

Normal Multivariada

Graciela Boente

Definición 1

- Sea $\mu \in \mathbb{R}^p$ y $\Sigma \in \mathbb{R}^{p \times p}$ simétrica y definida positiva
Se dice que $\mathbf{x} \sim N(\mu, \Sigma)$ si su densidad está dada por

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{p}{2}}} \frac{1}{|\Sigma|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right\}$$

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- En particular, si $\mathbf{x} \sim N(\mathbf{0}, \mathbf{I}_p)$, x_1, \dots, x_p son i.i.d. $N(0, 1)$.
- Si $\mathbf{x} \sim N(\mu, \Sigma)$ y $\mathbf{A} \in \mathbb{R}^{p \times p}$ es no singular \Rightarrow
 $\mathbf{Ax} + \mathbf{b} \sim N(\mathbf{A}\mu + \mathbf{b}, \mathbf{A}\Sigma\mathbf{A}^T)$

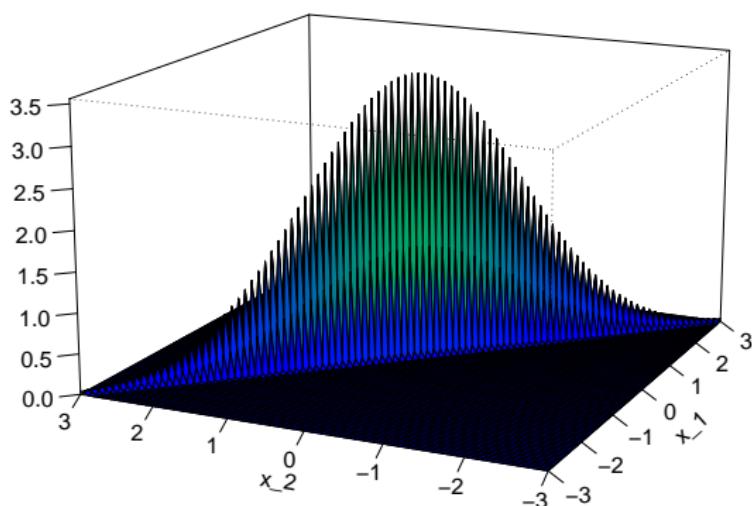
Caso p=2

- Sea $\mu \in \mathbb{R}^2$ y $\Sigma = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$ definida positiva ($|\rho| \neq 1$)

$$f(\mathbf{x}) = \frac{1}{2\pi \sigma_1 \sigma_2 (1 - \rho^2)^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2(1 - \rho^2)} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 - 2\rho \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) \right] \right\}$$

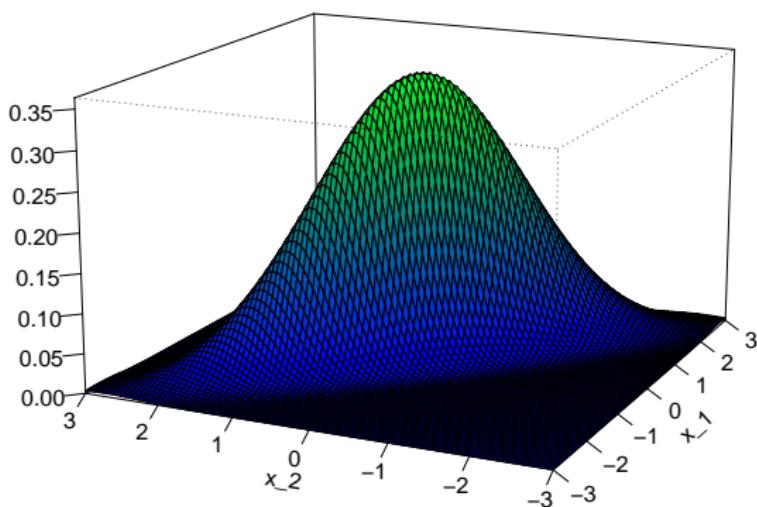
Caso $p=2$

rho = -0.999



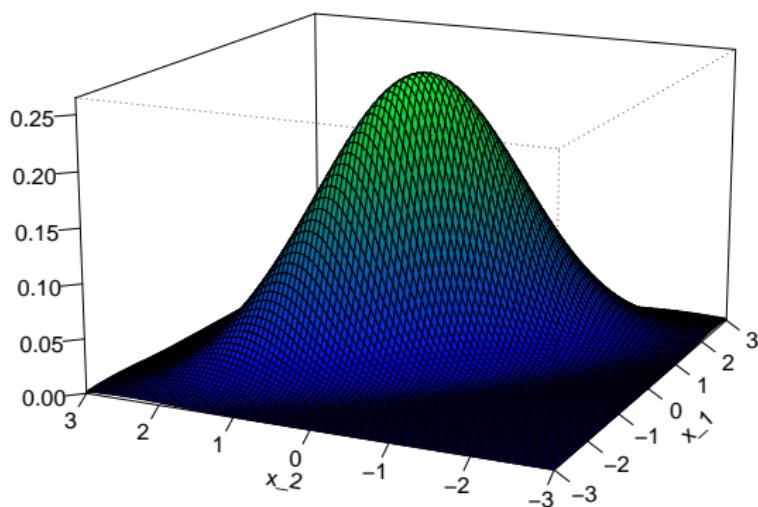
Caso p=2

rho = -0.8991



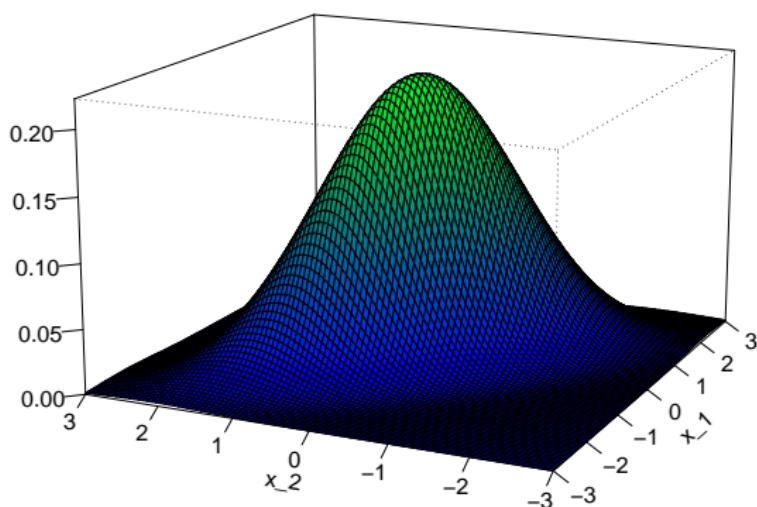
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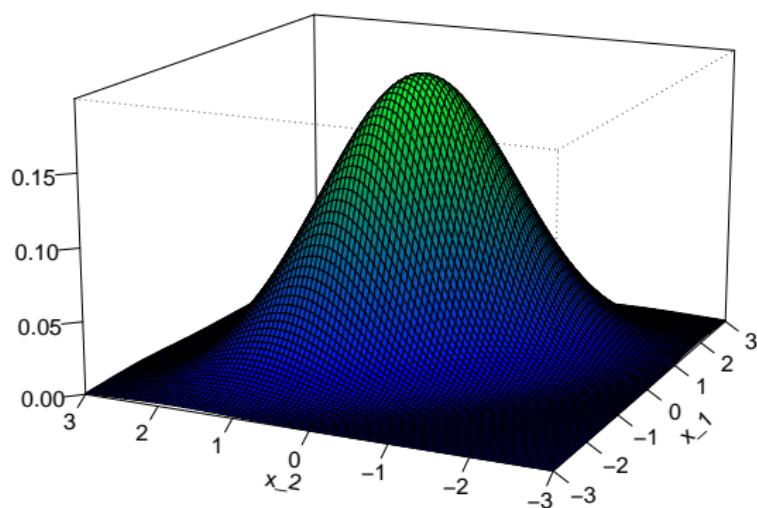
Caso p=2

rho = -0.6993



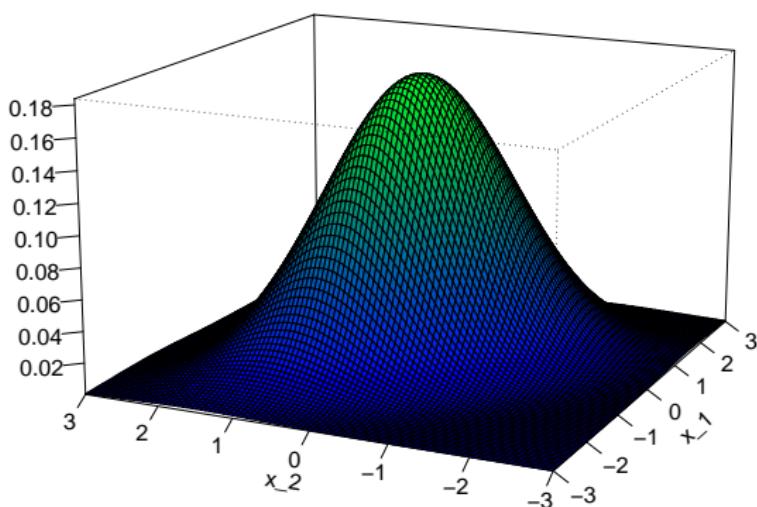
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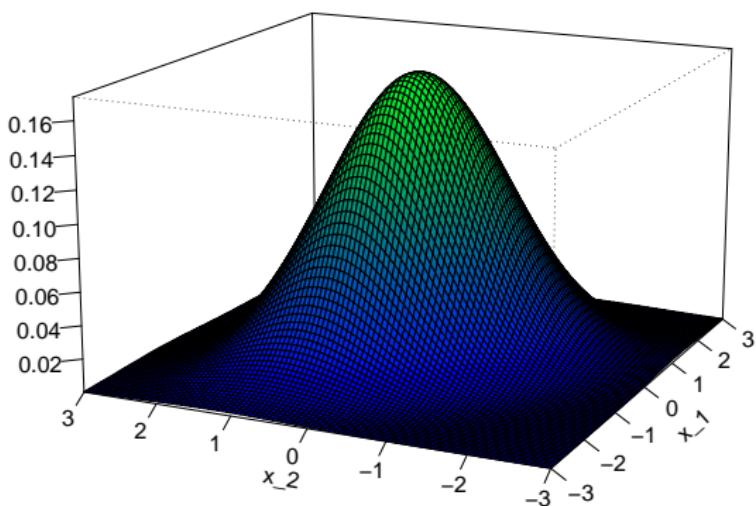
Caso p=2

