

Scaling and Multiscaling in Financial Indexes: a Simple Model

Francesco Caravenna

Università degli Studi di Padova

joint work with Alessandro Andreoli (Padova),
Paolo Dai Pra (Padova) and Gustavo Posta (Politecnico di Milano)

Nantes ~ June 7, 2010

Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
4. Main Results
5. Estimation and Simulations
6. Conclusions

Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
4. Main Results
5. Estimation and Simulations
6. Conclusions

Black & Scholes Model

Black & Scholes model for the price S_t of a stock price or index:

$$dS_t = S_t (r dt + \sigma dW_t)$$

- ▶ σ (the **volatility**) and r (the **interest rate**) are constant
- ▶ $(W_t)_{t \geq 0}$ is a standard Brownian motion.

Black & Scholes Model

Black & Scholes model for the price S_t of a stock price or index:

$$dS_t = S_t (r dt + \sigma dW_t)$$

- ▶ σ (the **volatility**) and r (the **interest rate**) are constant
- ▶ $(W_t)_{t \geq 0}$ is a standard Brownian motion.

Therefore $(S_t)_{t \geq 0}$ is a **geometric Brownian motion**.

Black & Scholes Model

Black & Scholes model for the price S_t of a stock price or index:

$$dS_t = S_t (r dt + \sigma dW_t)$$

- ▶ σ (the **volatility**) and r (the **interest rate**) are constant
- ▶ $(W_t)_{t \geq 0}$ is a standard Brownian motion.

Therefore $(S_t)_{t \geq 0}$ is a **geometric Brownian motion**.

The **detrended log-price** $X_t := \log S_t - r't$, with $r' := r - \sigma^2/2$, is then a standard Brownian motion

$$dX_t = \sigma dW_t \quad \implies \quad X_t = X_0 + \sigma W_t.$$

Black & Scholes Model

Black & Scholes model for the price S_t of a stock price or index:

$$dS_t = S_t (r dt + \sigma dW_t)$$

- ▶ σ (the **volatility**) and r (the **interest rate**) are constant
- ▶ $(W_t)_{t \geq 0}$ is a standard Brownian motion.

Therefore $(S_t)_{t \geq 0}$ is a **geometric Brownian motion**.

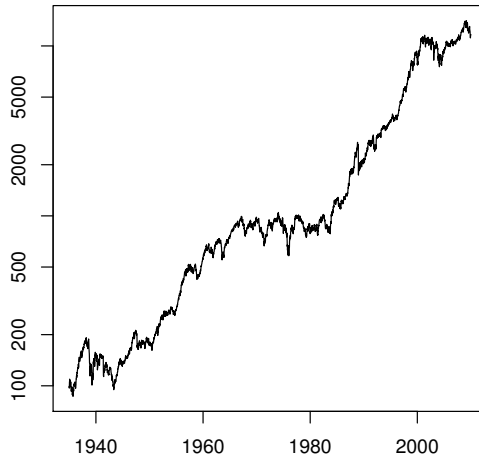
The **detrended log-price** $X_t := \log S_t - r't$, with $r' := r - \sigma^2/2$, is then a standard Brownian motion

$$dX_t = \sigma dW_t \quad \implies \quad X_t = X_0 + \sigma W_t.$$

Basic example: **Dow Jones Industrial Average (DJIA)**.

DJIA Time Series (1935-2009)

Exponential growth of the DJIA [log plot]:



Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
4. Main Results
5. Estimation and Simulations
6. Conclusions

Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

Beyond the Black & Scholes Model

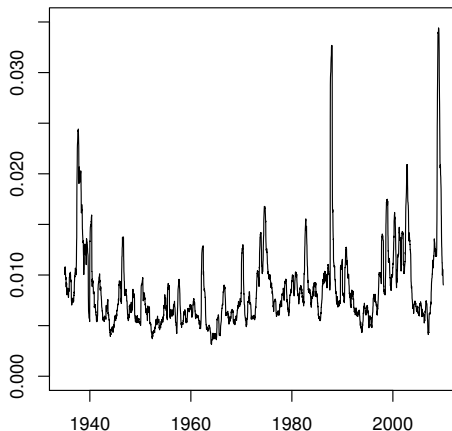
Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).

DJIA Time Series (1935-2009)

Empirical volatility



Local standard deviation of log-returns in a window of 100 days

Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).

Beyond the Black & Scholes Model

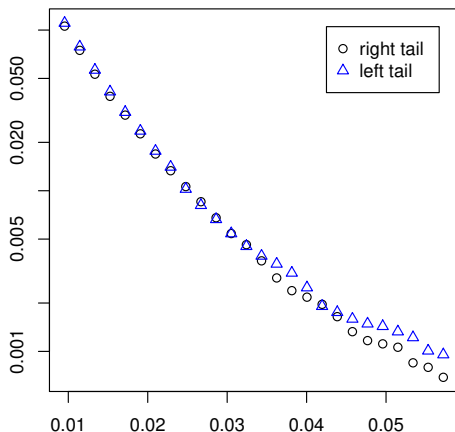
Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).
- ▶ The increments $(X_{t+h} - X_t)$, called **log-returns**, have a distribution with tails **heavier than Gaussian**.

DJIA Time Series (1935-2009)

Tails of daily log-return distribution [log plot]



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: 1 to 6 st. dev.

Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).
- ▶ The increments $(X_{t+h} - X_t)$, called **log-returns**, have a distribution with tails **heavier than Gaussian**.

Beyond the Black & Scholes Model

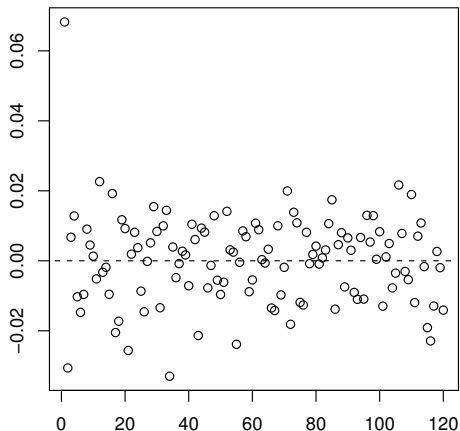
Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).
- ▶ The increments $(X_{t+h} - X_t)$, called **log-returns**, have a distribution with tails **heavier than Gaussian**.
- ▶ Log-returns corresponding to disjoint time intervals are **uncorrelated**...

DJIA Time Series (1935-2009)

Decorrelation of daily log-returns over 1–120 days



Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).
- ▶ The increments $(X_{t+h} - X_t)$, called **log-returns**, have a distribution with tails **heavier than Gaussian**.
- ▶ Log-returns corresponding to disjoint time intervals are **uncorrelated**...

Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).
- ▶ The increments $(X_{t+h} - X_t)$, called **log-returns**, have a distribution with tails **heavier than Gaussian**.
- ▶ Log-returns corresponding to disjoint time intervals are **uncorrelated**... but **not independent**!

Beyond the Black & Scholes Model

Despite its success, this model is not consistent with a number of **stylized facts** that are empirically detected in many real time series.

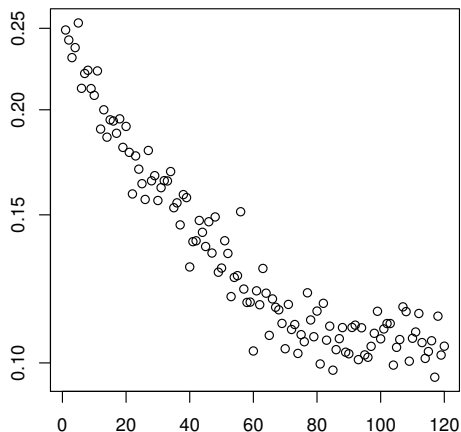
Detrended log-price $X_t := \log S_t - \bar{d}_t$ B&S: $dX_t = \sigma dW_t$

- ▶ The **volatility σ is not constant**: it may have high peaks (“**shocks**” in the market).
- ▶ The increments $(X_{t+h} - X_t)$, called **log-returns**, have a distribution with tails **heavier than Gaussian**.
- ▶ Log-returns corresponding to disjoint time intervals are **uncorrelated**... but **not independent**!

The correlation between $|X_{t+h} - X_t|$ and $|X_{s+h} - X_s|$, called **volatility autocorrelation**, has a **slow decay** in $|t - s|$, up to moderate values of $|t - s|$ (**clustering of volatility**).

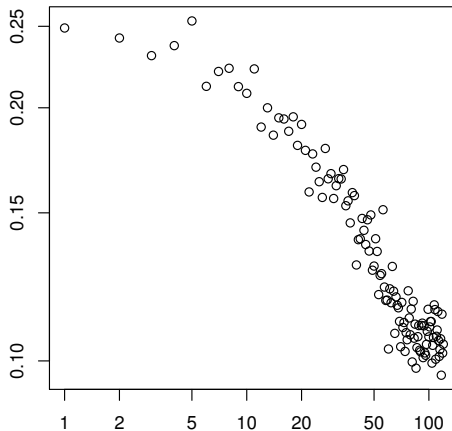
DJIA Time Series (1935-2009)

Volatility autocorrelation over 1–120 days [log plot]



DJIA Time Series (1935-2009)

Volatility autocorrelation over 1–120 days [log-log plot]



Some Alternative Models

Continuous time: the constant volatility σ is replaced by a stochastic process $(\sigma_t)_{t \geq 0} \longrightarrow$ **stochastic volatility model**

$$dX_t = \sigma_t dW_t$$

Some Alternative Models

Continuous time: the constant volatility σ is replaced by a stochastic process $(\sigma_t)_{t \geq 0} \longrightarrow$ **stochastic volatility model**

$$dX_t = \sigma_t dW_t$$

Example: generalized Ornstein-Uhlenbeck processes

$$d\sigma_t^2 = -\alpha \sigma_t^2 dt + dL_t,$$

where L_t is a **subordinator** (increasing Lévy process).

