

Online Appendix

OA-1 Summary Statistics

Table OA-1: Summary Statistics - Sellers and Buyers in 2016

	<i>Sellers</i>			<i>Buyers</i>		
	Mean	Median	SD	Mean	Median	SD
Total Sales (million USD)	14.95	8.26	24.33	2.35	0.20	24.33
Total Inputs (million USD)	10.58	5.31	18.94	1.92	0.15	24.13
Accounting Markup	1.21	1.20	0.20	1.21	1.10	0.61
Age	30.47	29.00	19.16	15.18	14.00	9.75
Import Share (%)	24.47	21.38	22.96	3.82	0.00	13.49
Export Share (%)	5.81	0.00	19.11	1.06	0.00	8.87
Observations	49			28,138		

Notes: This table reports summary statistics about the size, age, and trade exposure of buyers and sellers in the sample for the year 2016. Monetary values are in U.S. dollars for 2016.

Table OA-2: Industrial Composition of Buyers by Selling Sector

Seller Industry	Ranking	Buyer Industry	Average % Share Pairs
Textiles	1	Wholesale & Retail	40
Textiles	2	Manufacturing	15
Textiles	3	Professional Activities	8
Textiles	4	Agriculture	5
Textiles	5	Transporation & Storage	3
Textiles	6	Other	28
Pharmaceutical	1	Wholesale & Retail	46
Pharmaceutical	2	Human Health	17
Pharmaceutical	3	Manufacturing	10
Pharmaceutical	4	Construction	4
Pharmaceutical	5	Professional Activities	3
Pharmaceutical	6	Other	20
Cement-Products	1	Wholesale & Retail	25
Cement-Products	2	Construction	20
Cement-Products	3	Professional Activities	16
Cement-Products	4	Manufacturing	8
Cement-Products	5	Real Estate	5
Cement-Products	6	Other	26

Notes: This table provides a breakdown of the industrial composition of buyers for each selling sector.

Table OA-3: Summary Statistics - Electronic Invoice Database

	Mean	Median	SD
N. Buyers	8,028.41	613.50	25,078.11
N. Buyers (> USD250)	801.86	281.50	1,269.02
N. Buyers (> USD4,800)	160.49	73.00	265.19
Total Sales (million USD)	16.58	7.23	29.44
Total Q (million)	5.42	1.20	9.01
Q per Buyer	12,455.39	1,495.22	25,823.40
Bill per Buyer (USD)	43,490.37	9,067.65	105,840.28
Bill per Buyer (USD) (> USD250)	61,756.13	14,898.73	143,297.12
Bill per Buyer (USD) (> USD4,800)	176,632.43	46,899.44	444,389.15
Observations	49		

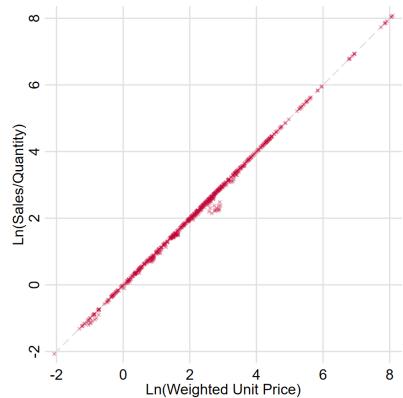
Notes: This table reports summary statistics of the electronic invoice database. N. buyers refers to the number of unique buyers each seller in the sample has on average over 2016 and 2017. Quantity is the sum of all quantities across products. Bill per buyer is the total value of the transactions between buyer and seller. The thresholds > 250 and > 4,800 correspond to the data selection thresholds in [Bernard et al. \(2019\)](#) and [Alfaro-Urena et al. \(2022\)](#), respectively.

Table OA-4: Example - Product Information, Prices, and Marginal Costs

Industry	Firm-ID	Product Description	Observed Unit Price	Imputed Marginal Cost
Textiles	1	Teddy King, Size 55, Brim 7CM, Color-B02 [Panama Hat]	33.9	11.96397
Textiles	2	Shirt, R:1931, Squares	19.34	9.851618
Textiles	3	Tank Undershirt, Male, Size M, White	10.26758	6.716691
Textiles	4	Betty K246	19.44	16.9426
Textiles	5	Bikini, Woman, 500306, Black, L	13.5	16.77684
Textiles	6	Ribbon, Black, 30 mm X 700	26.62	1.859854
Textiles	7	Skirt, Tropical Squares, Scottish	46.01	17.76808
Textiles	8	Boots, LLN NG AM, Size 39	7.091094	2.168639
Textiles	9	Elastic Socks, Nylon and Cotton	16.55658	8.476095
Textiles	10	Jacket, Kids, Spiderman Print, Hoodie	18.3	7.112451
Pharmaceutical	1	Nitazoxanida, 500mg X 6 tablets	5.27	4.831101
Pharmaceutical	1	Clopidogrel Tarbis 75 mg film-coated tablets	12.9	6.566845
Pharmaceutical	2	Losartan/Hydrochlorothiazide, 100mg X 28 tablets	5.04	.7821692
Pharmaceutical	3	B Complex, Syrup 120 ml	2.32	.8053265
Pharmaceutical	4	Sodium perborate, mint oil, saccharin	4.688695	1.814284
Pharmaceutical	5	Boldenone 50, Injectable, Bottle X 500 ml	123.12	3.014165
Pharmaceutical	6	Pinaver, Film-coated, 100 mg X 20 tables	10.32	2.623771
Pharmaceutical	7	Endobion X 60 tablets	14.83333	5.48487
Pharmaceutical	7	Prostageron X 60 capsules	14.75	7.036339
Pharmaceutical	8	Oral rehydration solution, cherry, 500ml	2.67	1.796762
Cement-Products	1	Gray French Pedestrian Paving Stone	11.27652	18.10799
Cement-Products	2	Corrugated Plate	23.73	9.56013
Cement-Products	3	Polymer-modified adhesive mortar for ceramics, 25kg	6.310047	2.988913
Cement-Products	4	Polymer-modified adhesive mortar for ceramics, 25kg	6.944	12.36252
Cement-Products	5	Polymer-modified adhesive mortar for ceramics, 25kg	6.650971	3.450306
Cement-Products	6	Straight Pole 21m x 1400kg, Reinforced Concrete	882	73.94785
Cement-Products	6	Straight Pole 21m x 2400kg, Reinforced Concrete	1362.73	73.94785
Cement-Products	7	Tile 50x50x2 cm (Color)	32	6.622508
Cement-Products	8	MFC Concrete, 300, XXXXX XXXX-XXXX	94	50.34137
Cement-Products	8	CFC Concrete, 240, XXXXX XXXX-XXXX	79.428	50.34137

Notes: This table presents a sample of ten random products from each of the studied sectors (textiles, pharmaceutical, and cement-products), with product descriptions translated into English and sensitive information, such as brand names, removed to ensure confidentiality. The observed average unit prices reflect the listed prices reported by the firms, while the imputed marginal costs are estimated using the firms' average cost as a proxy.

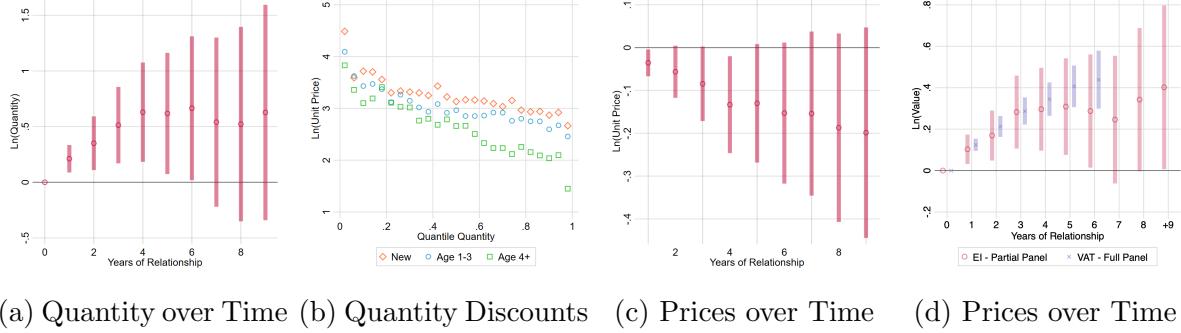
Figure OA-1: Average Price vs Weighted Price



Notes: This figure plots average unit prices (in logs) against weighted prices (in logs) across buyer-seller-year. Average unit prices are calculated by dividing total value of yearly transactions by total quantity purchased (pooling different products). Weighted prices are calculated by summing transaction-product-level unit prices by total expenditure share.

