

### Problem 1

$$H = X(X^T X)^{-1} X^T \quad \text{so} \quad \hat{Y} = X\hat{\beta} = HY$$

$$\begin{aligned} E(SS_E) &= E[Y^T (I-H)Y] = E[(X\beta + \epsilon)^T (I-H)(X\beta + \epsilon)] \\ &= \sigma^2 E[\text{trace}(I-H)] = \sigma^2 (n-p) \end{aligned}$$

### Problem 2

$$\begin{aligned} \textcircled{1} \quad \hat{Y} &= \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + \beta_5 X_5 + \beta_6 X_6 \\ &= 10.8 + .61X_1 - .077X_2 + .32X_3 + .082X_4 + .038X_5 - .21X_6 \end{aligned}$$

$$\textcircled{2} \quad \sum y_i = \sum \hat{y}_i = 1939$$

$\textcircled{3}$  If the model includes the intercept and we let

$$Q(\beta_0, \beta_1, \dots, \beta_p) = \sum (y_i - \hat{y}_i)^2 = \sum (y_i - \beta_0 - \beta_1 x_{i1} - \dots - \beta_p x_{ip})^2$$

then minimizing least squares guarantees

$$\frac{dQ}{d\beta_0} = 0 = 2 \sum (y_i - \hat{y}_i) \rightarrow \sum e_i = 0$$

$\textcircled{4}$  Classic Method gives  $\hat{\beta}_3 = .43$

Page 63 method also gives  $\hat{\beta}_4 = .43$  since:

$$\text{step 1: } \hat{Y} = 20 + .69X_4$$

residual is part of  $Y$  not explained by  $X_4$

$$\text{step 2: } \hat{X}_3 = 9.6 + .72X_4$$

residual is part of  $X_3$  not explained by  $X_4$

step 3: gives  $\hat{\beta}_3 = .43$ , which is the same since we are measuring the effect of  $X_3$  without  $X_4$  on  $Y$  without  $X_4$

### Problem 3

(a)

$$\hat{F} = \hat{\beta}_0 + \hat{\beta}_1 P_1 = -22.3 + 1.26 P_1 \quad (\text{Model I})$$

$$\hat{F} = \hat{\beta}_0 + \hat{\beta}_2 P_2 = -1.8 + 1.0 P_2 \quad (\text{Model II})$$

$$\hat{F} = \hat{\beta}_0 + \hat{\beta}_1 P_1 + \hat{\beta}_2 P_2 = -14.5 + .49 P_1 + .67 P_2 \quad (\text{Model III})$$

(b) Model I:  $|t^*| = 1.93 < 2.086 = t_{\alpha/2, n-2} \Rightarrow$  cannot reject

Model II:  $|t^*| = 0.24 < 2.086 = t_{\alpha/2, n-2} \Rightarrow$  cannot reject

Model III:  $|t^*| = 1.57 < 2.093 = t_{\alpha/2, n-3} \Rightarrow$  cannot reject

(c) Model II, which has predictor  $P_2$ , is better since it has a smaller SSE, or equivalently a larger  $R^2$ .

(d) Model III has the lowest SSE / highest  $R^2$  and predicts

$$\hat{F} = -14.5 + .49(78) + .67(85) = 80.7$$

### Problem 4

Note  $\sum (x_i - \bar{x})^2 = \sum (x_i^2 - 2\bar{x}x_i + \bar{x}^2) = \sum x_i^2 - 2n\bar{x}^2 + n\bar{x}^2 = \sum x_i^2 - n\bar{x}^2$

and  $\sum (x_i - \bar{x})(y_i - \bar{y}) = \sum (x_i y_i - \bar{x} y_i - \bar{y} x_i + \bar{x} \bar{y})$

$$= \sum x_i y_i - n\bar{x}\bar{y} - n\bar{x}\bar{y} + n\bar{x}\bar{y}$$

$$= \sum x_i y_i - n\bar{x}\bar{y}$$

$\therefore \hat{\theta} = \begin{pmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \end{pmatrix} = (X^T X)^{-1} X^T Y$  for  $X = \begin{bmatrix} 1 & x_{11} & \dots & x_{1p} \\ 1 & x_{21} & \dots & x_{2p} \\ \vdots & \vdots & & \vdots \\ 1 & x_{n1} & \dots & x_{np} \end{bmatrix}$