Some Alternative Models

Continuous time: the constant volatility σ is replaced by a stochastic process $(\sigma_t)_{t \geq 0} \rightarrow$ **stochastic volatility model**

$$dX_t = \sigma_t dW_t$$

Example: generalized Ornstein-Uhlenbeck processes

$$d\sigma_t^2 = -\alpha \sigma_t^2 dt + dL_t,$$

where L_t is a **subordinator** (increasing Lévy process).

Discrete time: autoregressive models such as the **GARCH** are used:

$$\varepsilon_t = \sigma_t z_t, \quad \sigma_t^2 = \omega + \beta \sigma_{t-1}^2 + \alpha \varepsilon_{t-1}^2$$

where $\varepsilon_t := X_{t+1} - X_t$ and $(z_t)_{t \in \mathbb{N}}$ are i.i.d. $N(0, 1)$.

Scaling Properties

More recently, some striking **scaling properties** of stock indexes of developed markets have been emphasized.

[Di Matteo, Aste & Dacorogna 2005] [Baldovin & Stella 2007 - 08]

Scaling Properties

More recently, some striking **scaling properties** of stock indexes of developed markets have been emphasized.

[Di Matteo, Aste & Dacorogna 2005] [Baldovin & Stella 2007 - 08]

- Diffusive scaling of log-regurns

Scaling Properties

More recently, some striking **scaling properties** of stock indexes of developed markets have been emphasized.

[Di Matteo, Aste & Dacorogna 2005] [Baldovin & Stella 2007 - 08]

- ▶ Diffusive scaling of log-regurns
- ▶ Multiscaling (or anomalous scaling) of moments

Diffusive Scaling of Log-Returns

Denote by $\hat{p}_h(\cdot)$ the **empirical distribution of the log-return** over h days, for an observed time series $(x_t)_{1 \leq t \leq T}$ of the detrended log-index X :

$$\hat{p}_h(\cdot) := \frac{1}{T-h} \sum_{t=1}^{T-h} \delta_{x_{t+h}-x_t}(\cdot),$$

where $\delta_x(\cdot)$ denotes the Dirac measure at $x \in \mathbb{R}$

Diffusive Scaling of Log>Returns

Denote by $\hat{p}_h(\cdot)$ the **empirical distribution of the log-return** over h days, for an observed time series $(x_t)_{1 \leq t \leq T}$ of the detrended log-index X :

$$\hat{p}_h(\cdot) := \frac{1}{T-h} \sum_{t=1}^{T-h} \delta_{x_{t+h}-x_t}(\cdot),$$

where $\delta_x(\cdot)$ denotes the Dirac measure at $x \in \mathbb{R}$

For various indexes (such as the DJIA) and for h within a suitable time scale, \hat{p}_h obeys approximately a **diffusive scaling relation**:

$$X_{t+h} - X_t \stackrel{d}{\approx} \sqrt{h}(X_{t+1} - X_t)$$

Diffusive Scaling of Log>Returns

Denote by $\hat{p}_h(\cdot)$ the **empirical distribution of the log-return** over h days, for an observed time series $(x_t)_{1 \leq t \leq T}$ of the detrended log-index X :

$$\hat{p}_h(\cdot) := \frac{1}{T-h} \sum_{t=1}^{T-h} \delta_{x_{t+h}-x_t}(\cdot),$$

where $\delta_x(\cdot)$ denotes the Dirac measure at $x \in \mathbb{R}$

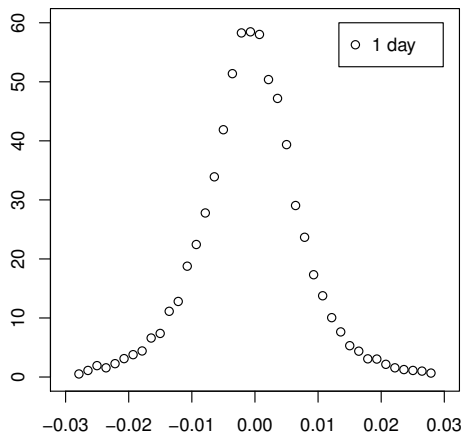
For various indexes (such as the DJIA) and for h within a suitable time scale, \hat{p}_h obeys approximately a **diffusive scaling relation**:

$$X_{t+h} - X_t \stackrel{d}{\approx} \sqrt{h} (X_{t+1} - X_t) \quad \rightarrow \quad \hat{p}_h(dr) \simeq \frac{1}{\sqrt{h}} g\left(\frac{r}{\sqrt{h}}\right) dr$$

where g is a **non-Gaussian** density.

DJIA Time Series (1935-2009)

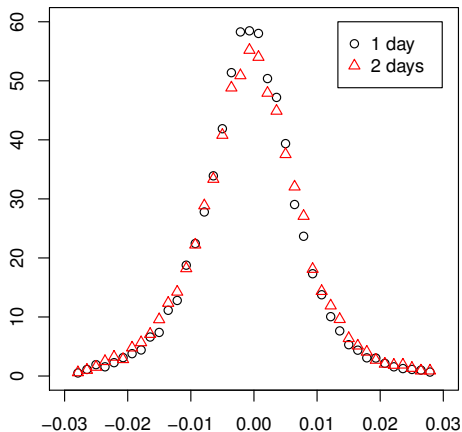
Rescaled empirical density of log-returns (1 day)



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to +3 st. dev.

DJIA Time Series (1935-2009)

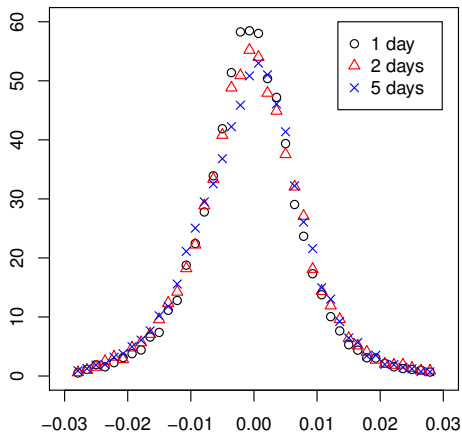
Rescaled empirical density of log-returns (1-2 days)



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to +3 st. dev.

DJIA Time Series (1935-2009)

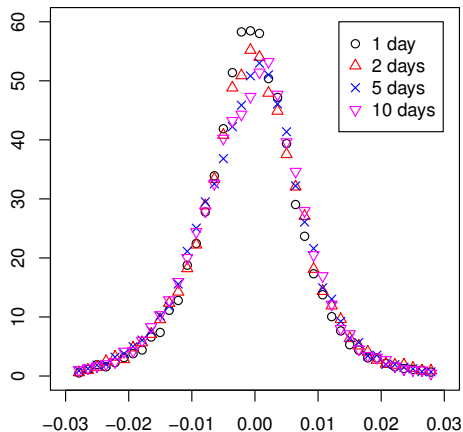
Rescaled empirical density of log-returns (1-2-5 days)



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to +3 st. dev.

DJIA Time Series (1935-2009)

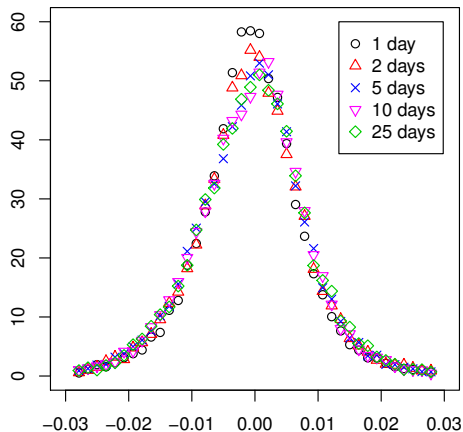
Rescaled empirical density of log-returns (1-2-5-10 days)



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to +3 st. dev.

DJIA Time Series (1935-2009)

Rescaled empirical density of log-returns (1-2-5-10-25 days)



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to +3 st. dev.

Multiscaling of Moments

Consider the empirical q -th moment of the log-return over h days:

$$\hat{m}_q(h) := \frac{1}{T-h} \sum_{i=1}^{T-h} |x_{i+h} - x_i|^q = \int |r|^q \hat{p}_h(dr)$$

Multiscaling of Moments

Consider the empirical q -th moment of the log-return over h days:

$$\hat{m}_q(h) := \frac{1}{T-h} \sum_{i=1}^{T-h} |x_{i+h} - x_i|^q = \int |r|^q \hat{p}_h(dr)$$

From the **diffusive scaling** $X_{t+h} - X_t \stackrel{d}{\approx} \sqrt{h}(X_{t+1} - X_t)$ it is natural to guess

$$\hat{m}_q(h) \approx h^{q/2} \quad \text{for } h \text{ small.}$$

Multiscaling of Moments

Consider the empirical q -th moment of the log-return over h days:

$$\hat{m}_q(h) := \frac{1}{T-h} \sum_{i=1}^{T-h} |x_{i+h} - x_i|^q = \int |r|^q \hat{p}_h(dr)$$

From the **diffusive scaling** $X_{t+h} - X_t \stackrel{d}{\approx} \sqrt{h}(X_{t+1} - X_t)$ it is natural to guess

$$\hat{m}_q(h) \approx h^{q/2} \quad \text{for } h \text{ small.}$$

This is true **only for $q \leq q^*$** (with $q^* \simeq 3$ for the DJIA).