OA-2 Motivating Evidence - Robustness

Figure OA-2: Motivating Facts - Robustness



(a) Quantity over Time (b) Quantity Discounts (c) Prices over Time (d) Prices over Time

Notes: Panel a) plots the coefficients of log total quantity on relationship age dummies controlling for pair fixed effects. Total quantity is obtained by aggregating all the reported units of sold goods. Standard errors are clustered at the pair-level. Panel b) shows the relationship between quantity purchased and average log unit price (right-panel) through binscatters of the measure of unit price against quantile of quantity by age of relationship. Quantiles of quantity are calculated for each seller-relationship age combination. Panel c) plots the regression coefficients of log unit prices on years of relationship, controlling for pair fixed effects. Standard errors are clustered at the pair-level. Panel d) plots regression coefficients for the value of total sales between buyer and supplier on age of relationship, controlling for pair fixed effects. The red figures use the electronic invoice database and are constructed using only a partial panel of two observations per pair for years 2016-2017. The purple marks are constructed using multiple observations of buyer-seller pairs from the VAT B2B database for years 2007-2015 for the sellers in the electronic invoice database.

Table OA-5: Benchmark: Quantity Discounts

VARIABLES	(1) ln(Price)
ln(Quantity)	-0.220*** (0.0238)
Constant	3.046*** (0.0718)
Seller-Year FE	Yes
Observations	76,473
R-squared	0.666

Notes: This table presents a regression of log average unit prices on log quantity, controlling for seller-year fixed effects. Standard errors are clustered at the seller-year level. *** p<0.01, ** p<0.05, * p<0.1

Table OA-6: **Robustness - Standardized Log Price**

VARIABLES	(1) Stdz. ln(Price)	(13) Stdz. ln(Price)
Stdz. ln(Quantity)	-0.0463*** (0.00722)	-0.0454*** (0.00631)
Age of Relationship	-0.00552*** (0.00146)	-0.00395*** (0.00134)
ln(Assets Buyer)	0.00131*** (0.000318)	0.000832*** (0.000230)
Supply Share	0.0262* (0.0157)	0.0143 (0.0145)
Demand Share	0.0119 (0.0486)	0.00402 (0.0454)
Observations	73,633	73,626
R-squared	0.082	0.091
Controls	Yes	Yes
Year FE	Yes	Yes
Buyer Sector FE	No	Yes

Notes: This table presents regressions regressions of standardized unit prices on age of relationship, standardized quantity, and different buyer characteristics. Controls include Exporter, Importer, Business Group, Multinational Dummies, as well as Distance between HQs, Sales, Age, and Number of Employees of Buyer. Standard errors are clustered at the seller-year level. *** p<0.01, ** p<0.05, * p<0.1

Table OA-7: Robustness - Seller's Sector

VARIABLES	(1)	(2)
	Stdz. ln(Price)	ln(Price)
<i>Textiles</i>		
Age of Relationship	-0.00319*	-0.0973***
	(0.00180)	(0.0305)
<i>Pharmaceuticals</i>		
Age of Relationship	-0.00691***	-0.0308**
	(0.00185)	(0.0139)
<i>Cements</i>		
Age of Relationship	-0.00447***	-0.0387***
	(0.00141)	(0.00728)
Seller-Year FE	No	Yes
Controls	Yes	Yes
Observations	73,633	73,633
R-squared	0.083	0.608

Notes: This table presents regression of prices on age of relationship by sector of the seller. Column (1) presents results for the standardized log prices. Column (2) presents results for log average price, controlling for seller-year fixed effects. Both columns control for standardized quantities as well as all variables in Table OA-6. Standard errors are clustered at the seller-year level. *** p<0.01, ** p<0.05, * p<0.1

Table OA-8: Price Dynamics by Payment Method

VARIABLES	(1)	(2)	(3)	(4)
	Stdz. ln(Price)	Stdz. ln(Price)	ln(Price)	ln(Price)
Payment Method	Pay-in-advance	Trade-Credit	Pay-in-advance	Trade-Credit
Age of Relationship	0.000360	-0.00248***	0.0314	-0.0240***
	(0.00215)	(0.000258)	(0.0266)	(0.00516)
Quantity Control	Yes	Yes	Yes	Yes
Seller-Year FE	Yes	Yes	No	No
Pair FE	No	No	Yes	Yes
Observations	8,777	68,730	1,548	33,680
R-squared	0.148	0.134	0.947	0.937

Notes: This table presents a regression of unit prices on age of relationship, controlling for quantity, by payment modality. The sample excludes buyers that switch between modalities. Columns (1) and (2) uses standardized prices and controls for seller-year fixed effects, while Columns (3) and (4) relies instead on average unit prices while controlling for seller-buyer fixed effects. Standard errors are clustered at the seller-buyer level. *** p<0.01, ** p<0.05, * p<0.1

OA-3 Existence and Non-Stationarity

To prove existence, I build on two results of the literature. First, I use the result of non-linear pricing of Jullien (2000) to prove the existence of a stationary optimal contract in the presence of heterogeneous participation constraints. I do so by showing the equivalence between the stationary contract with limited enforcement and a non-linear pricing problem with heterogeneous outside options. Then, similar to the argument in Martimort et al. (2017), I offer an simple non-stationary deviation that dominates the

stationary optimal contract.

Note that I will show existence results under the assumption of no exit, i.e., $X(\theta) = 0$ for all θ . To prove existence with exit, one must simply replace the discount factor δ for $\tilde{\delta} \equiv \min\{\delta(\theta)\}$, where $\delta(\theta) = \delta(1 - X(\theta))$ is the discount factor that accounts for heterogeneous breakups. This change will only affect one of the assumptions discussed below and set an upper bound in the worse-case exit rate.

OA-3.1 Existence of Stationary Contract

The model in [Jullien \(2000\)](#) solves the following problem:

$$\max_{\{t(\theta), q(\theta)\}} \int_{\underline{\theta}}^{\bar{\theta}} [t(\theta) - cq(\theta)]f(\theta)d\theta \quad \text{s.t.} \quad (\text{IR Problem})$$

$$v(\theta, q(\theta)) - t(\theta) \geq v(\theta, q(\hat{\theta})) - t(\hat{\theta}) \quad \forall \theta, \hat{\theta} \quad (\text{IC})$$

$$v(\theta, q(\theta)) - t(\theta) \geq \bar{u}(\theta) \quad \forall \theta. \quad (\text{IR})$$

Under a modified first-order approach, the seller's first-order condition is given by:

$$v_q(\theta, q(\theta)) - c = \frac{\gamma(\theta) - F(\theta)}{f(\theta)} v_{\theta q}(\theta, q(\theta)), \quad (27)$$

for each type θ , and the complementary slackness condition on the IR constraints:

$$\int_{\underline{\theta}}^{\bar{\theta}} [u(\theta) - \bar{u}(\theta)]d\gamma(\theta) = 0. \quad (28)$$

[Jullien \(2000\)](#) shows that under three assumptions there exists a unique optimal solution in which all consumers participates, which is characterized by the first-order conditions [27](#) and complementary slackness condition [28](#) with $q(\theta)$ increasing. The first-assumption is potential separation (PS), which requires that the optimal solution is non-decreasing in θ , and satisfied under weak assumptions on the distribution of θ and the curvature of the surplus relative to the return of the buyer. In particular, it requires that

$$\begin{aligned} \frac{d}{d\theta} \left(\frac{S_q(\theta, q)}{v_{\theta q}(\theta, q)} \right) &\geq 0 \\ \frac{d}{d\theta} \left(\frac{F(\theta)}{f(\theta)} \right) &\geq 0 \geq \frac{d}{d\theta} \left(\frac{1 - F(\theta)}{f(\theta)} \right). \end{aligned}$$

The second and *key* assumption is homogeneity (H), requiring that there exists a quantity profile $\{\bar{q}(\theta)\}$ such that the allocation with full participation $\{\bar{u}(\theta), \bar{q}(\theta)\}$ is implementable in that $\bar{u}'(\theta) = v_\theta(\theta, \bar{q}(\theta))$ and $\bar{q}(\theta)$ is weakly increasing. This assumption implies that the reservation return can be implemented as a contract without excluding any type, ensuring that incentive compatibility is not an issue when the individual rationality constraint is binding. Lastly, the assumption of full participation (FP) assumes all types participate, and is satisfied when (H) holds and the surplus generated in the reservation return framework is greater than the private return to the buyer, i.e. $s(\theta, \bar{q}(\theta)) \geq \bar{u}(\theta)$.