rho = -0.4995



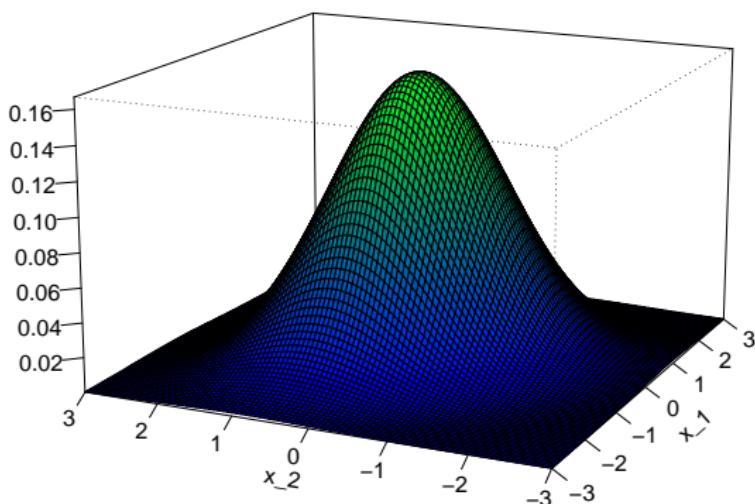
Caso p=2

rho = -0.3996



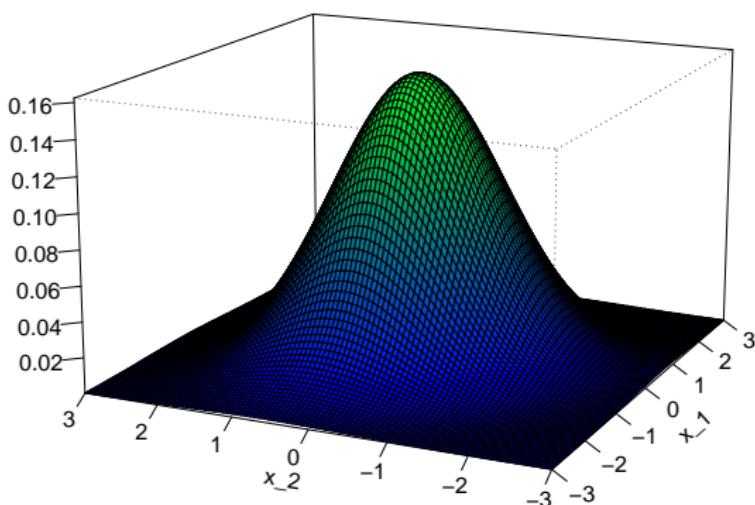
Caso p=2

rho = -0.2997



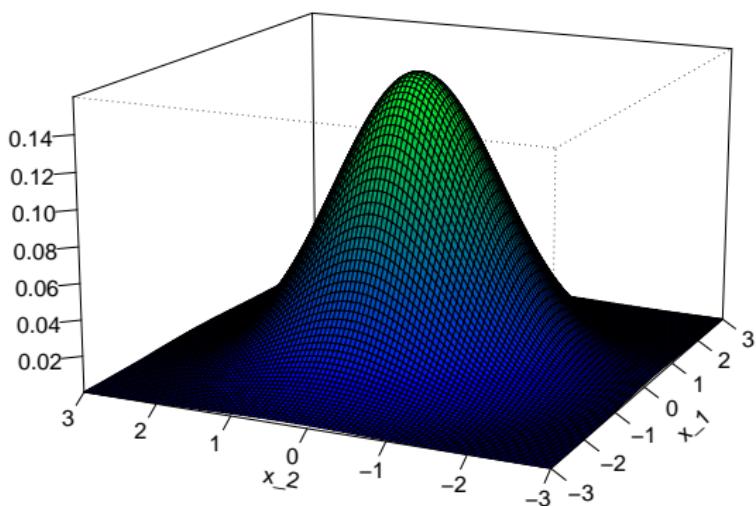
Caso p=2

rho = -0.1998



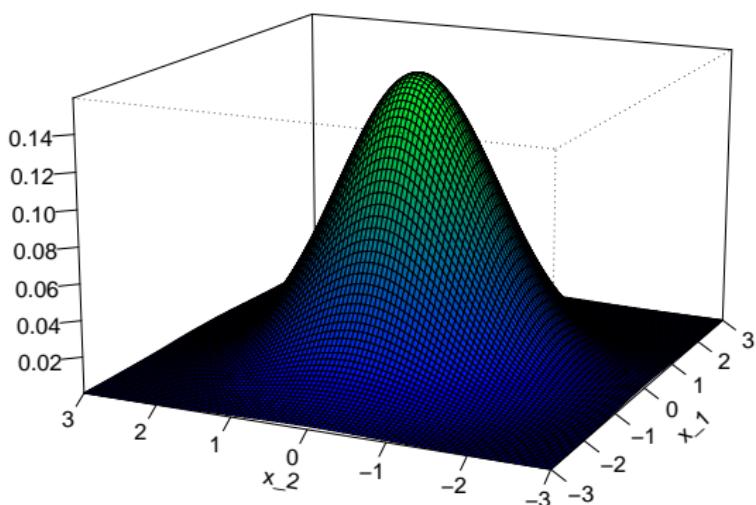
Caso p=2

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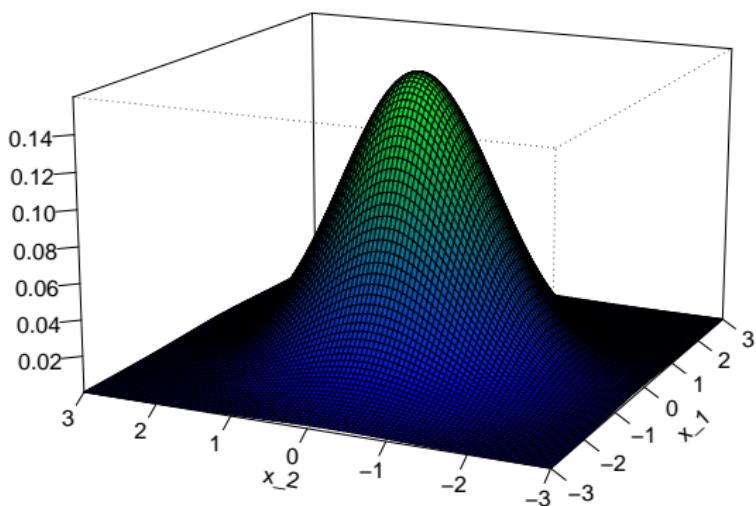
Caso $p=2$

rho = 0



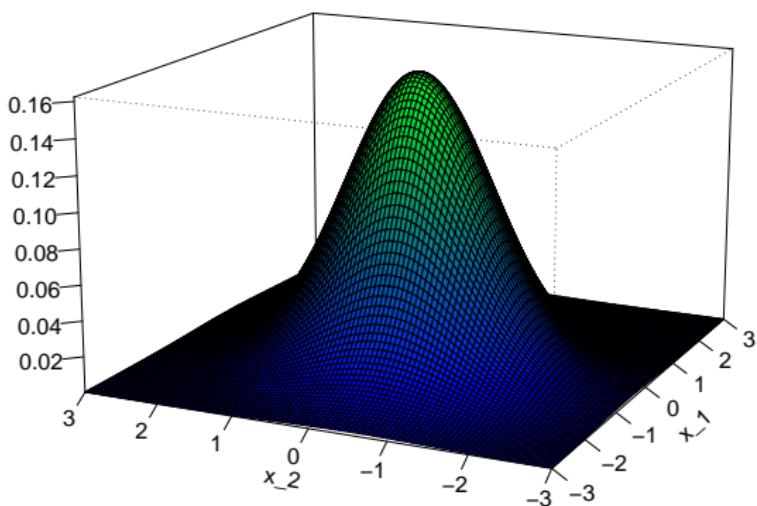
Caso p=2

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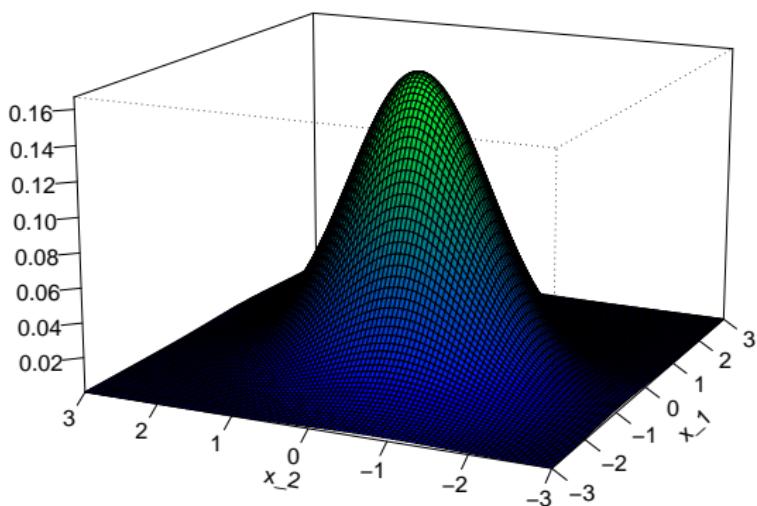
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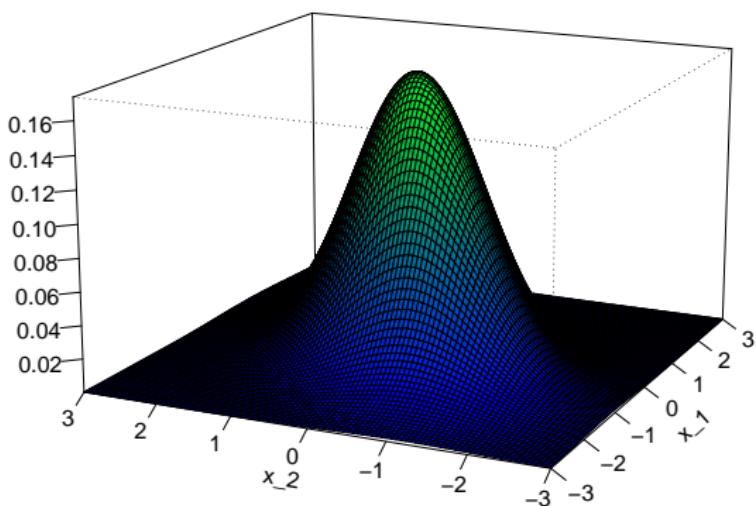
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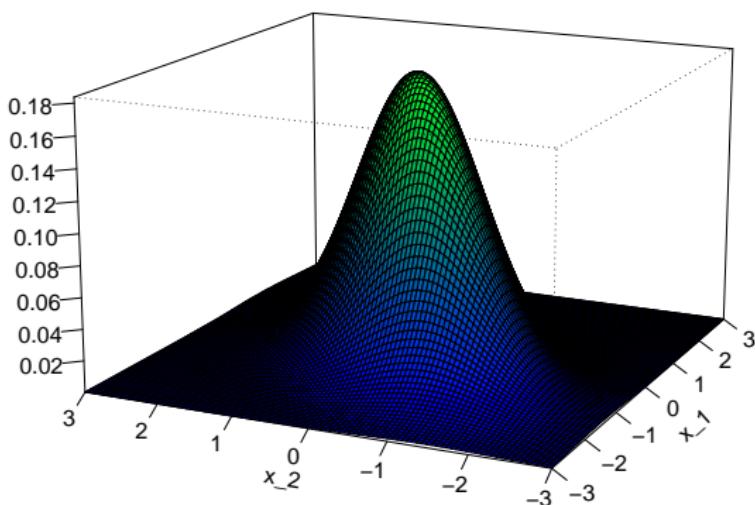
Caso p=2

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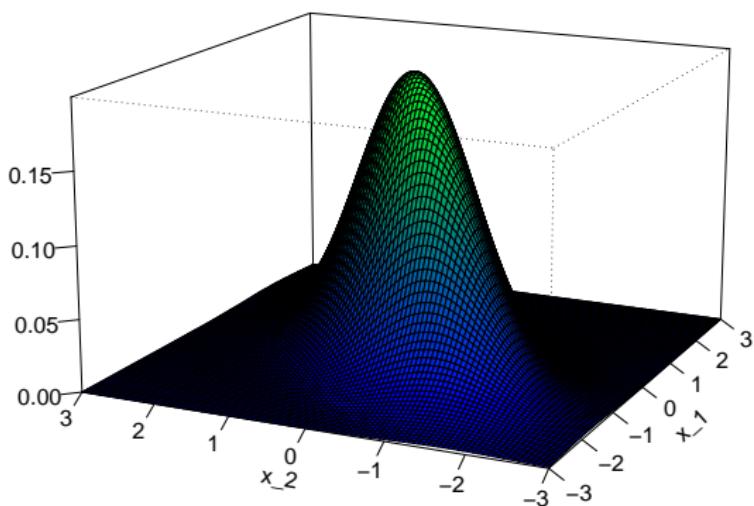
Caso p=2

rho = 0.4995



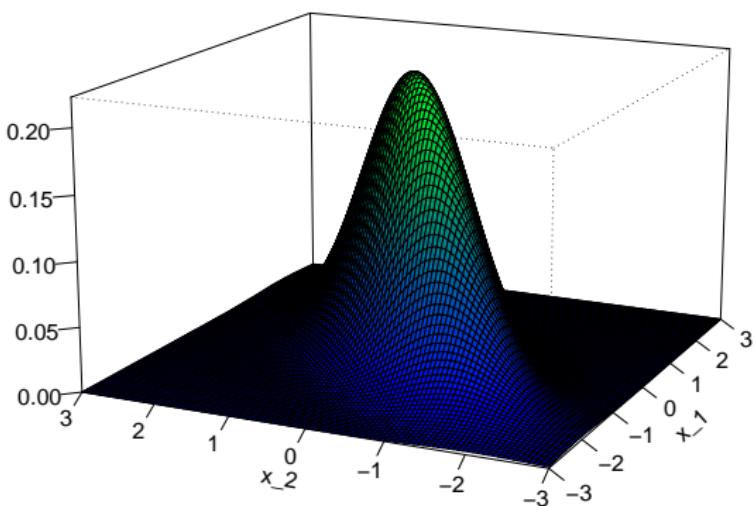
Caso $p=2$

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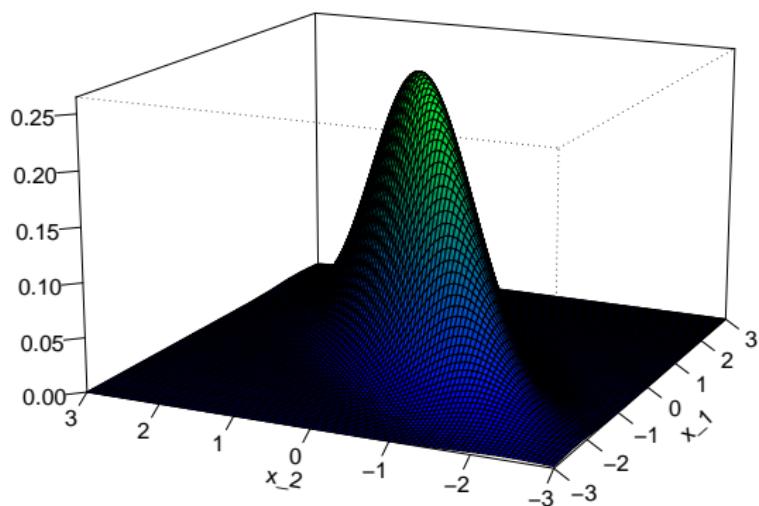
Caso p=2