Multiscaling of Moments

Consider the empirical q -th moment of the log-return over h days:

$$\hat{m}_q(h) := \frac{1}{T-h} \sum_{i=1}^{T-h} |x_{i+h} - x_i|^q = \int |r|^q \hat{p}_h(dr)$$

From the **diffusive scaling** $X_{t+h} - X_t \stackrel{d}{\approx} \sqrt{h}(X_{t+1} - X_t)$ it is natural to guess

$$\hat{m}_q(h) \approx h^{q/2} \quad \text{for } h \text{ small.}$$

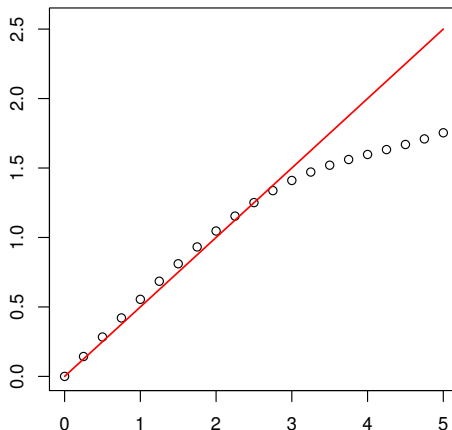
This is true **only for** $q \leq q^*$ (with $q^* \simeq 3$ for the DJIA).

For $q > q^*$ we have the **anomalous scaling** (or **multiscaling**)

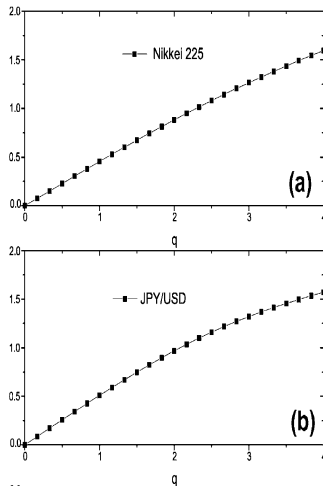
$$\hat{m}_q(h) \approx h^{A(q)}, \quad \text{with } A(q) < \frac{q}{2}.$$

DJIA Time Series (1935-2009)

Scaling exponent $A(q)$ (linear regression of $\log \hat{m}_q(h)$ vs. $\log h$)



Other Data Series (from [Di Matteo, Aste & Dacorogna, 2005])



Scaling Properties

It is nontrivial to identify a model in the literature which fits well all mentioned stylized facts.

Scaling Properties

It is nontrivial to identify a model in the literature which fits well
all mentioned stylized facts.

E.g., no multiscaling of moments is observed in GARCH.

Scaling Properties

It is nontrivial to identify a model in the literature which fits well [all mentioned stylized facts](#).

E.g., no multiscaling of moments is observed in GARCH.

Baldovin & Stella standpoint: [scaling properties](#) should primarily guide the construction of the model.

Baldovin & Stella's Model

Empirically:
$$\hat{p}_h(dr) \simeq \frac{1}{\sqrt{h}} g\left(\frac{r}{\sqrt{h}}\right) dr.$$

Assume g is symmetric and let g^* be its Fourier transform.

Baldovin & Stella's Model

Empirically: $\hat{p}_h(dr) \simeq \frac{1}{\sqrt{h}} g\left(\frac{r}{\sqrt{h}}\right) dr.$

Assume g is symmetric and let g^* be its Fourier transform.

Let $(Y_t)_{t \geq 0}$ be the process with finite dimensional densities

$$p(x_1, t_1; x_2, t_2; \dots; x_n, t_n) = h \left(\frac{x_1}{\sqrt{t_1}}, \frac{x_2 - x_1}{\sqrt{t_2 - t_1}}, \dots, \frac{x_n - x_{n-1}}{\sqrt{t_n - t_{n-1}}} \right),$$

Baldovin & Stella's Model

Empirically: $\hat{p}_h(dr) \simeq \frac{1}{\sqrt{h}} g\left(\frac{r}{\sqrt{h}}\right) dr.$

Assume g is symmetric and let g^* be its Fourier transform.

Let $(Y_t)_{t \geq 0}$ be the process with finite dimensional densities

$$p(x_1, t_1; x_2, t_2; \dots; x_n, t_n) = h \left(\frac{x_1}{\sqrt{t_1}}, \frac{x_2 - x_1}{\sqrt{t_2 - t_1}}, \dots, \frac{x_n - x_{n-1}}{\sqrt{t_n - t_{n-1}}} \right),$$

where $h : \mathbb{R}^n \rightarrow \mathbb{R}$ has Fourier transform h^* given by

$$h^*(u_1, u_2, \dots, u_n) := g^*\left(\sqrt{u_1^2 + \dots + u_n^2}\right).$$

Baldovin & Stella's Model

Empirically: $\hat{p}_h(dr) \simeq \frac{1}{\sqrt{h}} g\left(\frac{r}{\sqrt{h}}\right) dr.$

Assume g is symmetric and let g^* be its Fourier transform.

Let $(Y_t)_{t \geq 0}$ be the process with finite dimensional densities

$$p(x_1, t_1; x_2, t_2; \dots; x_n, t_n) = h \left(\frac{x_1}{\sqrt{t_1}}, \frac{x_2 - x_1}{\sqrt{t_2 - t_1}}, \dots, \frac{x_n - x_{n-1}}{\sqrt{t_n - t_{n-1}}} \right),$$

where $h : \mathbb{R}^n \rightarrow \mathbb{R}$ has Fourier transform h^* given by

$$h^*(u_1, u_2, \dots, u_n) := g^*\left(\sqrt{u_1^2 + \dots + u_n^2}\right).$$

- If g is standard Gaussian $\rightarrow (Y_t)_{t \geq 0}$ is Brownian motion.

Baldovin & Stella's Model

Empirically: $\hat{p}_h(dr) \simeq \frac{1}{\sqrt{h}} g\left(\frac{r}{\sqrt{h}}\right) dr.$

Assume g is symmetric and let g^* be its Fourier transform.

Let $(Y_t)_{t \geq 0}$ be the process with finite dimensional densities

$$p(x_1, t_1; x_2, t_2; \dots; x_n, t_n) = h \left(\frac{x_1}{\sqrt{t_1}}, \frac{x_2 - x_1}{\sqrt{t_2 - t_1}}, \dots, \frac{x_n - x_{n-1}}{\sqrt{t_n - t_{n-1}}} \right),$$

where $h : \mathbb{R}^n \rightarrow \mathbb{R}$ has Fourier transform h^* given by

$$h^*(u_1, u_2, \dots, u_n) := g^*\left(\sqrt{u_1^2 + \dots + u_n^2}\right).$$

- ▶ If g is standard Gaussian $\rightarrow (Y_t)_{t \geq 0}$ is Brownian motion.
- ▶ Is the definition well-posed? Conditions on g .

Baldovin & Stella's Model

- ▶ The increments of Y have **diffusive scaling**.
Their (rescaled) marginal density is $g(\cdot)$.

Baldovin & Stella's Model

- ▶ The increments of Y have **diffusive scaling**.
Their (rescaled) marginal density is $g(\cdot)$.
- ▶ The increments of Y are **uncorrelated** but **not independent**.

Baldovin & Stella's Model

- ▶ The increments of Y have **diffusive scaling**.
Their (rescaled) marginal density is $g(\cdot)$.
- ▶ The increments of Y are **uncorrelated** but **not independent**.
- ▶ However, they are **exchangeable**: **no decay of correlations**.

Baldovin & Stella's Model

- ▶ The increments of Y have **diffusive scaling**.
Their (rescaled) marginal density is $g(\cdot)$.
- ▶ The increments of Y are **uncorrelated** but **not independent**.
- ▶ However, they are **exchangeable**: **no decay of correlations**.

By De Finetti's theorem in continuous time [Freedman 1963] the process $(Y_t)_{t \geq 0}$ is a mixture of Brownian motions:

$$Y_t = \sigma W_t$$

where σ is **random and independent** of the BM $(W_t)_{t \geq 0}$.

Baldovin & Stella's Model

- ▶ The increments of Y have **diffusive scaling**.
Their (rescaled) marginal density is $g(\cdot)$.
- ▶ The increments of Y are **uncorrelated** but **not independent**.
- ▶ However, they are **exchangeable**: **no decay of correlations**.

By De Finetti's theorem in continuous time [Freedman 1963] the process $(Y_t)_{t \geq 0}$ is a mixture of Brownian motions:

$$Y_t = \sigma W_t$$

where σ is **random and independent** of the BM $(W_t)_{t \geq 0}$.

A sample path of $(Y_t)_{t \geq 0}$ **cannot be distinguished** from a sample path of a BM with constant volatility: **no ergodicity**.

Baldovin & Stella's Model

- ▶ The increments of Y have **diffusive scaling**.
Their (rescaled) marginal density is $g(\cdot)$.
- ▶ The increments of Y are **uncorrelated** but **not independent**.
- ▶ However, they are **exchangeable**: **no decay of correlations**.

By De Finetti's theorem in continuous time [Freedman 1963] the process $(Y_t)_{t \geq 0}$ is a mixture of Brownian motions:

$$Y_t = \sigma W_t$$

where σ is **random and independent** of the BM $(W_t)_{t \geq 0}$.

A sample path of $(Y_t)_{t \geq 0}$ **cannot be distinguished** from a sample path of a BM with constant volatility: **no ergodicity**.

Apart from this issue, there is still **no multiscaling** of moments.
This is solved introducing a **time inhomogeneity** in the model.