I show that my setting can be rewritten in terms of [Jullien \(2000\)](#), implying that an optimal separating stationary contract exists. The seller chooses the optimal stationary contract $\{t(\theta), q(\theta)\}$ that satisfy incentive-compatibility and the limited enforcement

constraint. Formally, the seller solves the problem:

$$\max_{\{t(\theta), q(\theta)\}} \frac{1}{1-\delta} \int_{\theta}^{\bar{\theta}} [t(\theta) - cq(\theta)] f(\theta) d\theta \quad \text{s.t.} \quad (\text{LE Problem})$$

$$v(\theta, q(\theta)) - t(\theta) \geq v(\theta, q(\hat{\theta})) - t(\hat{\theta}) \quad \forall \theta, \hat{\theta} \quad (\text{IC})$$

$$\frac{\delta}{1-\delta} (v(\theta, q(\theta)) - t(\theta)) \geq t(\theta) = v(\theta, q(\theta)) - u(\theta), \quad \forall \theta, \quad (\text{LC})$$

where $u(\theta)$ is the return obtained by type θ . The limited enforcement constraint can be easily written as the IR constraint in [Jullien \(2000\)](#):

$$u(\theta) \geq (1-\delta)v(\theta, q(\theta)) \equiv \bar{u}(\theta) \quad \forall \theta. \quad (\text{LE'})$$

In my model, with $v(\theta, q) = \theta v(q)$, the first condition of assumption PS is always satisfied as

$$\frac{d}{d\theta} \left(\frac{S_q(\theta, q)}{v_{\theta q}(\theta, q)} \right) = \frac{d}{d\theta} \left(\theta - \frac{c}{v'(q)} \right) \geq 0 \iff 1 \geq 0 \quad (\text{A1})$$

As stated earlier, the second condition of assumption PS is satisfied for a wide-range of distributions for θ . Therefore, assumption PS is satisfied for any of those distributions.

Then, consider Assumption H. It requires that an allocation $\{\bar{q}(\theta)\}$ exists such that $\bar{u}'(\theta) = v_\theta(\theta, \bar{q}(\theta))$ and $\bar{q}(\theta)$ is weakly increasing. Notice that under [LE'](#), we can define $\bar{q}(\theta)$ as $\bar{u}'(\theta) = (1-\delta)[\theta v'(q(\theta))q'(\theta) + v(q(\theta))] = v(\bar{q}(\theta))$. Define $G(\bar{q}, \theta) = v(\bar{q}) - (1-\delta)[\theta v'(q(\theta))q'(\theta) + v(q(\theta))] = 0$. By the implicit function theorem, $\bar{q}(\theta)$ is weakly increasing if

$$\begin{aligned} \bar{q}'(\theta) &= -\frac{dG/d\theta}{dG/d\bar{q}} \\ &= \frac{(1-\delta)[v'(q(\theta))q'(\theta) + \theta v''(q(\theta))(q'(\theta))^2 + \theta v'(q(\theta))q''(\theta) + v'(q(\theta))]}{v'(\bar{q})} \geq 0 \\ &\iff v'(q(\theta))[1 + q'(\theta) + \theta q''(\theta)] + \theta v''(q(\theta))(q'(\theta))^2 \geq 0 \\ &\iff \frac{q'(\theta) + \theta q''(\theta) + 1}{\theta(q'(\theta))^2} \geq A(q) \\ &\iff \left(\frac{T''(q)}{T'(q)} + A(q) \right) \left(1 + \theta(q)\theta'(q)r(q) + \theta'(q) \right) \geq A(q) \\ &\iff \frac{T''(q)}{T'(q)} \frac{M(q)}{M(q) - 1} \geq A(q), \end{aligned}$$

where $M(q) \equiv 1 + \theta(q)\theta'(q)r(q) + \theta'(q)$ and $r(q) = g^{-1}(q)$ for $g(\theta) \equiv q''(\theta)$. As we expect $T''(q) < 0$ and $T'(q) > 0$, it is necessary that $M(q)/(M(q) - 1) < 0$. Such condition will be satisfied if $M(q) < 1$ and $M(q) > 0$, which imply that

$$\begin{aligned} r(q)\theta(q) &< -1 \\ \text{and} \\ \theta'(q) &< \frac{1}{\theta(q)|r(q)| - 1}. \end{aligned} \quad (\text{A2})$$

The first condition sets restrictions on the rate of change of quantities, which requires $q''(\theta)$ to be negative, restricting how convex $u(\theta)$ can be. The second condition requires that quantities increase at a minimum rate. Moreover, the condition sets bounds on the price discounts offered relative to the buyers' return curvature at a given quantity.

Lastly, full participation requires H to hold as well as $s(\theta, \bar{q}(\theta)) \geq (1 - \delta)\theta v(\bar{q}(\theta))$. The condition becomes:

$$\delta \geq \frac{c\bar{q}(\theta)}{\theta v(\bar{q}(\theta))}, \quad (\text{A3})$$

which requires that agents value the future high enough, such that discount factor be greater than the ratio of average cost to average return.

Let $\{t^{st}(\theta), q^{st}(\theta)\}$ be the solution to the problem characterized by equations 27 and 28. Assuming that the $v(\cdot)$, $F(\theta)$, and δ are such that A1, A2, and A3 hold for $\{t^{st}(\theta), q^{st}(\theta)\}$, then $\{t^{st}(\theta), q^{st}(\theta)\}$ is uniquely optimal.

OA-3.2 Optimality of Non-Stationary Contracts

Having established the existence of an optimal stationary contract, I now show that a non-stationary contract exists, which dominates the stationary contract. A similar argument was briefly discussed in the working paper version of Martimort et al. (2017).

Consider the following deviation from the stationary contract, in which at tenure 0, the return obtained by the buyer is given by:

$$u_0(\theta) = u^{st}(\theta) - \varepsilon,$$

for some $\varepsilon > 0$ sufficiently small, $u_{st} = \theta v(q^{st}(\theta)) - t^{st}(\theta)$ and $t_0(\theta) = t^{st}(\theta)$. Define $q_0(\theta)$ to so it satisfies that the deviation defined above.. Under this deviation, the enforcement constraint at $\tau = 0$ is:

$$t^{st}(\theta) \leq \frac{\delta}{1 - \delta} [\theta v(q^{st}(\theta)) - t^{st}(\theta)],$$

which is identical to the one in the stationary contract, which we know $\{t^{st}(\theta), q^{st}(\theta)\}$ satisfy. Moreover, the incentive compatibility constraint is still satisfied as $\hat{\theta}$ maximizes

$$u_0(\theta, \hat{\theta}) + \frac{\delta}{1 - \delta} u^{st}(\theta, \hat{\theta}) = \frac{\delta}{1 - \delta} u^{st}(\theta, \hat{\theta}) - \varepsilon,$$

where $u_\tau(\theta, \hat{\theta}) \equiv \theta v(q_\tau(\hat{\theta})) - t_\tau(\hat{\theta})$.

Under this alternative scheme, the seller obtains additional payoff ε while still satisfying both the incentive compatibility and limited enforcement constraints. Therefore, the optimal contract is non-stationary.

OA-4 Proof that Gamma Equals One for Highest Type

I prove that $\Gamma_\tau(\bar{\theta}) = 1$ for all τ . To begin, recall we assumed the outside option $\bar{u}_\tau(\theta)$ was equal to zero for all τ and all θ . Suppose instead that at some k , the outside option is uniformly shifted downward by > 0 for all θ , that is, $\bar{u}_k(\theta) = -\varepsilon$. The enforcement constraint at k is now given by:

$$\delta \left[\sum_{s=1}^{\infty} \delta^{s-1} u_{k+s}(\theta) \right] - \bar{u}_k(\theta) = \sum_{s=1}^{\infty} \delta^s u_{k+s}(\theta) + \varepsilon \geq t_k(\theta) = \theta v(q_k(\theta)) - u_k(\theta). \quad (29)$$

The seller's problem in the Lagrangian-form is

$$W(\varepsilon) = \max_{\{q_\tau(\theta), u_\tau(\theta)\}} \sum_{\tau=0}^{\infty} \delta^\tau \left\{ \int_{\underline{\theta}}^{\bar{\theta}} [\theta v(q_\tau(\theta)) - cq_\tau - u_\tau(\theta)] f(\theta) d\theta + \right. \quad (30)$$

$$\left. \int_{\underline{\theta}}^{\bar{\theta}} \left[\sum_{s=1}^{\infty} \delta^s u_{\tau+s} + \varepsilon * 1\{\tau = k\} - t_\tau(\theta) \right] d\Gamma_\tau(\theta) \right\} \quad (31)$$

such that $u'_\tau(\theta) = \theta v'(q_\tau(\theta))$ for all τ, θ . The change in the value of the problem of the seller given the uniform change in outside options is:

$$\frac{dW(\varepsilon)}{d\varepsilon} = \delta^k \int_{\underline{\theta}}^{\bar{\theta}} d\Gamma_k(\theta), \quad (32)$$

where the integral is the cumulative multiplier.