rho = 0.6993



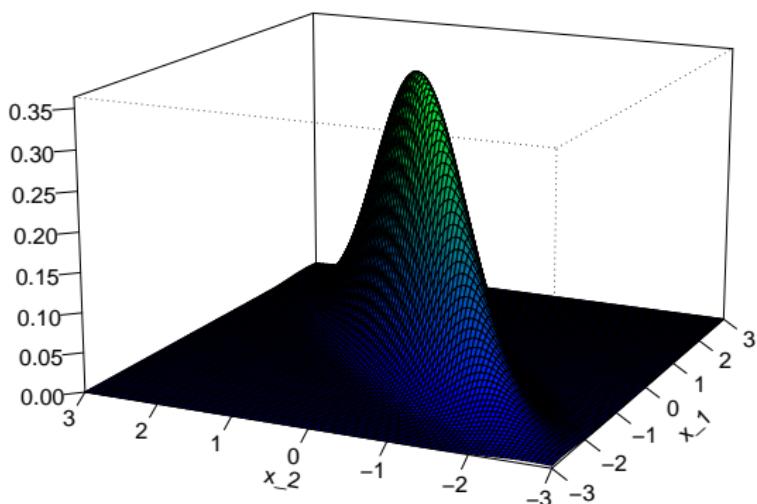
Caso p=2

rho = 0.7992



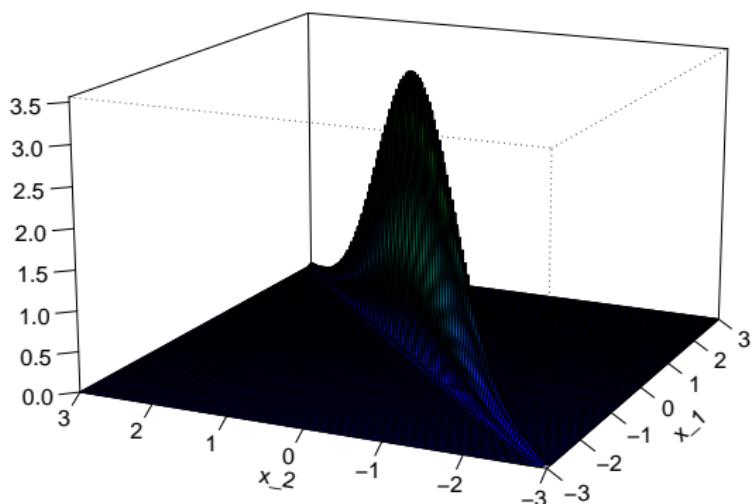
Caso p=2

rho = 0.8991



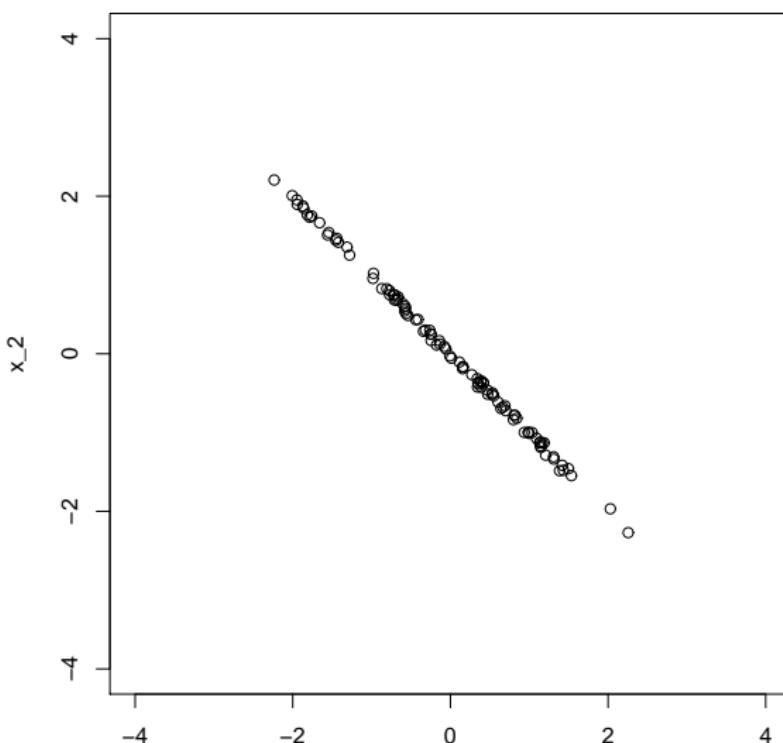
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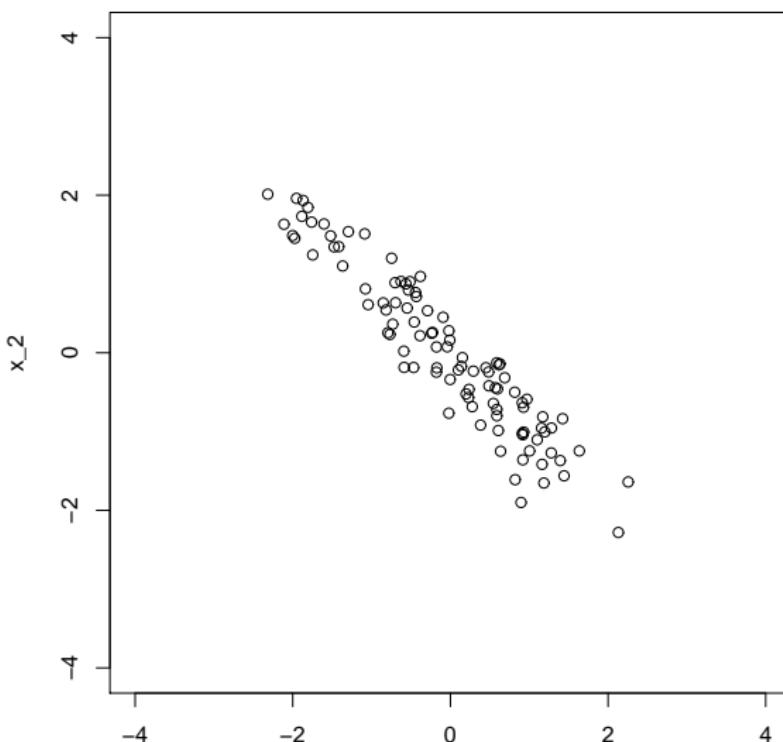
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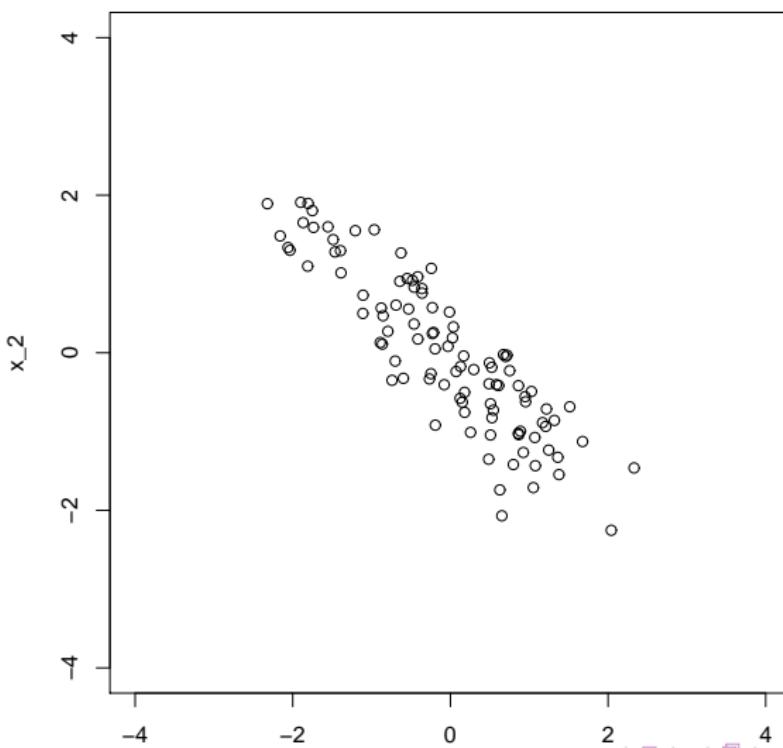
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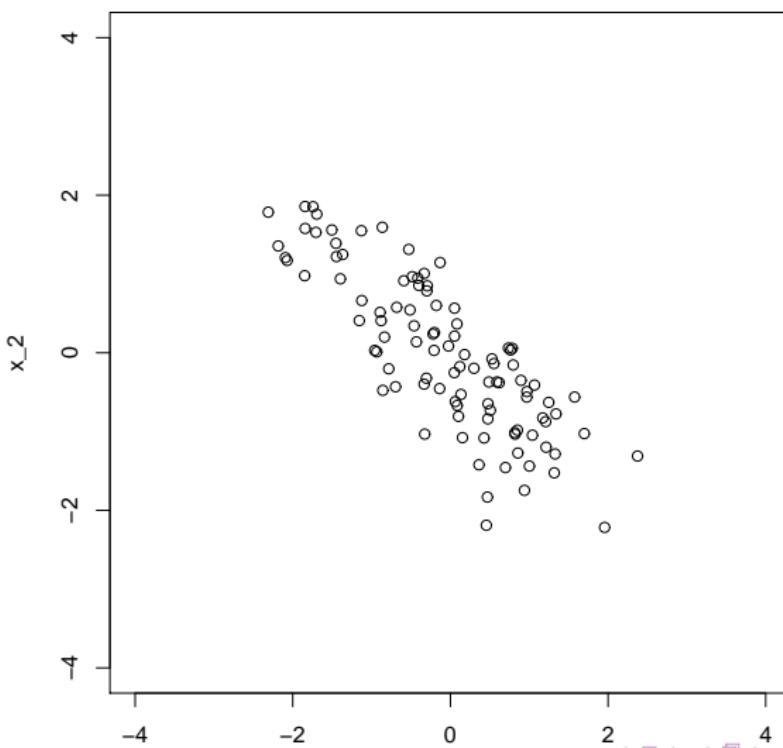
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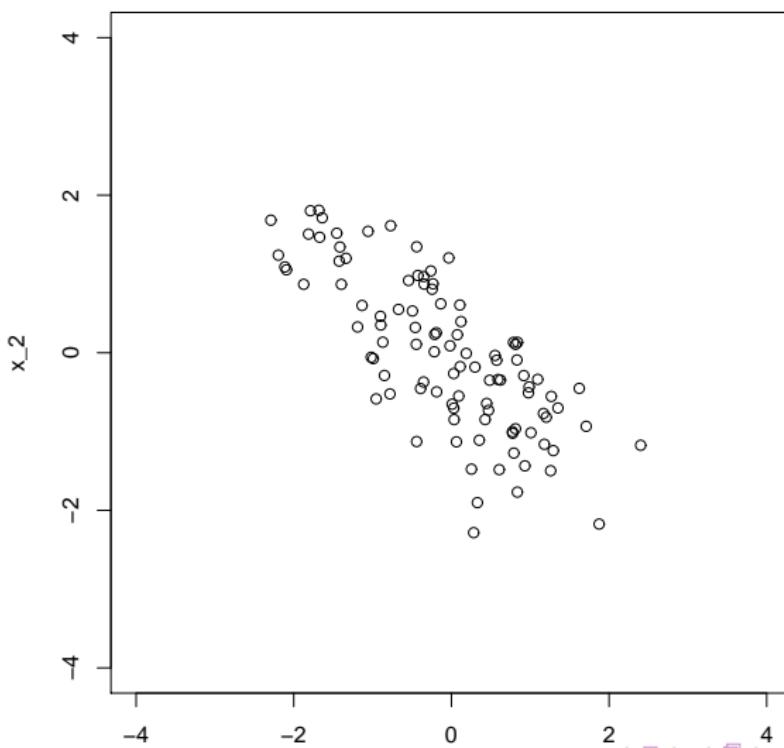
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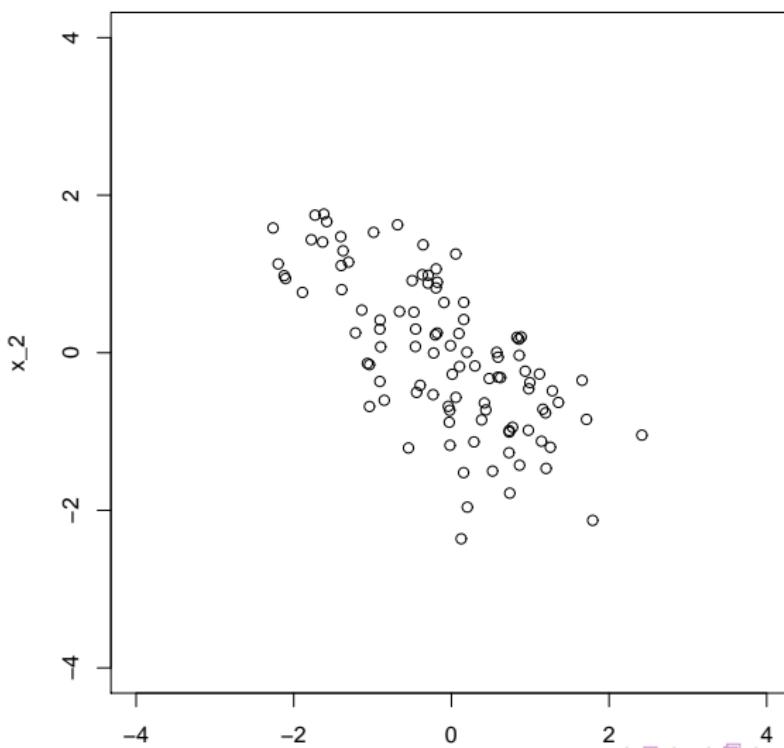
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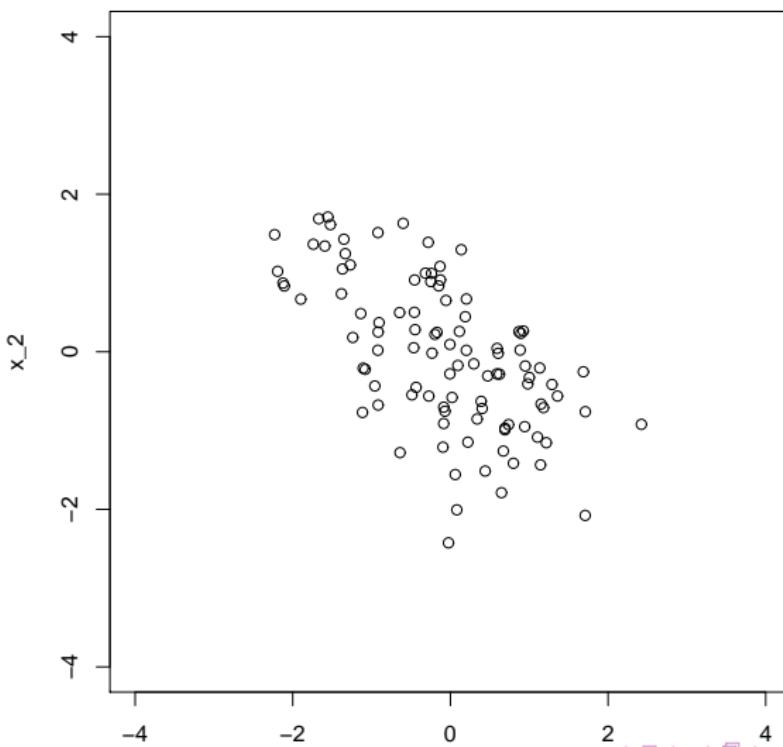
Caso p=2