Baldovin & Stella's Model

Fix a (periodic) sequence of epochs $0 < \tau_1 < \tau_2 < \dots < \tau_n \uparrow +\infty$ and a parameter $0 < D \leq 1/2$. Define a new process $(X_t)_{t \geq 0}$ by

$$X_t := Y_{t^{2D}} \quad \text{for } t \in [0, \tau_1),$$

$$X_t := Y_{(t-\tau_n)^{2D} + \sum_{k=1}^n (\tau_k - \tau_{k-1})^{2D}} \quad \text{for } t \in [\tau_n, \tau_{n+1}).$$

Baldovin & Stella's Model

Fix a (periodic) sequence of epochs $0 < \tau_1 < \tau_2 < \dots < \tau_n \uparrow +\infty$ and a parameter $0 < D \leq 1/2$. Define a new process $(X_t)_{t \geq 0}$ by

$$X_t := Y_{t^{2D}} \quad \text{for } t \in [0, \tau_1),$$

$$X_t := Y_{(t-\tau_n)^{2D} + \sum_{k=1}^n (\tau_k - \tau_{k-1})^{2D}} \quad \text{for } t \in [\tau_n, \tau_{n+1}).$$

- For $D = 1/2$ we have the old model $X_t \equiv Y_t$.

Baldovin & Stella's Model

Fix a (periodic) sequence of epochs $0 < \tau_1 < \tau_2 < \dots < \tau_n \uparrow +\infty$ and a parameter $0 < D \leq 1/2$. Define a new process $(X_t)_{t \geq 0}$ by

$$X_t := Y_{t^{2D}} \quad \text{for } t \in [0, \tau_1),$$

$$X_t := Y_{(t-\tau_n)^{2D} + \sum_{k=1}^n (\tau_k - \tau_{k-1})^{2D}} \quad \text{for } t \in [\tau_n, \tau_{n+1}).$$

- ▶ For $D = 1/2$ we have the old model $X_t \equiv Y_t$.
- ▶ For $D < 1/2$, the process $(X_t)_{t \geq 0}$ is obtained from $(Y_t)_{t \geq 0}$ by a **nonlinear time-change**, refreshed at each time τ_n .

Baldovin & Stella's Model

Fix a (periodic) sequence of epochs $0 < \tau_1 < \tau_2 < \dots < \tau_n \uparrow +\infty$ and a parameter $0 < D \leq 1/2$. Define a new process $(X_t)_{t \geq 0}$ by

$$X_t := Y_{t^{2D}} \quad \text{for } t \in [0, \tau_1),$$

$$X_t := Y_{(t-\tau_n)^{2D} + \sum_{k=1}^n (\tau_k - \tau_{k-1})^{2D}} \quad \text{for } t \in [\tau_n, \tau_{n+1}).$$

- ▶ For $D = 1/2$ we have the old model $X_t \equiv Y_t$.
- ▶ For $D < 1/2$, the process $(X_t)_{t \geq 0}$ is obtained from $(Y_t)_{t \geq 0}$ by a **nonlinear time-change**, refreshed at each time τ_n .
- ▶ Increments are amplified immediately after the times $(\tau_n)_{n \geq 1}$ and then progressively damped out.

Baldovin & Stella's Model

Fix a (periodic) sequence of epochs $0 < \tau_1 < \tau_2 < \dots < \tau_n \uparrow +\infty$ and a parameter $0 < D \leq 1/2$. Define a new process $(X_t)_{t \geq 0}$ by

$$X_t := Y_{t^{2D}} \quad \text{for } t \in [0, \tau_1),$$

$$X_t := Y_{(t-\tau_n)^{2D} + \sum_{k=1}^n (\tau_k - \tau_{k-1})^{2D}} \quad \text{for } t \in [\tau_n, \tau_{n+1}).$$

- ▶ For $D = 1/2$ we have the old model $X_t \equiv Y_t$.
- ▶ For $D < 1/2$, the process $(X_t)_{t \geq 0}$ is obtained from $(Y_t)_{t \geq 0}$ by a **nonlinear time-change**, refreshed at each time τ_n .
- ▶ Increments are amplified immediately after the times $(\tau_n)_{n \geq 1}$ and then progressively damped out.
- ▶ Interpretation: $(\tau_n)_{n \geq 1}$ linked to “shocks” in the market.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**.
However, Baldovin & Stella show by simulations that this model
(with $(\tau_n)_n$ a periodic sequence) fits **all mentioned stylized facts**.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**. However, Baldovin & Stella show by simulations that this model (with $(\tau_n)_n$ a periodic sequence) fits **all mentioned stylized facts**. They actually simulate **a different model**: an **autoregressive version** of $(X_t)_{t \geq 0}$.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**.

However, Baldovin & Stella show by simulations that this model (with $(\tau_n)_n$ a periodic sequence) fits **all mentioned stylized facts**.

They actually simulate **a different model**: an **autoregressive version** of $(X_t)_{t \geq 0}$.

Other issue: the density of Y_1 is $g(\cdot)$ by construction. However, the density of X_1 **is not $g(\cdot)$** and depends on the choice of $(\tau_n)_n$.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**.

However, Baldovin & Stella show by simulations that this model (with $(\tau_n)_n$ a periodic sequence) fits **all mentioned stylized facts**.

They actually simulate **a different model**: an **autoregressive version** of $(X_t)_{t \geq 0}$.

Other issue: the density of Y_1 is $g(\cdot)$ by construction. However, the density of X_1 **is not $g(\cdot)$** and depends on the choice of $(\tau_n)_n$.

Our aims

- ▶ Define a simple model capturing the essential features of Baldovin & Stella's construction.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**.

However, Baldovin & Stella show by simulations that this model (with $(\tau_n)_n$ a periodic sequence) fits **all mentioned stylized facts**.

They actually simulate **a different model**: an **autoregressive version** of $(X_t)_{t \geq 0}$.

Other issue: the density of Y_1 is $g(\cdot)$ by construction. However, the density of X_1 **is not $g(\cdot)$** and depends on the choice of $(\tau_n)_n$.

Our aims

- ▶ Define a simple model capturing the essential features of Baldovin & Stella's construction.
- ▶ Easy to describe and to **simulate**.

Baldovin & Stella's Model

Despite the time-change, the process $(X_t)_{t \geq 0}$ remains **not ergodic**. However, Baldovin & Stella show by simulations that this model (with $(\tau_n)_n$ a periodic sequence) fits **all mentioned stylized facts**. They actually simulate **a different model**: an **autoregressive version** of $(X_t)_{t \geq 0}$.

Other issue: the density of Y_1 is $g(\cdot)$ by construction. However, the density of X_1 **is not $g(\cdot)$** and depends on the choice of $(\tau_n)_n$.

Our aims

- ▶ Define a simple model capturing the essential features of Baldovin & Stella's construction.
- ▶ Easy to describe and to **simulate**.
- ▶ **Rigorous proofs** of the mentioned stylized facts.

Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
4. Main Results
5. Estimation and Simulations
6. Conclusions

Our Model

Fix $D \in (0, 1/2]$, $\lambda \in (0, \infty)$ and a probability ν on $(0, \infty)$ (the law of σ). These are our “parameters”.

Our Model

Fix $D \in (0, 1/2]$, $\lambda \in (0, \infty)$ and a probability ν on $(0, \infty)$ (the law of σ). These are our “parameters”.

We need the following independent sources of randomness:

- ▶ a standard Brownian motion $W = (W_t)_{t \geq 0}$;

Our Model

Fix $D \in (0, 1/2]$, $\lambda \in (0, \infty)$ and a probability ν on $(0, \infty)$ (the law of σ). These are our “parameters”.

We need the following independent sources of randomness:

- ▶ a standard Brownian motion $W = (W_t)_{t \geq 0}$;
- ▶ a Poisson point process $\mathcal{T} = (\tau_n)_{n \in \mathbb{Z}}$ on \mathbb{R} with intensity λ ;

Our Model

Fix $D \in (0, 1/2]$, $\lambda \in (0, \infty)$ and a probability ν on $(0, \infty)$ (the law of σ). These are our “parameters”.

We need the following independent sources of randomness:

- ▶ a standard Brownian motion $W = (W_t)_{t \geq 0}$;
- ▶ a Poisson point process $\mathcal{T} = (\tau_n)_{n \in \mathbb{Z}}$ on \mathbb{R} with intensity λ ;
- ▶ an i.i.d. sequence of r.v.s $\Sigma = (\sigma_n)_{n \geq 0}$ with marginal law ν .

Our Model

Fix $D \in (0, 1/2]$, $\lambda \in (0, \infty)$ and a probability ν on $(0, \infty)$ (the law of σ). These are our “parameters”.

We need the following independent sources of randomness:

- ▶ a standard Brownian motion $W = (W_t)_{t \geq 0}$;
- ▶ a Poisson point process $\mathcal{T} = (\tau_n)_{n \in \mathbb{Z}}$ on \mathbb{R} with intensity λ ;
- ▶ an i.i.d. sequence of r.v.s $\Sigma = (\sigma_n)_{n \geq 0}$ with marginal law ν .

For $t \geq 0$ we set

$$i(t) := \sup\{n \geq 0 : \tau_n \leq t\},$$

so that $\tau_{i(t)}$ is the location of the last point in \mathcal{T} before t .