I argue that the quantities that solve the original problem still maximize the current one but that the transfers are all shifted upward by the constant ε . That is, if $q_\tau(\theta)$ is the solution for the problem with $\bar{u}_\tau(\theta) = 0$ for all θ and all τ with associated $t_\tau(\theta)$, $q_\tau(\theta)$ is also the solution for the problem with outside options $\bar{u}_\tau(\theta) = -\varepsilon 1\{\tau = k\}$ for all θ and all τ with associated transfers equal to $t_\tau(\theta) + \varepsilon 1\{\tau = k\}$. The value of the problem for the seller is:

$$W(\varepsilon) = \sum_{\tau=0}^{\infty} \delta^\tau \left\{ \int_{\underline{\theta}}^{\bar{\theta}} [t_\tau(\theta) + \varepsilon 1\{\tau = k\} - cq_\tau] f(\theta) d\theta \right\} \quad (33)$$

$$= \sum_{\tau=0}^{\infty} \delta^\tau \left\{ \int_{\underline{\theta}}^{\bar{\theta}} [t_\tau(\theta) - cq_\tau] f(\theta) d\theta \right\} + \delta^k \varepsilon. \quad (34)$$

So

$$\frac{dW(\varepsilon)}{d\varepsilon} = \delta^k. \quad (35)$$

Therefore, the cumulative multiplier for any k will satisfy the following property:

$$\Gamma_k(\bar{\theta}) \equiv \int_{\underline{\theta}}^{\bar{\theta}} d\Gamma_k(\theta) = \frac{dW(\varepsilon)}{d\varepsilon} \frac{1}{\delta^k} = 1. \quad (36)$$

OA-5 Additional Theoretical Results

OA-5.1 Model Dynamics

Proofs are available in Supplemental Material Section [SM3](#).

Quantity Discounts

Define $T_\tau(q_\tau(\theta)) \equiv t_\tau(\theta_\tau(q))$, $\Lambda_\tau(\theta) \equiv \Gamma_\tau(\theta) - \sum_{s=0}^{\tau-1} (1 - \Gamma_s(\theta)) + \theta \gamma_\tau(\theta)$, and $\lambda_\tau(\theta) \equiv d\Lambda_\tau/d\theta$. The price schedule is said to feature quantity discounts if $T_\tau''(q) < 0$.

Proposition 1. *Assume strict monotonicity of quantity $q'_\tau(\theta) > 0$ and that $\lambda_\tau(\theta) < f_\tau(\theta)$. If the densities $f_\tau(\theta)$ satisfy log-concavity and $d(F_\tau(\theta)/f_\tau(\theta))/d\theta \geq F_\tau(\theta)/[(\theta - 1)f_\tau(\theta)]$, then the tariff schedule exhibits quantity discounts, $T_\tau''(q) \leq 0$ for each $q = q_\tau(\theta)$, $\theta \in (\underline{\theta}, \bar{\theta})$ and τ .*

Intuitively, the condition states that for a general class of distributions, as long as the incentive-compatibility marginal effects dominate those of the limited enforcement, the

seller finds it optimal to offer quantity discounts at any relationship age. This is likely to be satisfied if the limited enforcement constraint is slack for some buyers already at their first interaction. Moreover, it also requires the enforcement constraint is slack for all buyers in the long run. This last requirement is in line with the model of [Martimort et al. \(2017\)](#), where buyers reach a *mature* phase in which the constraints no longer bind, as well as Proposition 2 below, which also finds that trade reaches a mature phase.

In terms of generality, the usual monopolist screening problem requires (or uses) log-concavity of $f(\theta)$.³³ I am strengthening the requirement that the evolution of the distribution also satisfies log-concavity, implicitly placing bounds on the distribution of exit rates over types.

The second condition strengthens the conditions on the dynamic distribution of types, in order to guarantee that the seller has the desire of price discriminating across types.

An alternative way to consider this property is to use [*t*-RULE](#) to obtain that the tariff schedule is concave if and only if $q'_\tau(\theta) > v'(q_\tau(\theta))/-v''(q_\tau(\theta)\theta)$. As long as quantities increase by types fast enough, then the seller will offer quantity discounts. The rate at which the quantities have to increase is determined by the level of the type and the curvature of the return function.

Evolution of Quantities

Next, I discuss how quantities evolve in Proposition 2.

Proposition 2. *For each θ , quantity increases monotonically in τ (i.e., $q_\tau(\theta) \leq q_{\tau+1}(\theta)$) if and only if the limited enforcement constraint is relaxed over time ($\gamma_\tau(\theta) \geq \gamma_{\tau+1}(\theta)$). Moreover, there is a time τ^* such that $\forall \tau \geq \tau^*$, $\gamma_{\tau^*}(\theta) = 0$ for all θ and $q_{\tau^*}(\theta) \geq q_\tau(\theta)$ for all $\tau < \tau^*$ and all θ .*

In the model, quantities go hand-in-hand with enforcement constraints. Although the exact path depends on further assumptions on the return function and the distribution of types, the model predicts that quantities will reach a mature phase in which constraints no longer bind. At this mature phase, quantities will be at their highest level in the relationship.

Discounts over time

The model also offers conditions under which discounts over time are observed.

Proposition 3. *If $M_{\tau+1}(\theta) \equiv q_{\tau+1}(\theta) - q_\tau(\theta) \geq 0$ for all θ and with strict inequality for $\underline{\theta}$, then $p_{\tau+1}(q) \equiv T_{\tau+1}(q)/q < T_\tau(q)/q \equiv p_\tau(q)$.*

As long as quantities (weakly) increase from τ to $\tau + 1$, unit prices at any given q decrease. The intuition behind this result is that marginal prices match marginal returns. A right-ward shift in quantities for (some) buyers further lowers marginal returns, requiring a decrease in marginal prices as well. As such, average prices will be lower at each q as well.

To further understand the dynamics in the model, I present a solved two-type example in Supplemental Material Section [SM4](#). The example illustrates the backloading of prices and quantities together with quantity discounts as a way to maximize lifetime profits for the seller while preventing opportunistic behavior from the buyer.

³³Log-concavity of a density function $g(x)$ is equivalent to $g'(x)/g(x)$ being monotone decreasing. Families of density functions satisfying log-concavity include: uniform, normal, extreme value, exponential, amongst others.

OA-5.2 Static Efficiency of Limited Enforcement

We now turn to analyzing the efficiency of contracts with limited enforcement. Relationship-specific total surplus (and thus efficiency) is determined by the total quantity transacted at a point in time. I concentrate on static (period-by-period) efficiency, as it is common in the relational contracting literature (e.g., as in [Fong and Li, 2017](#); [Kostadinov and Kuvalekar, Forthcoming](#)), rather than total lifetime efficiency.

For simplicity, suppose that $\theta\gamma_\tau(\theta)$ is small enough so the quantities allocated in the limited enforcement contract with no exit ($X(\theta) = 0$) and the assumed parametrization of $v(\cdot)$ are given by:

$$q_\tau^{LEM}(\theta)^{1-\beta} = \frac{k\beta}{c} \left[\theta - \frac{\Gamma_\tau(\theta) - F_\tau(\theta) - \sum_{s=0}^{\tau-1}(1 - \Gamma_s(\theta))}{f(\theta)} \right].$$

With some abuse of notation, define the modified value of the cumulative multiplier at time τ as $\tilde{\Gamma}_\tau(\theta) = \Gamma_\tau(\theta) - \sum_{s=0}^{\tau-1}(1 - \Gamma_s(\theta))$, so the allocation is given by:

$$q_\tau^{LEM}(\theta)^{1-\beta} = \frac{k\beta}{c} \left[\theta - \frac{\tilde{\Gamma}_\tau(\theta) - F(\theta)}{f(\theta)} \right].$$

Moreover, recall that the first-best outcome is given by:

$$q_\tau^{FB}(\theta)^{1-\beta} = \frac{k\beta}{c}\theta.$$

If $\tilde{\Gamma}_\tau(\theta) < F(\theta)$, there is overconsumption relative to first best. If $\tilde{\Gamma}_\tau(\theta) > F(\theta)$, there is underconsumption. If $\tilde{\Gamma}_\tau(\theta) = F(\theta)$, trade is fully efficient. Therefore, this limited enforcement model allows for the possibility of efficient trade, as well as inefficient trade either through underconsumption or overconsumption.

For the case with underconsumption, i.e., $\tilde{\Gamma}_\tau(\theta) > F(\theta)$, efficiency increases over time if $\tilde{\Gamma}_\tau(\theta) < \tilde{\Gamma}_{\tau-1}(\theta)$. By reordering and eliminating repeated terms, the condition becomes $\Gamma_\tau(\theta) < 1$. Thus, under the case with no exit and underconsumption, we expect efficiency to increase until pair-wise trade becomes unconstrained. Note, however, that quantities may converge at inefficient levels.