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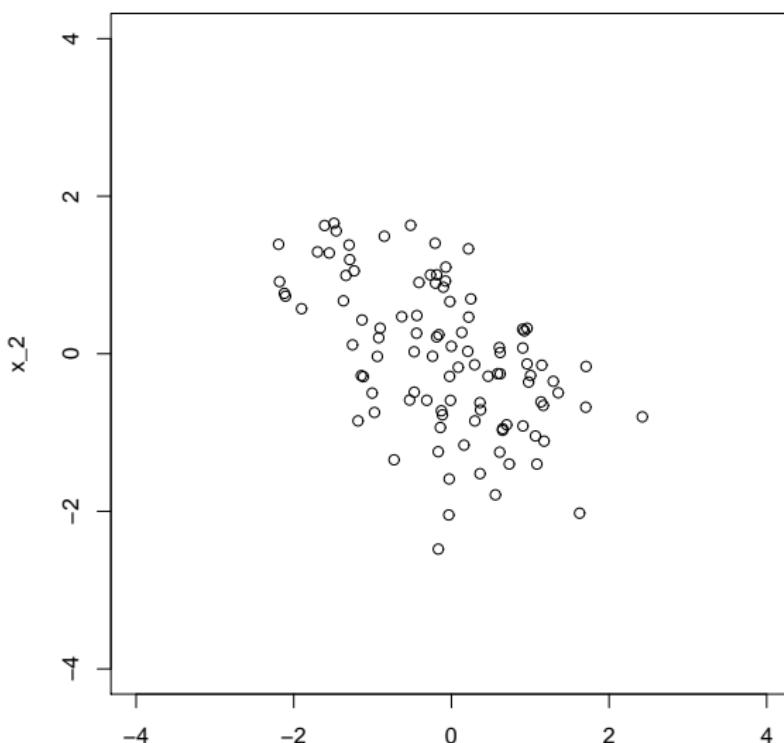
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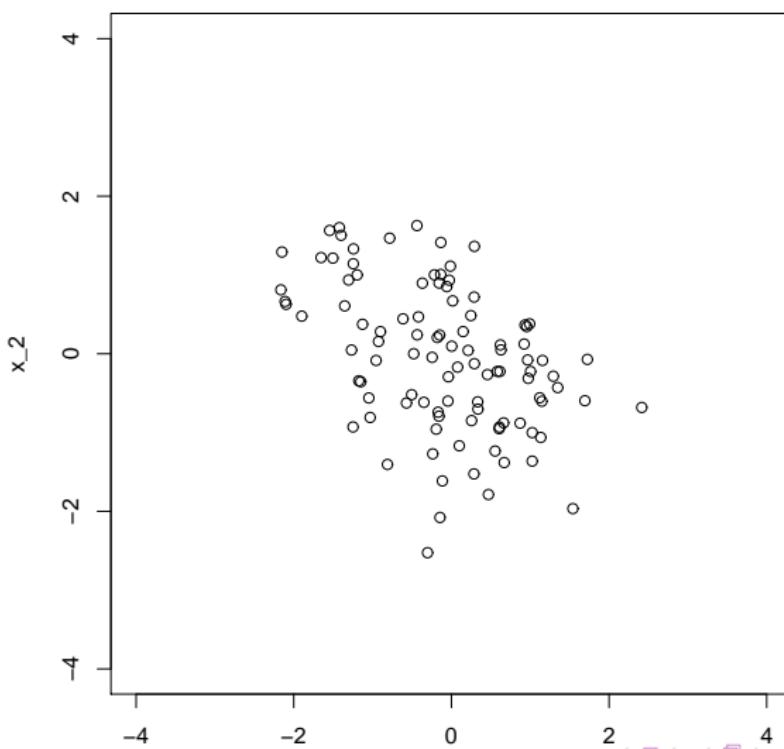
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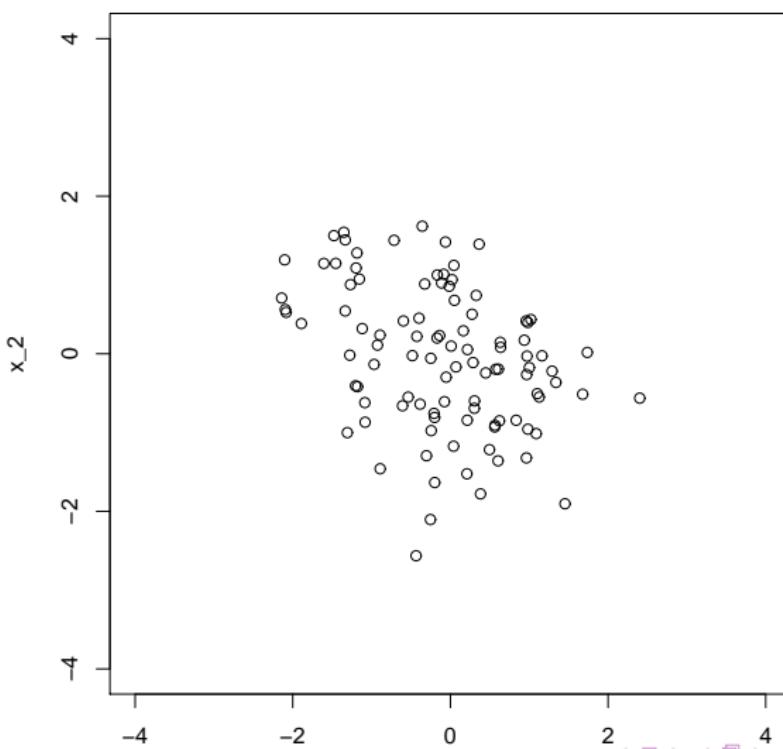
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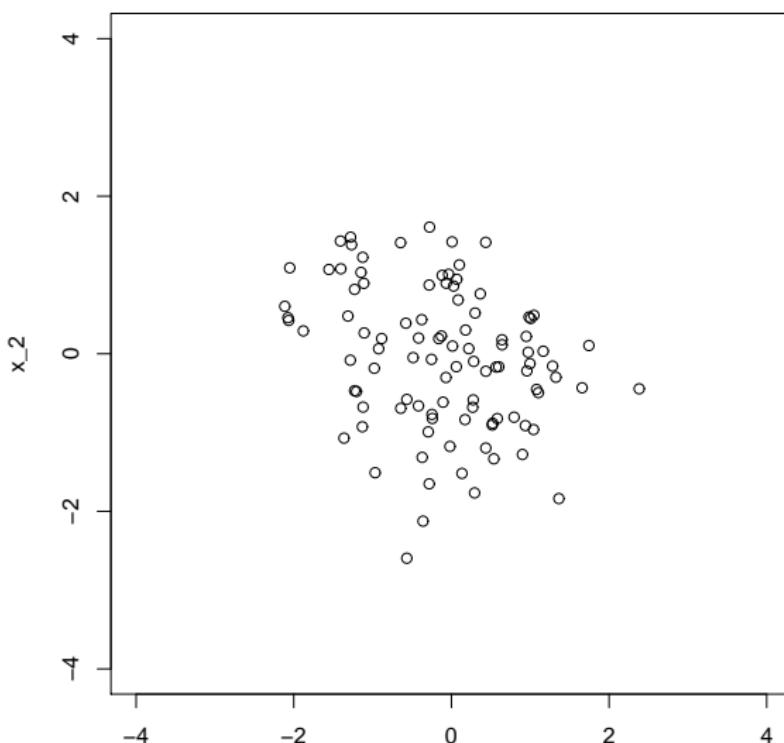
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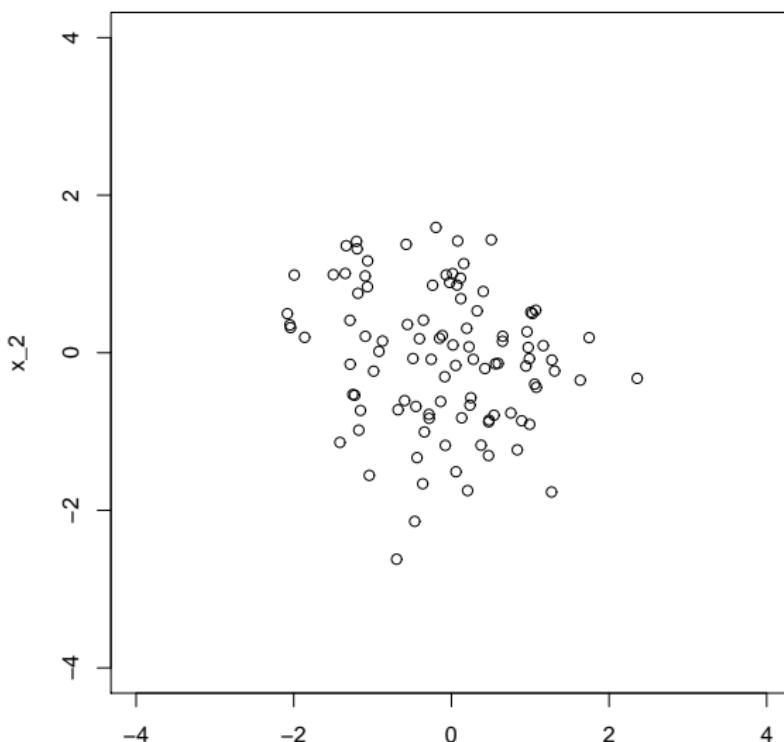
Caso $p=2$