Our Model

Our process

Define the auxiliary process $I = (I_t)_{t \geq 0}$ by

$$I_t := \sigma_{i(t)}^2 (t - \tau_{i(t)})^{2D} + \sum_{k=1}^{i(t)} \sigma_{k-1}^2 (\tau_k - \tau_{k-1})^{2D} - \sigma_0^2 (-\tau_0)^{2D}.$$

Our Model

Our process

Define the auxiliary process $I = (I_t)_{t \geq 0}$ by

$$I_t := \sigma_{i(t)}^2 (t - \tau_{i(t)})^{2D} + \sum_{k=1}^{i(t)} \sigma_{k-1}^2 (\tau_k - \tau_{k-1})^{2D} - \sigma_0^2 (-\tau_0)^{2D}.$$

We define our basic process $X = (X_t)_{t \geq 0}$ by setting

$$X_t := W_{I_t}.$$

Our Model

Our process

Define the auxiliary process $I = (I_t)_{t \geq 0}$ by

$$I_t := \sigma_{i(t)}^2 (t - \tau_{i(t)})^{2D} + \sum_{k=1}^{i(t)} \sigma_{k-1}^2 (\tau_k - \tau_{k-1})^{2D} - \sigma_0^2 (-\tau_0)^{2D}.$$

We define our basic process $X = (X_t)_{t \geq 0}$ by setting

$$X_t := W_{I_t}.$$

- $(I_t)_{t \geq 0}$ is a **strictly increasing process** with absolutely continuous paths, independent of the BM $(W_t)_{t \geq 0}$.

Our Model

Our process

Define the auxiliary process $I = (I_t)_{t \geq 0}$ by

$$I_t := \sigma_{i(t)}^2 (t - \tau_{i(t)})^{2D} + \sum_{k=1}^{i(t)} \sigma_{k-1}^2 (\tau_k - \tau_{k-1})^{2D} - \sigma_0^2 (-\tau_0)^{2D}.$$

We define our basic process $X = (X_t)_{t \geq 0}$ by setting

$$X_t := W_{I_t}.$$

- ▶ $(I_t)_{t \geq 0}$ is a **strictly increasing process** with absolutely continuous paths, independent of the BM $(W_t)_{t \geq 0}$.
- ▶ Thus our model $X = (X_t)_{t \geq 0}$ may be viewed as an **independent random time change** of a Brownian motion.

Basic Properties

- ▶ The process X has stationary **ergodic** increments.

Basic Properties

- ▶ The process X has stationary **ergodic** increments.
- ▶ The process X is a **stochastic volatility process**:

$$dX_t = v_t dB_t,$$

where

$$B_t := \int_0^t \frac{1}{\sqrt{I'(I^{-1}(u))}} dW_u, \quad v_t := \sqrt{I'(t)} = \frac{\sqrt{2D} \sigma_{i(t)}}{(t - \tau_{i(t)})^{\frac{1}{2}-D}},$$

and $(B_t)_{t \geq 0}$ is a standard Brownian motion.

Basic Properties

- ▶ The process X has stationary **ergodic** increments.
- ▶ The process X is a **stochastic volatility process**:

$$dX_t = v_t dB_t,$$

where

$$B_t := \int_0^t \frac{1}{\sqrt{I'(I^{-1}(u))}} dW_u, \quad v_t := \sqrt{I'(t)} = \frac{\sqrt{2D} \sigma_{i(t)}}{(t - \tau_{i(t)})^{\frac{1}{2}-D}},$$

and $(B_t)_{t \geq 0}$ is a standard Brownian motion.

- ▶ The process X is a **zero-mean, square-integrable martingale**, provided $E(\sigma^2) = \int \sigma^2 \nu(d\sigma) < \infty$.

Basic Properties

- ▶ The process X has stationary **ergodic** increments.
- ▶ The process X is a **stochastic volatility process**:

$$dX_t = v_t dB_t,$$

where

$$B_t := \int_0^t \frac{1}{\sqrt{I'(I^{-1}(u))}} dW_u, \quad v_t := \sqrt{I'(t)} = \frac{\sqrt{2D} \sigma_{i(t)}}{(t - \tau_{i(t)})^{\frac{1}{2}-D}},$$

and $(B_t)_{t \geq 0}$ is a standard Brownian motion.

- ▶ The process X is a **zero-mean, square-integrable martingale**, provided $E(\sigma^2) = \int \sigma^2 \nu(d\sigma) < \infty$.
- ▶ $E[|X_t|^q] < +\infty$ iff $E(\sigma^q) < +\infty$.

Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
- 4. Main Results**
5. Estimation and Simulations
6. Conclusions

Approximate Diffusive Scaling

Theorem

As $h \downarrow 0$ we have the convergence in distribution

$$\frac{(X_{t+h} - X_t)}{\sqrt{h}} \xrightarrow[h \downarrow 0]{d} f(x) dx,$$

where $f(\cdot)$ is an explicit mixture of Gaussian densities.

Approximate Diffusive Scaling

Theorem

As $h \downarrow 0$ we have the convergence in distribution

$$\frac{(X_{t+h} - X_t)}{\sqrt{h}} \xrightarrow[h \downarrow 0]{d} f(x) dx ,$$

where $f(\cdot)$ is an explicit mixture of Gaussian densities.

$$f(x) = \int_0^\infty \nu(d\sigma) \int_0^\infty ds \lambda e^{-\lambda s} \frac{s^{1/2-D}}{\sigma \sqrt{4D\pi}} \exp\left(-\frac{s^{1-2D} x^2}{4D\sigma^2}\right) .$$

Approximate Diffusive Scaling

Theorem

As $h \downarrow 0$ we have the convergence in distribution

$$\frac{(X_{t+h} - X_t)}{\sqrt{h}} \xrightarrow[h \downarrow 0]{d} f(x) dx,$$

where $f(\cdot)$ is an explicit mixture of Gaussian densities.

Approximate Diffusive Scaling

Theorem

As $h \downarrow 0$ we have the convergence in distribution

$$\frac{(X_{t+h} - X_t)}{\sqrt{h}} \xrightarrow[h \downarrow 0]{d} f(x) dx,$$

where $f(\cdot)$ is an explicit mixture of Gaussian densities.

$f(x)$ has always **polynomial tails**: $\int |x|^q f(x) dx < \infty$ iff

$$q < q^* \quad \text{where} \quad q^* = q^*(D) := \frac{1}{\frac{1}{2} - D}$$

Therefore $f(\cdot)$ is not the density of X_t .

Approximate Diffusive Scaling

Theorem

As $h \downarrow 0$ we have the convergence in distribution

$$\frac{(X_{t+h} - X_t)}{\sqrt{h}} \xrightarrow[h \downarrow 0]{d} f(x) dx,$$

where $f(\cdot)$ is an explicit mixture of Gaussian densities.

$f(x)$ has always **polynomial tails**: $\int |x|^q f(x) dx < \infty$ iff

$$q < q^* \quad \text{where} \quad q^* = q^*(D) := \frac{1}{\frac{1}{2} - D}$$

Therefore $f(\cdot)$ is not the density of X_t .

Heavy tails of f related to **multiscaling** of $E[|X_{t+h} - X_t|^q]$.

Multiscaling of Moments

Theorem

Let $q > 0$, and assume $E(\sigma^q) := \int \sigma^q \nu(d\sigma) < +\infty$.

The moment $m_q(h) := E(|X_{t+h} - X_t|^q) = E(|X_h|^q)$ is finite and has the following asymptotic behavior as $h \downarrow 0$:

$$m_q(h) \sim \begin{cases} C_q h^{\frac{q}{2}} & \text{if } q < q^* \\ C_q h^{\frac{q}{2}} \log(\frac{1}{h}) & \text{if } q = q^* \\ C_q h^{Dq+1} & \text{if } q > q^* \end{cases}, \quad \text{where } q^* := \frac{1}{(\frac{1}{2} - D)}.$$

Multiscaling of Moments

Theorem

Let $q > 0$, and assume $E(\sigma^q) := \int \sigma^q \nu(d\sigma) < +\infty$.

The moment $m_q(h) := E(|X_{t+h} - X_t|^q) = E(|X_h|^q)$ is finite and has the following asymptotic behavior as $h \downarrow 0$:

$$m_q(h) \sim \begin{cases} C_q h^{\frac{q}{2}} & \text{if } q < q^* \\ C_q h^{\frac{q}{2}} \log(\frac{1}{h}) & \text{if } q = q^* \\ C_q h^{Dq+1} & \text{if } q > q^* \end{cases}, \quad \text{where } q^* := \frac{1}{(\frac{1}{2} - D)}.$$

- C_q **explicit function** of D , λ and $E(\sigma^q)$ (used in estimation)

Multiscaling of Moments

Theorem

Let $q > 0$, and assume $E(\sigma^q) := \int \sigma^q \nu(d\sigma) < +\infty$.

The moment $m_q(h) := E(|X_{t+h} - X_t|^q) = E(|X_h|^q)$ is finite and has the following asymptotic behavior as $h \downarrow 0$:

$$m_q(h) \sim \begin{cases} C_q h^{\frac{q}{2}} & \text{if } q < q^* \\ C_q h^{\frac{q}{2}} \log(\frac{1}{h}) & \text{if } q = q^* \\ C_q h^{Dq+1} & \text{if } q > q^* \end{cases}, \quad \text{where } q^* := \frac{1}{(\frac{1}{2} - D)}.$$

- ▶ C_q **explicit function** of D , λ and $E(\sigma^q)$ (used in estimation)
- ▶ We can write $m_q(h) \approx h^{A(q)}$ with **scaling exponent** $A(q)$

$$A(q) := \lim_{h \downarrow 0} \frac{\log m_q(h)}{\log h} = \begin{cases} q/2 & \text{if } q \leq q^* \\ Dq + 1 & \text{if } q \geq q^* \end{cases}.$$

Decay of Correlations

Theorem

The correlations of the absolute values of the increments of the process X have the following asymptotic behavior as $h \downarrow 0$:

$$\lim_{h \downarrow 0} \rho(|X_{s+h} - X_s|, |X_{t+h} - X_t|)$$

$$=: \rho(t - s) = \frac{2}{\pi \operatorname{Var}(\sigma | W_1| S^{D-1/2})} e^{-\lambda|t-s|} \phi(\lambda|t-s|).$$

where

$$\phi(x) := \operatorname{Cov}(\sigma S^{D-1/2}, \sigma (S+x)^{D-1/2})$$

and $\sigma \sim \nu$, $S \sim \operatorname{Exp}(1)$ are independent and independent of W .