OA-5.3 Static Efficiency Relative to Perfect Enforcement

Comparing equations [SFOC](#) and [PE](#), in the case with no exit $X(\theta) = 0$ for all θ , the total quantity transacted is greater under full enforcement than under limited enforcement if:

$$(1 - \Gamma_\tau(\theta)) + \sum_{s=0}^{\tau-1}(1 - \Gamma_s(\theta)) - \theta\gamma_\tau(\theta) < 0. \quad (37)$$

For the types for which the limited enforcement constraint is not binding (so $\gamma_\tau(\theta) = 0$), except for the highest type, the inequality does not hold, and pair-wise welfare decreases under full enforcement. This will likely matter for middle/high types early on. Moreover, it might apply too for lower types in the long-term that started with binding constraints at the beginning for the contract but that grew over time to become unconstrained. Therefore, welfare can be greater under a long-term relational contract with limited enforcement than under perfect enforcement.

For types with $\gamma_\tau(\theta) > 0$, the inequality can be written as:

$$\theta - \frac{1 - \Gamma_\tau(\theta)}{\gamma_\tau(\theta)} > \frac{\sum_{s=0}^{\tau-1}(1 - \Gamma_s(\theta))}{\gamma_\tau(\theta)}.$$

The inequality above reminds us of a modified virtual surplus, where instead of the distribution of types we use the distribution of enforcement constraints. For perfect enforcement to be welfare increasing, the virtual surplus accounting for contemporaneous information rents of limited enforcement has to be greater than the information rents (promises to increase quantity) stemming from past enforcement constraints. Of course, early on, perfect enforcement could be more efficient, yet, as relationships age this might be more difficult to sustain.

In contrast with the arguments set forward in past literature, I have shown that in the interaction of market power and enforcement constraints could imply that weak legal enforcement is actually efficiency *increasing* at some points in time, and particularly so in the long-run. Intuitively, absent enforcement constraints, the seller is able to offer the profit-maximizing menu of quantities and prices. The buyer's ability to act opportunistically restricts how much the seller can extract and changes the surplus in favor of the buyer.

OA-6 Estimation

This section discusses the details of the estimation procedure.

OA-6.1 Tariff Function

In identification, I treated the tariff function $T_\tau(\cdot)$ as given. However, I observe only pairs of payments and quantities $(t_{i\tau}, q_{i\tau})$ for $i = 1, 2, \dots, N_\tau$ for each tenure. The pricing model discussed in section 4 implies that observed transfers lie on the curve $t = T(q)$, as they are both functions of the type $\theta_{i\tau}$ in a given tenure. As noted by [Luo et al. \(2018\)](#), observed prices and quantities may not lie on the curve, if there is measurement error or unobserved heterogeneity, introducing additional randomness beyond $\theta_{i\tau}$.

To deal with this additional randomness, I follow [Perrigne and Vuong \(2011\)](#), which show that the tariff function is nonparametrically identified under the assumption that observed tariffs differ from optimal tariffs due to random measurement error. In particular, observed tariffs are a function of optimal tariffs $t_{i\tau} = T(q_{i\tau})e^{v_{i\tau}}$, such that $v_{i\tau}$ is independent of $q_{i\tau}$.

I consider a parametric version of the model, in which $T_\tau(q) = e^{\beta_0\tau}q^{\beta_1\tau}$. This leads to the estimation model with measurement error:

$$\ln(t_{i\tau}) = \beta_{0\tau} + \beta_{1\tau}\ln(q_{i\tau}) + v_{i\tau}, \quad (38)$$

where $t_{i\tau}$ is the observed tariff and $q_{i\tau}$ is the observed quantities for buyer i with tenure τ . Under the given assumption of independence, the tariff schedule can be estimated via ordinary least squares. The estimated tariff schedule linking observed quantities is $\hat{T}_\tau(q_{i\tau}) = e^{\hat{\beta}_{0\tau}}q_{i\tau}^{\hat{\beta}_{1\tau}}$, while the marginal tariff is $\hat{T}'_\tau(q_{i\tau}) = \hat{\beta}_{1\tau}t_{i\tau}/q_{i\tau}$. Note that I allow for differences in tariff schedules across τ , responding to the dynamic treatment of the problem, i.e. the same level of quantity q may have different associated tariffs if the buyer-seller relationship is new or have been sustained for some years.

OA-6.1.1 Heterogenous Hazard Rates

I estimate heterogenous hazard rates at the percentile-tenure level. In particular, I rank buyers in percentiles of quantity for each tenure in 2016. I then calculate the share of buyers in each percentile that survived until 2017. To reduce the noise and preserve a monotonicity of hazard rate, I then approximate the estimated nonparametric hazard

rates as a logistic function of percentiles:

$$S_\tau(r) = \frac{\exp(a_\tau + b_\tau r)}{1 + \exp(a_\tau + b_\tau r)} + \varepsilon_\tau^s(r), \quad (39)$$

where $S_\tau(r)$ is the share of buyers surviving from 2016 until 2017 in percentile rank r for tenure τ and $\varepsilon_\tau^s(r)$ is Gaussian noise orthogonal to r .

OA-6.2 Marginal Cost

Marginal cost is estimated directly from the data under the assumption that marginal cost is equal to average variable cost. As defined in Section 2, average variable cost is defined as total expenditures and total wages divided by total quantity sold.

OA-6.3 LE Multipliers

Recall that the LE multiplier $\Gamma_\tau(\alpha)$ has the properties of a cumulative distribution function. Following Attanasio and Pastorino (2020), I parametrize the multiplier as a logistic distribution:³⁴

$$\Gamma_\tau(\alpha) = \frac{\exp(\phi_\tau(q_\tau(\alpha)))}{1 + \exp(\phi_\tau(q_\tau(\alpha)))}, \quad (40)$$

where $\phi_\tau(q_\tau(\alpha))$ is a polynomial up to the second degree. Under this parametrization, the derivative of the multiplier is $\gamma_\tau(\alpha) = \phi'_\tau(q_\tau(\alpha))\Gamma_\tau(\alpha)(1 - \Gamma_\tau(\alpha))$.

Moreover, I parametrize $\theta'(\alpha)/\theta(\alpha)$ as a inverse quadratic function of quantity:

$$\frac{\theta'(\alpha)}{\theta(\alpha)} = \frac{1}{d_0 + d_1 q_\tau(\alpha) + d_2 q_\tau(\alpha)^2}. \quad (41)$$

The key identification equation I:EQ provides the following estimating equation:

$$\begin{aligned} \frac{\hat{\beta}_{1\tau} p_\tau(\alpha) - \hat{c}}{\hat{\beta}_{1\tau} p_\tau(\alpha)} &= && \text{(Main Est. Eq.)} \\ &= && \\ &\frac{1}{d_0 + d_1 q_\tau(\alpha) + d_2 q_\tau(\alpha)^2} \left[\frac{\exp(\phi_\tau(q_\tau(\alpha)))}{1 + \exp(\phi_\tau(q_\tau(\alpha)))} - \alpha - \widehat{M}_\tau(\alpha) \right] \\ &+ \phi'_\tau(q_\tau(\alpha)) \frac{\exp(\phi_\tau(q_\tau(\alpha)))}{1 + \exp(\phi_\tau(q_\tau(\alpha)))} \left(1 - \frac{\exp(\phi_\tau(q_\tau(\alpha)))}{1 + \exp(\phi_\tau(q_\tau(\alpha)))} \right) + \varepsilon_\tau^g(\alpha), \end{aligned}$$

where I have used $p_{i\tau} = t_{i\tau}/q_{i\tau}$ and where ε^g is measurement error coming from the misspecification of Γ , the tariff function, or the marginal cost. Moreover, past multipliers are captured by $\widehat{M}_\tau(\alpha) \equiv \sum_{s=0}^{\tau-1} \widehat{S}_s^{\tau-s} (1 - \widehat{\Gamma}_s(\alpha))$, with \widehat{S}_s and $\widehat{\Gamma}_s(\alpha)$ for $s < \tau$ estimated in earlier stages and taken in τ as given. The equation is estimated via maximum likelihood under the assumption that ε^g is drawn from a Gaussian with parameters $(0, \sigma^{\varepsilon^g})$. This step in the estimation process recovers the parameters $\{\phi_\tau, d_0, d_1, d_2, \sigma^{\varepsilon^g}\}$.

To match previously estimated LE multipliers $\Gamma_s(\theta)$ to $\theta(\alpha)$ at tenure τ , I use the estimated hazard rates to generate a percentile-percentile transition matrix. Then, I can match percentiles matching α_s for $s < \tau$ to percentiles matching α_τ . Moreover, I use the estimated hazard rates for τ corresponding to α to properly discount past promises captured in past multipliers.