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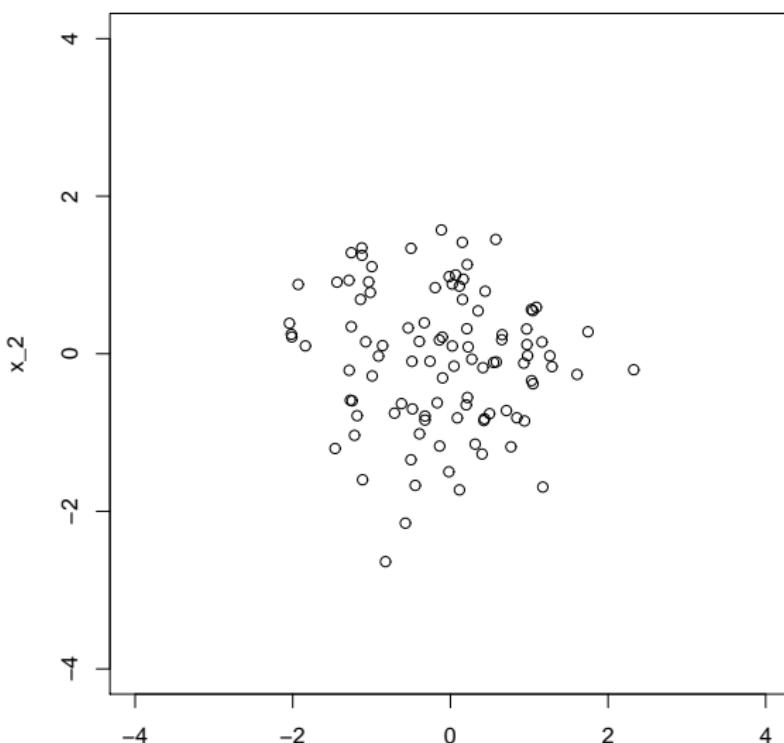
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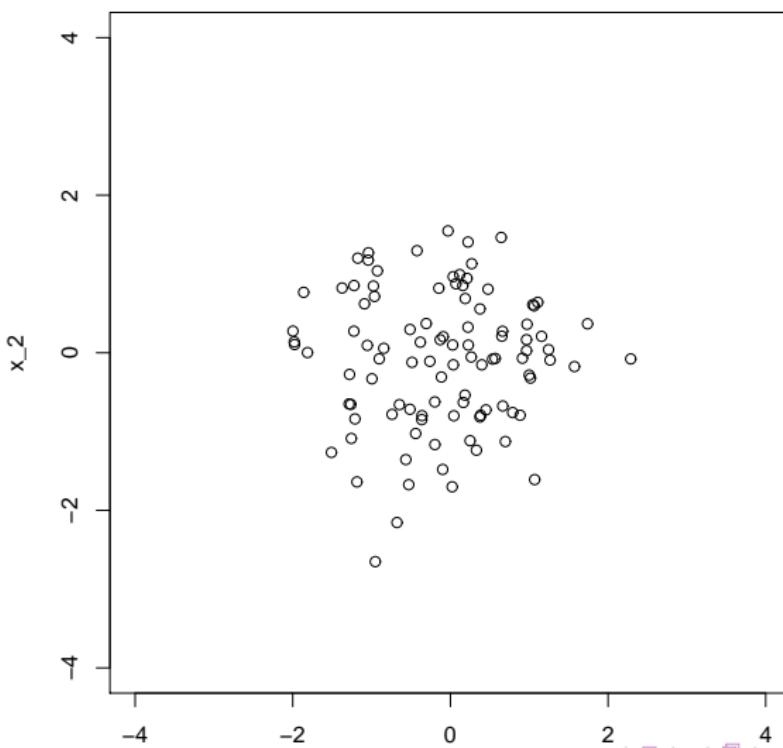
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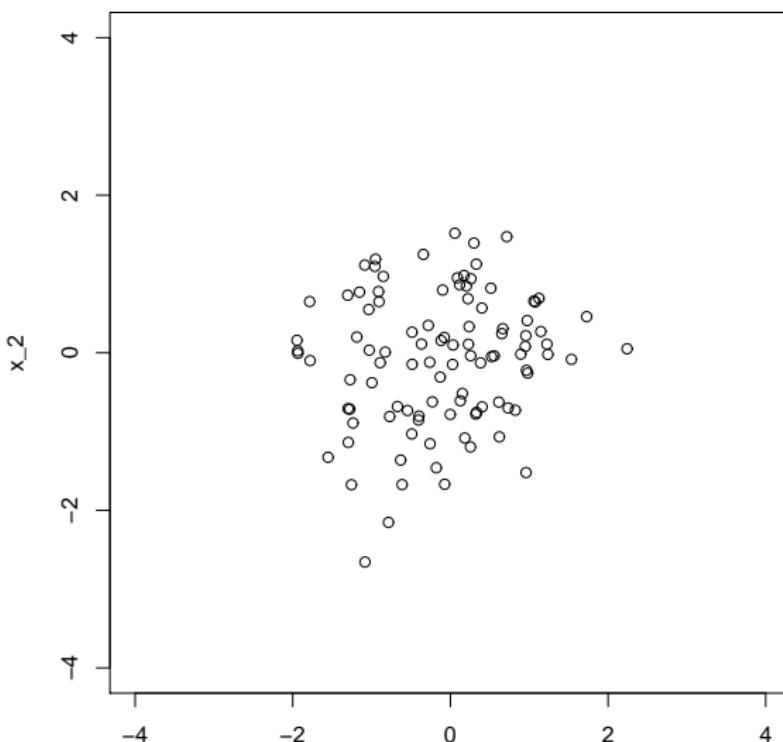
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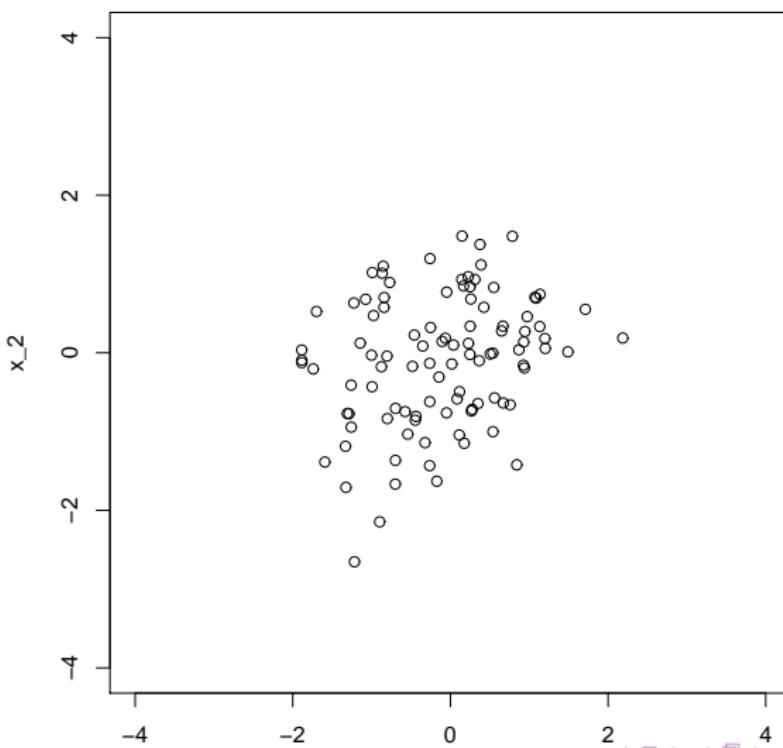
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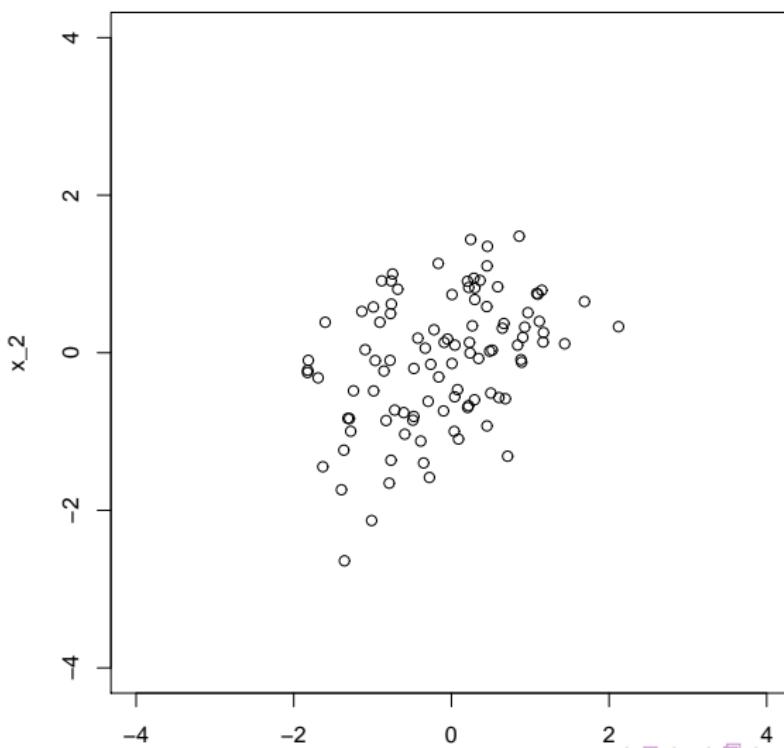
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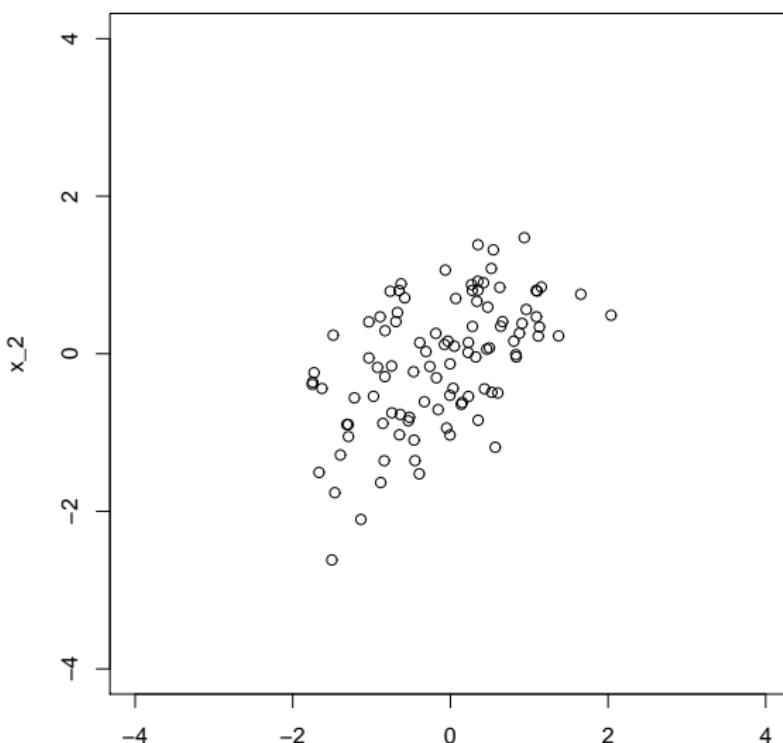
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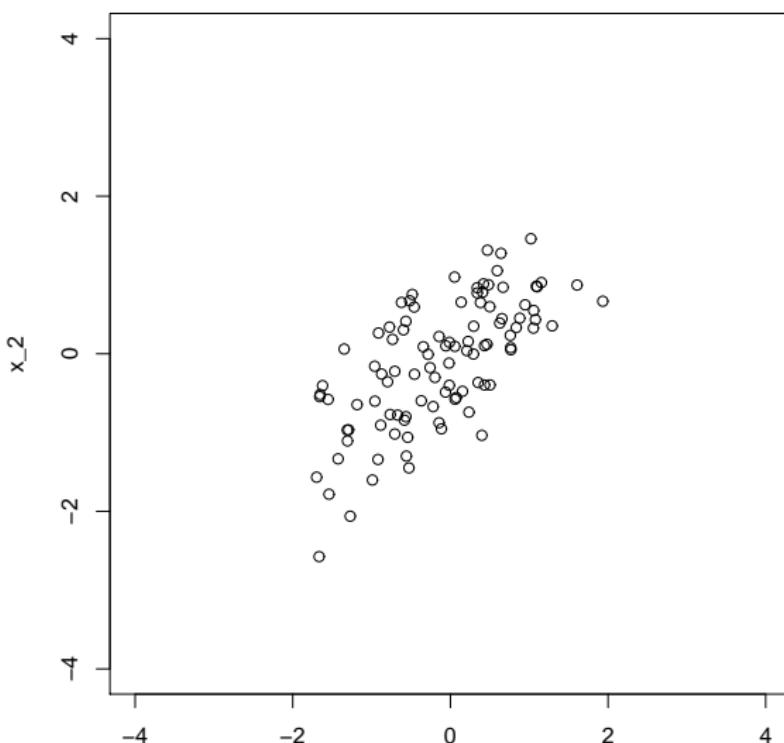
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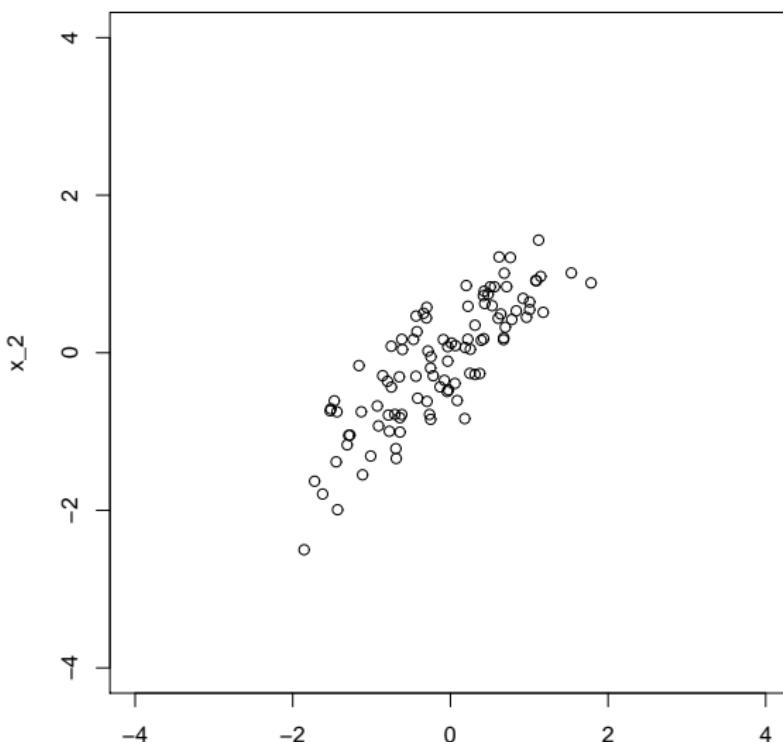
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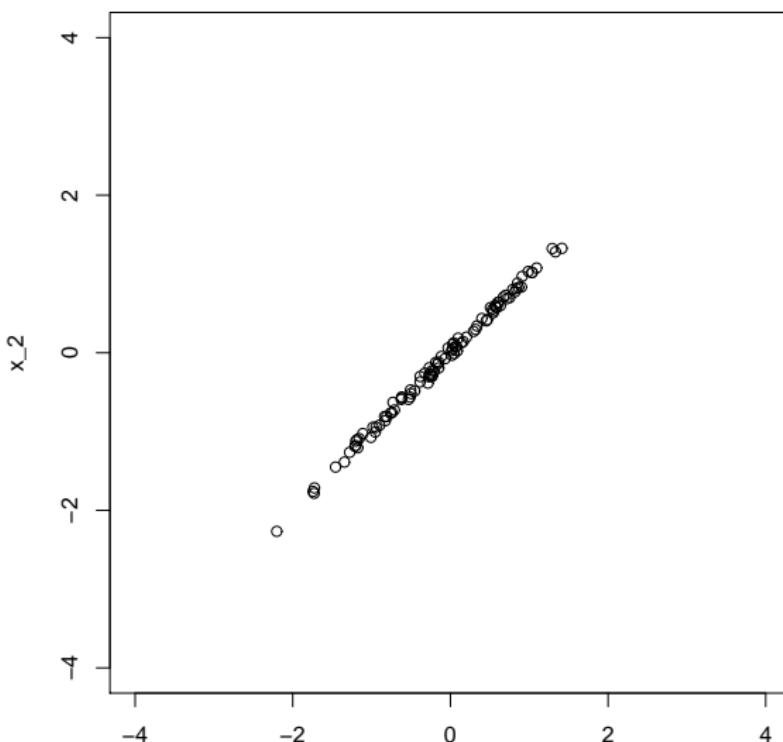
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Definición 2

Se dice que \mathbf{x} es normal multivariada si y sólo si $\forall \mathbf{t} \in \mathbb{R}^p$ se tiene que $\mathbf{t}^T \mathbf{x}$ es normal univariada.