Decay of Correlations

Theorem

The correlations of the absolute values of the increments of the process X have the following asymptotic behavior as $h \downarrow 0$:

$$\lim_{h \downarrow 0} \rho(|X_{s+h} - X_s|, |X_{t+h} - X_t|)$$

$$=: \rho(t - s) = \frac{2}{\pi \operatorname{Var}(\sigma | W_1| S^{D-1/2})} e^{-\lambda|t-s|} \phi(\lambda|t - s|).$$

where

$$\phi(x) := \operatorname{Cov}(\sigma S^{D-1/2}, \sigma (S + x)^{D-1/2})$$

and $\sigma \sim \nu$, $S \sim \operatorname{Exp}(1)$ are independent and independent of W .

- The function $\phi(\cdot)$ has a slower than exponential decay.

Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
4. Main Results
5. Estimation and Simulations
6. Conclusions

Estimation of the Parameters

The parameters of our model are D , λ and the law of σ , that we want to estimate on the DJIA time series (1935–2009).

Estimation of the Parameters

The parameters of our model are D , λ and the law of σ , that we want to estimate on the DJIA time series (1935–2009).

We start focusing on D , λ , $E(\sigma)$ and $E(\sigma^2)$.

Estimation of the Parameters

The parameters of our model are D , λ and the law of σ , that we want to estimate on the DJIA time series (1935–2009).

We start focusing on D , λ , $E(\sigma)$ and $E(\sigma^2)$.

Multiscaling of moments: $m_q(h) = E(|X_h|^q) \sim C_q h^{A(q)}$

Estimation of the Parameters

The parameters of our model are D , λ and the law of σ , that we want to estimate on the DJIA time series (1935–2009).

We start focusing on D , λ , $E(\sigma)$ and $E(\sigma^2)$.

Multiscaling of moments: $m_q(h) = E(|X_h|^q) \sim C_q h^{A(q)}$

1. Scaling exponent $A(q)$ function of D :

$$A(q) = \begin{cases} q/2 & \text{if } q \leq q^* \\ Dq + 1 & \text{if } q \geq q^* \end{cases}.$$

Estimation of the Parameters

The parameters of our model are D , λ and the law of σ , that we want to estimate on the DJIA time series (1935–2009).

We start focusing on D , λ , $E(\sigma)$ and $E(\sigma^2)$.

Multiscaling of moments: $m_q(h) = E(|X_h|^q) \sim C_q h^{A(q)}$

1. Scaling exponent $A(q)$ function of D :

$$A(q) = \begin{cases} q/2 & \text{if } q \leq q^* \\ Dq + 1 & \text{if } q \geq q^* \end{cases}.$$

2. Constants C_1 and C_2 functions of D , λ , $E(\sigma)$ and $E(\sigma^2)$:

$$C_1 = \frac{2}{\sqrt{\pi}} \sqrt{D} \Gamma(\tfrac{1}{2} + D) E(\sigma) \lambda^{1/2-D} \quad C_2 = 2D \Gamma(2D) E(\sigma^2) \lambda^{1-2D}.$$

Estimation of the Parameters

3. Volatility autocorrelation $\rho(t)$ function of D , λ , $E(\sigma)$, $E(\sigma^2)$:

$$\rho(t) = \frac{2}{\pi \operatorname{Var}(\sigma | W_1 | S^{D-1/2})} e^{-\lambda t} \phi(\lambda t)$$

with $\phi(\cdot)$ (quite) easily computable.

Estimation of the Parameters

3. Volatility autocorrelation $\rho(t)$ function of D , λ , $E(\sigma)$, $E(\sigma^2)$:

$$\rho(t) = \frac{2}{\pi \operatorname{Var}(\sigma | W_1| S^{D-1/2})} e^{-\lambda t} \phi(\lambda t)$$

with $\phi(\cdot)$ (quite) easily computable.

We evaluate the corresponding statistics $\hat{A}(q)$, \hat{C}_1 , \hat{C}_2 , $\hat{\rho}(t)$ on the (detrended log-)DJIA time series $(x_i)_{1 \leq i \leq T=18849}$

Estimation of the Parameters

3. Volatility autocorrelation $\rho(t)$ function of D , λ , $E(\sigma)$, $E(\sigma^2)$:

$$\rho(t) = \frac{2}{\pi \operatorname{Var}(\sigma | W_1| S^{D-1/2})} e^{-\lambda t} \phi(\lambda t)$$

with $\phi(\cdot)$ (quite) easily computable.

We evaluate the corresponding statistics $\hat{A}(q)$, \hat{C}_1 , \hat{C}_2 , $\hat{\rho}(t)$ on the (detrended log-)DJIA time series $(x_i)_{1 \leq i \leq T=18849}$

$$\log \hat{m}_q(h) \sim \hat{A}(q) (\log h) + \log \hat{C}_q \quad \hat{m}_q(h) := \frac{1}{T-h} \sum_{i=1}^{T-h} |x_{i+h} - x_i|^q$$

Estimation of the Parameters

3. Volatility autocorrelation $\rho(t)$ function of D , λ , $E(\sigma)$, $E(\sigma^2)$:

$$\rho(t) = \frac{2}{\pi \operatorname{Var}(\sigma | W_1 | S^{D-1/2})} e^{-\lambda t} \phi(\lambda t)$$

with $\phi(\cdot)$ (quite) easily computable.

We evaluate the corresponding statistics $\hat{A}(q)$, \hat{C}_1 , \hat{C}_2 , $\hat{\rho}(t)$ on the (detrended log-)DJIA time series $(x_i)_{1 \leq i \leq T=18849}$

$$\log \hat{m}_q(h) \sim \hat{A}(q) (\log h) + \log \hat{C}_q \quad \hat{m}_q(h) := \frac{1}{T-h} \sum_{i=1}^{T-h} |x_{i+h} - x_i|^q$$

$$\hat{\rho}(t) := \operatorname{Corr}((x_{i+1} - x_i)_{1 \leq i \leq T-1-t}, (x_{i+t+1} - x_{i+t})_{1 \leq i \leq T-1-t})$$

Estimation of the Parameters

Loss function: $(T = 40)$

$$L(D, \lambda, E(\sigma), E(\sigma^2)) = \frac{1}{2} \left\{ \left(\frac{\hat{C}_1}{C_1} - 1 \right)^2 + \left(\frac{\hat{C}_2}{C_2} - 1 \right)^2 \right\} \\ + \int_0^5 \left(\frac{\hat{A}(q)}{A(q)} - 1 \right)^2 \frac{dq}{5} + \sum_{t=1}^{400} \frac{e^{-t/T}}{\sum_{s=1}^{400} e^{-s/T}} \left(\frac{\hat{\rho}(t)}{\rho(t)} - 1 \right)^2$$

Estimation of the Parameters

Loss function: ($T = 40$)

$$L(D, \lambda, E(\sigma), E(\sigma^2)) = \frac{1}{2} \left\{ \left(\frac{\hat{C}_1}{C_1} - 1 \right)^2 + \left(\frac{\hat{C}_2}{C_2} - 1 \right)^2 \right\} \\ + \int_0^5 \left(\frac{\hat{A}(q)}{A(q)} - 1 \right)^2 \frac{dq}{5} + \sum_{t=1}^{400} \frac{e^{-t/T}}{\sum_{s=1}^{400} e^{-s/T}} \left(\frac{\hat{\rho}(t)}{\rho(t)} - 1 \right)^2$$

Estimator: minimization constrained on $E(\sigma^2) \geq E(\sigma)^2$.

$$(\hat{D}, \hat{\lambda}, \widehat{E(\sigma)}, \widehat{E(\sigma^2)}) = \arg \min L(D, \lambda, E(\sigma), E(\sigma^2))$$

Estimation of the Parameters

Loss function: ($T = 40$)

$$L(D, \lambda, E(\sigma), E(\sigma^2)) = \frac{1}{2} \left\{ \left(\frac{\hat{C}_1}{C_1} - 1 \right)^2 + \left(\frac{\hat{C}_2}{C_2} - 1 \right)^2 \right\} \\ + \int_0^5 \left(\frac{\hat{A}(q)}{A(q)} - 1 \right)^2 \frac{dq}{5} + \sum_{t=1}^{400} \frac{e^{-t/T}}{\sum_{s=1}^{400} e^{-s/T}} \left(\frac{\hat{\rho}(t)}{\rho(t)} - 1 \right)^2$$

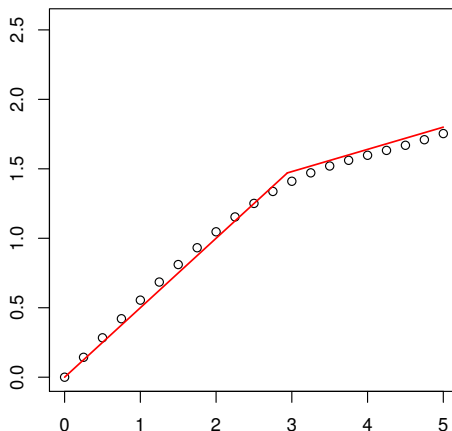
Estimator: minimization constrained on $E(\sigma^2) \geq E(\sigma)^2$.