³⁴The multiplier function is the solution to a differential equation. As shown in Supplemental Material Section SM2, it is a function of the cumulative distribution of types θ , the marginal cost, and the expected base marginal return (i.e., depends on the curvature of the return function).

OA-6.4 Buyer Types and Type Distribution

Once Γ_τ and γ_τ are estimated, the consumer type $\theta_\tau(\alpha)$ is obtained from

$$\ln(\widehat{\theta}_\tau(\alpha)) = \quad (42)$$

$$\frac{1}{N_\tau} \sum_{k=1}^{N_\tau} \frac{1\{\alpha \geq k/N_\tau\}}{\widehat{\Gamma}_\tau(k/N_\tau) - k/N_\tau - \widehat{M}_\tau(k/N_\tau)} \left[1 - \frac{\widehat{c}}{\widehat{\beta}_{1\tau} p_\tau(k/N_\tau)} - \widehat{\gamma}_\tau(k/N_\tau) \right], \quad (43)$$

for $\alpha \in [0, (N_\tau - 1)/N_\tau]$ and where N_τ is the total count of buyers of tenure τ . The estimator for $\theta'_\tau(\alpha)$ is

$$\widehat{\theta}'_\tau(\alpha) = \frac{\widehat{\theta}_\tau(\alpha)}{\widehat{\Gamma}_\tau(\alpha) - \alpha - \widehat{M}_\tau(k/N_\tau)} \left[1 - \frac{\widehat{c}}{\widehat{\beta}_{1\tau} p_\tau(\alpha)} - \widehat{\gamma}_\tau(\alpha) \right]. \quad (44)$$

Finally, the density function $\widehat{f}_\tau(\theta(\alpha))$ is $1/\widehat{\theta}'_\tau(\alpha)$.

OA-6.4.1 Base Marginal Return and Return Function

The derivative of the transfer rule links the base marginal return with the marginal tariff and the consumer type: $v'(q_\tau(\alpha)) = T'_\tau(q_\tau(\alpha))/\theta_\tau(\alpha)$. Therefore, an estimator for the base marginal return is

$$\widehat{v'(q_\tau(\alpha))} = \frac{\widehat{\beta}_{1\tau} p_\tau(\alpha)}{\widehat{\theta}_\tau(\alpha)}. \quad (45)$$

Following the discussion in the identification section, $v(\cdot)$ is estimated by

$$v(q_\tau(\alpha)) = \widehat{T}_\tau(q_\tau(0)) + \frac{1}{N_\tau} \sum_{k=1}^{N_\tau} \widehat{v'(q_\tau(k/N_\tau))} 1\{\alpha \geq k/N_\tau\}. \quad (46)$$

OA-6.5 Parametrization of $v(\cdot)$ for Counterfactual Analysis

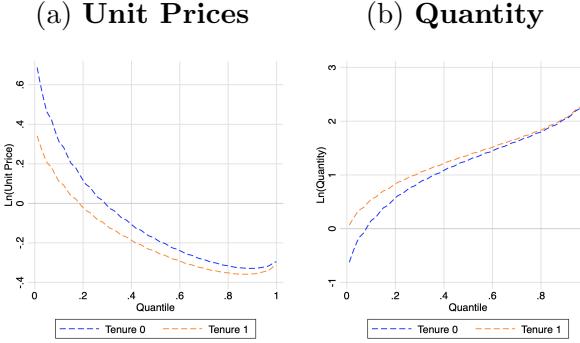
To calculate pair-specific efficient (first-best) quantities, I require estimated buyer types θ , base marginal returns $v'(\cdot)$ and seller marginal costs c . The range of optimal quantities may not be covered by the range of realized quantities, and thus, base marginal returns may be undefined for some quantities. For that reason, during counterfactual analysis, I parametrize the seller-specific marginal return functions $v(\cdot)$ as $v(q) = kq^\beta$, for $k > 0$ and $\beta \in (0, 1)$ and estimate these functions for each seller using linear least squares and the values of estimated marginal returns $\widehat{v'(\cdot)}$.

OA-7 Monte Carlo Study

The Monte Carlo studies the behavior of my estimators for two periods of a dynamic contract without breakups. I use the following design. The return function is $v(\theta, q) = \theta q^{1/2}$. The type distribution is Weibull with scale parameter equal to 1 and shape parameter equal to 2, $F(\theta) = 1 - \exp(-(\theta - 1)^2)$, normalized so $\underline{\theta} = 1$.³⁵ Marginal cost is 0.45. Although the multiplier function $\Gamma_\tau(\theta)$ is the solution to a differential equation linking the type distribution $F(\theta)$, the marginal cost, and the average base marginal return of types $\widehat{\theta} \leq \theta$, I parametrize it as a logistic distribution. In tenure 0, $\Gamma_0(\theta)$ has

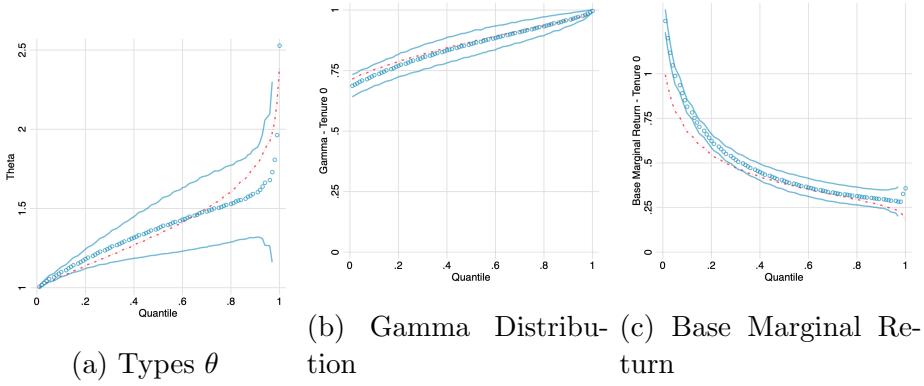
³⁵Recall that the model requires the type distribution to verify the monotone hazard rate condition, $\frac{d}{d\theta} \frac{F(\theta)}{f(\theta)} \geq 0 \geq \frac{d}{d\theta} \frac{1-F(\theta)}{f(\theta)}$. Distributions that satisfy the monotone hazard rate condition include: Uniform, Normal, Logistic, Extreme Value (including Frechet), Weibull (shape parameter ≥ 1), Exponential, and Power functions.

Figure OA-3: Prices and Quantities by Quantile



Notes: These figures show the level of prices and quantities by quantile of quantity for tenure 0 and tenure 1 in the Monte Carlo simulation.

Figure OA-4: Monte Carlo Results for Tenure 0



Notes: Panel (a) plots the true (red) and estimated distribution of types (in blue) by quantile of quantity. Panel (b) plots the true (red) and estimated value (blue) of the LE multiplier for tenure 0 by quantile of quantity. Panel (c) plots the true (red) and estimated value (blue) of the base marginal return for tenure 0 by quantile of quantity. Error margins indicate ± 1.96 variation around estimated mean from 300 simulations.

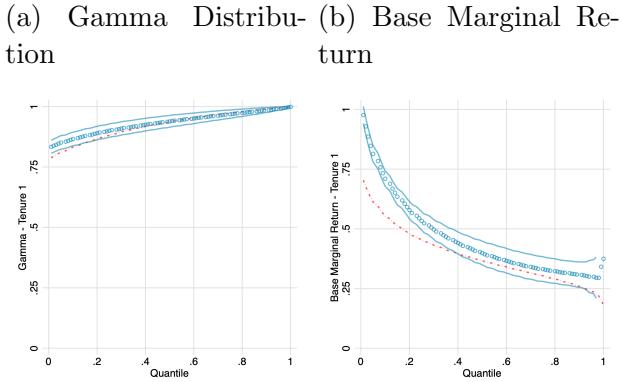
location parameter equal to 1 and scale parameter equal to 0.5. Instead, in tenure 1, $\Gamma_1(\theta)$ has location parameter 1 and scale 0.35. The lower scale parameter at tenure 1 reflects the idea that over time, the limited enforcement constraint is less binding. I construct the tariffs following [Pavan et al. \(2014\)](#): $t_\tau(\theta) = \theta q_\tau(\theta)^{1/2} - \int_\theta^\theta q_\tau(x)^{1/2} dx$.

I randomly draw 1000 values of θ using $F(\theta)$ and obtain corresponding quantities $q_0(\theta)$ and $q_1(\theta)$ using the first-order condition of the seller and the assumed parametrizations of the return function, marginal cost, and multiplier at tenure 0 and 1. Then, I obtain the corresponding tariffs and I apply my estimator as defined in the previous sections to estimate $\{\theta, U(\cdot), \Gamma_\tau(\cdot)\}$. I repeat this 300 times to construct the dispersion for my estimates.

