y(t)$ at the output of RC filter (LTI system)
 - $Y(t)$ is also zero-mean **Gaussian** as the filter is LTI system
 - We have, $H(f) = \frac{1}{1 + j2\pi fRC}$
 $\Rightarrow S_{YY}(f) = \frac{\eta}{2} \frac{1}{(1 + 4\pi^2 f^2 R^2 C^2)}$ (**no longer White**)
 - Power $E\{|Y(t)|^2\} = \frac{\eta}{2} \int_{-\infty}^{\infty} \frac{1}{(1 + 4\pi^2 f^2 R^2 C^2)} df = \frac{\eta}{4RC} = R_{YY}(0) = \sigma_Y^2$
 - \Rightarrow PDF $F_{Y(t)}(y) = \sqrt{\frac{2RC}{\pi\eta}} e^{-\frac{2RCy^2}{\eta}}$

Example: In Wireless Comm. System

- Consider a signal $X(t) = \alpha \cos(2\pi f_c t + \theta)$, where Phase θ , is **RANDOM** and uniformly distributed in $[-\pi, \pi]$ i.e., $F_{\Theta}(\theta) = \frac{1}{2\pi}$, for $-\pi \leq \theta < \pi$ and the parameters α, f_c are constant with time.
- To examine if $X(t)$ is WSS
 - $E\{X(t)\} = E\{\alpha \cos(2\pi f_c t + \theta)\} = \int_{-\pi}^{\pi} \alpha \cos(2\pi f_c t + \theta) F_{\Theta}(\theta) d\theta = \frac{\alpha}{2\pi} \int_{-\pi}^{\pi} \cos(2\pi f_c t + \theta) d\theta = 0$
 $\mu_X(t) = 0 = \mu_x$ is a constant and is **stationary in the mean**.
 - $E\{X(t)X(t + \tau)\} = E\{\alpha \cos(2\pi f_c t + \theta) \times \alpha \cos(2\pi f_c(t + \tau) + \theta)\} = \frac{\alpha^2}{2\pi} \int_{-\pi}^{\pi} \cos(2\pi f_c t + \theta) \times \alpha \cos(2\pi f_c(t + \tau) + \theta) d\theta = \frac{\alpha^2}{2} \cos(2\pi f_c \tau) = R_{XX}(\tau)$, does NOT depend on t
 \therefore **stationary in the autocorrelation**

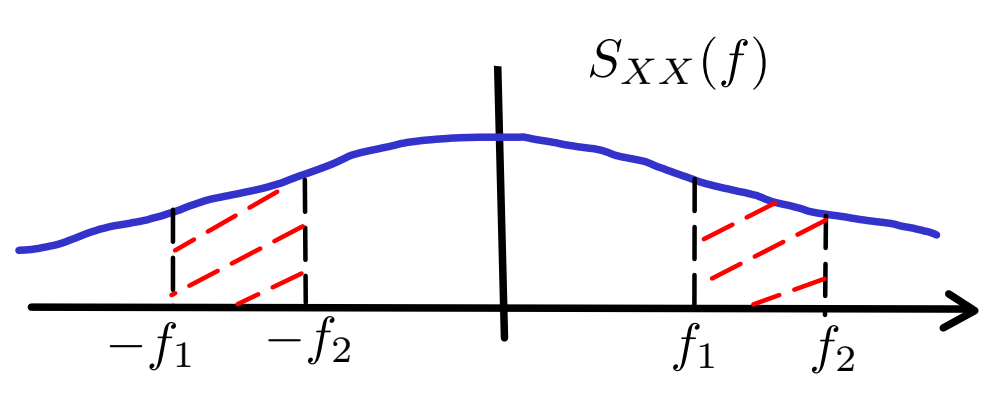
\Rightarrow Random process $X(t)$ is **WSS** with avg. power $R_{XX}(0) = \frac{\alpha^2}{2}$

Power Spectral Density (PSD)

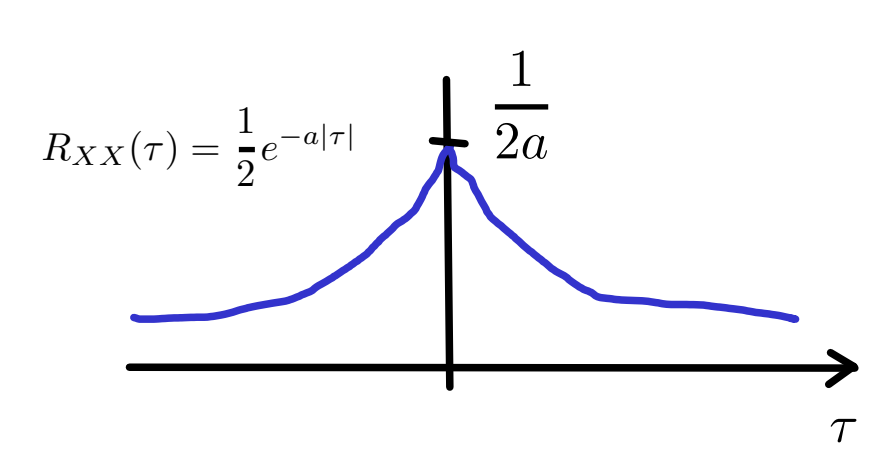
- PSD $S_{XX}(f)$, of a WSS random process \triangleq Fourier Transform of $R_{XX}(\tau)$
 $\Rightarrow S_{XX}(f) \triangleq \int_{-\infty}^{\infty} R_{XX}(\tau) e^{-j2\pi f \tau} d\tau$
 $\Rightarrow R_{XX}(\tau) = \int_{-\infty}^{\infty} S_{XX}(f) e^{j2\pi f \tau} df$ (Inverse Fourier Transform)