$$(\hat{D}, \hat{\lambda}, \widehat{E(\sigma)}, \widehat{E(\sigma^2)}) = \arg \min L(D, \lambda, E(\sigma), E(\sigma^2))$$

$$\hat{D} \simeq 0.16 \quad \hat{\lambda} \simeq 0.00097 \quad \widehat{E(\sigma)} \simeq 0.108 \quad \widehat{E(\sigma^2)} \simeq (\widehat{E(\sigma)})^2$$

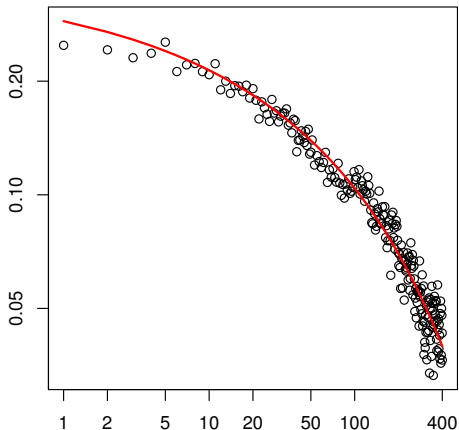
DJIA Time Series (1935-2009)

Empirical (circles) and theoretical (line) scaling exponent $A(q)$



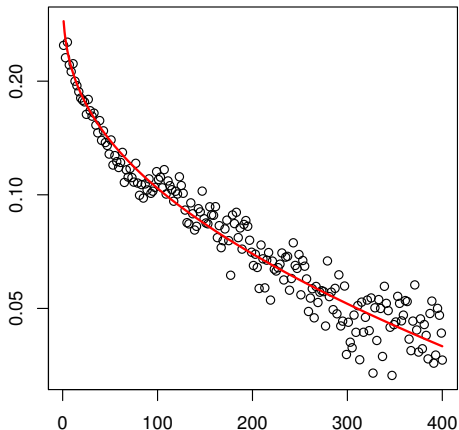
DJIA Time Series (1935-2009)

Empirical (circles) and theoretical (line) volatility autocorrelation [log plot]



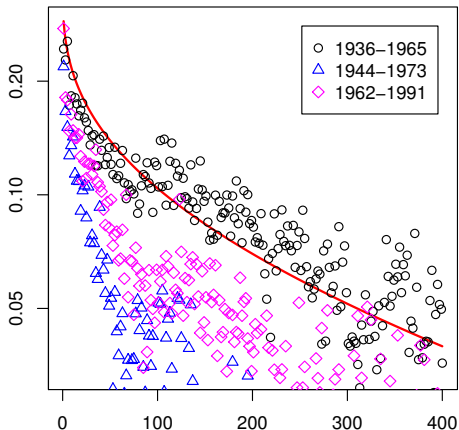
DJIA Time Series (1935-2009)

Empirical (circles) and theoretical (line) volatility autocorrelation [log-log plot]



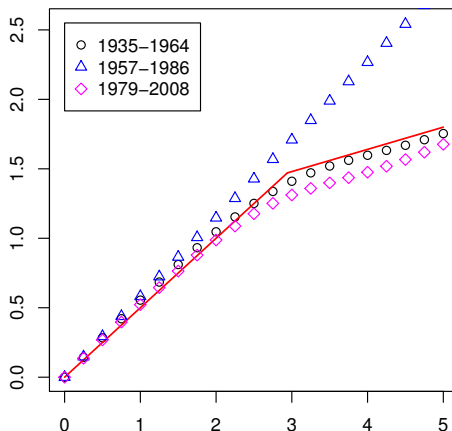
DJIA Time Series (1935-2009)

Volatility autocorrelation over sub-periods of 30 years [log plot]



DJIA Time Series (1935-2009)

Empirical scaling exponent $A(q)$ over sub-periods of 30 years.



Estimation of the Law of σ

The estimated values give $E(\sigma^2) \simeq E(\sigma)^2 \rightarrow \text{Var}(\sigma) \simeq 0$

Estimation of the Law of σ

The estimated values give $E(\sigma^2) \simeq E(\sigma)^2 \rightarrow \text{Var}(\sigma) \simeq 0$

The model is therefore **completely specified**.

Estimation of the Law of σ

The estimated values give $E(\sigma^2) \simeq E(\sigma)^2 \rightarrow \text{Var}(\sigma) \simeq 0$

The model is therefore **completely specified**.

We then compare the law of X_1 (daily log-return) predicted by our model with the empirical one evaluated on the DJIA time series.

Estimation of the Law of σ

The estimated values give $E(\sigma^2) \simeq E(\sigma)^2 \rightarrow \text{Var}(\sigma) \simeq 0$

The model is therefore **completely specified**.

We then compare the law of X_1 (daily log-return) predicted by our model with the empirical one evaluated on the DJIA time series.

No further parameter has to be estimated.

Estimation of the Law of σ

The estimated values give $E(\sigma^2) \simeq E(\sigma)^2 \rightarrow \text{Var}(\sigma) \simeq 0$

The model is therefore **completely specified**.

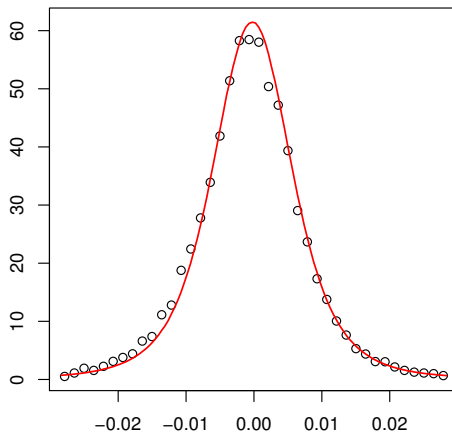
We then compare the law of X_1 (daily log-return) predicted by our model with the empirical one evaluated on the DJIA time series.

No further parameter has to be estimated.

The agreement is **remarkably good** (both **bulk** and **tails**).

DJIA Time Series (1935-2009)

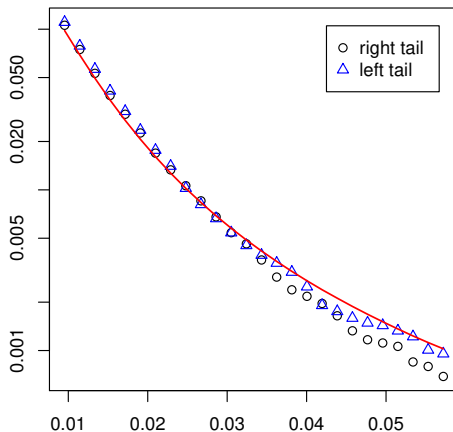
Empirical (circles) and theoretical (line) distribution of daily log return



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to 3 st. dev.

DJIA Time Series (1935-2009)

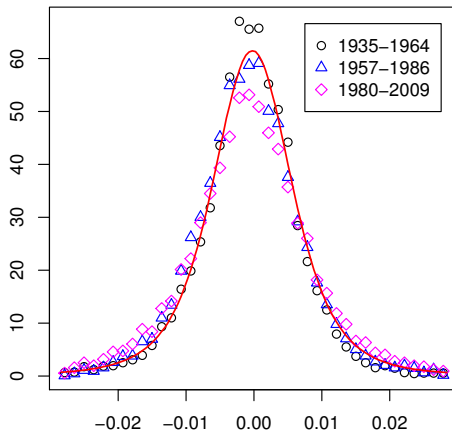
Empirical and theoretical tails of daily log return [log plot]



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: 1 to 6 st. dev.

DJIA Time Series (1935-2009)

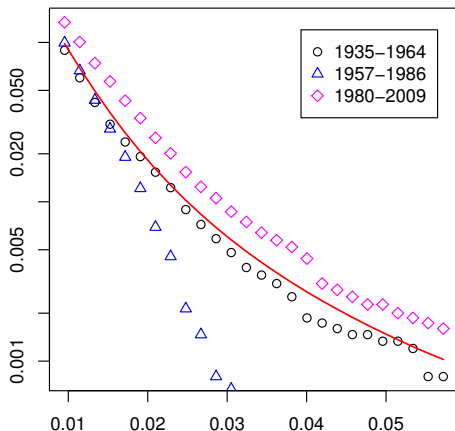
Variability of the distribution in sub-periods of 30 years



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: -3 to 3 st. dev.

DJIA Time Series (1935-2009)

Variability of the left tail in sub-periods of 30 years



Daily log-return **standard deviation ≈ 0.01** \rightarrow Range: 1 to 6 st. dev.

On the Law of σ

Even if we had found $\text{Var}(\sigma) > 0$, the details of the law of σ **would not be relevant**.

On the Law of σ

Even if we had found $\text{Var}(\sigma) > 0$, the details of the law of σ **would not be relevant**.

In fact $1/\lambda \simeq$ **1000 working days** \longrightarrow in 75 years we sample **only** $18849/1000 \simeq$ **18 different variables** σ_k .

On the Law of σ

Even if we had found $\text{Var}(\sigma) > 0$, the details of the law of σ **would not be relevant**.

In fact $1/\lambda \simeq 1000$ **working days** \longrightarrow in 75 years we sample **only** $18849/1000 \simeq 18$ **different variables** σ_k .