Propiedades

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- $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \iff \mathbf{a}^T \mathbf{x} \sim N(\mathbf{a}^T \boldsymbol{\mu}, \mathbf{a}^T \boldsymbol{\Sigma} \mathbf{a}), \forall \mathbf{a} \in \mathbb{R}^p$

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- $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \iff \mathbf{a}^T \mathbf{x} \sim N(\mathbf{a}^T \boldsymbol{\mu}, \mathbf{a}^T \boldsymbol{\Sigma} \mathbf{a}), \forall \mathbf{a} \in \mathbb{R}^p$
- $\mathbf{x} = (x_1, \dots, x_p)^T \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \implies \mathbb{E}(\mathbf{x}) = \boldsymbol{\mu} \text{ y}$
 $\text{COV}(\mathbf{x}) = (\text{COV}(x_i, x_j))_{1 \leq i, j \leq p} = \boldsymbol{\Sigma}$

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- $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \iff \mathbf{a}^T \mathbf{x} \sim N(\mathbf{a}^T \boldsymbol{\mu}, \mathbf{a}^T \boldsymbol{\Sigma} \mathbf{a}), \forall \mathbf{a} \in \mathbb{R}^p$
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 $\text{COV}(\mathbf{x}) = (\text{COV}(x_i, x_j))_{1 \leq i, j \leq p} = \boldsymbol{\Sigma}$
- Sea $\mathbf{x} = (x_1, \dots, x_p)^T \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, entonces,
 x_1, \dots, x_p son independientes $\iff \boldsymbol{\Sigma}$ es diagonal.

Propiedades

Vimos que si $\mathbf{x} \sim N(\mathbf{0}, \boldsymbol{\Sigma})$, entonces

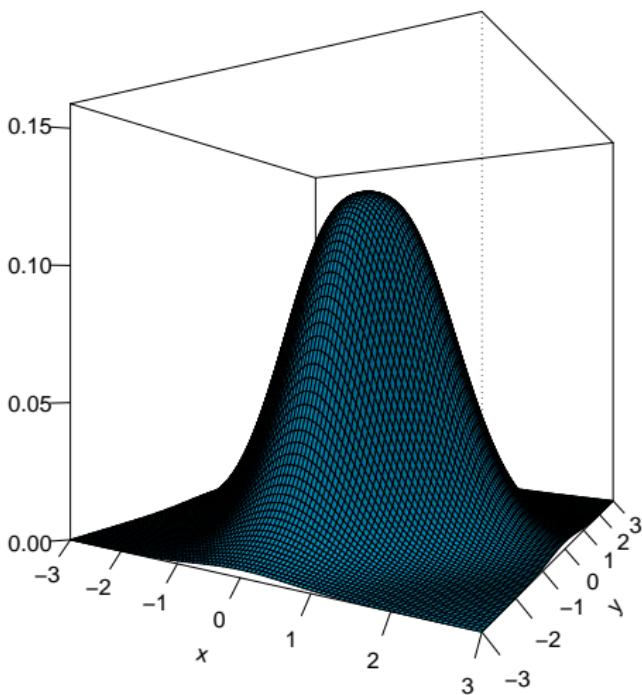
- x_1, \dots, x_p son independientes $\iff \boldsymbol{\Sigma}$ es diagonal.
- $x_j \sim N(0, \sigma_{jj})$ donde $\boldsymbol{\Sigma} = (\sigma_{ij})_{1 \leq i, j \leq p}$

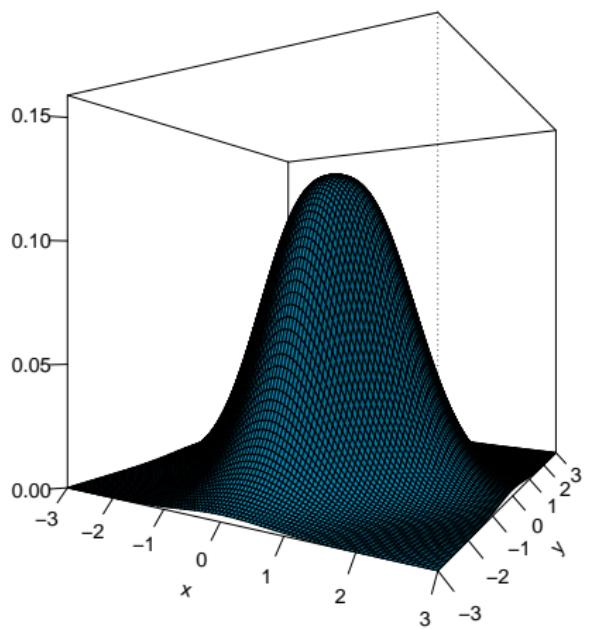
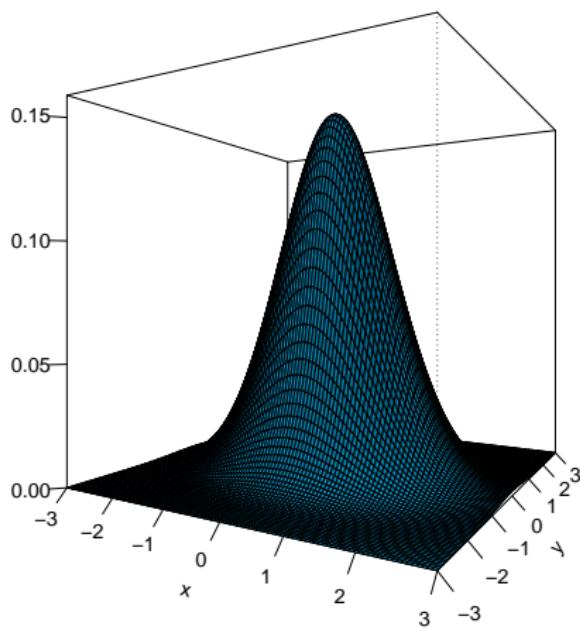
El siguiente ejemplo muestra que existen vectores aleatorios $\mathbf{x} = (x_1, x_2)$ tales que

- $\text{COV}(\mathbf{x}) = \mathbf{I}_p$.
- $x_j \sim N(0, 1)$
- x_1 y x_2 no son independientes.

Sea x con densidad

$$h(x, y) = \frac{1}{2\pi} \left\{ \left(\sqrt{2} e^{-\frac{x^2}{2}} - e^{-x^2} \right) e^{-y^2} + \left(\sqrt{2} e^{-\frac{y^2}{2}} - e^{-y^2} \right) e^{-x^2} \right\}$$



Densidad h Densidad $N(\mathbf{0}, \mathbf{I}_2)$ 

Propiedades

Sea $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ con $\boldsymbol{\Sigma} > 0$. Definamos $\mathbf{x} = \begin{pmatrix} \mathbf{x}^{(1)} \\ \mathbf{x}^{(2)} \end{pmatrix}$, $\boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}^{(1)} \\ \boldsymbol{\mu}^{(2)} \end{pmatrix}$ y $\boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix}$ con $\mathbf{x}^{(i)}, \boldsymbol{\mu}^{(i)} \in \mathbb{R}^{p_i}$, $\boldsymbol{\Sigma}_{ii} \in \mathbb{R}^{p_i \times p_i}$, $p_1 + p_2 = p$.

Entonces,

- a) $\mathbf{x}^{(1)} \sim N(\boldsymbol{\mu}^{(1)}, \boldsymbol{\Sigma}_{11})$ y $\mathbf{x}^{(2)} \sim N(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$.
- b) Más aún, $\mathbf{x}^{(1)}$ y $\mathbf{x}^{(2)}$ son independientes $\iff \boldsymbol{\Sigma}_{21} = 0$.
- c) Dada $\mathbf{A} \in \mathbb{R}^{q \times p}$, $rg(\mathbf{A}) = q \implies \mathbf{Ax} \sim N(\mathbf{A}\boldsymbol{\mu}, \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^T)$

En particular, si $\mathbf{x} \sim N(\mathbf{0}, \mathbf{I}_p)$ y $\mathbf{H} = (\mathbf{h}_1, \dots, \mathbf{h}_q) \in \mathbb{R}^{p \times q}$, es ortogonal incompleta, o sea, $\mathbf{H}^T \mathbf{H} = \mathbf{I}_q$, entonces

$$\mathbf{y} = \mathbf{H}^T \mathbf{x} \sim N(\mathbf{0}, \mathbf{I}_q).$$

Propiedades

- d) Sea $\Sigma = \mathbf{H}\Lambda\mathbf{H}^T$, con \mathbf{H} ortogonal y $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$, $\lambda_1 \geq \dots, \lambda_p$. Si $\mathbf{x} \sim N(\boldsymbol{\mu}, \Sigma) \implies \mathbf{H}^T(\mathbf{x} - \boldsymbol{\mu}) \sim N(\mathbf{0}, \Lambda)$.

Propiedades

- d) Sea $\Sigma = \mathbf{H}\Lambda\mathbf{H}^T$, con \mathbf{H} ortogonal y $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$, $\lambda_1 \geq \dots, \lambda_p$. Si $\mathbf{x} \sim N(\mu, \Sigma) \implies \mathbf{H}^T(\mathbf{x} - \mu) \sim N(\mathbf{0}, \Lambda)$.
- e) Si $\mathbf{x} \sim N(\mu, \Sigma) \iff \mathbf{x} = \mathbf{A}\mathbf{z} + \mu$ con $\mathbf{z} \sim N(\mathbf{0}, \mathbf{I}_p)$ y $\mathbf{A}\mathbf{A}^T = \Sigma$.

Propiedades

- d) Sea $\Sigma = \mathbf{H}\Lambda\mathbf{H}^T$, con \mathbf{H} ortogonal y $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$, $\lambda_1 \geq \dots, \lambda_p$. Si $\mathbf{x} \sim N(\mu, \Sigma) \implies \mathbf{H}^T(\mathbf{x} - \mu) \sim N(\mathbf{0}, \Lambda)$.
- e) Si $\mathbf{x} \sim N(\mu, \Sigma) \iff \mathbf{x} = \mathbf{Az} + \mu$ con $\mathbf{z} \sim N(\mathbf{0}, \mathbf{I}_p)$ y $\mathbf{A}\mathbf{A}^T = \Sigma$.
- f) Si $\mathbf{x} \sim N(\mu, \Sigma) \implies (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \sim \chi_p^2$
- g) Si $\mathbf{x} \sim N(\mu, \Sigma) \implies \mathbf{x}^T \Sigma^{-1} \mathbf{x} \sim \chi_p^2(\delta^2)$ con $\delta^2 = \mu^T \Sigma^{-1} \mu$

Propiedades

Sea $\mathbf{x} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}_p)$

y sea $\mathbf{H}_1 = (\mathbf{h}_1, \dots, \mathbf{h}_q) \in \mathbb{R}^{p \times q}$ ortogonal incompleta, o sea,
 $\mathbf{H}_1^T \mathbf{H}_1 = \mathbf{I}_q$.

Sea $\mathbf{H} = (\mathbf{H}_1, \mathbf{H}_2)$ ortogonal, o sea, $\mathbf{H}^T \mathbf{H} = \mathbf{H} \mathbf{H}^T = \mathbf{I}_p$

Entonces

- a) $\mathbf{z} = \mathbf{H}_1^T \mathbf{x} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}_q)$
- b) \mathbf{z} es independiente de $\mathbf{x}^T \mathbf{x} - \mathbf{z}^T \mathbf{z}$
- c) $\frac{\mathbf{x}^T \mathbf{x} - \mathbf{z}^T \mathbf{z}}{\sigma^2} \sim \chi_{p-q}^2$

Propiedades

- Las regiones de densidad constante son los elipsoides

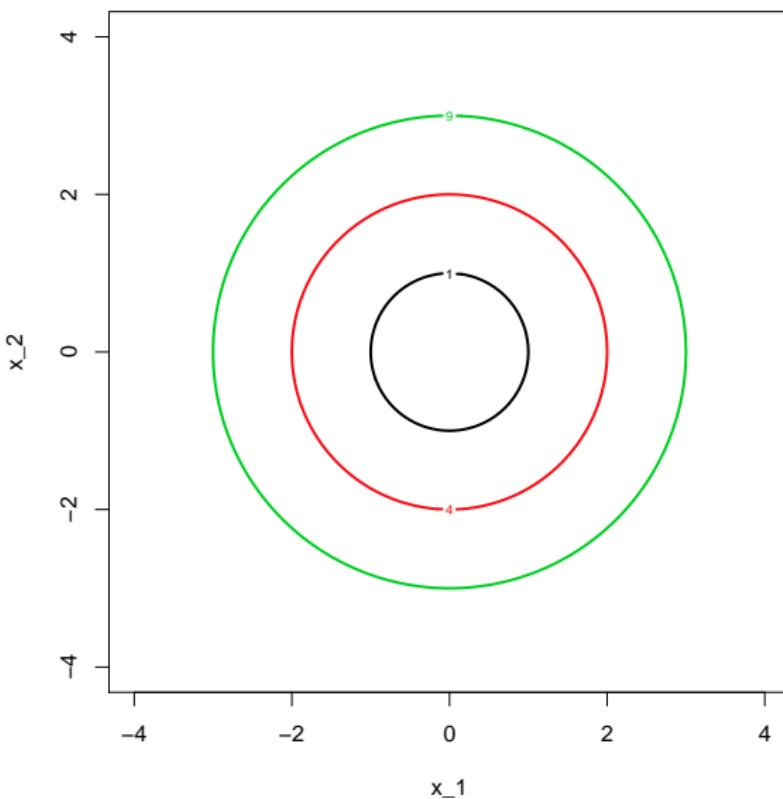
$$(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) = c^2$$

- Si $\boldsymbol{\Sigma} = \mathbf{H}\boldsymbol{\Lambda}\mathbf{H}^T$, con \mathbf{H} ortogonal y $\boldsymbol{\Lambda} = \text{diag}(\lambda_1, \dots, \lambda_p)$, definamos $\mathbf{y} = \mathbf{H}^T \mathbf{x}$ y $\boldsymbol{\nu} = \mathbf{H}^T \boldsymbol{\mu}$.