Online Appendix Figure OA-3 shows the (log) average prices and average quantities generated by the model for the two types of tenure. The model delivers quantity discounts (decreasing unit prices in θ), strict monotonicity of quantity (increasing quantities in θ), and backloading in the dynamic model, namely, further discounts and larger quantities in tenure 1 for each θ .

Online Appendix Figure OA-4 shows the results of the estimated Gamma distribution

Figure OA-5: Monte Carlo Results for Tenure 1



Notes: Panel (a) plots the true (red) and estimated value (blue) of the LE multiplier for tenure 1 by quantile of quantity, with error margins indicating ± 1.96 variation around the estimated mean. Panel (b) plots the true (red) and estimated value (blue) of the base marginal return for tenure 1 by quantile of quantity, with error margins indicating ± 1.96 variation around the estimated mean from 300 simulations.

and the base marginal return, again in blue the estimated results and in red the true values. Both cases indicate good fit. Subfigure (a) shows the estimated $\hat{\theta}$ in blue and true θ in red by quantile. Dispersion at the 95 percent level are included for all except the top 2 quantiles, as they start to diverge. Overall, the figure shows a good fit, with most sections of including the true θ within their dispersion.

Next, I show the tenure 1's results estimates. Recall that the first-order condition of the seller now includes a backward-looking variable $1 - \Gamma_0(\theta)$ that keeps track of whether the limited commitment constraint was binding in the past. This variable is used by seller as a promise-keeping constraint that guarantees the seller delivers higher quantities and return in the future to prevent buyers from defaulting in the past. In my estimation, I use the tenure 0's predicted $\widehat{\Gamma}_0(\theta(\alpha))$ for each quantile α . Online Appendix Figure OA-5 shows the estimated Gamma distribution and the base marginal return. Although the fit is worse than in tenure 0, the dispersion of both gamma and the base marginal return include tend to include their true values.

With respect to the differences between true and estimated functions, I find that the slight upward bias in the Gamma function for tenure 1 disappears if I use the true $\Gamma_0(\theta)$ function instead of the estimated $\widehat{\Gamma}_0$, suggesting that the bias is generated by sampling error in the tenure 0 estimates. Moreover, differences in the base marginal return for both tenure 0 and tenure 1 come from approximating the tariff function as log-linear. In the Monte-Carlo, the change in unit price is very steep for low-types, and this generates some approximation error for low-types in terms of the base marginal return function. Despite this error, the coefficient of the base return function is correctly estimated when using the assumed parametrization, observations of quantity, and the nonparametric estimates of $v'(\cdot)$ as target. In particular, the estimated coefficient cannot be rejected to be different from 0.5 (the assumed value in simulation).

OA-8 Additional Estimation Results and Model Fit

OA-8.1 Distribution of t-Statistics against Standard Model Null

Online Appendix Table OA-8.1 show the distribution of t-statistics for tests against a standard model null.

Table OA-9: Distribution of t-Statistics

	p10	p25	p50	p75	p90
Tenure 0	0.31	4.64	11.55	30.08	109.27

Notes: This table reports distribution of t-statistics for tests against a standard model null (e.g., $\Gamma_0(\cdot) = 1$).

OA-8.2 Parametrization of the Base Return Function

To conduct counterfactual experiments that consider quantities beyond those observed in the data, I parametrize the seller-specific buyer's return function $v(q) = kq^\beta$ for $k > 0$ and $\beta \in (0, 1)$. This return function satisfies modelling assumptions $v'(\cdot) > 0$ and $v''(\cdot) < 0$. To estimate parameters, I consider tenure 0 transactions between buyer i and seller j at year t and perform the following uniform least squares regression:

$$\ln(\hat{v}'_{ijt}) = \ln(k) + \ln(\beta) + (\beta - 1)\ln(q_{ijt}) + \varepsilon_{ijt},$$

using $v'(q) = k\beta q^{\beta-1}$, the estimated base marginal returns \hat{v}'_{ijt} and under the assumption that ε_{ijt} is Gaussian error. Online Appendix Table OA-10 present the distribution of k and β .

Table OA-10: Parameters of Return Function

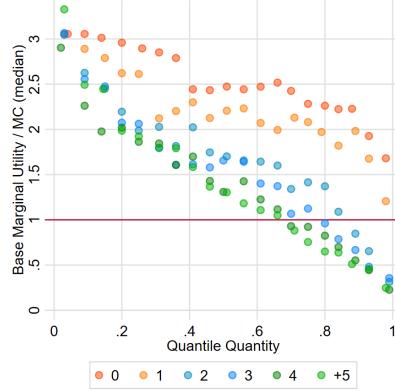
	mean	p10	p25	p50	p75	p90
β	0.56	0.30	0.48	0.61	0.76	0.82
k	171.23	9.00	17.24	39.64	86.61	282.40

Notes: This table reports distribution of estimated values for the ex-post parametrization of the return function.

OA-8.3 Economic Magnitudes: Base Marginal Return

Online Appendix Figure OA-6 presents a binscatter of the ratio marginal revenue product (base marginal return) over marginal costs against the quantile of quantity, across sellers for tenure 0. It shows that the return of the input for the buyer is greater than the private marginal cost of providing it for the seller, for a majority of the buyers. For instance, the median buyer obtains 1.5 dollars of revenue for each dollar spent by the seller to produce the product.

Figure OA-6: Base Marginal Return over Marginal Costs

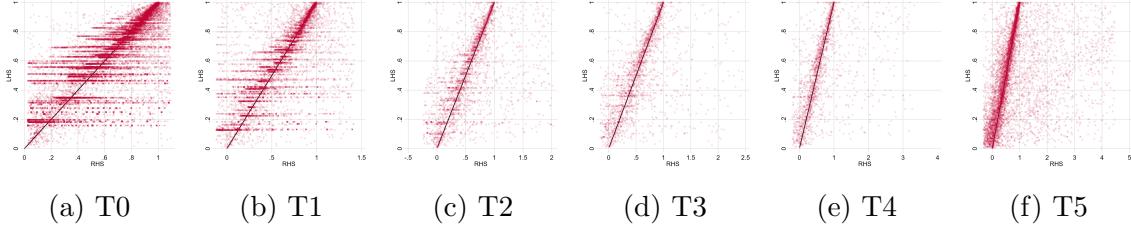


Notes: This figure plots the median of the ratio of base marginal return by marginal costs across sellers by quantile of quantity for each tenure.

OA-8.4 Model Fit

Online Appendix Figure OA-7 presents the statistical fit of the model across tenures. It plots a reordered equation I:EQ's left-hand side on the X-axis and the model's prediction using estimated coefficients of the right-hand side on the Y-axis.³⁶ Fit generally worsens for higher tenures; the results from Monte Carlo studies in Online Appendix OA-7 suggest that the decrease in statistical fit is driven by noise from using estimates for limited enforcement multipliers $\Gamma_s(\cdot)$ for earlier tenures s .

Figure OA-7: Model Fit - Statistical



Notes: These figures show bincscatters of statistical fit of the model across tenures as implied by identification equation I:EQ. On the X-axis, it shows the predicted cumulative distribution function for the observation while on the Y-axis it plots the observed value.

Online Appendix Figure OA-8 shows the fit in terms of quantities. To obtain quantities, I use the parametrization $v(q) = kq^\beta$, for $k > 0$ and $\beta \in (0, 1)$ and the closed-form formula in Q-CES.

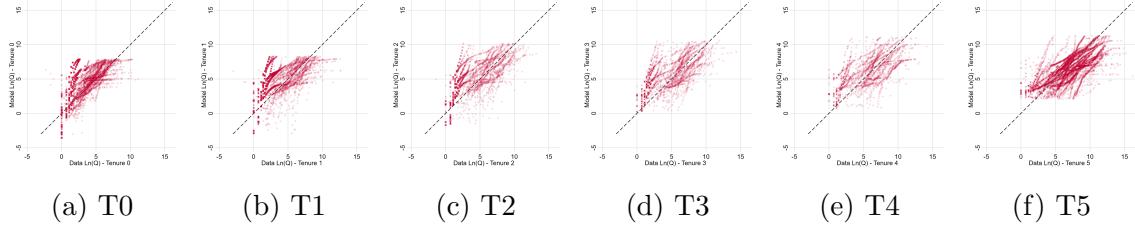
Online Appendix Figure OA-9 shows the fit of tariffs. To generate tariffs in the model, I use the empirical equivalent of equation t-RULE.