$$= \begin{pmatrix} n & \sum x_i \\ \sum x_i & \sum x_i^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum y_i \\ \sum x_i y_i \end{pmatrix}$$

$$= \frac{1}{n \sum x_i^2 - (\sum x_i)^2} \begin{pmatrix} \sum x_i^2 & -\sum x_i \\ -\sum x_i & n \end{pmatrix} \begin{pmatrix} \sum y_i \\ \sum x_i y_i \end{pmatrix}$$

$$= \frac{1/n}{\sum (x_i - \bar{x})^2} \begin{pmatrix} \sum x_i^2 & -n\bar{x} \\ -n\bar{x} & n \end{pmatrix} \begin{pmatrix} n\bar{y} \\ \sum x_i y_i \end{pmatrix}$$

$$\text{So } \hat{\beta}_1 = \frac{\sum x_i y_i - n\bar{x}\bar{y}}{\sum (x_i - \bar{x})^2} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} \quad \text{and}$$

$$\bar{y} - \hat{\beta}_1 \bar{x} = \frac{\bar{y} \sum (x_i - \bar{x})^2 - \bar{x} (\sum x_i y_i - n\bar{x}\bar{y})}{\sum (x_i - \bar{x})^2}$$

$$= \frac{\bar{y} \sum x_i^2 - n\bar{y}\bar{x}^2 - \bar{x} \sum x_i y_i + n\bar{x}^2 \bar{y}}{\sum (x_i - \bar{x})^2}$$

$$= \frac{\bar{y} \sum x_i^2 - \bar{x} \sum x_i y_i}{\sum (x_i - \bar{x})^2}$$

which we see equals  $\beta_0$ , the first element in the  $\beta$  vector.

$$\begin{aligned} \text{Now, } \text{Var}(\hat{\beta}) &= \text{Var}((X^T X)^{-1} X^T Y) = (X^T X)^{-1} X^T \text{Var}(Y) X (X^T X)^{-1} \\ &= \sigma^2 (X^T X)^{-1} \quad \text{for } \text{Var}(Y) = \sigma^2 \end{aligned}$$

For SLR:

$$\begin{pmatrix} \text{Var}(\hat{\beta}_0) & \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) \\ \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) & \text{Var}(\hat{\beta}_1) \end{pmatrix} = \frac{\sigma^2/n}{\sum (x_i - \bar{x})^2} \begin{pmatrix} \sum x_i^2 & -n\bar{x} \\ -n\bar{x} & n \end{pmatrix}$$

$$\text{So } \text{Var}(\hat{\beta}_1) = \frac{\sigma^2}{\sum (x_i - \bar{x})^2} \quad \text{and}$$

$$\begin{aligned} \text{Var}(\hat{\beta}_0) &= \frac{\sigma^2}{n} \cdot \frac{\sum x_i^2}{\sum (x_i - \bar{x})^2} = \frac{\sigma^2}{n} \cdot \frac{\sum x_i^2}{\sum x_i^2 - n\bar{x}^2} = \frac{\sigma^2}{n} \cdot \left( 1 + \frac{n\bar{x}^2}{\sum x_i^2 - n\bar{x}^2} \right) \\ &= \sigma^2 \left( \frac{1}{n} + \frac{\bar{x}^2}{\sum x_i^2 - n\bar{x}^2} \right) \end{aligned}$$

Problem 5

$$\begin{aligned} F^* &= \left[ \frac{SSE(R) - SSE(F)}{df_R - df_F} \right] / \left[ \frac{SSE(F)}{df_F} \right] \\ &= \left[ \frac{SSE(R) - SSE(F)}{3} \right] / \left[ \frac{SSE(F)}{n-4-1} \right] \\ &\approx 20.5 \end{aligned}$$

$F_{\alpha, 3, n-5} = 2.71$  and we reject since  $F^* > F_{\alpha}$