

DeMorgan’s Laws: <div> <ul style="list-style-type: none"> $\overline{A \cap B} = \overline{A} \cup \overline{B}$ $\overline{A \cup B} = \overline{A} \cap \overline{B}$ </div>
Countable if elements can be put in one-to-one with $\mathbb{N} = \{1, 2, 3, \dots\}$
Uncountable sets: <div> <ul style="list-style-type: none"> $\mathbb{R} = (-\infty, \infty)$ $[0, 1]$ and $(0, 1)$ $(a, b), \forall a, b \in \mathbb{R}$ such that $a < b$ </div>

A family of sets $\{\mathcal{A}_i, i \in I\}$ is a **partition of \mathcal{S}** if it is disjoint and collectively exhaustive over \mathcal{S} .

Axioms of Probability

- $\mathcal{P}(A) \geq 0, \forall A \in \mathcal{F}(\mathcal{S})$
- $\mathcal{P}(\mathcal{S}) = 1$
- $\mathcal{P}(\cup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} \mathcal{P}(A_i)$

Probability Equations

- $\mathcal{P}(A \cup B) = \mathcal{P}(A) + \mathcal{P}(B) - \mathcal{P}(A \cap B)$
- $\mathcal{P}(A|B) = \frac{\mathcal{P}(A \cap B)}{\mathcal{P}(B)} = \frac{\mathcal{P}(B|A)\mathcal{P}(A)}{\mathcal{P}(B)}$

Given \mathcal{S} and a family of subsets $\mathcal{G} = \{\mathcal{A}_i, i \in I\}$ of \mathcal{S} , the **σ -field generated by \mathcal{G}** , denoted $\sigma(\mathcal{G})$, is the smallest σ field containing all the subsets in \mathcal{G} .
By “smallest” σ -field, we mean that for any σ -field \mathcal{F}_0 containing all the sets in \mathcal{G} :

$$\sigma(\mathcal{G}) \subset \mathcal{F}_0$$

Given \mathbb{R} , the **Borel field of \mathbb{R}** is defined as the σ -field generated by the family of all open intervals

$$\mathcal{G} = \{(a, b) : \forall (a, b) \in \mathbb{R} \text{ such that } a < b\}.$$

We denote the Borel field of \mathbb{R} by $\mathcal{B}(\mathbb{R})$.

pmfs uniform $p(\omega) = \frac{1}{n}, \forall \omega \in \mathcal{S}$ binomial $p(k) = \binom{n}{k} a^k (1-a)^{n-k}$ geometric $p(k) = (1-a)a^k, a \in (0, 1)$ poisson $p(k) = \frac{\lambda^k e^{-k}}{k!}; k = 0, 1, 2, \dots; \lambda > 0$
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Properties of the pdf:

- $f(r) \geq 0, \forall r \in \mathbb{R}$
- $\int_{-\infty}^{\infty} f(r)dr = 1$

Given a valid pdf, we can get a valid probability measure $\mathcal{P}(\cdot)$:

$$\mathcal{P}(A) = \int_A f(r)dr = \int_{-\infty}^{\infty} \cdot 1_A(r)dr$$

where

$$1_A(r) = \begin{cases} 1, r \in A \\ 0, r \notin A \end{cases}$$

is called the **indicator function** of the set A.

pdfs uniform $f(r) = \frac{1}{b-a} 1_{[a,b]}(r)$ exponential $f(r) = \lambda e^{-\lambda r} \cdot 1_{[0,\infty]}(r), \lambda > 0$ $\hookrightarrow = \begin{cases} \lambda e^{-\lambda r}, r \geq 0 \\ 0, r < 0 \end{cases}$ gaussian $f(r) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(\frac{-(r-\mu)^2}{2\sigma^2}\right), r \in \mathbb{R}, \mu \in \mathbb{R}, \sigma > 0$

