## Lord's paradox

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## Causality

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One way in which researchers have tried to study this is through a **pre-post test design**, in which the (potential) cause x is measured once, and the outcome is measured twice ( $y_1$  and  $y_2$ ).

#### Two broad classes of models

#### 1: Change score method

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#### 2: Regressor variable method:

$$y_2 = \gamma_0 + \gamma_1 x_1 + \gamma_2 y_1 + v$$

also known as: pretest-postest covariance or covariance adjusted score (when  $x_1$  is dichotomous), or as cross-lagged panel analysis.

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#### An alternative expression of the second model is:

#### 2: Baseline-adjusted gain scores

$$y_2 - y_1 = \gamma_0 + \gamma_1 x_1 + (\gamma_2 - 1)y_1 + v$$

also known as: residualized gain scores or residual change.

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Corrective action	$\beta$ for W2 to W3 longitudinal net effects $\!\!^a$	r between W2 & W2 to W3 gains			
	Antisocial behavior				
Professional intervention	ıs				
Psychotherapy visits	.07**	.00			
Ritalin	.07**	.04			
Parental disciplinary action	ons				
Non-physical punishm	ent .03	08**			
Physical punishment	.07**	05			
Scolding/yelling	.06*	−.08**			
"Hostile/ineffective" so	ale .09**	15**			

#### This shows that:

- regressor variable method (first column): adverse effect (or no effect)
- change score method (second column): beneficial effect (or no effect)

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#### So what is the truth?

# Bad reputation of change scores

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Allison (1990) indicates that psychometricians have claimed that the **change score method** is **problematic** because of:

1. **Unreliability**:  $y_2 - y_1$  tends to be (much) less reliable than  $y_1$  and  $y_2$ 

## Bad reputation of change scores

Allison (1990) indicates that psychometricians have claimed that the **change score method** is **problematic** because of:

1. **Unreliability**:  $y_2 - y_1$  tends to be (much) less reliable than  $y_1$  and  $y_2$ 

#### 2. Regression towards the mean:

- y<sub>2</sub> y<sub>1</sub> is typically negatively correlated with y<sub>1</sub> (people high on y<sub>1</sub> will decrease and those low on y<sub>1</sub> will increase)
- if  $x_1$  is correlated with  $y_1$ , it will have a spurious relationship with  $y_2 y_1$

## However, here is Lord's paradox

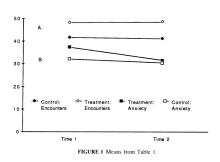
#### Allison (1990) gives this example of a quasi-experiment:

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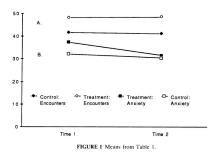


**Example A (encounters):** Regressor variable method **erroneously detects a difference in change** between the groups (suggesting treatment had a detrimental effect)

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**Example A (encounters):** Regressor variable method **erroneously detects a difference in change** between the groups (suggesting treatment had a detrimental effect)

**Example B (anxiety):** Regressor variable method **fails to detect a difference in change** between the groups (suggesting that treatment does not decrease anxiety)

## Repeated measures models

Both models can also be expressed as **repeated measures models** (in which case the Regressor variable method is a special case of Change score method!).

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- main issue is: did separate groups exist at the pre-measurement

#### Advice for four scenarios by van Breukelen:

- random assignment: use Regressor variable method (more power)
- assignment (entirely!) dependent on pretest score: use Regressor variable method (Change score is biased)
- assignment based on preexisting/natural groups: do not use Regressor variable method; Change score method might be right (requires the assumption that both groups change by the same amount when there is no treatment)
- self-assignment: unclear

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Note that this all (seems to) generalize to the case where x is a **continuous variable**, measured **simultaneously** with  $y_1$ .

#### Model = truth?

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There is a large body of literature based on the idea that there is **unobserved heterogeneity** (i.e., stable between-person, trait-like differences), like:

$$y_{it} = \beta_0 + \alpha_i + \beta_1 x_{it} + \varepsilon_{it}$$

where  $\alpha_i$  captures unobserved omitted variables that are invariant over time.

## Lord's paradox

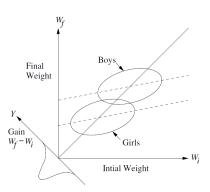
Question: What is the effect of the diet provided by university dinning halls on students' weight, and are there sex differences in these effects?

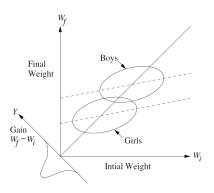
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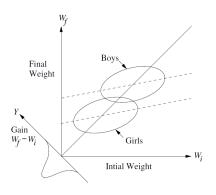
**Basics**: This is a **pre-post test design** with **two existing groups** (boys and girls).