Properties of PSD

- $S_{XX}(f)$ is a real quantity and $S_{XX}(f) \geq 0$, as it related to power
- $S_{XX}(f) = S_{XX}(-f)$ is symmetric
- Setting $\tau = 0$, $R_{XX}(0) = \int_{-\infty}^{\infty} S_{XX}(f) df$
 \Rightarrow Power of WSS RP $P_X = \text{Integral of PSD over entire freq.}$
- Power in WSS RP in the band $[f_1, f_2]$ is
 $P_X[f_1, f_2] = \int_{-f_1}^{-f_2} S_{XX}(f) df + \int_{f_1}^{f_2} S_{XX}(f) df = 2 \int_{f_1}^{f_2} S_{XX}(f) df$

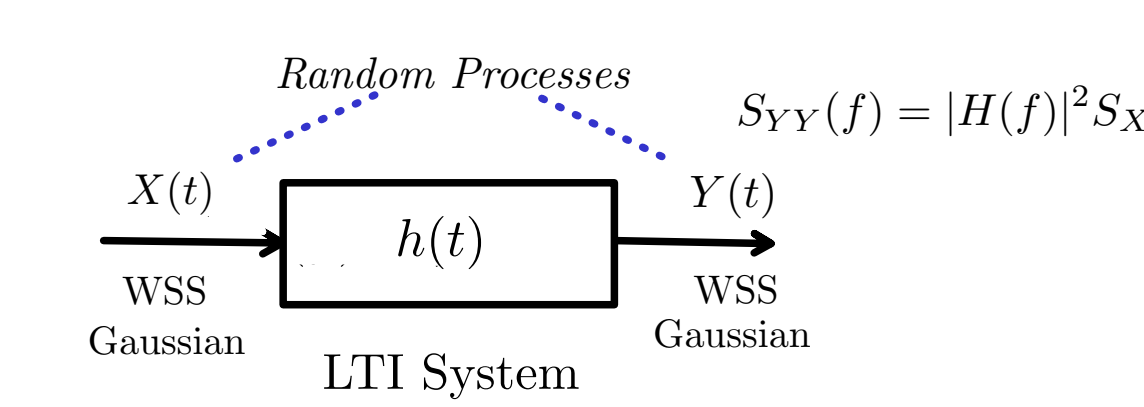


Example: Wireless context



- Consider Wireless signal $X(t)$ with $R_{XX}(\tau) = \frac{1}{2} e^{-a|\tau|}$, and $a = 5 \text{ KHz}$
- Then, power in $X(t)$ is $E\{X^2(t)\} = R_{XX}(0) = \frac{1}{2} a$
- Power spectral density, PSD is
 $S_{XX}(f) = \int_{-\infty}^{\infty} \frac{1}{2} e^{-a|\tau|} e^{-j2\pi f \tau} d\tau = \frac{1}{2a} \left\{ \int_0^{\infty} e^{-a\tau} e^{-j2\pi f \tau} d\tau + \int_{-\infty}^0 \frac{1}{2a} e^{a\tau} e^{-j2\pi f \tau} d\tau \right\} = \frac{1}{a^2 + 4\pi^2 f^2}$
- Band $[-w, w]$ with 90% of the energy can be obtained as
 $\frac{0.9}{2a} = \int_{-w}^w \frac{1}{a^2 + 4\pi^2 f^2} df = \frac{1}{4\pi^2} \int_{-w}^w \frac{1}{f^2 + \frac{a^2}{4\pi^2}} df$
Further solving we get, $w = 1.005a$
 \Rightarrow Required bandwidth = $[-1.005a, 1.005a]$, i.e., $[-5 \text{ KHz}, 5 \text{ KHz}]$ and passband BW is 10 KHz.

Transmission of Random Process through LTI



- Consider $X(t)$ is input to LTI system with impulse response $h(t)$, then $Y(t) = X(t) * h(t) = \int_{-\infty}^{\infty} X(t - \alpha) h(\alpha) d\alpha$, where $Y(t)$ is the output
- $Y(t)$ is a random process
- We have Mean of $Y(t)$ as
 $\mu_Y(t) = E\{Y(t)\} = E\{\int_{-\infty}^{\infty} X(t - \alpha) h(\alpha) d\alpha\} = \int_{-\infty}^{\infty} E\{X(t - \alpha)\} h(\alpha) d\alpha = \mu_X \int_{-\infty}^{\infty} h(\alpha) d\alpha = \mu_Y$
 $\therefore Y(t)$ is **stationary** in the mean
- We have Auto-correlation of $Y(t)$
 $R_{YY}(t) = E\{Y(t)Y(t + \tau)\} = E\{\int_{-\infty}^{\infty} X(t - \alpha) h(\alpha) d\alpha \times \int_{-\infty}^{\infty} X(t + \tau - \beta) h(\beta) d\beta\} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} E\{X(t - \alpha)X(t + \tau - \beta)\} h(\alpha) h(\beta) d\alpha d\beta = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_{XX}(\tau - \beta + \alpha) h(\alpha) h(\beta) d\alpha d\beta$
 \Rightarrow depends only on τ
 $\Rightarrow Y(t)$ is **stationary** in the auto-correlation
 $\Rightarrow Y(t)$ is also **WSS**
- Further $R_{YY}(\tau) = R_{XX}(\tau) * h(\tau) * h(-\tau)$ (in τ domain)
- And $S_{YY}(f) = S_{XX}(f) |H(f)|^2$, where $H(f) \leftrightarrow h(t)$

References:

[1] NPTEL Course: Probability and Random Variables for Wireless Communications by Prof. Aditya Jaganathan