Conditional Probability

$$\mathcal{P}(A|B) = \frac{\mathcal{P}(A \cap B)}{\mathcal{P}(B)}$$

Bayes Formula

$$\mathcal{P}(A|B) = \frac{\mathcal{P}(B|A)\mathcal{P}(A)}{\mathcal{P}(B)}$$

Total Probability Law

$$\mathcal{P}(B) = \sum_{i=1}^n \mathcal{P}(B|A_i)\mathcal{P}(A_i)$$

Bayes Theorem

$$\mathcal{P}(A_m|B) = \frac{\mathcal{P}(B|A_m)\mathcal{P}(A_m)}{\sum_{i=1}^n \mathcal{P}(B|A_i)\mathcal{P}(A_i)}$$

Events A and B are **statistically independent** if and only if

$$\mathcal{P}(A \cap B) = \mathcal{P}(A)\mathcal{P}(B)$$

Bernoulli Trials

$$p_n(k) = \mathcal{P}(\mathcal{B}_k) = \binom{n}{k} p^k (1-p)^{n-k}$$

where $p = \mathcal{P}_0(A)$.

Properties of the cdf

$$F_{\mathbb{X}}(\alpha) = \mathcal{P}_x((-\infty, \alpha)) = \mathcal{P}(\{\mathbb{X} \leq \alpha\})$$

where

$$\{\mathbb{X} \leq \alpha\} = \{\omega \in \mathcal{S} : \mathbb{X}(\omega) \leq \alpha\} \in \mathcal{F}$$

Properties:

- $F_{\mathbb{X}}(\infty) = 1$ and $F_{\mathbb{X}}(-\infty) = 0$
- If $x_1 < x_2$, then $F_{\mathbb{X}}(x_1) \leq F_{\mathbb{X}}(x_2)$
- $\mathcal{P}(\{\mathbb{X} > \alpha\}) = 1 - F_{\mathbb{X}}(\alpha)$
- If $x_1 < x_2$, then $\mathcal{P}(\{x_1 < \mathbb{X} \leq x_2\}) = F_{\mathbb{X}}(x_2) - F_{\mathbb{X}}(x_1)$
- $\mathcal{P}(\{\mathbb{X} = x_0\}) = F_{\mathbb{X}}(x_o) - F_{\mathbb{X}}(x_0^-)$ where $F_{\mathbb{X}}(x_0^-) = \lim_{\epsilon \downarrow 0} F_{\mathbb{X}}(x_0 - \epsilon)$

Properties of the pdf

$$f_{\mathbb{X}}(x) = \frac{dF_{\mathbb{X}}(x)}{dx}$$

Properties:

- $f_{\mathbb{X}}(x) \geq 0, \forall x \in \mathbb{R}$
- $F_{\mathbb{X}}(x) = \int_{-\infty}^x f_{\mathbb{X}}(\alpha)d\alpha$
- $\int_{-\infty}^{\infty} f_{\mathbb{X}}(x)dx = F_{\mathbb{X}}(\infty) - F_{\mathbb{X}}(-\infty) = 1$
- $\mathcal{P}(\{x_1 < \mathbb{X} \leq x_2\}) = \int_{x_1}^{x_2} f_{\mathbb{X}}(x)dx = F_{\mathbb{X}}(x_2) - F_{\mathbb{X}}(x_1)$

Dirac δ-functions: $\delta(x)$ <div> <ul style="list-style-type: none"> $\delta(x) = 0, \forall x \neq 0$ $\int_{-\infty}^{\infty} \delta(x)dx = \int_{-\epsilon}^{\epsilon} \delta(x)dx = 1, \forall \epsilon > 0$ </div> Sifting Property of Dirac δ-functions $\int_{-\infty}^{\infty} g(x)\delta(x-x_0)dx = g(x_0)$
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Random Variable Forms

gaussian

$$f_{\mathbb{X}}(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp[\frac{-(x-\mu)^2}{2\sigma^2}], \forall x \in \mathbb{R}$$

n.b.

$$F_{\mathbb{X}}(x) = \int_{-\infty}^x f_{\mathbb{X}}(\alpha)d\alpha = \Phi(\frac{x-\mu}{\sigma})$$

$$\text{where } \Phi(r) = \int_{-\infty}^r \frac{1}{\sqrt{2\pi}} e^{\frac{-z^2}{2}} dz$$