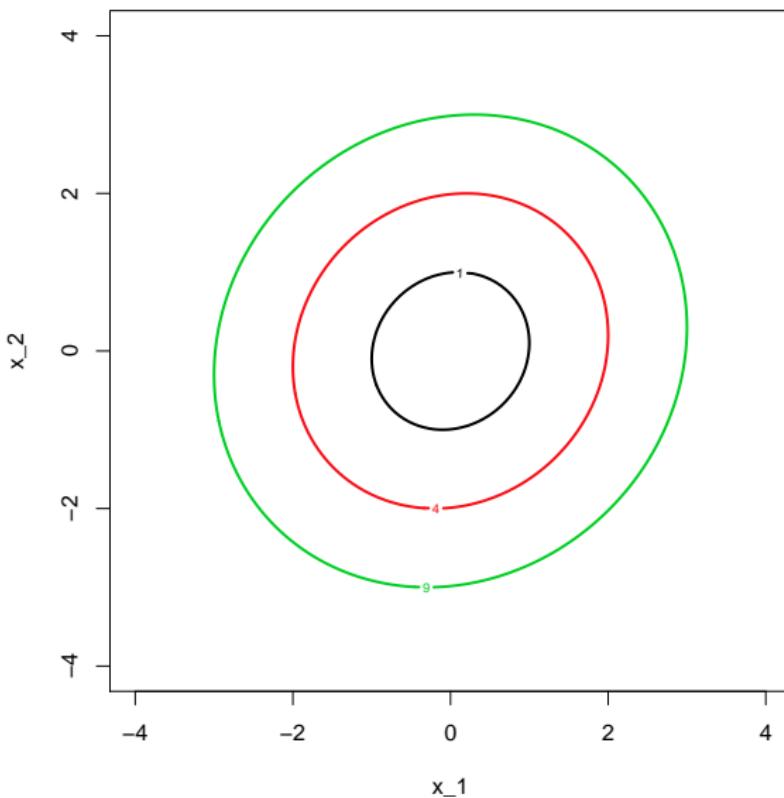
Los contornos de densidad constante son los elipsoides

centrados en $\boldsymbol{\nu}$ con ejes principales de longitud $2c\lambda_j^{\frac{1}{2}}$
soportados en los autovectores, $\mathbf{h}_1, \dots, \mathbf{h}_p$, o sea,

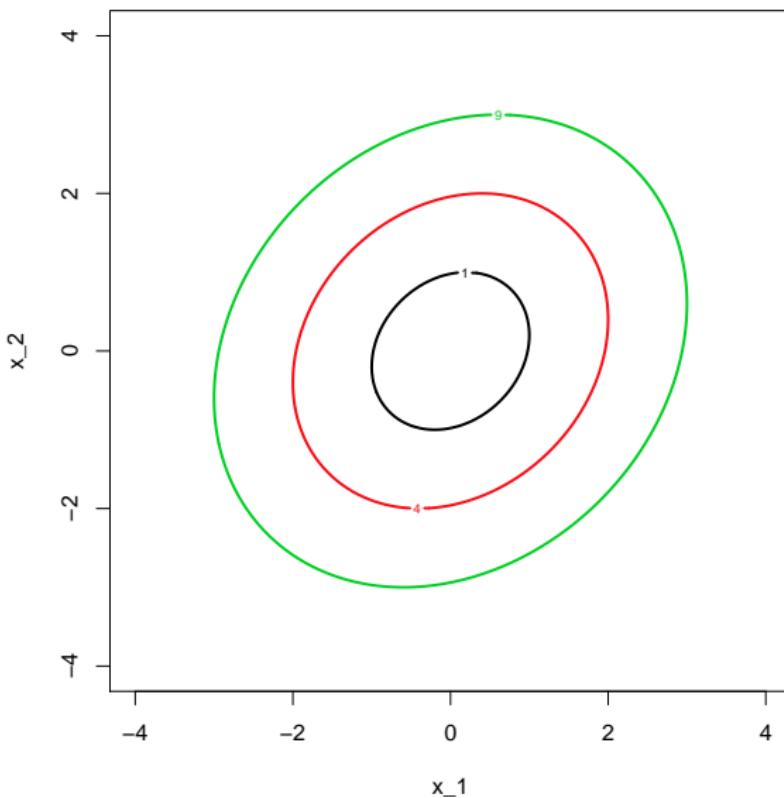
$$\sum_{j=1}^p \frac{(y_j - \nu_j)^2}{\lambda_j} = c^2$$