³⁶Reorder equation I:EQ to obtain:

$$\alpha = \Gamma_\tau(\alpha) - \sum_{s=0}^{\tau-1} (1 - \Gamma_s(\alpha)) - \left[\frac{T'_\tau(q_\tau(\alpha)) - c_\tau}{T'_\tau(q_\tau(\alpha))} - \gamma_\tau(\alpha) \right] \frac{\theta_\tau(\alpha)}{\theta'_\tau(\alpha)},$$

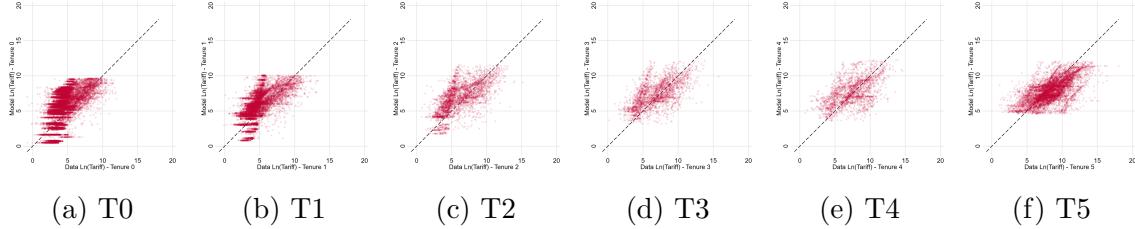
and use the estimated analogues of the right-hand side to make the predictions.

Figure OA-8: Model Fit - Quantities



Notes: These figures display binscatters of model fit according to quantities. Estimated quantities use the close-form formula under the CES parametrization of the return function, as discussed in Appendix OA-10. The X-axis plots the observed (log) quantities and Y-axis model predicted (log) quantities.

Figure OA-9: Model Fit - Tariffs



Notes: These figures display binscatters of model fit according to tariffs. Estimated tariffs are generated by using the empirical analogue of the transfer rule **t-RULE**, taking as inputs estimated parameters θ , the parametrized return function $v(\cdot)$, and model generated quantities. The X-axis plots the observed (log) tariffs and Y-axis model predicted (log) tariffs.

OA-9 Additional Counterfactual Results

This subsection presents comparisons of different counterfactual models relative to baseline nonlinear pricing regime with limited enforcement. The tables present the share of observations in each percentile group for which each reported category (e.g., buyer's net return) is greater under the baseline than under the alternative. Online Appendix Table OA-11 shows all the results.

The main takeaways are the following.

Buyers. Small-quantity buyers tend to prefer limited enforcement of contracts over perfect enforcement. They can effectively use the threat of default to reap higher returns. Instead, the median and top buyers prefer perfect enforcement in the short-term but limited enforcement in the long-term.

Under weak enforcement of contract, buyers prefer price discrimination than uniform pricing, as otherwise they would be excluded from trade (only median and top buyers prefer uniform pricing in the long-term). However, if exclusion and default are restricted, most buyers prefer uniform pricing.

Sellers. They prefer limited enforcement in the short-term but perfect enforcement in the long-term. Under weak enforcement of contract, they enjoy the ability to price discriminate, as it allows them to sell to buyers that would otherwise be excluded from trade. Instead, if enforcement is strong, they prefer uniform pricing in the short-term but price discrimination in the long-term. This is driven by the rapid increase in quantities, despite the decrease in unit prices offer to most buyers as incentive not to default.

Table OA-11: Counterfactual Policies

		Nonlinear + Perfect					Uniform + Limited					Uniform + Perfect				
		10%	25%	50%	75%	100%	10%	25%	50%	75%	100%	10%	25%	50%	75%	100%
Buyer Return	Tenure 0	24.2	24.2	10.9	4.8	6.8	97.3	96.5	96.0	94.3	91.6	0.1	0.2	0.6	6.9	40.6
	Tenure 1	68.3	55.3	23.0	9.4	11.8	94.6	92.2	88.6	88.1	87.4	0.1	0.1	0.2	14.0	55.9
	Tenure 2	62.5	45.3	30.6	25.9	28.4	82.1	79.2	70.2	66.5	63.2	0.9	0.4	0.9	10.0	31.8
	Tenure 3	65.3	59.9	40.2	32.3	38.0	76.9	71.6	58.3	54.4	55.9	3.1	0.8	1.6	11.2	28.3
	Tenure 4	63.5	43.4	36.9	33.9	44.2	74.6	62.1	44.8	41.7	39.8	5.4	1.2	4.9	8.7	17.4
	Tenure 5	52.7	59.6	65.7	60.8	68.9	65.5	53.7	37.3	33.2	30.9	0.7	1.6	2.8	7.0	19.9
Seller Profit	Tenure 0	34.1	41.6	88.3	95.0	93.1	92.7	92.6	96.4	98.0	98.8	7.1	7.4	11.1	35.0	47.9
	Tenure 1	53.8	55.0	83.3	90.6	88.2	99.1	96.7	94.8	97.1	90.5	28.4	19.1	29.8	45.9	51.3
	Tenure 2	49.0	50.4	72.0	74.1	71.6	95.0	97.0	98.2	99.5	97.6	34.3	35.1	50.9	70.2	85.7
	Tenure 3	45.2	48.1	61.7	68.1	62.0	96.5	99.0	97.0	99.3	93.7	50.2	50.0	61.9	76.9	86.6
	Tenure 4	48.6	53.2	65.3	66.1	55.8	92.0	98.5	95.1	95.2	94.5	54.0	64.2	71.5	86.5	93.5
	Tenure 5	61.1	44.5	37.7	39.3	31.1	94.0	87.8	95.1	97.4	95.8	63.2	66.3	81.9	92.8	94.6
Surplus	Tenure 0	18.6	18.9	9.0	3.8	2.6	98.4	98.1	98.8	98.5	99.5	3.8	4.1	5.2	12.0	65.0
	Tenure 1	40.5	41.7	30.3	12.6	29.6	97.5	96.2	97.3	99.2	100.0	6.0	7.4	11.0	31.9	76.1
	Tenure 2	47.2	50.7	48.3	63.2	72.6	90.0	92.2	91.7	98.6	99.7	17.0	18.8	27.6	57.0	95.0
	Tenure 3	60.9	57.6	69.7	76.6	70.1	90.6	92.5	88.1	98.4	99.6	24.3	26.0	37.4	69.1	98.4
	Tenure 4	66.7	71.9	74.5	77.1	67.3	86.2	89.6	85.9	98.3	99.5	28.1	40.1	53.8	79.2	97.9
	Tenure 5	74.2	79.6	88.6	91.4	84.8	82.0	78.3	80.3	96.9	100.0	30.5	34.2	53.7	86.1	99.9
Unit Prices	Tenure 0	76.3	75.7	89.1	94.6	93.1	93.6	93.1	95.4	90.3	43.6	93.6	93.1	95.4	90.3	43.6
	Tenure 1	56.0	55.7	77.8	90.5	88.1	98.6	96.8	87.9	68.2	25.0	98.6	96.8	87.9	68.2	25.0
	Tenure 2	44.0	56.0	68.3	74.0	71.7	92.6	94.8	90.9	64.0	18.3	92.3	94.7	90.9	64.0	18.3
	Tenure 3	37.3	41.5	58.8	67.9	61.8	91.2	96.7	89.2	55.1	13.3	90.8	96.7	89.2	55.1	13.3
	Tenure 4	38.2	56.4	63.1	67.6	55.6	88.3	95.5	88.0	65.3	20.0	87.7	95.5	88.0	65.2	20.0
	Tenure 5	43.5	38.2	35.9	38.2	31.1	90.6	91.3	87.7	53.5	10.5	89.8	91.1	87.6	53.2	10.0
% Excluded	Tenure 0	-	-	-	-	-	97.3	96.4	95.8	94.1	90.4	-	-	-	-	-
	Tenure 1	-	-	-	-	-	93.4	91.9	88.6	87.3	85.8	-	-	-	-	-
	Tenure 2	-	-	-	-	-	79.8	77.4	70.0	65.6	61.3	-	-	-	-	-
	Tenure 3	-	-	-	-	-	74.4	69.2	58.0	51.4	50.0	-	-	-	-	-
	Tenure 4	-	-	-	-	-	71.3	60.1	44.7	38.5	37.4	-	-	-	-	-
	Tenure 5	-	-	-	-	-	62.3	50.7	36.4	30.3	25.8	-	-	-	-	-

Notes: This table reports the % share of observations for which the reported category (e.g., Buyer's Net Return) is greater under the observed nonlinear pricing regime than under the alternative policy. The values are reported across different tenures and percentiles in the distribution of types. The policies considered are 1) Nonlinear pricing with perfect enforcement, 2) Uniform monopolist pricing with limited enforcement, and 3) Uniform monopolist pricing with perfect enforcement. The reported categories are Buyer's Net Return, Seller's Profits, Total Surplus, Unit Prices, and percentage of Excluded Buyers.

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