So if \mathbb{X} is a Gaussian RV with parameters $\mu \in \mathbb{R}$ and $\sigma > 0$, then

$$\mathcal{P}(\{a < \mathbb{X} \leq b\}) = \Phi(\frac{b-\mu}{\sigma}) - \Phi(\frac{a-\mu}{\sigma})$$

uniform

$$\mathbb{X} = \text{u}[a, b], a < b$$

$$f_{\mathbb{X}}(x) = \frac{1}{b-a} \cdot 1_{[a,b]}(x)$$

$$F_{\mathbb{X}}(x) = \int_{-\infty}^x f_{\mathbb{X}}(\alpha)d\alpha$$

exponential

An RV \mathbb{X} with pdf of the form

$$f_{\mathbb{X}}(x) = \alpha e^{-\alpha x} \cdot 1_{[0,\infty]}(x)$$

where $\alpha > 0$ is called an exponential RV with parameter α .

binomial

Discrete RV taking on values in the set $\{0, 1, 2, \dots, n\} \subset \mathbb{R}$ with pmf

$$\mathcal{P}_{\mathbb{X}}(k) = \binom{n}{k} p^k (1-p)^{n-k},$$

$$F_{\mathbb{X}}(x) = \sum_{k=0}^{m(x)} \binom{n}{k} p^k (1-p)^{n-k}$$

$$\text{where } m(x) \leq x < m(x) + 1,$$

$$f_{\mathbb{X}}(x) = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} \delta(x-k)$$

Conditional cdf and pdf

$$F_{\mathbb{X}}(x|M) = \mathcal{P}(\{\mathbb{X} \leq x\}|M) = \frac{\mathcal{P}(\{\mathbb{X} \leq x\} \cap M)}{\mathcal{P}(M)}$$

$$f_{\mathbb{X}}(x|M) = \frac{dF_{\mathbb{X}}(x|M)}{dx}$$

Total Probability Law

$$F_{\mathbb{X}}(x) = F_{\mathbb{X}}(x|A_1)P(A_1) + \dots + F_{\mathbb{X}}(x|A_n)P(A_n)$$

$$f_{\mathbb{X}}(x) = f_{\mathbb{X}}(x|A_1)P(A_1) + \dots + f_{\mathbb{X}}(x|A_n)P(A_n)$$

Bayes Formula

$$\mathcal{P}(A|\{\mathbb{X} \leq x\}) = \frac{F_{\mathbb{X}}(x|A)\mathcal{P}(A)}{F_{\mathbb{X}}(x)}$$

Now consider $\mathcal{P}(A|B) = \frac{\mathcal{P}(B|A)\mathcal{P}(A)}{\mathcal{P}(B)}$ when $B = \{x_1 < \mathbb{X} \leq x_2\}$:

$$\begin{aligned} \mathcal{P}(A|\{x_1 < \mathbb{X} \leq x_2\}) &= \\ \hookrightarrow &= \frac{F_{\mathbb{X}}(x_2|A) - F_{\mathbb{X}}(x_1|A)\mathcal{P}(A)}{F_{\mathbb{X}}(x_2) - F_{\mathbb{X}}(x_1)} \end{aligned}$$

For $\mathbb{X} = x$,

$$\begin{aligned} \mathcal{P}(A|\{\mathbb{X} = x\}) &= \frac{f_{\mathbb{X}}(x|A)}{f_{\mathbb{X}}(x)}\mathcal{P}(A) \\ \mathcal{P}(A) &= \int_{-\infty}^{\infty} \mathcal{P}(A|\{\mathbb{X} = x\})f_{\mathbb{X}}dx \end{aligned}$$

Bayes’ Theorem

$$f_{\mathbb{X}}(x|A) = \frac{\mathcal{P}(A|\{\mathbb{X} = x\})f_{\mathbb{X}}(x)}{\int_{-\infty}^{\infty} \mathcal{P}(A|\{\mathbb{X} = \alpha\})f_{\mathbb{X}}(\alpha)d\alpha}$$