rho = 0

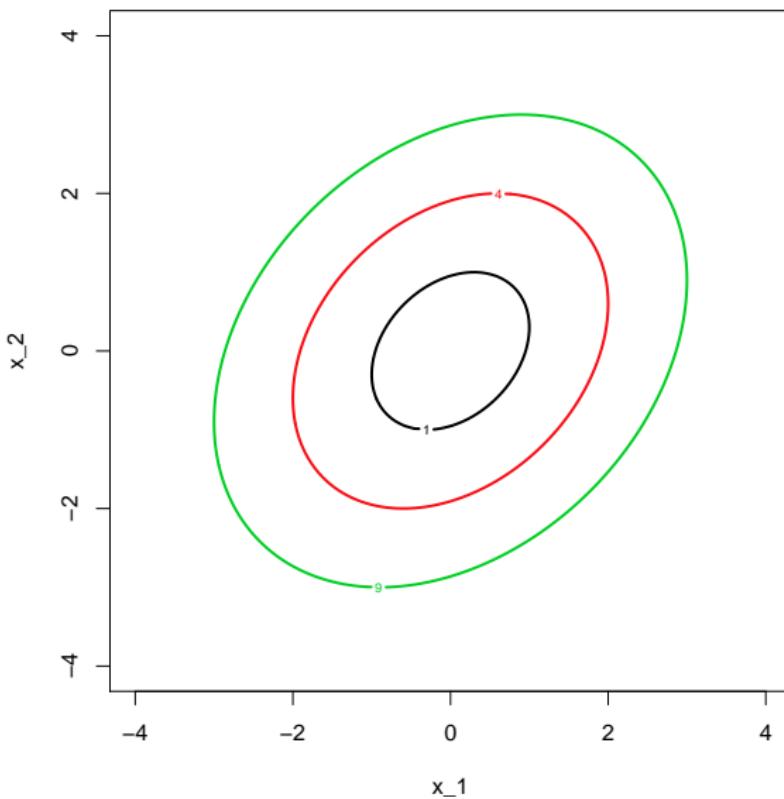
rho = 0.0999



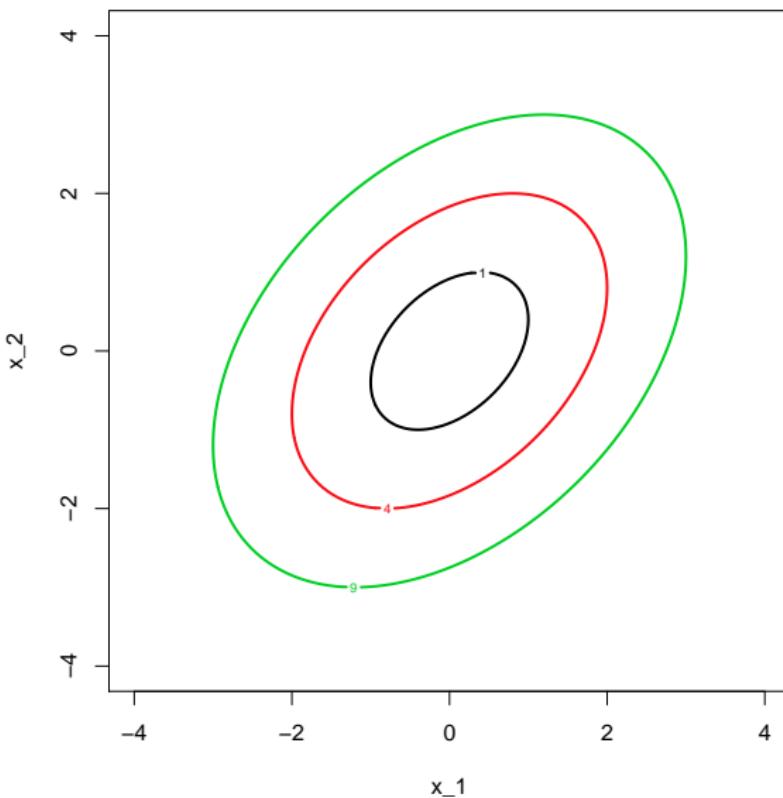
rho = 0.1998



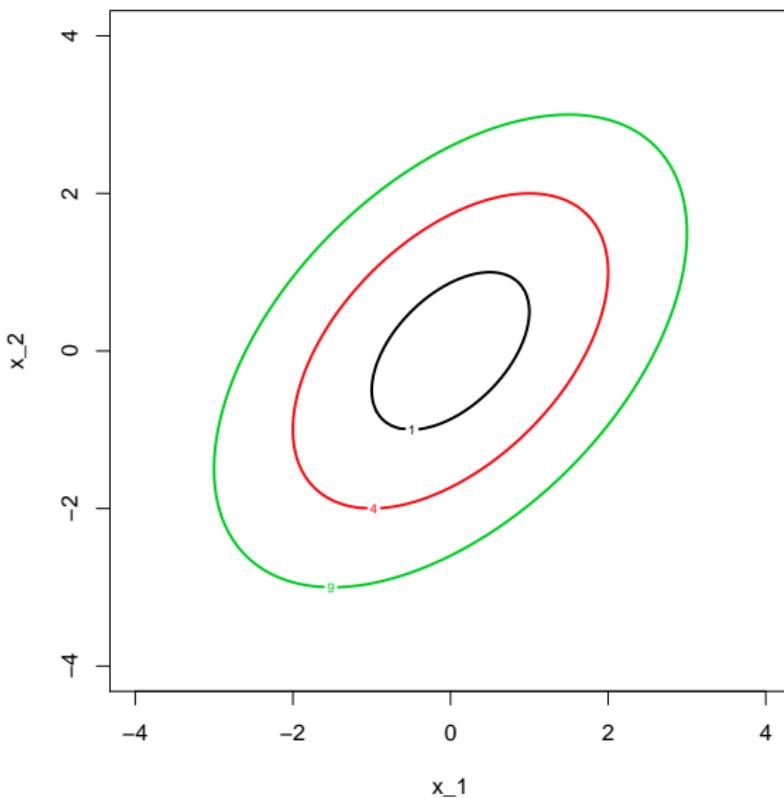
rho = 0.2997



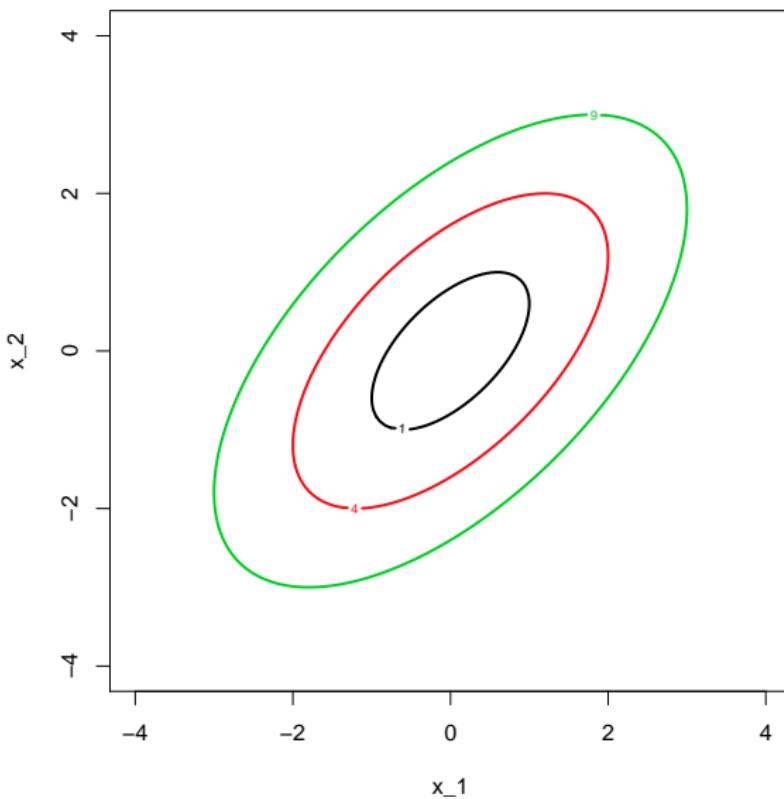
rho = 0.3996



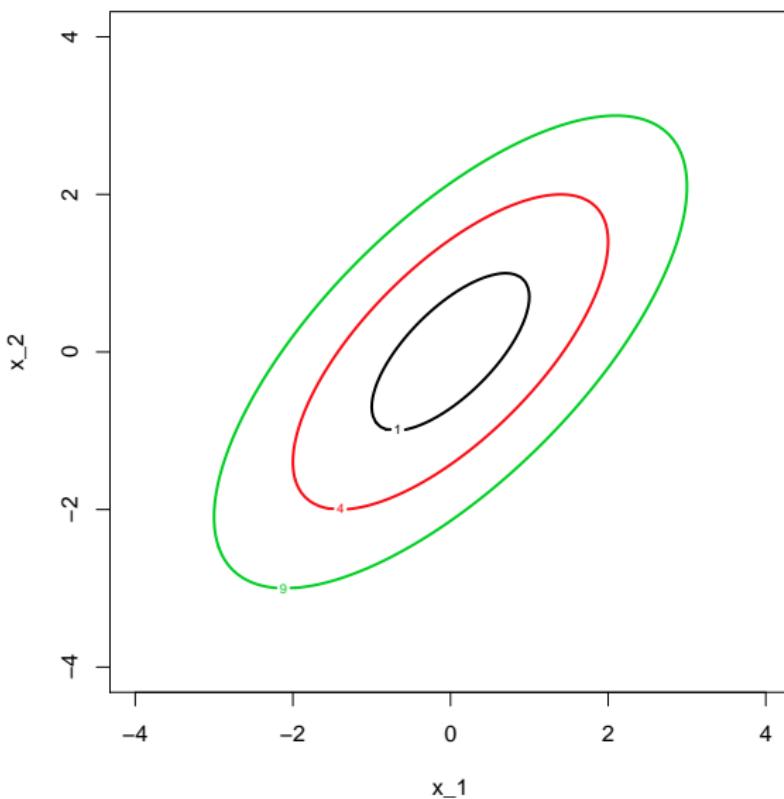
rho = 0.4995



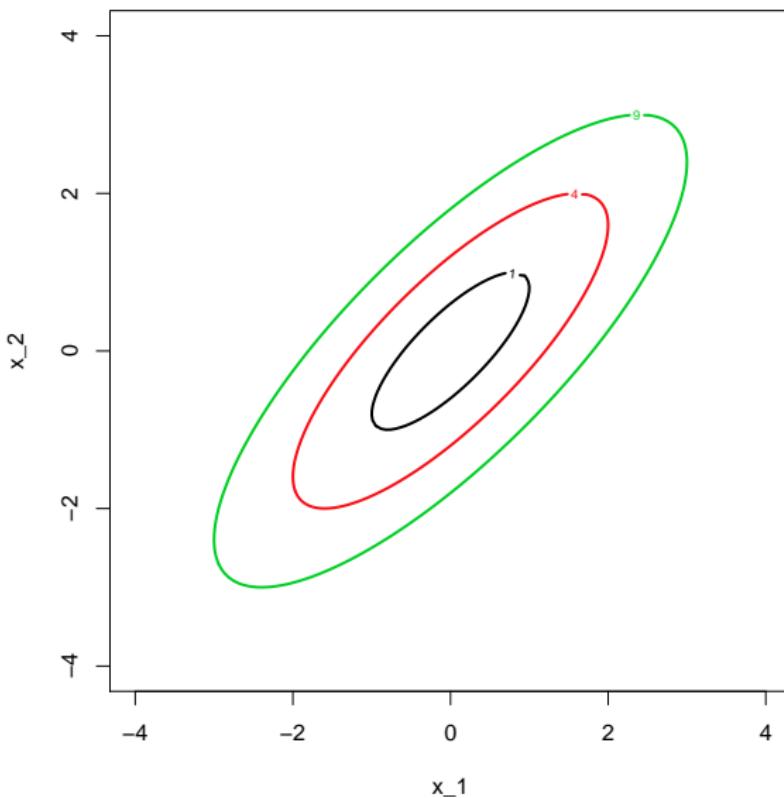
rho = 0.5994



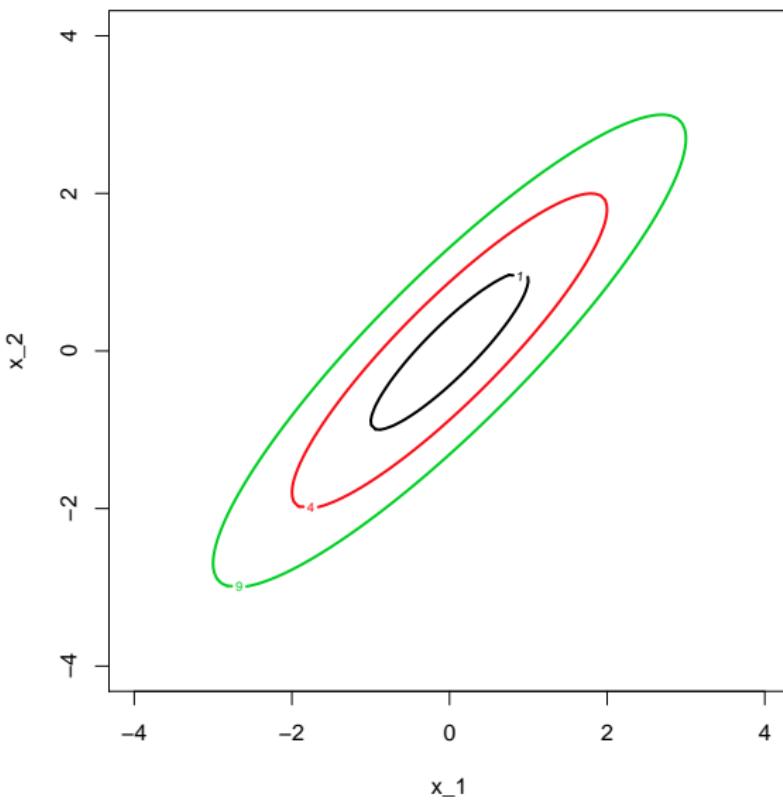
rho = 0.6993



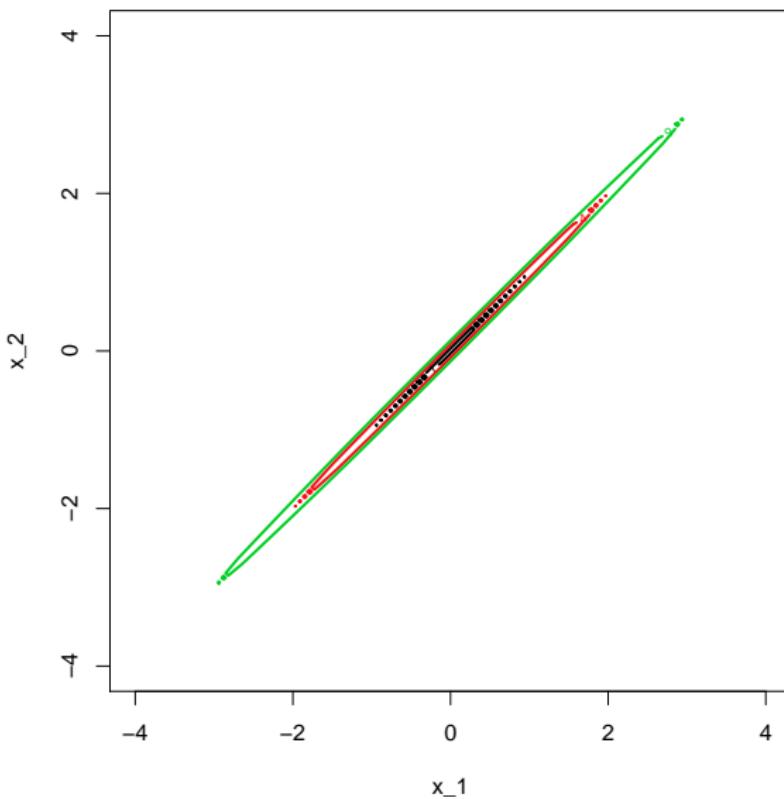
rho = 0.7992



rho = 0.8991



rho = 0.999



Teorema

Sea $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ con $\boldsymbol{\Sigma} > 0$. Definamos $\mathbf{x} = \begin{pmatrix} \mathbf{x}^{(1)} \\ \mathbf{x}^{(2)} \end{pmatrix}$,
 $\boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}^{(1)} \\ \boldsymbol{\mu}^{(2)} \end{pmatrix}$ y $\boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix}$ con $\mathbf{x}^{(i)}, \boldsymbol{\mu}^{(i)} \in \mathbb{R}^{p_i}$,
 $\boldsymbol{\Sigma}_{ii} \in \mathbb{R}^{p_i \times p_i}$, $p_1 + p_2 = p$.

Entonces,

$$\mathbf{x}^{(1)} | \mathbf{x}^{(2)} = \mathbf{x}_0 \sim N\left(\boldsymbol{\mu}^{(1)} + \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}(\mathbf{x}_0 - \boldsymbol{\mu}^{(2)}), \boldsymbol{\Sigma}_{11.2}\right)$$

donde $\boldsymbol{\Sigma}_{11.2} = \boldsymbol{\Sigma}_{11} - \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\boldsymbol{\Sigma}_{21}$.

Observación

- $\Sigma_{11.2} = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$

- $\Sigma_{22.1} = \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12}$

$$\Sigma^{-1} = \begin{pmatrix} \Sigma_{11.2}^{-1} & -\Sigma_{11}^{-1}\Sigma_{12}\Sigma_{22.1}^{-1} \\ -\Sigma_{22}^{-1}\Sigma_{21}\Sigma_{11.2}^{-1} & \Sigma_{22.1}^{-1} \end{pmatrix}$$

- $\beta = \Sigma_{12}\Sigma_{22}^{-1}$: la matriz de regresión de $x^{(1)}$ en $x^{(2)}$.
- $\Sigma_{11.2}$: matriz de covarianza parcial de $x^{(1)}$ dado $x^{(2)}$
- $\mu_{1.2} = \mu^{(1)} - \beta\mu^{(2)}$
- $x_{1.2} = x^{(1)} - \mathbb{E}(x^{(1)}|x^{(2)}) = x^{(1)} - \beta x^{(2)} - \mu_{1.2}$ es el residuo y

$$\text{COV}(x_{1.2}) = \Sigma_{11.2}$$

Caso $q = 1$

Sea $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ con $\boldsymbol{\Sigma} > 0$. Definamos $\mathbf{x} = \begin{pmatrix} \mathbf{x}_1 \\ \mathbf{x}^{(2)} \end{pmatrix}$, $\boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \boldsymbol{\mu}^{(2)} \end{pmatrix}$ y $\boldsymbol{\Sigma} = \begin{pmatrix} \sigma_{11} & \boldsymbol{\sigma}_{12} \\ \boldsymbol{\sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix}$ con $\mathbf{x}^{(2)}, \boldsymbol{\mu}^{(2)} \in \mathbb{R}^{p-1}$.

Entonces,

$$x_1 | \mathbf{x}^{(2)} = \mathbf{x}_0 \sim N\left(\mu_1 + \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}(\mathbf{x}_0 - \boldsymbol{\mu}^{(2)}), 1/\sigma^{11}\right)$$

donde $\boldsymbol{\Sigma}^{-1} = (\sigma^{ij})_{1 \leq i, j \leq p}$.

Teorema

Sea $\mathbf{x}_1, \dots, \mathbf{x}_n$ i.i.d. $\mathbf{x}_i \sim N(\mathbf{0}, \boldsymbol{\Sigma})$ con $\boldsymbol{\Sigma} > 0$. Definamos

$$\mathbf{X}^T = (\mathbf{x}_1, \dots, \mathbf{x}_n), \text{ o sea, } \mathbf{X} = \begin{pmatrix} \mathbf{x}_1^T \\ \vdots \\ \mathbf{x}_n^T \end{pmatrix} = (\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(p)}).$$

Se tiene,

- a) $\mathbf{x}^{(j)} \sim N(\mathbf{0}, \sigma_{jj} \mathbf{I}_n)$
- b) Dado $\mathbf{a} \in \mathbb{R}^n$, $\mathbf{X}^T \mathbf{a} = \sum_{i=1}^n a_i \mathbf{x}_i \sim N(\mathbf{0}, \|\mathbf{a}\|^2 \boldsymbol{\Sigma})$
- c) Dados $\mathbf{a}_\ell = (a_{\ell,1}, \dots, a_{\ell,n})^T \in \mathbb{R}^n$, $1 \leq \ell \leq r$ con $r \leq n$ ortogonales, entonces $\mathbf{X}^T \mathbf{a}_\ell = \sum_{i=1}^n a_{\ell,i} \mathbf{x}_i$ son independientes.
- d) Dado $\mathbf{b} \in \mathbb{R}^p$, $\mathbf{X} \mathbf{b} = \sum_{j=1}^p b_j \mathbf{x}^{(j)} \sim N(\mathbf{0}, \sigma_{\mathbf{b}}^2 \mathbf{I}_n)$