Hence, the "treatment" is not the diet (as this is the same for everyone), but gender: Do gender differences in metabolism have a different effect on the weight of boys than on the weight of girls?

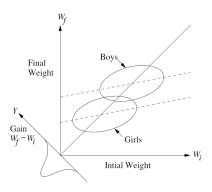




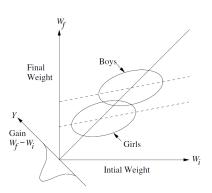
Mean of girls has not changed; mean of boys has not changed

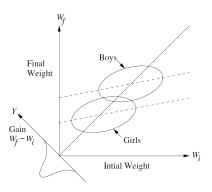


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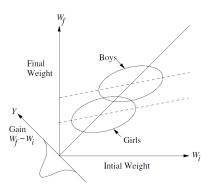


- Mean of girls has not changed; mean of boys has not changed
- Frequency distributions within groups has not changed
- Conclusion: while there are individual changes, overall there are no changes for either boys or girls



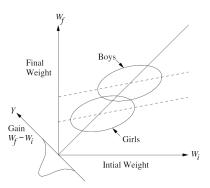


ANCOVA with initial weight as covariate and gender as the factor



- ANCOVA with initial weight as covariate and gender as the factor
- Conclusion: the weight gain for boys is larger than that for girls, when proper allowance for initial weight is made (see the difference in intercepts)

#### **Paradox**



When the question is: Is there differential gain?

- there are no changes in mean for either group; hence NO differential gain
- when boys and girls start with the same weight, the boys will gain more than the girls; so there is differential gain

You can think of this as:

- weight gain  $(\Delta W_i = W_{f,j} W_{i,j})$  is the outcome
- gender is the predictor (cause!)
- initial weight  $(W_{i,j})$  is the mediator

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Statistician 1 looks at the **total effect of gender** (with dummy variable  $M_j$  for males) on weight gain:

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Statistician 2 looks at the **direct effect of gender** on weight gain:

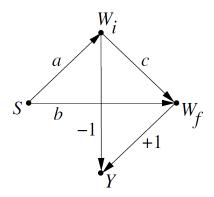
$$\Delta W_i = b_0 + b_1 M_i + b_2 W_{i,j} + e_j$$

which can be expressed as the ANCOVA model:

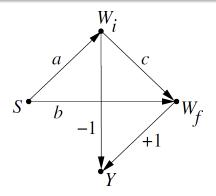
$$W_{f,j} = b_0 + b_1 M_j + (b_2 + 1) W_{i,j} + e_j$$

## And now with a DAG

- Cause is sex (S)
- Outcome is weight gain  $(Y = W_f W_i)$
- Mediator is initial weight (W<sub>i</sub>)



## The two answers based on the DAG



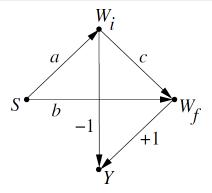
**Total effect**: multiply all coefficients of a path from S to Y, and sum these

$$TE = b * 1 + a * c * 1 + a * (-1) = b - a(1 - c)$$

**Direct effect**: consider only paths that do not contain the mediator

$$DE = b * 1$$

#### In words

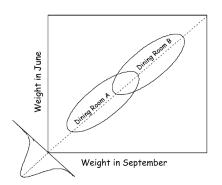


No total effect: b - a(1 - c) = 0Positive direct effect: b > 0

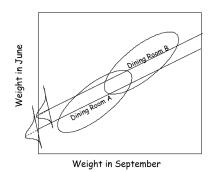
On average a boy gains more than a girl **of equal initial weight** (b>0), but since there are more heavy-weight boys than girls and we subtract a portion of this difference, overall the gain for boys is the same as the gain for girls.

#### Conclusion: There is no paradox!

# Different diets (instead of sex)



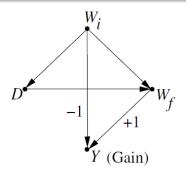
Group means are again on the 45-degree line: no mean changes over time in either group.



ANCOVA results in different intercepts for the two groups: More weight gain in Dining Room B.

Critical here is that **heavier students tended to choose dining room B** more often.

# Pearl: Now it is confounding, not mediation



- Initial weight is **no longer the mediator**; it is now the first variable in the causal sequence.
- It is a common cause or confounder of the relationships between (potential) cause (dinning room) and outcome (final weight or weight gain).
- We need to control for this; failing to do so biases the results

## The role of the pre-test score in the DAG

So the critical distinction is: Is the pre-test score a mediator (affected by the potential cause of interest), or a confounder (affecting the potential cause of interest)?

#### Draw the **DAGs for these scenarios**:

- Larzelere: Pre- and post-test measures of deviant behavior; potential cause is parental discipline
- Allison: Pre- and post-test of number of social encounters; groups are children with facial abnormalities and controls; first group is treated between pre-test and post-test