Not enough to see the details of the law of σ .

On the Law of σ

Even if we had found $\text{Var}(\sigma) > 0$, the details of the law of σ **would not be relevant**.

In fact $1/\lambda \simeq 1000$ **working days** \longrightarrow in 75 years we sample **only** $18849/1000 \simeq 18$ **different variables** σ_k .

Not enough to see the details of the law of σ .

The law of the log-returns is completely determined by the t^{2D} time scaling at the points of the Poisson process \mathcal{T} .

Variability of estimators

We have noticed that a **considerable variability** in different time windows is present for the **multiscaling of moments**, for the **decay of correlations** and for the **empirical distribution** (bulk and tails).

Variability of estimators

We have noticed that a **considerable variability** in different time windows is present for the **multiscaling of moments**, for the **decay of correlations** and for the **empirical distribution** (bulk and tails).

It is relevant to show that this is also consistent with our model.

Variability of estimators

We have noticed that a **considerable variability** in different time windows is present for the **multiscaling of moments**, for the **decay of correlations** and for the **empirical distribution** (bulk and tails).

It is relevant to show that this is also consistent with our model.

We have therefore simulated 75 years of data and computed the above quantities in different time windows.

Variability of estimators

We have noticed that a **considerable variability** in different time windows is present for the **multiscaling of moments**, for the **decay of correlations** and for the **empirical distribution** (bulk and tails).

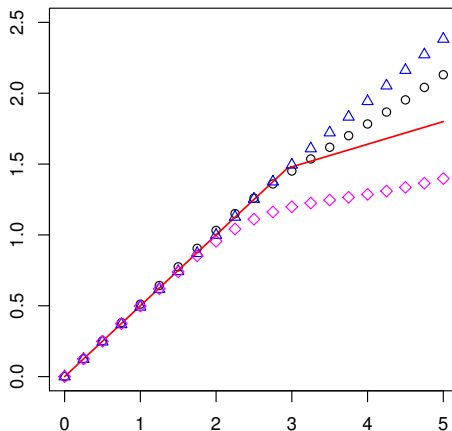
It is relevant to show that this is also consistent with our model.

We have therefore simulated 75 years of data and computed the above quantities in different time windows.

A **significant variability** is present also in our model.

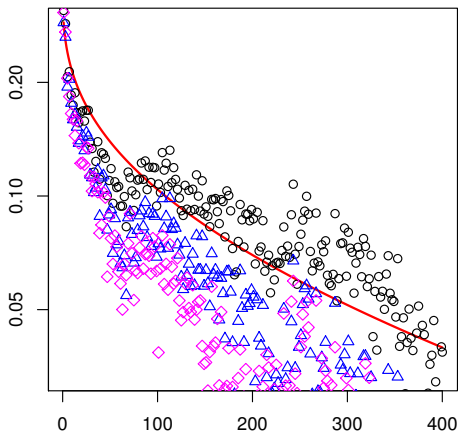
Simulated Data (75 years)

Simulated scaling exponent of our model over sub-periods of 30 years



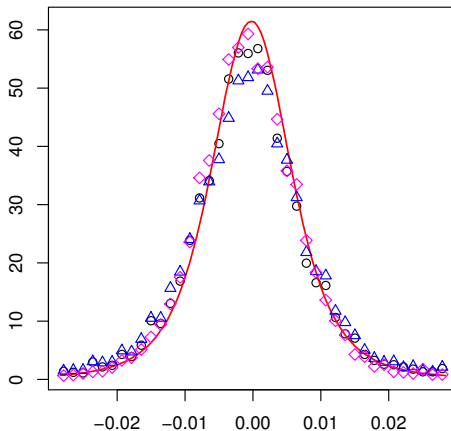
Simulated Data (75 years)

Simulated volatility autocorrelation of our model over sub-periods of 30 years



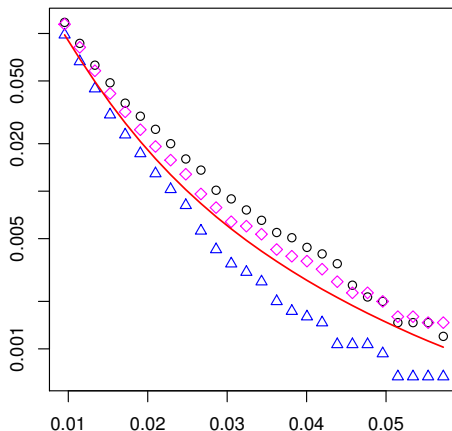
Simulated Data (75 years)

Simulated distribution of our model over sub-periods of 30 years



Simulated Data (75 years)

Simulated tails of our model over sub-periods of 30 years

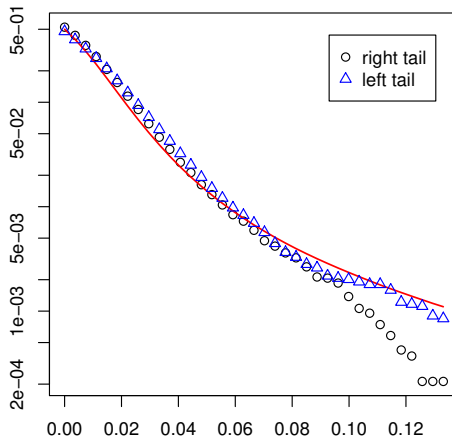


Other observables

Is everything going as expected?

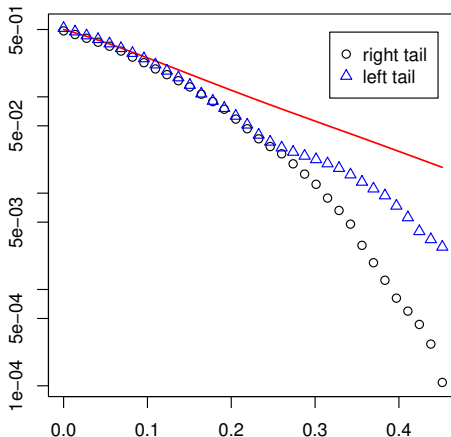
DJIA Time Series (1935-2009)

Empirical and theoretical tails of 5-day log return [log plot]



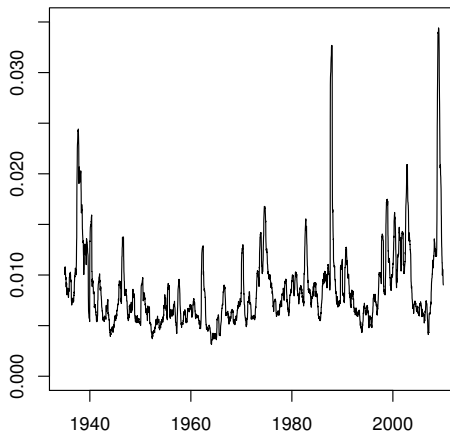
DJIA Time Series (1935-2009)

Empirical and theoretical tails of 400-day log return [log plot]



DJIA Time Series (1935-2009)

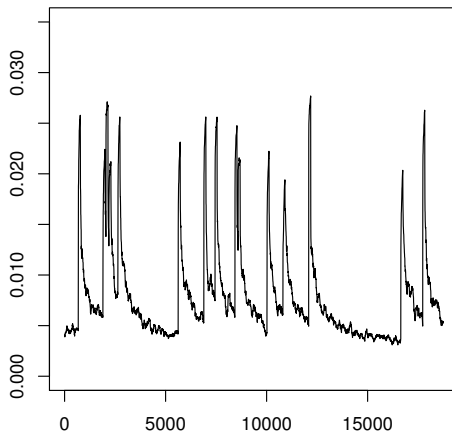
Empirical volatility



Local standard deviation of log-returns in a window of 100 days

Simulated Data (75 years)

Empirical volatility



Local standard deviation of log-returns in a window of 100 days

Outline

1. Introduction: the Black & Scholes Model
2. Beyond Black & Scholes
3. Our Model
4. Main Results
5. Estimation and Simulations
- 6. Conclusions**

Conclusions

We have proposed a model with the following features:

- ▶ It is analytically tractable. In particular, sharp asymptotics for scaling relations and correlations are obtained.

Conclusions

We have proposed a model with the following features:

- ▶ It is analytically tractable. In particular, sharp asymptotics for scaling relations and correlations are obtained.
- ▶ It is easy to simulate.

Conclusions

We have proposed a model with the following features:

- ▶ It is analytically tractable. In particular, sharp asymptotics for scaling relations and correlations are obtained.
- ▶ It is easy to simulate.
- ▶ Despite of the few parameters, it accounts for various phenomena observed in real time series.

Conclusions

We have proposed a model with the following features:

- ▶ It is analytically tractable. In particular, sharp asymptotics for scaling relations and correlations are obtained.
- ▶ It is easy to simulate.
- ▶ Despite of the few parameters, it accounts for various phenomena observed in real time series.
- ▶ Several generalizations can be considered: correlations between Σ , \mathcal{T} and W can be introduced, or the nonlinear time change $t \mapsto t^{2D}$ can be modified in many ways.

Conclusions

We have proposed a model with the following features:

- ▶ It is analytically tractable. In particular, sharp asymptotics for scaling relations and correlations are obtained.
- ▶ It is easy to simulate.
- ▶ Despite of the few parameters, it accounts for various phenomena observed in real time series.
- ▶ Several generalizations can be considered: correlations between Σ , \mathcal{T} and W can be introduced, or the nonlinear time change $t \mapsto t^{2D}$ can be modified in many ways.

Next steps:

- ▶ Solve specific problems by using this model: pricing of options, portfolio management, . . .

Thanks.